# Long-Range OASDI Projection Methodology 

## Intermediate Assumptions of the 2016 Trustees Report

June 2016
Office of the Chief Actuary
Social Security Administration

## A. <br> Flow Charts

## Chart 1:

Overview of Long-Range OASDI Projection Methodology


## Chart 2: Demography - Process 1



## Chart 3: Economics - Process 2



## Chart 4: Beneficiaries - Process 3



Note: Insured widow refers to widow beneficiaries who are insured for OAIB benefits, but not receiving those benefits

## Chart 5: Trust Fund Operations and Actuarial Status - Process 4



# B. Process Descriptions 

The long-range programs used to make projections for the annual Trustees Report are grouped into four major processes. These include Demography, Economics, Beneficiaries, and Trust Fund Operations and Actuarial Status. Each major process consists of a number of subprocesses. Each subprocess is described in terms of three elements:

- This overview attempts to provide a general description of the purpose of each subprocess. This element introduces key projected variables used in the subprocess. Some variables are represented as being dependent in an equation, where the dependent variable is defined in terms of one or more independent variables. Independent variables may include previously calculated dependent variables or data provided from outside the subprocess. Other key variables are referenced by " $(\cdot)$ " following the variable name. This symbol indicates that the calculation of this variable cannot easily be communicated by an equation and, thus, requires a more complex discussion.
- Input Data - Data used in the subprocess are described. These data include those from other subprocesses, ultimate long-range assumptions provided by the Board of Trustees of the OASDI Trust Funds, data from other offices of the Social Security Administration, and data from outside the Social Security Administration (e.g., estimates of the U.S. population). Data description includes data source and data detail (e.g., define age detail of data). In addition, this element includes how often additional data are expected to be received.
- Development of Output - The key variables are described in greater detail, including the level of disaggregation of the data.


## Process 1:

## Demography

## 1. Demography

The primary purpose of the Demography Process is to provide estimates of the projected Social Security area population ${ }^{1}$ for each year of the 75 -year projection period in the Trustees Report. For the 2016 report, the projection period covers the years 2016 through 2090. The Demography Process receives input data mainly from other government agencies, and provides output data to the Economics, Beneficiaries, and Trust Fund Operations and Actuarial Status processes.

The Demography Process is composed of eight subprocesses: FERTILITY, MORTALITY, LEGAL IMMIGRATION, HISTORICAL POPULATION, OTHER IMMIGRATION, MARRIAGE, DIVORCE, and PROJECTED POPULATION. The following chart displays the key outputs of each subprocess:

| Subprocess | Key Outputs |
| :---: | :---: |
| FERTILITY | - Birth rates, by age of mother |
| MORTALITY | - Probabilities of death, by age and sex |
| LEGAL IMMIGRATION | - Legal immigrants, by age and sex <br> - Legal emigrants, by age and sex <br> - Adjustments of status from other than legal (OTL) status to legal status, by age and sex |
| HISTORICAL POPULATION | - Historical estimates of the Social Security area total population, by single year of age, sex, and marital status <br> - Historical estimates of the OTL population, by single year of age and sex |
| OTHER IMMIGRATION | - OTL immigrants, by age and sex <br> - OTL emigrants, by age and sex <br> - Projected OTL populations, by single year of age, sex, and OTL type (never-authorizeds, nonimmigrants, visa-overstayers) <br> - Historical estimates of the OTL population, by single year of age, sex, and OTL type |
| MARRIAGE | - Marriage rates, by age-of-husband crossed with age-of-wife |
| DIVORCE | - Divorce rates, by age-of-husband crossed with age-of-wife |
| PROJECTED POPULATION | - Projected total populations, by age, sex, and marital status |

[^0]
### 1.1. FERTILITY

## 1.1.a. Overview

The National Center for Health Statistics (NCHS) collects data on annual numbers of births by single year of age of mother and the U.S. Census Bureau produces estimates of the resident population by single year of age for females. Age-specific birth rates $\left(b_{x}^{z}\right)$ for a given year $z$ are defined as the ratio of (1) births ( $B_{x}^{z}$ ) during the year to mothers at the specified age $x$ to (2) the midyear female population $\left(P_{x}^{z}\right)$ at that age. The total fertility rate $T F R^{z}$ summarizes the age-specific fertility rates for a given year $z$. The total fertility rate for a given year $z$ equals the sum of the age-specific birth rates for all ages $x$ during the year. One can also interpret the total fertility rate as the number of children born to a woman if she were to survive her childbearing years and experience the age-specific fertility rates of year $z$ throughout her childbearing years.

The FERTILITY subprocess combines the historical values of $b_{x}^{z}$ and $T F R^{z}$ and assumed future values of the TFR to develop projections of $b_{x}^{z}$. The primary equations of this subprocess are given below:

$$
\begin{align*}
& b_{x}^{z}=b_{x}^{z}(\cdot)  \tag{1.1.1}\\
& T F R^{z}=\sum_{x} b_{x}^{z} \tag{1.1.2}
\end{align*}
$$

## 1.1.b. Input Data

## Trustees Assumptions -

Each year the Board of Trustees of the OASDI Trust Funds sets the ultimate assumed values for the TFR. Under the intermediate assumptions underlying the 2016 Trustees Report, the ultimate TFR is 2.0 and it is assumed to be reached in 2027.

## Other input data -

- From the NCHS, annual numbers of births by age of mother ${ }^{2}(10-14,15,16,17, \ldots, 48,49-54)$ for years 1980-2014. In general, the NCHS provides an annual update including one additional year of final birth data and the previous historical years are only updated if the NCHS makes a historical revision to their data.
- From the NCHS, the preliminary TFR for 2014.
- From the U.S. Census Bureau, estimates of the July $1^{\text {st }}$ female resident population by single year of age for ages 14-49 for 1980-2014. In general, each year, Census provides updated data for years after the most recent decennial census.
- From the NCHS, historical birth rates, by single year of age of mother (14-49) for the period 19171979. No updates of these data are needed.

[^1]
## 1.1.c. Development of Output

## Equation 1.1.1 - Age-specific birth rates

The FERTILITY subprocess produces the age-specific birth rates, by childbearing ages 14 through 49, for years 1941 through the end of the 75 -year projection period. For historical years prior to 1980, age-specific birth rates come from the NCHS. For years 1980 through the remaining historical period, age-specific birth rates are calculated as: $b_{x}^{z}=\frac{B_{x}^{z}}{P_{x}^{z}}$, using birth data from the NCHS and estimates of the July $1^{\text {st }}$ female resident population from the U.S. Census Bureau.

The age-specific birth rates are projected using a process that is consistent with both the observed trends in recent data and the ultimate assumed total fertility rate. This process consists of the following steps:

1. Averaged birth rates by age group, ${ }^{3}$ designated as ${ }_{5} b_{x}^{2}$, are calculated from the age-specific birth rates $b_{x}^{z}$ for each year during the period 1980-2014.
2. To calculate the starting values of the projection process, the ${ }_{5} b_{x}^{z}$ values from the last five years of historical data are averaged using weights of $5,4,3,2$, and 1 for years 2014, 2013, 2012, 2011, and 2010, respectively.
3. For each ${ }_{5} b_{x}^{z}$ age group series, the slope of the least squares line is calculated based on a regression over the period 1989-2014.
4. For 2015, each of the seven starting values of ${ }_{5} b_{x}^{z}$ (from Step 2) is projected forward by adding 100 percent of their respective slope (from Step 3).
5. Then, a preliminary total fertility rate for $2015, T F R_{p}^{2015}$, is calculated such that, in general, it is equal to 5 times the sum of each ${ }_{5} b_{x}^{2015}$. For the age group 14-19, 6 times the sum is used since this age group actually contains one additional age.
6. For 2015, the Trustees assume this is $0.5 \%$ higher than the preliminary 2014 TFR from NCHS and thus, the Trustees set $T F R^{2015}$ equal to 1.8708 .
7. To ensure the assumed total fertility rate is achieved for 2015, each value of ${ }_{5} b_{x}^{2015}$ (from Step 4) is now multiplied by the ratio of the assumed $T F R^{2015}$ (from Step 6) and the respective value of $T F R_{p}^{2015}$ (from Step 5).
8. For 2016, each final ${ }_{5} b_{x}^{2}$ for 2015 is projected forward by adding 92.3 percent of the respective slope (from Step 3). For subsequent projection years (2016-2040), an arithmetically decreasing portion of the slopes, until $0 \%$ is reached, ${ }^{4}$ is added to the previous year's final values of ${ }_{5} b_{x}^{z}$ to get preliminary values of ${ }_{5} b_{x}^{z}$.

[^2]9. For years 2016 and later, a preliminary total fertility rate, $T F R_{p}^{z}$, is calculated from the preliminary values of ${ }_{5} b_{x}^{z}$ in Step 8 and is calculated in the same manner as in Step 5.
10. Then, for each year, an adjustment is made so that the annual $T F R^{z}$ is consistent with the Trustees’ assumed TFRs. As mentioned in Step 6, $T F R^{2015}$ is assumed to be 1.8708 . From $2016-2023, T F R^{2}$ is assumed to increase consistent with the economic recovery. For years after 2023, TFR ${ }^{2}$ is assumed to decrease linearly from $T F R^{2023}$ until reaching the ultimate value in 2027.
11. To ensure the assumed total fertility rate is achieved, each value of ${ }_{5} b_{x}^{z}$ (Step 8 ) is multiplied by the ratio of the assumed $T F R^{z}$ (Step 10) and the respective value of $T F R_{p}^{z}$ (Step 9).
12. The final step of the projection method disaggregates the adjusted ${ }_{5} b_{x}^{z}$ into single age birth rates by multiplying the final ${ }_{5} b_{x}^{z}$ values for each year (Steps 7 or 11) by the ratio of the single year $b_{x}^{z}$ to the ${ }_{5} b_{x}^{2}$ for each of the respective ages and age groups as calculated in the last year of complete historical data (Step 1). Then, minor adjustments correct rounding issues.

### 1.2 MORTALITY

## 1.2.a. Overview

The National Center for Health Statistics (NCHS) collects data on annual numbers of deaths and the U.S. Census Bureau produces estimates of the U.S. resident population. Central death rates ( ${ }_{y} M_{x}$ ) are defined as the ratio of (1) the number of deaths occurring during the year to persons between exact ages $x$ and $x+y$ to (2) the midyear population between exact ages $x$ and $x+y$. For historical years prior to 1968, ${ }_{y} M_{x}$ are calculated from NCHS and Census data by sex. For historical years beginning in 1968, the same data are used in the calculations for ages under 65, but data from the Centers for Medicare and Medicaid Services (CMS) are used for ages 65 and over. Based on death by cause data from the NCHS, the ${ }_{y} M_{x}$ are distributed by cause of death for years 1979 and later. ${ }^{5}$

Over the last century, death rates have decreased substantially. The historical improvement in mortality is quantified by calculating the average annual percentage reduction ( ${ }_{y} A A_{x}$ ) in the central death rate. In order to project future ${ }_{y} M_{x}$, the Board of Trustees of the OASDI Trust Funds determines the ultimate average annual percentage reduction that will be realized during the projection period ( ${ }_{y} A A_{x}^{u}$ ) for each sex and cause of death.

The basic mortality outputs of the MORTALITY subprocess that are used in projecting the population are probabilities of death by age and sex $\left(q_{x}\right)$. The probability that a person age x will die within one year $\left(q_{x}\right)$ is calculated from the central death rates (the series of ${ }_{y} M_{x}$ ).

Period life expectancy ( $\stackrel{\circ}{\mathrm{e}}_{\mathrm{x}}$ ) is defined as the average number of years of life remaining for people who are age x and are assumed to experience the assumed probabilities of death throughout their lifetime. It is generated from the probabilities of death for a given year and is a summary statistic of overall mortality for that year.

Age-adjusted death rates ( $A D R$ ) are also used to summarize the mortality experience of a single year, making different years comparable to each other. Age-adjusted death rates are a weighted average of the ${ }_{y} M_{x}$, where the weights used are the numbers of people in the corresponding age groups of the standard population, the 2010 U.S. Census resident population ( ${ }_{y} S P_{x}$ ). Thus, if the age-adjusted death rate for a particular year and sex is multiplied by the total 2010 U.S. Census resident population, the result gives the number of deaths that would have occurred in the 2010 U.S. Census resident population if the ${ }_{y} M_{x}$ for that particular year and sex had been experienced. Age-sex-adjusted death rates (ASDR) are calculated to summarize death rates for both sexes combined and are calculated as a weighted average of the ${ }_{y} M_{x}$, where each weight is the number of people in the corresponding age and sex group of the 2010 U.S. Census resident population.

MORTALITY projects annual ${ }_{y} \mathrm{M}_{\mathrm{x}}$, which are then used to calculate the program's additional outputs. The equations for this subprocess, 1.2.1 through 1.2.6, are given below:

$$
\begin{equation*}
{ }_{y} M_{x} \quad=\quad{ }_{y} M_{x}(\cdot) \tag{1.2.1}
\end{equation*}
$$

[^3]\[

$$
\begin{gather*}
y_{y} A A_{x}={ }_{y} A A_{x}(\cdot)  \tag{1.2.2}\\
q_{x}=q_{x}(\cdot)  \tag{1.2.3}\\
\stackrel{\circ}{\mathrm{e}}_{\mathrm{x}}={ }_{\mathrm{e}}^{\mathrm{e}}(\mathrm{P} \cdot)  \tag{1.2.4}\\
A D R_{s}^{z}=\frac{\sum_{x} S P_{x} \cdot{ }_{y} M_{x, s}^{z}}{\sum_{x}{ }_{y} S P_{x}}  \tag{1.2.5}\\
A S D R^{z}=  \tag{1.2.6}\\
\sum_{s} \sum_{x}{ }_{y} S P_{x, s} \cdot{ }_{y} M_{x, s}^{z} \\
\sum_{s} \sum_{x} S P_{x, s}
\end{gather*}
$$
\]

where ${ }_{y} M_{x, S}^{z}$ refers to the central death rate between exact age $x$ and $x+y$, by sex, in year $z$;
${ }_{y} S P_{x}$ denotes the number of people in the standard population (male and female combined) who are between exact age $x$ and $x+y$; and ${ }_{y} S P_{x, S}$ denotes the number of people, by sex, in the standard population who are between exact age $x$ and $x+y$.

## 1.2.b. Input Data

## Trustees Assumptions -

Each year the Board of Trustees of the OASDI Trust Funds sets the ultimate assumed values for the ${ }_{y} A A_{x}$ by sex, age group, ${ }^{6}$ and cause of death. ${ }^{7}$ The average annual percentage reductions reach their ultimate values in the $25^{\text {th }}$ year of the 75 -year projection period. The ultimate rates of reduction by sex, age group, and cause of death can be found in Appendix 1.2-1.

## NCHS Data -

- Annual numbers of registered deaths by sex and age group for the period 1900-1978. These data are not updated. Registered deaths refer to deaths in the Death Registration area. Since 1933, the Death Registration area has included all of the U.S.
- Annual numbers of deaths by sex, age group, and cause for the period 1979-2013. Generally, a new year of data is received each year. In addition, revised data are often available for years beginning with 1999. (1999 was the starting year of the latest international classification of diseases - ICD10.)
- The monthly number of births, by sex, for years 1938-2013. These data are updated annually, when the NCHS provides an additional year of data.
- The number of infant deaths, by age, sex, and age group, ${ }^{8}$ for years 1938-2013. These data are updated annually, when the NCHS provides an additional year of data.

[^4]- Deaths for 1995 and 1996 by sex, 4 marital statuses, and 21 age groups. The age groups are generally 5 -year age groups and are as follows: $0,1-4,5-9,10-14, \ldots$, and $95+$ ). These data are updated as resources are available.
- The population of states in the Death Registration area by age group ${ }^{9}$ and sex, for years 1900-1939. These data are not updated.
- The number of registered deaths, by sex and age groups (85-89, 90-94, and 95+), for the years 19001967. These data are not updated.


## U.S. Census Bureau Data -

- Estimates of the July 1 resident population by single year of age ( 0 through $100+$ ) for years 1980 2013. Each year, Census provides an additional year of data and updated data for years after the most recent decennial census.
- From the Current Population Survey (CPS), the population by sex, marital status, and age group ${ }^{10}$ for the years 1995 and 1996. These data are updated as resources are available.
- The resident population by sex, marital status, and age group, ${ }^{11}$ as of as of July 1, 1995 and 1996. These data are updated when new NCHS death data by marital status are incorporated.
- The resident population at ages 75-79 and 80-84, by sex, for years 1900-1940 (at ten year intervals). These data are not updated.
- The resident population, by sex and age group, ${ }^{12}$ for 1940-2000. These data are not updated.


## CMS Data -

- Annual numbers of deaths, by sex and single year of age (ages 65 and over), for the period 19682013. These data are updated annually, when the CMS provides an additional year of preliminary data and replaces the prior year’s preliminary with final data.
- Annual numbers of Medicare enrollments (who are insured for Social Security benefits), by sex and single year of age (ages 65 and over), for the period 1968-2014. These data are updated annually, when the CMS provides an additional year of preliminary data and replaces the prior year's preliminary with final data.


## Other input data -

- From a previous year’s Trustees Report, the July 1, 1995 and 1996, Social Security area population by sex, marital status, and single year of age (5 through 100+). These data are updated when new NCHS death data by marital status are incorporated.

[^5]
## 1.2.c. Development of Output

Equation 1.2.2 - Average Annual Percentage Reduction in the Central Death Rates ( ${ }_{y} A A_{x}$ )

The ${ }_{y} A A_{x}$, by sex and cause, are calculated based on the decline in the ${ }_{y} M_{x}$ for the period 2003 through 2013, and distributed by 21 age groups, ${ }^{13} 2$ sexes, and 5 causes of death. ${ }^{14}$ The values are calculated as the complement of the exponential of the slope of the least-squares line through the logarithms of the ${ }_{y} M_{x}$.

The ultimate assumed values for the central death rates ( ${ }_{y} A A_{x}^{u}$ ), as set by the Board of Trustees of the OASI and DI Trust Funds, are assumed to be reached in the $25^{\text {th }}$ year of the 75 -year projection period. The assumed ultimate values are specified by five causes of death for the following five age groups: under age 15, 15-49, 5064, 65-84, and 85 and older. Male and female values are set equal.

The starting values of ${ }_{y} A A_{x}$, by the 21 age groups, sex, and cause, are assumed to equal the average ${ }_{y} A A_{x}$ based on the decline in the ${ }_{y} M_{x}$ for the period 2003-2013 when that average is non-negative. If the historical average is negative, then the starting values are assumed to be 75 percent of the average. Available Medicare preliminary data is used for overall levels with the last available NCHS data year cause of death percentages carrying forward. A method of graduation is used that causes the absolute difference between the starting values of ${ }_{y} A A_{x}$ and the ultimate values of ${ }_{y} A A_{x}^{u}$ to decrease rapidly until it reaches the Trustees' ultimate assumed values, ${ }_{y} A A_{x}^{u}$. This is accomplished by repeating the following steps for each historical year after 2013, and the first 25 years of the projection:

1. The absolute value of the distance between the prior year's calculated ${ }_{y} A A_{x}$ and the ultimate assumed ${ }_{y} A A_{x}^{u}$ is calculated.
2. If the ultimate assumed ${ }_{y} A A_{x}^{u}$ is greater than the prior year's ${ }_{y} A A_{x}$, then 80 percent of the difference is subtracted from the ultimate assumed ${ }_{y} A A_{x}^{u}$. If the ultimate assumed ${ }_{y} A A_{x}^{u}$ is less than the prior year's ${ }_{y} A A_{x}$, then 80 percent of the difference is added to the ultimate assumed ${ }_{y} A A_{x}^{u}$.
3. These steps are repeated until the $25^{\text {th }}$ year at which time the ${ }_{y} A A_{x}$ are set equal to their ultimate assumed values, ${ }_{y} A A_{x}^{u}$.
[^6]
## Equation 1.2.1 - Central Death Rates ( ${ }_{y} M_{x}$ )

Values of ${ }_{y} M_{x}$ are determined for each historical and projected year by the 21 age groups, 2 sexes, and 5 causes of death. The starting year for the projections of the ${ }_{y} M_{x}$ is 2013 , and is the most recent data year in the historical period. However, instead of using the historical data for ${ }_{y} M_{x}$ in this year as the starting point for mortality projections, starting $y_{x} M_{x}$ values are calculated to be consistent with the trend inherent in the last 12 years of available data. Each starting value for the ${ }_{y} M_{x}$, by sex and cause of death, is computed as the value for the most recent year falling on a weighted least square line, where ${ }_{y} M_{x}$ is regressed on year, over the last 12 years. The weights are $0.2,0.4,0.6$, and 0.8 for the earliest four years of the 12 years and are 1.0 for all other years.

For years after 2013, ${ }_{y} M_{x}$ are projected, by sex and cause of death, by applying the respective ${ }_{y} A A_{x}$ to the prior year ${ }_{y} M_{x}$.

## Equations 1.2.3 - Probabilities of death $\left(q_{x}\right)$

In order to project population by age and sex, probabilities of death are applied to determine the projected number of deaths that will occur in the population. These probabilities, denoted as $q_{x}$, reflect the probability a person age x will die within one year, where $x$ refers to age last birthday as of the beginning of each year. For each year in the historical and projection period, separate $q_{x}$ series are estimated by sex.

Different methods of projecting $q_{x}$ are used for age 0 , for ages 1 through 4 , for ages 5 through 94 , and for ages 95 and above. The following descriptions provide a brief discussion of these different methods. Additional detail is provided in Actuarial Study number 120. This study, titled Life Tables for the United States Social Security area 1900-2100, can be accessed at the following internet site: http://www.socialsecurity.gov/OACT/NOTES/s2000s.html. (Choose study number 120.)

- Values for $q_{x}$ at Age 0: During the first year of life, mortality starts at an extremely high level, which becomes progressively lower. This is unlike mortality at other ages, which does not change very much within a single year of age. Thus, it is particularly important at age 0 to estimate accurately the pattern of mortality throughout the year of age, as described above, for the calculation of $q_{0}$. For the period 1940 through the last historical year, $q_{0}$ is calculated directly from tabulations of births by month and from tabulations of deaths at ages $0,1-2,3-6,7-28$ days, 1 month, 2 months, $\ldots$, and 11 months. After the last historical year, $q_{0}$ is calculated from ${ }_{1} M_{0}$, assuming that the ratio of $q_{0}$ to ${ }_{1} M_{0}$ measured for the last historical year would remain constant thereafter.
- Values for $q_{x}$ at Ages 1-4: For the period 1940 through the last year of historical data, probabilities of death at each age 1 through $4\left(q_{x}, \mathrm{x}=1,2,3,4\right)$ are calculated from tabulations of births by year and from tabulations of deaths at ages $1,2,3$, and 4 years. After the last historical year, each $q_{x}$ (where $\mathrm{x}=1,2,3$, 4) is calculated from ${ }_{4} M_{1}$ assuming that the ratio of $q_{x}$ to ${ }_{4} M_{1}$ measured for the last historical year would remain constant thereafter.
- Values for $q_{x}$ at Ages 5-94: Probabilities of death for these ages are calculated from the projected central death rates, ${ }_{5} M_{x}$. As mentioned above, the calculations are discussed in detail in Actuarial Study number 120.
- Values for $q_{x}$ at Ages 95+: It has been observed that the mortality rates of women, though lower than those of men, tend to increase faster with advancing age than those of men. An analysis of Social Security charter Old-Age Insurance beneficiaries has shown that at the very old ages mortality increases about five
percent per year of age for men and about six percent per year for women. For men, probabilities of death at each ages 95 and older are calculated as follows:

$$
\begin{array}{ll}
q_{x}=q_{x-1} \cdot\left(\frac{q_{94}}{q_{93}} \cdot \frac{99-x}{5}+1.05 \cdot \frac{x-94}{5}\right) & x=95,96,97,98,99 \\
q_{x}=1.05 \cdot q_{x-1} & x=100,101,102, \ldots
\end{array}
$$

For women, the same formulas are used, except that 1.06 is substituted for 1.05. The larger rate of growth in female mortality would eventually, at a very high age, cause female mortality to be higher than male mortality. At the point where this crossover would occur, female mortality is set equal to male mortality.

The values of $q_{x}$ used in projecting the population are based on age last birthday and are calculated by sex for $1 / 2 q_{0}$ (neonatal) and for $q_{x}$, where $x$ represents age last birthday for ages 0 through 100 (with 100 representing the age group 100 and older). Because life table values of probabilities of death are based on exact ages, values for $q_{x}$ representing age last birthday are derived as follows:

$$
\begin{array}{ll}
1 / 2 q_{0}=1-L_{0} / l_{0} & \text { for neonatal } \\
\mathrm{q}_{\mathrm{x}}=1-\mathrm{L}_{\mathrm{x}+1} / \mathrm{L}_{\mathrm{x}} & \text { for ages } 0 \text { to } 99 \\
\mathrm{q}_{100}=1-\mathrm{T}_{101} / \mathrm{T}_{100} & \text { for age group } 100 \text { and older }
\end{array}
$$

See Actuarial Study number 120 for the definitions of the life table terms. This study can be accessed at the following internet site: http://www.socialsecurity.gov/OACT/NOTES/s2000s.html. (Choose study number 120; then section IV.A in the table of contents.)

In addition, probabilities of death are broken down further into marital status. Historical data indicate differential in mortality by marital status is significant. To reflect this, projected relative differences in death rates by marital status are projected to be the same as observed during calendar years 1995 and 1996.

## Equation 1.2.4 -Life expectancy

Actuarial Study number 120 presents background information on the calculation of life expectancy, ${ }^{\circ}{ }_{\mathrm{e}}^{\mathrm{x}}$, from the probabilities of death ( $q_{x}$ ). This study can be accessed at the following internet site: http://www.socialsecurity.gov/OACT/NOTES/s2000s.html. (Choose study number 120; then IV.A in the table of contents.)

## Appendix: 1.2-1

The Board of Trustees of the OASDI Trust Funds sets the ultimate rates of mortality reduction by age group and cause of death. For comparison purposes, rates are also presented for two historical periods. Note that although the ultimate rates are the same for males and females, the historical rates differ.

|  | Historical |  | Alternative II* |  | Historical |  | Alternative II* |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | 2015 TR | 2016 TR |  |  | 2015 TR | 2016 TR |
|  | 1979 to 2013 | 2003 to 2013 | 2039-2089 | 2040-2090 | 1979 to 2013 | 2003 to 2013 | 2039-2089 | 2040-2090 |
| Under Age 15 | Male |  |  |  | Female |  |  |  |
| Cardiovascular Disease | 2.73 | 2.08 | 2.3 | 2.3 | 2.63 | 2.16 | 2.3 | 2.3 |
| Cancer | 2.37 | 1.90 | 1.5 | 1.5 | 1.93 | 1.46 | 1.5 | 1.5 |
| Violence | 2.84 | 2.44 | 1.0 | 1.0 | 2.37 | 2.67 | 1.0 | 1.0 |
| Respiratory Disease | 2.59 | 1.50 | 2.0 | 2.0 | 2.62 | 1.25 | 2.0 | 2.0 |
| Other | 2.34 | 2.36 | 1.7 | 1.7 | 2.22 | 1.98 | 1.7 | 1.7 |
| Resulting Total ** | 2.46 | 2.32 | 1.56 | 1.54 | 2.26 | 2.04 | 1.58 | 1.58 |
| Ages 15-49 | Male |  |  |  | Female |  |  |  |
| Cardiovascular Disease | 1.99 | 1.84 | 1.5 | 1.5 | 1.22 | 1.71 | 1.5 | 1.5 |
| Cancer | 1.87 | 2.36 | 1.5 | 1.5 | 1.66 | 1.80 | 1.5 | 1.5 |
| Violence | 0.83 | 0.89 | 0.7 | 0.7 | 0.05 | -0.11 | 0.7 | 0.7 |
| Respiratory Disease | 0.84 | 0.71 | 0.5 | 0.5 | -0.21 | -0.44 | 0.5 | 0.5 |
| Other | 0.53 | 3.08 | 0.8 | 0.8 | -0.37 | 1.42 | 0.8 | 0.8 |
| Resulting Total ** | 1.15 | 1.71 | 0.88 | 0.88 | 0.58 | 1.04 | 0.94 | 0.95 |
| Ages 50-64 | Male |  |  |  | Female |  |  |  |
| Cardiovascular Disease | 3.07 | 2.09 | 2.2 | 2.2 | 2.65 | 2.44 | 2.2 | 2.2 |
| Cancer | 1.64 | 1.58 | 1.5 | 1.5 | 1.39 | 1.80 | 1.5 | 1.5 |
| Violence | 0.09 | -2.21 | 0.5 | 0.5 | -0.57 | -3.07 | 0.5 | 0.5 |
| Respiratory Disease | 1.09 | 0.11 | 0.7 | 0.7 | -0.47 | -0.13 | 0.7 | 0.7 |
| Other | -0.43 | -0.42 | 0.6 | 0.6 | -0.45 | -0.14 | 0.6 | 0.6 |
| Resulting Total ** | 1.64 | 0.83 | 1.06 | 1.07 | 1.15 | 1.03 | 1.06 | 1.07 |
| Ages 65-84 | Male |  |  |  | Female |  |  |  |
| Cardiovascular Disease | 3.25 | 3.84 | 2.2 | 2.2 | 2.90 | 4.01 | 2.2 | 2.2 |
| Cancer | 0.88 | 1.94 | 0.9 | 0.9 | -0.02 | 1.31 | 0.9 | 0.9 |
| Violence | 0.60 | -0.03 | 0.5 | 0.5 | -0.08 | -0.25 | 0.5 | 0.5 |
| Respiratory Disease | 0.64 | 1.55 | 0.3 | 0.3 | -1.95 | 0.51 | 0.3 | 0.3 |
| Other | -0.78 | -0.43 | 0.3 | 0.3 | -1.50 | -0.59 | 0.3 | 0.3 |
| Resulting Total ** | 1.59 | 1.98 | 0.76 | 0.77 | 0.76 | 1.57 | 0.71 | 0.72 |
| Ages 85 and older | Male |  |  |  | Female |  |  |  |
| Cardiovascular Disease | 1.71 | 3.15 | 1.2 | 1.2 | 1.88 | 3.54 | 1.2 | 1.2 |
| Cancer | -0.23 | 1.22 | 0.5 | 0.5 | -0.50 | 0.55 | 0.5 | 0.5 |
| Violence | -0.64 | -0.43 | 0.3 | 0.3 | -1.21 | -1.83 | 0.3 | 0.3 |
| Respiratory Disease | -0.56 | 2.03 | 0.2 | 0.2 | -1.76 | 1.33 | 0.2 | 0.2 |
| Other | -2.31 | -2.04 | 0.2 | 0.2 | -3.20 | -2.27 | 0.2 | 0.2 |
| Resulting Total ** | 0.30 | 1.23 | 0.49 | 0.51 | 0.10 | 0.99 | 0.47 | 0.48 |
| Total | Male |  |  |  | Female |  |  |  |
| Cardiovascular Disease | 2.65 | 3.22 |  |  | 2.39 | 3.56 |  |  |
| Cancer | 0.93 | 1.74 |  |  | 0.47 | 1.36 |  |  |
| Violence | 0.60 | 0.04 |  |  | -0.12 | -0.83 |  |  |
| Respiratory Disease | 0.31 | 1.57 |  |  | -1.58 | 0.69 |  |  |
| Other | -0.84 | -0.56 |  |  | -1.58 | -0.96 |  |  |
| Resulting Total ** | 1.22 | 1.54 | 0.73 | 0.74 | 0.63 | 1.26 | 0.68 | 0.69 |

* Alternative 1 is $1 / 2$ times Alternative 2; Alternative 3 is $5 / 3$ times Alternative 2.
**Resulting total represents average annual percent reduction in age-adjusted death rates for the last 50 years of the 75 year projection period.


### 1.3. LEGAL IMMIGRATION

## 1.3.a. Overview

Legal immigration is defined as those persons who have been admitted into the United States and been granted legal permanent resident (LPR) status. Legal emigration consists of legal permanent residents and U.S. Citizens who depart the Social Security area population to reside elsewhere.

For each year z of the projection period, the LEGAL IMMIGRATION subprocess produces estimates of legal immigration ( $L^{\mathrm{z}}$ ) and legal emigration ( $E^{\mathrm{z}}$ ), by age and sex, based on assumptions set by the Trustees for each category. In addition, the LEGAL IMMIGRATION subprocess disaggregates the estimates of $L^{Z}$ into those who have been admitted into the United States during the year ( $N E W^{2}$ ) and those who adjusted from the otherimmigrant population to LPR status $\left(A O S^{\mathrm{Z}}\right)$.

Each fiscal year, ${ }^{15}$ the Department of Homeland Security (DHS) collects data on the number of persons granted LPR status by age, sex, and class of admission. The U.S Census Bureau provided OCACT with an unpublished estimate of the annual number of legal emigrants, by sex and age, based on the change between the 1980 and 1990 censuses. These historical data are used as a basis for developing age-sex distributions that are applied to the Trustees' aggregate immigration assumptions to produce annual legal immigration and emigration estimates by age and sex.

The primary equations of LEGAL IMMIGRATION, by age (x) and sex (s), for each year (z) of the 75-year projection period are summarized below:

$$
\begin{align*}
& N E W_{x, s}^{z}=N E W_{x, s}^{z}(\cdot)  \tag{1.3.1}\\
& A O S_{x, s}^{z}=A O S_{x, s}^{z}(\cdot)  \tag{1.3.2}\\
& L_{x, s}^{z}=N E W_{x, s}^{z}+A O S_{x, s}^{z}  \tag{1.3.3}\\
& E_{x, s}^{z}=E_{x, s}^{z}(\cdot)  \tag{1.3.4}\\
& N L_{x, s}^{z}=L_{x, s}^{z}-E_{x, s}^{z} \tag{1.3.5}
\end{align*}
$$

## 1.3.b. Input Data

## Trustees Assumptions -

Each year the Board of Trustees of the OASDI Trust Funds specifies the total annual assumed values for legal immigration and legal emigration. For the 2016 Trustees Report, the ultimate values for legal immigration and emigration are 1,060,000 and 265,000, respectively (both reached in 2015).

## Department of Homeland Security -

- Historical legal immigration by fiscal year (1941-1973), 5-year age group (0-4, 5-9, ..., and 80-84), and sex. These data will not be updated.
- Legalizations due to IRCA by type (pre-1982s and SAWs), single year of age (0-98 and unknown age), sex (including unknown) and month for the years 1989-1996. These data will not be updated.
- Historical legal immigration by fiscal year (1973-2014), single year of age (0 through 99, and unknown

[^7]age), sex (including unknown), and class of admission (New Arrival, Adjustment of Status, Refugee, and Asylee). These data are updated annually, with the DHS providing an additional year of data each year.

- Total adjustments of status for the years 1966 to 1995 (OCACT further estimates total adjustments of status for 1963-1965). These data will not be updated.


## U.S. Census Bureau -

- Unpublished estimates of annual legal emigration by five-year age groups (0-4, 5-9, $\ldots$, and $80-84$ ) and sex for 1990 based on the change between the 1980 and 1990 censuses. These data are updated occasionally (based on having new data from an outside source and on OCACT resource time constraints).


## Other input data -

- Legal emigration conversion factors. These estimates were developed internally by five-year age groups ( $0-4,5-9, \ldots$, and $80-84$ ) and sex to reflect the fact that the estimated number of people leaving the United States is not equivalent to the number of people leaving the Social Security area. These data are updated when annual legal emigration estimates are updated (see above).


## 1.3.c. Development of Output

## Equations 1.3.1 and 1.3.2 - Legal Immigration

The Trustees specify the aggregate amount of legal immigration for each year of the 75-year projection period. In order to incorporate the numbers of new immigrants into the Social Security area population projections, the total level of new immigrants is disaggregated by age and sex.

There are two ways for an immigrant to be admitted into the U.S. for lawful permanent residence:
(1) New arrivals, such as persons living abroad who are granted an LPR visa and then enter the U.S. through a port of entry. Refugees and asylees that are granted LPR status are also treated as new arrivals in the OCACT model.
(2) Adjustments of status, who are people already residing in the U.S. as other immigrants and have an application for adjustment to LPR status approved by the DHS.

The DHS provides data on legal immigrants by sex, single year of age, classification of admission, and fiscal year of entry. The 10 most recent years of data are used to calculate separate age-sex distributions for both new arrivals and adjustments of status by taking the following steps:

1. Refugee and Asylee LPR admissions are subtracted from the adjustment of status data and added into the new arrival category.
2. The data are converted from fiscal year data to calendar year data.
3. For each class of admission, new arrival and adjustment of status, the historical data for the last 10 years (from 2005-2014) are combined into an average age-sex distribution.
$N E W_{x, s}^{z}$, the expected number of new arrival legal immigrants by age ( x ) and sex ( s ), is calculated by applying the age-sex distribution for new arrivals to the Trustees assumed level of new arrivals. The Trustees' assumed
number of adjustments of status is multiplied by the age-sex distribution of adjustments of status to calculate $A O S_{x, s}^{z}$.

## Equation 1.3.4 - Legal Emigration

The Trustees specify the aggregate amount of legal emigration for each year of the projection period. This is done by setting the ratio of emigration to legal immigration. For the 2016 Trustees Report, the ratio is set at 25 percent.

In order to produce the number of emigrants from the Social Security area population, the total level of emigrants is disaggregated by age and sex. The disaggregation is based on a distribution of emigrants, by sex and five-year age groups, provided to OCACT in unpublished estimates by Census that are based on changes between the 1980 and 1990 censuses. Since the emigration numbers estimated by Census are for all people leaving the United States, they are adjusted downward by a series of conversion factors so the data correspond to the number of people leaving the Social Security area population.

For each sex (s), the Beers formula is used to interpolate and distribute each five-year age group into a single year of age ( x ) distribution, EDIST $_{x, s}$. For each projection year, this distribution is used to distribute the assumed level of total legal emigrants by age and sex using the following equation:

$$
E_{x, s}^{z}=.25\left(\sum_{s=m}^{f} \sum_{x=0}^{84} L_{\chi, s}^{z}\right) * E D I S T_{x, s}
$$

### 1.4. HISTORICAL POPULATION

## 1.4.a. Overview

For each historical year, the HISTORICAL subprocess provides estimates of the Social Security area population for the period December 31, 1940, through December 31, 2013. The Social Security area population consists of:

- U.S. resident population and armed forces overseas plus
- Net census undercount plus
- Civilian residents of Puerto Rico, the Virgin Islands, Guam, the Northern Mariana Islands, and American Samoa plus
- Federal civilian employees overseas plus
- Dependents of armed forces and federal civilian employees overseas plus
- Residual beneficiaries living abroad plus
- Other citizens overseas

The U.S. Census Bureau collects population data and tabulates it by age, sex, and marital status every ten years for the decennial census. Generally, each subsequent year, the Census Bureau publishes an estimate of the postcensal population. This subprocess combines these census and post-censal estimates, along with the estimates of the other components of the Social Security area population listed above, and components of change described in sections 1.1 to 1.3 to develop historical estimates of the total Social Security area population ( $P_{x, s}^{z}$ ) and other than legal population ( $O_{x, s}^{z}$ ). Combining the total populations by single year of age and sex with an estimated marital status matrix provides the total Social Security area historical population by single year of age, sex, and marital status ( $P_{x, s, m}^{z}$ ). These estimates are then used as the basis for the PROJECTED POPULATION subprocess described in section 1.8. The primary equations for this subprocess, 1.4.1, 1.4.2, and 1.4.3, are given below:

$$
\begin{align*}
& P_{x, s}^{z}=P_{x, s}^{z}(\cdot)  \tag{1.4.1}\\
& P_{x, s, m}^{z}=P_{x, s, m}^{z}(\cdot)  \tag{1.4.2}\\
& O_{x, s}^{z}=O_{x, s}^{z}(\cdot) \tag{1.4.3}
\end{align*}
$$

## 1.4.b. Input Data

## Long-Range OASDI Projection Data -

## Demography

- Probabilities of death from MORTALITY, by age last birthday and sex, for years 1941-2014. These data are updated every year.
- The number of new legal immigrants by age and sex for years 1941-2013. These data are from the LEGAL IMMIGRATION subprocess and are updated each year.
- The number of legal emigrants by age and sex for years 1941-2013. These data are from the LEGAL IMMIGRATION subprocess and are updated each year.
- The number of adjustments of status by age and sex for years 1941-2013. These data are from the LEGAL IMMIGRATION subprocess and are updated each year.
- The number of "other immigrants" legalized under the Immigration Reform and Control Act of 1986 (IRCA) from LEGAL IMMIGRATION. These data are reproduced each year and updated if new data is available.
- Birth rates by single year of age of mother (14-49) for the years 1941-2014 from the FERTILITY subprocess. These data are updated each year.


## U.S. Census Bureau Data -

- Estimates of U.S resident population and Armed Forces population overseas as of each July 1 (1940-1979) by sex and single-year of age through 84, and for the group aged 85 and older. These data are generally not updated.
- Estimates of the U.S. resident population for each Census (April 1) 1970-2010 by sex and single year of age 0 through $85+$. New decennial Census estimates come out about every ten years.
- Estimates of total U.S. residential population and total residential population plus Armed Forces overseas population for each January of each Census year from 1990 and 2010. New decennial Census estimates come out about every ten years.
- Estimates of U.S resident population, and U.S. resident plus Armed Forces population overseas as of each July 1 (1980-2014) by sex and single-year of age 0 through 99, and ages 100 and older. Generally, the U.S. Census Bureau restates the data back to the most recent decennial census and includes one additional year of data.
- Estimates of U.S resident population, and U.S. resident plus Armed Forces population overseas as of each January 1 for each Census year starting in 1990 and the final data year (2014 for the 2016 TR) by sex and single-year of age 0 through 99, and ages 100 and older.
- Estimates of the population by marital status, sex, and age from the American Community Survey (ACS) public use microdata samples (PUMS) for years 2000 - 2014. In general, an additional year of data is available each year.
- Undercount factors by single year of age (0-85+) and sex, estimated using post-censal survey data. These data are updated after each decennial census.
- The total annual population estimates for Puerto Rico, Virgin Islands, Guam, Northern Marianas, and American Samoa for years 1951-2014. For each Trustees Report, an additional data year is downloaded from the U.S. Census Bureau's international database. Historical data back to 1951 is also obtained if any changes have occurred.
- Decennial Census population estimates, by varying degree of age detail and sex, for Censuses from 1950 - 2000 for territories and components outside the 50 states, D.C., and armed forces overseas. Most data is aggregated into 18 age groups for each sex though single year of age data is available for young ages in the territories for 1960 and 1970 and all ages starting in 1980. New estimates are added as they become available.
- July populations of the territories by single year of age and sex from 2000-2014. An additional year of data is available each year.
- From ACS PUMS, number of existing marriages from 2000 - 2014 by age of husband crossed with age of wife. In addition, starting with the 2012 ACS, number of existing marriages by age of each spouse is available for same-sex couples. Final grids for same-sex couples are adjusted based on reported same-sex marriages from the states. Generally, an additional year of data is available each year.
- From decennial Census PUMS via the University of Minnesota's IPUMS website, number of existing marriages for Census years 1940 - 2000 by age group of husband crossed with age group of wife. New estimates are added as they become available.
- From decennial Census PUMS via the University of Minnesota's IPUMS website, estimates of the population by marital status, age, and sex for Census years 1940 - 2000. New estimates are added as they become available.
- Estimates of net immigration by age and sex of the U.S. residential plus armed forces overseas (USAF) population from April 1, 2000, through July 1, 2014. An additional year of data is available each year.
- Total Americans overseas estimate based on international data sources and estimates of federal employees and military in Iraq and Afghanistan. The data from the various international sources are derived from different years but center around the year 2003. Additional data will be updated as they become available.


## Other input data -

- From the Department of State, old historical total estimates of outside area populations (federal employees overseas, overseas dependents of federal employees and military, and other Americans overseas.
- The SSA Annual Statistical Supplement provides estimates of the total number of OASDI Beneficiaries living abroad for most years December 31, 1953-2013. Age group data is also available. For each Trustees Report, an additional year of data is available. Age group data is
updated for each decennial Census year.
- From the National Center for Health Statistics (NCHS), the sex ratio (number of males born per female) for years 1941 - 2013. Each year, NCHS provides another year of data. For 2014, the sex ratio is assumed to be 1.048 , the same as the assumed projected sex ratio.
- From the NCHS, National Survey of Family Growth (NSFG) public-use data that helps split up the population eligible for same-sex marriage into marital statuses starting in 2013.
- From the Department of Homeland Security (DHS), the total number of other immigrants (unauthorized immigrants plus nonimmigrants) from January 1, 2005 - 2012.
- From the Office of Personnel Management (OPM), total estimates of the number of federal employees overseas from July 1, 1998 - 2013. These estimates are updated as they become available on the OPM website.
- From the OPM, the number of federal employees overseas by single year of age and sex from a subset of the OPM data source above. Years 1980 - 2014 are available. These estimates are updated as they become available.
- From the Department of Defense, total numbers of armed forces in Puerto Rico, Virgin Islands, Guam, and American Samoa each Census year starting in 1990. These data are updated as they become available.
- Using 2015 TR death rates and historical populations, an assumed December 31, 1940, 85+ distribution. These data are not updated.
- Assumed January total populations added to the Social Security for years 1951, 1957, and 1961 when new territories were added to the Social Security area.


## 1.4.c. Development of Output

## Equation 1.4.1 - Historical Population by age and sex $\left(P_{x, s}^{z}\right)$

The Census Bureau's estimate of the residents of the 50 States, D.C., and U.S. Armed Forces overseas is used as a basis for calculating $P_{x, s}^{z}$. The base estimate is adjusted for net census undercount and increased for other U.S. citizens living abroad (including residents of US territories) and for non-citizens living abroad who are insured for Social Security benefits.

The estimates of the number of residents of the fifty States and D.C. and Armed Forces overseas, as of July 1 of each year, by sex for single years of age through 84, and for the group aged 85 or older, are obtained from the Census Bureau. January 1 and April 1 estimates by sex for single years of age through 84, and for the group aged 85 or older for selected years starting in 1990 and 1970, respectively, are also obtained for from the Census Bureau. Adjustments for net census undercount are estimated using post-censal survey data from the Census Bureau. Population counts over age 65 after the last Census year are modified to be consistent with OCACT mortality and Census USAF net immigration data. The numbers of persons in the other components of the Social Security area as of July 1 are estimated by sex for single years of age through 84, and for the group aged 85 or older, from data of varying detail. Numbers of people residing in Puerto Rico, the Virgin Islands, Guam, American Samoa, and the Northern Mariana Islands are estimated from data obtained from the Census

Bureau. Numbers of Federal civilian employees overseas are based on estimates from the Office of Personnel Management (OPM). Dependents of Federal civilian employees and Armed Forces overseas are based on the stock of Federal civilian employees from OPM and the stock of armed forces overseas from the Census Bureau. Other citizens overseas covered by Social Security are also based on estimates compiled by the Census Bureau. The overlap among the components, believed to be small, is ignored.

The first step of the process is to estimate $P_{x, s}^{z}$ as of January $1^{\text {st }}$ for certain "tab years" (1941, 1951, 1957, 1961, each decennial Census year [1970 through 2010], and the last year of historical data [2014 for the 2016 Trustees Report]). For ages $0-84, P_{x, s}^{z}$ for each tab year, is set equal to an undercount adjustment plus other component populations plus:

- The averaged surrounding July 1 U.S. population and armed forces overseas counts from the Census Bureau prior to 1970
- Modified April 1 U.S. populations from the Census Bureau for Census years from 1970 through 2000.
- The January 1 U.S. population and armed forces overseas counts from the Census Bureau modified by OCACT mortality rates and Census USAF net immigration for ages over 65 for 2010 and 2014. For ages 85 and over, $P_{x, s}^{z}$ for each tab year is set equal to [Built Up Pops Age x, Sex s] * [Total 85+ for Sex s]/[Total Built Up 85+ for Sex s], where the built up estimates are created by taking into account deaths and immigration data from the previous tab year and [Total 85+ for Sex s] is the sum of the calculated U.S. population and armed forces overseas calculated using the same method listed above for each year for ages 0 84.

For years between the tab years, populations are estimated taking into account the components of changes due to births, deaths, legal emigration, adjustments of status, and net legal immigration (or net immigration, if known) during that time period. These estimates are then multiplied by the appropriate age-sex-specific ratios so that the error of closure at the tab years is eliminated.

## Equation 1.4.2 - Historical Population by age, sex, and marital status ( $P_{x, s, m}^{z}$ )

Since eligibility for auxiliary benefits is dependent on marital status, the Social Security area population is disaggregated by marital status. The four marital states are defined as single (having never been married), married, widowed, and divorced.

The distribution of the number of existing marriages are available for Census years 1940 - 2000 from Census public use microsample data (PUMS) samples. These data are aggregated by age group of husband crossed with age group of wife. Additional tabulations from the American Community Survey from 2000 - 2014 are incorporated to adjust these marital prevalence grids for changes since 2000. The grids are transformed from age grouped numbers to single year of age figures from ages 14 to $100+$ for husband and wife using the two dimensional H.S. Beers method of interpolation.

Percentages of single, married, widowed, and divorced persons are calculated by taking the estimate for each marital status category and dividing them by the total number of people for each age group and sex based on Census and/or ACS PUMS. Then, for each sex, if one age group has a higher or lower percentage than the surrounding age groups, an average of the surrounding groups replaces the original value. After verifying the percentages are close to the original data (and adjusted if needed), these percentages are multiplied by the total populations calculated in Equation 1.4.1 for each age, sex, and year to get a preliminary population for each age, sex, and marital status.

To keep the marriage prevalence grids and the marital status percentages smooth and consistent, several algorithms are used. First, the married population is adjusted so that the number of married males equals the number of married females. Then, the number of married persons for each age and sex is set equal to the marginal total of the associated year's marital prevalence grid. Finally, the other marital statuses population totals are adjusted to keep the total number of people in all marital statuses the same as calculated before splitting into marital statuses.

The population is modeled to include the following population statuses for December 31, 2013, and later: heterosexual, gay, and lesbian. The initial December 31, 2013, gay and lesbian populations are broken out assuming $3.0 \%$ of the male population is gay and $4.0 \%$ of the female population is lesbian. Marriage grids of age of older spouse crossed with age of younger spouse for same-sex couples are needed starting December 31, 2013. The grids and populations were produced using data from the American Community Survey, National Survey of Family Growth, and state-level same-sex marriage data.

## Equation 1.4.3 - Historical Other Than Legal (OTL) Population by age and sex ( $O_{x, s}^{z}$ )

This subprocess also estimates historical levels of "other immigrants" in the population, by age and sex. For each year, an initial net residual estimate by single year of age and sex is backed out from estimates of beginning and end of year populations, births, deaths, legal immigrants, adjustments of status, and legal emigrants. This net residual equals the implied initial other in minus other out. These residuals are then modified to ensure reasonableness. Next, using these modified net residuals, along with adjustments of status and OTL deaths (using the same death rates as for the total population), an initial OTL stock is built. These stocks are then modified to ensure reasonableness. After 2000, one further adjustment is done to the stocks. From January 2001 through January 2004, the total OTL populations are set equal to the values that linearly grade from the final OCACT January 2000 total OTL population to a DHS-based January 2005 total OTL population. From January 2005 through January 2012, the total OTL population is forced to match DHS-based total OTL population estimates. The total OTL population from the DHS is equal to the sum of their estimates for unauthorized immigrants and nonimmigrants. Nonimmigrants include categories such as students, temporary workers, and exchange visitors.

### 1.5. OTHER IMMIGRATION

## 1.5.a. Overview

The term "other immigration" refers to persons entering the U.S. in a manner other than being lawfully admitted for permanent residence and who reside in the U.S. for at least 6 months. This includes temporary immigrants (persons legally admitted for a limited period of time, such as temporary workers and foreign students), also called nonimmigrants, in addition to undocumented immigrants living in the U.S. These undocumented immigrants can be split into those that were never authorized or those that were nonimmigrants but overstayed their visas (visa-overstayers).

For each year z of the projection period, the OTHER IMMIGRATION subprocess produces estimates of other-than-legal (OTL) immigration ( $O I_{x, s, t}^{z}$ ), by age ( x ), sex, and OTL type ( t ) based on assumptions set by the Trustees. Estimates of projected other emigration (those leaving the Social Security area, denoted as $O E_{\chi, s, t}^{z}$ ), by age, sex and OTL type are also developed in this subprocess.

The Department of Homeland Security (DHS) estimated the stock of nonimmigrants by age group and sex for April 2008 and for the end of 2010. The HISTORICAL POPULATION program already produces historical estimates of other-than-legal immigrants. These historical data are used to develop recent estimates of the OTL stock by age, sex, and immigration type where immigration type is never-authorizeds, nonimmigrants, and visaoverstayers.

The primary equations of OTHER IMMIGRATION, by age (x), sex (s), and OTL type ( t ) for each year ( z ) of the 75-year projection period are summarized below:

$$
\begin{align*}
& O I_{x, s, t}^{z}=O I_{x, s, t}^{z}(\cdot)  \tag{1.5.1}\\
& O E_{x, s, t}^{z}=O E_{x, s, t}^{z}(\cdot)  \tag{1.5.2}\\
& N O_{x, s, t}^{z}=O I_{x, s, t}^{z}-O E_{x, s, t}^{z}-A O S_{x, s, t}^{z} \tag{1.5.3}
\end{align*}
$$

where $N O_{x, s, t}^{z}$ are the number of net other immigrants, by age (x), sex ( s ), and OTL type ( t ) for year z , and $A O S_{x, s, t}^{z}$ are the number of adjustments to legal status by age (x), sex (s), and OTL type(t) for year $z$;

$$
\begin{equation*}
O P_{x, s, t}^{z}=O P_{x-1, s, t}^{z-1}+O I_{x-1, s, t}^{z}-O E_{x-1, s, t}^{z}-A O S_{x-1, s, t}^{z}-O D_{x-1, s, t}^{z} \tag{1.5.4}
\end{equation*}
$$

where, $O P_{x, s, t}^{z}$ is equal to the other immigrant population, by age ( x ), sex ( s ), and OTL type ( t ) as of December $31^{\text {st }}$ of year $\mathrm{z}, O D_{x, s, t}^{z}$ are the number of deaths in the other immigrant population by age ( x ), sex ( s ), and OTL type (t) for year z , and $A O S_{x, s, t}^{z}$ are the number of adjustments to legal status by age (x), sex (s), and OTL type ( t ) for year z .

## 1.5.b. Input Data

## Trustees Assumptions -

Each year the Board of Trustees of the OASDI Trust Funds specifies the total annual assumed values for other immigration. The ultimate annual level was set at $1,350,000$ persons per year for each year beginning in 2022.

Due to the lingering effects of the recent recession, the level of other immigration is estimated at $1,200,000$ in 2014. The Trustees assume other immigration will gradually increase above the ultimate level to $1,550,000$ in 2018 and 2019 followed by a gradual decrease, reaching the ultimate value in 2022.

## Long-Range OASDI Projection Data -

## Demography

- Historical and projected probabilities of death by age last birthday (including a neonatal mortality factor, single year of age for ages 0-99, and age group 100+) and sex, for years 1941-2100. These data are updated each year from the MORTALITY program.
- Historical net other immigration by single year of age (0-100+) and sex for years 1961-2013. These data are updated each year from the HISTORICAL program.
- Historical December 31 OTLs by single year of age (0-100+) and sex for years 1963-2013. These data are updated each year from the HISTORICAL program.
- Historical July 1 OTLs by single year of age (0-100+) and sex for years 1964-2013. These data are updated each year from the HISTORICAL program.
- Historical new arrivals by single year of age (0-100+) and sex for years 1941-2013. These data are updated each year from the LEGAL IMMIGRATION program.
- Historical and projected adjustments of status by single year of age (0-100+) and sex for years 19412100. These data are updated each year from the LEGAL IMMIGRATION program.


## Department of Homeland Security -

- Components of the Unauthorized Immigrant Population by year for 2005-2012. These data are updated as new data become available.
- Nonimmigrant stock in April 2008 and December 2010 by age-group and sex. These data are updated as new data become available and internal resources are sufficient to examine and interpret such new data.
- Nonimmigrant admissions by class of admission for various fiscal years 1981 - 2012. These data are updated as new data become available and internal resources are sufficient to examine and interpret such new data.
- Total nonimmigrants from January 1, 2005 - 2012 Unauthorized Reports. These data are updated as new data become available.


## U.S. Census Bureau -

- From the American Community Survey (ACS), foreign-born new persons by ACS year (2000-2014), entry year (1900-2014), age ( $0-100$ ) and sex. These data are updated as new data become available.
- From the ACS, total foreign-born population and total population for 2000 - 2014, and total population in Puerto Rico for 2005 - 2014, used to calculate undercount factors. These data are updated as new data become available.
- From the 2012 ACS, persons, by entry year (1900-2012), age (0-100) and sex, that are:

1. Foreign-born citizens
2. Foreign-born non-citizens that are in school or are a high-school graduate
3. Non-citizen parents of citizen children
4. Non-citizen parents of citizen children that are in school or are a high school graduate These data are not updated.

- From the 2012 ACS, persons, by entry year (1900-2012), age (0-100) and sex, that are eligible for
temporary protected status (TPS) based on originating from various countries by certain dates and:

1. Foreign-born citizens
2. Foreign-born non-citizens that are in school or are a high-school graduate
3. Non-citizen parents of citizen children
4. Non-citizen parents of citizen children that are in school or are a high school graduate These data are not updated.

## Other input data -

- The number of those potentially eligible under the 2012 Deferred Action for Childhood Arrivals (DACAs) initiative by age group and an overall gender split, from the Migration Policy Institute. These data are updated as new data become available and internal resources are sufficient to examine and interpret such new data.
- Internally developed numbers of those potentially eligible under the 2014 DACA initiative that were not eligible under the 2012 DACA initiative by age and sex, using various estimates from the 2012 ACS data listed above, the non-executive action 2015 TR run OTHER IMMIGRATION program, and assumptions including foreign-born undercounts in the ACS and percentage felons in the unauthorized population. These data are updated as new data become available and internal resources are sufficient to examine and interpret such new data.
- Internally developed numbers of those potentially eligible under the Deferred Action for Parents of Americans and LPRs (DAPAs) initiative, by age and sex, using various estimates from the 2012 ACS data listed above, the non-executive action 2015 TR run OTHER IMMIGRATION program, and assumptions including foreign-born undercounts in the ACS, percentage felons in the unauthorized population, an inflation factor to ratio up parents of citizens to include parents of LPRs as well, and an inflation factor to ratio up counts of parents of citizens and LPRs that do not live with their child. These data are updated as new data become available and internal resources are sufficient to examine and interpret such new data.
- Internally developed factors of potential DACA stock attaining DACA status by sex and ages 5-100 for the initial and ultimate DACA years. These data are updated as new data become available and internal resources are sufficient to examine and interpret such new data.
- Internally developed factors of potential DAPA stock attaining DAPA status by sex and ages 5-100 for the initial and ultimate DAPA years. These data are updated as new data become available and internal resources are sufficient to examine and interpret such new data.
- Internally developed factors to apply to other immigrants that enter as nonimmigrants. These data will not be updated.
- Overstay percentages by age. These data are based off a RAND Corporation document using data from the 1980s, and are adjusted based on insights from the DHS. These final rates were increased a small amount to factor in the effects of the recent 2014 executive actions. These data are updated as new data become available and internal resources are sufficient to examine and interpret such new data.
- Internally developed rates of departure for the non-DACA/DAPA potential/actual never-authorizeds for non-recent arrivals. These data are updated as new data become available and internal resources are sufficient to examine and interpret such new data.
- Internally developed rates of departure for the non-DACA/DAPA potential/actual visa overstayers. These data are updated as new data become available and internal resources are sufficient to examine and interpret such new data.
- Internally developed rates of departure for non-DACA potential/actual never authorizeds for non-recent arrivals to use prior to 2015 (when the Executive Actions went into effect including decreased deportation of non-felons). These data are updated as new data become available and internal resources are sufficient to examine and interpret such new data.
- Internally developed rates of departure for the non-DACA potential/actual visa overstayers to use prior to 2015 (when the Executive Actions went into effect including decreased deportation of non-felons). These data are updated as new data become available and internal resources are sufficient to examine and interpret such new data.


## 1.5.c. Development of Output

The ACS provides the number of foreign-born new arrivals, which is then used to separate the historical net other than legal immigration into those entering and those leaving. There are several other key inputs that go into this calculation, including an estimated undercount factor. This factor accounts for (1) differences between the foreign-born data from the ACS and the component pieces obtained from DHS, (2) differences between the ACS (Public Use Microdata Sample) and Census' total population, and (3) the foreign-born residing in Puerto Rico. The estimated other than legal immigration is calculated by taking the Beers'd foreign born from the ACS (after applying the undercount factors) and subtracting the legal new arrivals. The estimated historical other than legal emigration is then calculated as the difference between the net other than legal immigration (calculated in the HISTORICAL subprocess) and the estimated historical other than legal immigration. A series of steps are then taken to smooth the two categories. Assumptions split the historical other immigrants into those who arrive or depart the Social Security area as a never-authorized immigrant, nonimmigrant, and visaoverstayer immigrant.

## Equation 1.5.1 - Other Immigration

For each projection year, an age-sex-type distribution is used to distribute the aggregate number of other immigrants by age, sex, and OTL type. This age-sex distribution is denoted as ODIST ${ }_{x, s, t}$ and is developed by using average historical estimates of other immigrants entering the country from 2000 through 2007.

The assumed total level of other immigration is denoted by $T O^{z}$. Thus, for each year ( z ) other immigration is defined by the following equation:

$$
O I_{x, s, t}^{z}=T O^{z} * O D I S T_{x, s, t}
$$

## Equation 1.5.2 - Other Emigration

$O E_{x, s, t}^{z}$ is equal to the annual number of other immigrants who depart the Social Security area by age ( x ), sex (s), and OTL type ( t ) they left from. These estimates are based on 2014 TR build-up of stocks from 2008 through 2010 including other immigration discussed above, deaths, adjustments of status (from the LEGAL IMMIGRATION subprocess), and assumptions about how many departures come from each OTL type. Deaths for the other immigrant population use the same death probabilities as the total population:

$$
O D_{x, s, t}^{z}=q_{\chi, s}^{z} O P_{x, s, t}^{z}
$$

Then, rates are calculated by dividing $O E_{x, s, t}^{z}$ by $O P_{x, s, t}^{z}$ for this $2008-2010$ period at each age, sex, and OTL type. After smoothing and adjusting for the effects of the recent recession and the 2014 executive actions, these rates are used to calculate $O E_{x, s, t}^{z}$ in projected years by being applied to the OTL stock populations $O P_{x, s, t}^{z-1}$ for the overstayer and nonimmigrant stocks. For the never authorized stock, these rates are further adjusted and split into two categories so that recent arrivals are exposed to twice the rates as the residual never authorized stock.

This subprocess also splits historical OTLs into the various categories. It is assumed that all OTL immigrants were nonimmigrants as of December 31, 1963. Between December 31, 1963, and December 31, 2010, the percentage of total OTL by age and sex in each OTL category is linearly interpolated from the percentages at those two points in time. A final adjustment ensures the total nonimmigrants are appropriate, based on DHS nonimmigrant admissions or, if available, stock estimates.

Finally, this subprocess also projects the Deferred Action for Childhood Arrivals (DACA) and the Deferred Action for Parents of Americans and LPR's (DAPA) populations, both a subset of the other immigrant population, by age ( x ) and sex (s). The DACA and DAPA populations consists of other than legal immigrants who meet the criteria for either initiative and will be granted authorization to work. The eligible DACA populations are estimated separately by those that meet the age, residency and educational requirements of the initial DACA program implemented in 2012, and also those that meet the requirements of the more recent executive actions in 2014, which expand the eligibility requirements. Rates are then applied to the eligible population to estimate the net number of individuals who actually apply and obtain DACA status. A similar process is done to estimate the eligible and actual DAPA populations for each year in the projection period. Both of these populations further divide the never-authorized and visa-overstayer populations.

### 1.6. MARRIAGE

## 1.6.a Overview

The National Center for Health Statistics (NCHS) collected detailed data on the annual number of new marriages in the Marriage Registration Area (MRA), by age of husband crossed with age of wife, for the period 1978 through 1988 (excluding 1980). In 1988, the MRA consisted of 42 States and D.C. and accounted for 80 percent of all marriages in the U.S. Estimates of the unmarried population in the MRA, by single year of age (or age group if single year of age was not available) and sex, were obtained from the NCHS. Marriage rates for this period are calculated from these data. The age-of-husband crossed with age-of-wife marriage grid rates are transformed from age grouped numbers to single year of age figures from ages 14 to 100+ for husband and wife using the two dimensional H.S. Beers method of interpolation.

The NCHS stopped collecting data on the annual number of new marriages in the MRA in 1989. Less detailed data on new marriages from a subset of the MRA were obtained for the years 1989-1995. The American Community Survey (ACS) public use microdata samples (PUMS) started asking if a person was married in the last 12 months in 2008. Using this question along with ages of spouses, new marriages grids by age-group-ofhusband crossed with age-group-of-wife were developed. For the years in between the NCHS and ACS data, the marriage grids were linearly interpolated. These NCHS and ACS data are used to adjust the more detailed age-of-husband crossed with age-of-wife rates from the earlier years.

Age-specific marriage rates ( $\hat{m}_{x, y}^{z}$ ) for a given year $z$ are defined as the ratio of (1) number of marriages for a given age-of-husband $(x)$ crossed with age-of-wife $(y)$ to (2) a theoretical midyear unmarried population at those ages ( $P_{x, y}^{z}$ ). The theoretical midyear population is defined as the geometric mean ${ }^{16}$ of the midyear unmarried males and unmarried females.

An age-adjusted central marriage rate ( $A \hat{M} R^{z}$ ) summarizes the $\hat{m}_{x, y}^{z}$ for a given year. The standard population chosen for age adjusting is the unmarried males and unmarried females in the Social Security area population as of July 1, 2010. The first step in calculating the total age-adjusted central marriage rate for a particular year is to determine an expected number of marriages by applying the age-of-husband-age-of-wife specific central marriage rates for that year to the geometric mean of the corresponding age groups in the standard population. The $A \hat{M} R^{z}$ is then obtained by dividing:

- The expected number of marriages by
- The geometric mean of (a) the number of unmarried males, ages 15 and older, and (b) the unmarried females, ages 15 and older, in the standard population.

The MARRIAGE subprocess projects annual $\hat{m}_{x, y}^{z}$ by age-of-husband crossed with age-of-wife. The equations for this subprocess, 1.6.1 and 1.6.2, are given below:

$$
\begin{align*}
& \hat{m}_{x, y}^{z}=\hat{m}_{x, y}^{z}(\cdot)  \tag{1.6.1}\\
& A \hat{M} R^{z}=\frac{\sum_{x, y} P_{x, y}^{S} * \hat{m}_{x, y}^{z}}{\sum_{x, y} P_{x, y}^{S}} \tag{1.6.2}
\end{align*}
$$

[^8]where and $x$ and $y$ refer to the age of males and females, respectively, and $P_{x, y}^{s}$ is the theoretical unmarried population in the Social Security area population as of July 1, 2010 (the geometric mean of the corresponding age groups in the standard population).

## 1.6.b. Input Data

## Long-Range OASDI Projection Data -

## Demography

- Estimates of the Social Security area population as of December 31, by age, sex, and marital status for years 1978-2012, excluding 1980. These data are updated each year based on output of the HISTORICAL POPULATION subprocess.


## Assumptions -

For each Trustees Report, ultimate values for the $\hat{A M} R^{z}$ are assumed. The $A \hat{M} R^{z}$ reaches its ultimate value in the 25th year of the 75 -year projection period. For the 2016 report, the intermediate ultimate $A \hat{M} R^{z}$ assumption is 4,000 per 100,000 unmarried couples.

## NCHS Data -

- Number of new marriages in the MRA, by age-of-husband crossed with age-of-wife, for calendar years 1978 through 1988, excluding 1980. These data are no longer available for years after 1988. The data vary in detail by year. They are broken out by single year age-of-husband crossed with single year age-of-wife for many ages (particularly younger ages).
- Number of unmarried males and females in the MRA for calendar years 1978 through 1988, excluding 1980. These data are no longer available for years after 1988. The data are generally broken out by single year age for ages under 40 and by age groups $40-44,45-49,50-54,55-59,60-64,65-74$, and $75+$.
- Number of new marriages, in a subset of the MRA, by age-group-of-husband crossed with age-group-ofwife (age groups include 15-19, 20-24, 25-29, 30-34, 35-44, 45-54, 55-64, and 65+), for calendar years 1989-1995. These data are updated as new data become available and internal resources are sufficient to examine and interpret such new data.
- The total number of new marriages in the MRA less marriages in those states not included in the MRA unmarried population for the period 1957-1988. These data are not updated.
- The total number of new marriages in the United States for the period 1989-2012. Normally, each year, the NCHS publishes the total number of marriages for one more year.
- Number of new marriages in the MRA for years 1979 and 1981-1988 by age group (age groups include 14-19, 20-24, 25-29, 30-34, 35-44, 45-54, 55-64, and 65+), sex, and prior marital status (single, widowed, and divorced). These data are no longer available for years after 1988.
- Number of unmarried people in the MRA for years 1979 and 1981-1988 by age group (age groups include 14-19, 20-24, 25-29, 30-34, 35-44, 45-54, 55-64, and 65+), sex, and prior marital status (single, widowed, and divorced). These data are no longer available for years after 1988.


## U.S. Census Bureau Data -

- Estimates of new marriages by age-group-of-husband crossed with age-group-of-wife from the American Community Survey (ACS) public use microdata samples (PUMS) occurring, on average, end of years 2007 - 2013. An additional year of data is available each year.


## Other Input Data -

- From the vital statistics offices in various states, number of same-sex marriages from 2004 - 2012. These data are updated as they become available.


## 1.6.c. Development of Output

## Equation 1.6.1 -

Age-specific marriage rates are determined for a given age-of-husband crossed with age-of-wife, where ages range from 14 through 100+. The historical period includes years of complete NCHS data on the number of marriages and the unmarried population in the MRA for the period 1978 through 1988, excluding 1980. Data for a subset of the MRA, available by age group only, are used for the period 1989 through 1995, and ACS new married grids by age group are used for the period 2008 through 2012. The marriage grids by age group for the years 1996 through 2007 are linearly interpolated. The total number of marriages from NCHS are also used in the age-specific marriage rate calculations for the provisional period 1989 - 2012. The projection period of the MARRIAGE subprocess begins in 2013.

The historical age-specific marriage rates are calculated for each year in the historical period based on NCHS data of the number of new marriages by age-of-husband crossed with age-of-wife and the number of unmarried persons by age and sex. The formula used in the calculations is given below:

$$
\hat{m}_{x, y}^{z}=\frac{\hat{M}_{x, y}^{z}}{P_{x, y}^{z}} \text {, where }
$$

- $\quad x$ refers to the age of males and $y$ refers to the age of females;
- $\quad \hat{M}_{x, y}^{z}$ is the number of marriages in year $z$; and
- $\quad\left(P_{x, y}^{z}\right)$ is the geometric mean of the midyear unmarried males and unmarried females in year $z$.

The rates for the period 1978 through $1988^{17}$ are then averaged, graduated, and loaded into an 87 by 87 matrix (age-of-husband crossed with age-of-wife for ages 14 through 100+), denoted as MarGrid. This matrix is used in the calculation of the age-specific marriage rates for all later provisional years and the years in the projection period.

For the provisional period, 1989-2012, the NCHS and ACS provided data on the number of marriages by age-group-of-husband crossed with age-group-of-wife (age groups include 15-19, 20-24, 25-29, 30-34, 35-44, 4554, 55-64, and 65+). These data are used to change the distribution of MarGrid by these age groups. For each age-group-of-husband crossed with age-group-of-wife, the more detailed marriage rates in MarGrid that are contained within this group are adjusted so that the number of marriages obtained by using the rates in MarGrid match the number implied in the subset.

[^9]For each year of the entire provisional period (1989 - 2012), an expected total number of marriages is calculated by multiplying the rates in the MarGrid (or the adjusted MarGrid for years 1989-2012) by the corresponding geometric mean of the unmarried males and unmarried females in the Social Security area population. All rates in MarGrid (or the adjusted MarGrid for years 1989-2012) are then proportionally adjusted to correspond to the total number of marriages estimated in the year for the Social Security area population. This estimate is obtained by increasing the number of marriages reported in the U.S. to reflect the difference between the Social Security area population and the U.S. population. In addition, we also subtract out same-sex marriages from the NCHS data, as we handle those in a later step. The provisional period agespecific rates are then graduated using the Whittaker-Henderson method and are used to calculate the ageadjusted rates for each year.

The age-adjusted marriage rates are expected to reach their ultimate value in the $25^{\text {th }}$ year of the 75 -year projection period. Rather than use the last year of provisional data to calculate the starting rate, we calculate the weighted average of the rates for the past five historical data years to derive the starting value. The annual rate of change decreases in absolute value as the ultimate year approaches.

To obtain the age-of-husband-age-of-wife-specific rates for a particular year from the age-adjusted rate projected for that year, the age-of-husband-age-of-wife-specific rates in MarGrid are proportionally scaled so as to produce the age-adjusted rate for the particular year. The MarGrid rates are then adjusted to produce two sets of marriage rates: opposite-sex marriage rates and same-sex marriage rates.

A complete projection of age-of-husband-age-of-wife-specific marriage rates was not done separately for each previous marital status. However, data indicate that the differential in marriage rates by prior marital status is significant. Thus, future relative differences in marriage rates by prior marital status are assumed to be the same as the average of those experienced during 1979 and 1981-1988.

### 1.7. DIVORCE

## 1.7.a. Overview

For the period 1979 through 1988, the National Center for Health Statistics (NCHS) collected data on the annual number of divorces in the Divorce Registration Area (DRA), by age-group-of-husband crossed with age-group-of-wife. In 1988, the DRA consisted of 31 States and accounted for about 48 percent of all divorces in the U.S. These data are then inflated to represent an estimate of the total number of divorces in the Social Security area. This estimate for the Social Security area is based on the total number of divorces in the 50 States, the District of Columbia, Puerto Rico, and the Virgin Islands. Divorce rates for this period are calculated using this adjusted data on number of divorces and estimates of the married population by age and sex in the Social Security area.

An age-of-husband ( $x$ ) crossed with age-of-wife ( $y$ ) specific divorce rate ( $\hat{d}_{x, y}^{z}$ ) for a given year $z$ is defined as the ratio of (1) the number of divorces in the Social Security area for the given age of husband and wife ( $\hat{D}_{x, y}^{z}$ ) to (2) the corresponding number of married couples in the Social Security area ( $P_{x, y}^{z}$ ) with the given age of husband and wife. An age-adjusted central divorce rate ( $A \hat{D} R_{x, y}^{z}$ ) summarizes the $\hat{d}_{x, y}^{z}$ for a given year.

The $A \hat{D} R^{z}$ is calculated by determining the expected number of divorces by applying:

- The age-of-husband crossed with age-of-wife specific divorce rates to
- The July 1, 2010, population of married couples in the Social Security area by corresponding age-ofhusband and age-of-wife.

The DIVORCE subprocess projects annual $\hat{d}_{x, y}^{z}$ by age-of-husband crossed with age-of-wife. The primary equations, 1.7.1 and 1.7.2, are given below:

$$
\begin{align*}
& \hat{d}_{x, y}^{z}=\hat{d}_{x, y}^{z}(\cdot)  \tag{1.7.1}\\
& A \hat{D} R^{z}=\frac{\sum_{x, y} P_{x, y}^{S} * \hat{d}_{x, y}^{z}}{\sum_{x, y} P_{x, y}^{S}} \tag{1.7.2}
\end{align*}
$$

where $x$ and $y$ refer to the age of husband and age of wife, respectively, and $P_{x, y}^{s}$ is the number of married couples in the Social Security area population as of July 1, 2010.

## 1.7.b. Input Data

## Long-Range OASDI Projection Data -

Demography

- Social Security area population of married couples by age-of-husband crossed with age-of-wife as of December 31 for years 1978-2013. These data are updated each year from the HISTORICAL POPULATION subprocess.
- The total July 1 population in the Social Security area for years 1979-2012. An additional year of data is added for each additional year of divorce data from the NCHS.
- The total July 1 population in the residential population plus armed forces overseas for years 1979-2012. An additional year of data is added for each additional year of divorce data from the NCHS.
- The total July 1 population in Puerto Rico and the Virgin Islands for years 1979-2012. An additional year of data is added for each additional year of divorce data from the NCHS.


## Assumptions -

Each year, the ultimate assumed value for the age-adjusted divorce rate is established. The rate reaches its ultimate value in the $25^{\text {th }}$ year of the 75 -year projection period. For the 2016 report, the ultimate assumed $A \hat{D} R^{z}$ is 1,700 per 100,000 married couples.

## NCHS Data -

- The number of divorces in the DRA, by age-of-husband crossed with age-of-wife, for calendar years 1979 through 1988. These data are no longer available for years after 1988. The data are broken out by single year age-of-husband crossed with single year age-of-wife for many ages (particularly younger ages).
- The total number of divorces in the United States for the period for 1989-2012. Data is updated when it becomes available.
- The total number of divorces in Puerto Rico and the Virgin Islands for years 1989-2007. The most recent year of data was obtained in 2000; the 2000 figures are used as a proxy for 2001-2007. ${ }^{18}$ New data are incorporated as they become available and resources are sufficient to validate their use.


## State Divorce Data -

Since NCHS stopped collecting state specific divorce data by age of husband crossed with age of wife, we directly contacted various state health departments for their most recent data. We were able to get this data from 18 states. The years and age groups available vary by state. In general, the years were from 2009 2012. States that had these data available online, or that sent us the data via email, are Alabama, Alaska, Idaho, Kansas, Kentucky, Michigan, Missouri, Montana, Nebraska, New Hampshire, Tennessee, Texas, Utah, Vermont, Virginia, Washington, West Virginia, and Wyoming.

## Census Bureau Data -

The number of Divorces for years 2008 - 2012 in Puerto Rico, estimated using the 2008 - 2014 American Community Survey (ACS) public use microdata samples (PUMS).

## 1.7.c. Development of Output

## Equation 1.7.1 -

[^10]Age-specific divorce rates are calculated for ages 14 through 100+. Detailed NCHS data on the number of divorces by age-group-of-husband crossed with age-group-of-wife are available for the period 1979 through 1988. Provisional data on the total number of divorces in the United States are used for the period 1989 through 2012. With the data from the various states, we developed an age-group-of-husband crossed with age-group-ofwife grid for 2011.

First, the detailed NCHS data on divorces by age group is disaggregated into single year of age of husband ( $x$ ) and age of wife $(y)$, for ages $14-100+$, using the H.S. Beers method of interpolation. Then, the age-specific divorce rates ( $\hat{d}_{x, y}^{z}$ ), for each year, $z$, are calculated for the period 1979-1988 by taking the number of divorces (inflated to represent the Social Security area, $D_{x, y}^{z}$ ) and dividing by the married population in the Social Security area at that age-of-husband and age-of-wife $\left(P_{x, y}^{z}\right)$. The formula for this calculation is given below:

$$
\begin{equation*}
\hat{d}_{x, y}^{z}=\frac{\hat{D}_{x, y}^{z}}{P_{x, y}^{z}} \tag{1.7.3}
\end{equation*}
$$

These rates are then averaged, graduated, ${ }^{19}$ and loaded into an 87 by 87 matrix (age-of-husband crossed with age-of-wife for ages 14 through 100+), denoted as DivGrid. DivGrid is then adjusted using the state data grid developed for 2011. DivGrid after 1988 will be equal to a weighted average of the 1988 DivGrid and the state data single year grid. This state data single year of age grid is derived by ratioing the 1988 DivGrid cells using the original state age-group data. DivGrid will be used in the calculation of the age-specific divorce rates for all later years including the projection period.

For each year in the provisional period (1989-2012), an expected number of total divorces in the Social Security area is obtained by applying the age-of-husband crossed with age-of-wife rates in DivGrid to the corresponding married population in the Social Security area. The rates in DivGrid are then proportionally adjusted so that they would yield an estimate of the total number of divorces in the Social Security area. The estimate of total divorces is obtained by adjusting the reported number of divorces in the U.S. for (1) the differences between the total divorces in the U.S. and in the combined U.S., Puerto Rico, and Virgin Islands area, and (2) the difference between the population in the combined U.S., Puerto Rico, and Virgin Islands area and in the Social Security area.

The starting age-adjusted divorce rate is set to a weighted average of the past five years of data. This ageadjusted rate is assumed to reach its ultimate value in the $25^{\text {th }}$ year of the 75 -year projection period. The annual rate of change decreases in absolute value as the ultimate year approaches.

To obtain age-specific rates for use in the projections, the age-of-husband-age-of-wife-specific rates in DivGrid are adjusted proportionally so as to produce the age-adjusted rate assumed for that particular year.

[^11]
### 1.8. PROJECTED POPULATION

## 1.8.a. Overview

For the 2016 Trustees Report, the starting population for the population projections is the December 31, 2013, Social Security area population, by age, sex, and marital status, produced by the HISTORICAL POPULATION subprocess. (For this section, section 1.8, the term "starting year" refers to the year 2013.) The Social Security area population is then projected using a component method. The components of change include births, deaths, net legal immigration, and net other immigration. The components of change are applied to the starting population by age and sex to prepare estimated populations as of December 31, 2014 and 2015, and to project the population through the 75-year projection period (years 2016-2090).

Beginning with December 31, 2013, the historical and projected populations are modeled using the following population statuses: heterosexual, gay, and lesbian. The initial December 31, 2013, gay and lesbian populations in the HISTORICAL POPULATION program are broken out assuming 3.0\% of the male population and $4.0 \%$ of the female population is gay or lesbian, and the same is true for cohorts born in the PROJECTED POPULATION program.

There is a separate equation for each of the components of change as follows:

$$
\begin{equation*}
B_{s, p}^{z}=B_{s, p}^{z}(\cdot) \tag{1.8.1}
\end{equation*}
$$

where $B_{s, p}^{z}$ is the number of births of each sex (s) by population status (p) born in year z;

$$
\begin{equation*}
D_{x, s, p}^{z}=D_{x, s, p}^{z}(\cdot) \tag{1.8.2}
\end{equation*}
$$

where $D_{x, s, p}^{z}$ is the number of deaths by age (x), sex (s), and population status (p) that occurs in year z; and

$$
\begin{equation*}
N I_{x, s, p}^{z}=N L_{x, s, p}^{z}+N O_{x, s, p}^{z} \tag{1.8.3}
\end{equation*}
$$

where $N I_{x, s, p}^{z}$ is the net total immigration (both legal and other) by age(x), sex (s), and population status (p), $N L_{x, s}^{z}$ is the net number of legal immigrants (produced by the LEGAL IMMIGRATION subprocess), and $N O_{x, s}^{z}$ is the net number of other immigrants (produced by the OTHER IMMIGRATION subprocess). The population program further disaggregates these variables by population status.

Once the components of change are calculated, the following equation is used to calculate the Social Security area population by age and sex:

$$
\begin{array}{ll}
P_{0, s, p}^{z}=B_{s, p}^{z}-D_{0, s, p}^{z}+N I_{0, s, p}^{z} & \text { for age }=0 \\
P_{x, s, p}^{z}=P_{x-1, s, p}^{z-1}-D_{x, s, p}^{z}+N I_{x, s, p}^{z} & \text { for ages }>0 \tag{1.8.4}
\end{array}
$$

where $P_{x, s, p}^{z}$ is the population, by age (x), sex (s), and population status (p), as of December $31^{\text {st }}$ of year z .

The population is further disaggregated into the following four marital statuses: single (never married), married, widowed, and divorced. The following equation shows the population by age (x), sex (s), population status (p), and marital status (m) for each year z :

$$
\begin{equation*}
P_{x, s, p, m}^{z}=P_{x, s, p, m}^{z}(\cdot) \tag{1.8.5}
\end{equation*}
$$

The children (ages 0-18) population is further disaggregated into the following four parent statuses (i.e., fates): both parents survive, only father survives, only mother survives, and both parents deceased. The following equation shows the children population by age of child ( x ), sex of parent ( s ), age group of parent (g), and fate of parent (f) for each year z :

$$
\begin{equation*}
C_{x, s, g, f}^{z}=C_{x, s, g, f}^{z}(\cdot) \tag{1.8.6}
\end{equation*}
$$

## 1.8.b. Input Data

## Long-Range OASDI Projection Data -

## Demography

## FERTILITY

- Historical birth rates by single year of age of mother (14-49) for the years beginning with 1941 and ending with the year prior to the starting year. These data are updated each year.
- Projected birth rates by single year of age of mother (14-49) for the years beginning with the starting year and ending with 2100. These data are updated each year.


## MORTALITY

- Historical probabilities of death by age last birthday (including neonatal mortality factor, single year of age for ages 0-99, and age group 100+) and sex for years beginning with 1941 and ending with the year prior to the starting year. These data are updated each year.
- Projected probabilities of death by age last birthday (including neonatal mortality factor, single year of age for ages 0-99, and age group 100+) and sex for the years beginning with the starting year and ending with 2100. These data are updated each year.
- Factors to distribute probabilities of death by marital status. They are dimensioned by sex, single year of age (ages 14-100+), and marital status. These data are updated each year.


## LEGAL IMMIGRATION

- Projected numbers of legal immigrants who are new arrivals, by single year of age (-1-100) and sex for years beginning with the starting year and ending with 2100 . These data are updated each year. Note that age -1 represents births that occur during the year.
- Projected numbers of legal emigrants by single year of age (-1-100) and sex for years beginning with the starting year and ending with 2100 . These data are updated each year. Note that age -1 represents births that occur during the year.
- Projected numbers of legal immigrants who are adjustments of status, by single year of age (-1100) and sex for years beginning with the starting year and ending with 2100 . These data are updated each year. Note that age -1 represents births that occur during the year.


## HISTORICAL POPULATION

- Social Security area population by single year of age (0-99 and 100+), sex, and marital status for the years beginning with 1940 and ending with the starting year in total and broken down by population status. These data are updated each year.
- Married couples by single year of age of husband (ages 14-100+) crossed with single year of age of wife (ages 14-100+) for the years beginning with 1940 and ending with the starting year in total and by marriage type (opposite-sex, same-sex male, and same-sex female). These data are updated each year.
- Other than legal population by age and sex for the years beginning with 1963 and ending with the starting year. These data are updated each year.


## OTHER IMMIGRATION

- Projected numbers of other immigrants entering the country by age ( $-1-100$ ) and sex for years beginning with the starting year and ending with 2100. These data are updated each year.
- Projected numbers of other immigrants leaving the country by age ( $-1-100$ ) and sex for years beginning with the starting year and ending with 2100. These data are updated each year.
- Other than legal population by age and sex for the years beginning with the starting year and ending with 2100. These data are updated each year.


## MARRIAGE

- Projected central marriage rates by single year of age of husband (ages 14-100+) crossed with single year of age of wife (ages 14-100+) for each year of the projection period. These data are updated each year.
- Projected central same-sex marriage rates by single year of age of spouse1 (ages 14-100+) crossed with single year of age of spouse2 (ages 14-100+) for each year of the projection period. These data are updated each year.
- Averaged and graduated marriage rates for the period 1979 and 1981-1988 by single year of age (ages $14-100+$ ), sex, and prior marital status (single, divorced, and widowed). These data are updated each year.
- Total number of marriages for the years beginning with 1989 and ending the year prior to the starting date. These data are updated each year.


## DIVORCE

- Projected central divorce rates by single year of age of husband (14-100+) crossed with single year of age of wife (14-100+) for each year of the projection period. These data are updated each year.


## U.S. Census Bureau Data -

- CPS data on the average number of children per married couple with children by age group of householder (age groups 20-24, 25-29, 30-34, 35-39, 40-44, 45-49, 50-54, and 55-64) for 19602014. (Note that the program splits the last age group, which is a 10 -year age group, into two 5 -year age groups.) An additional year of data is added each year.


## 1.8.c. Development of Output

## Equation 1.8.1-Births

The number of births in the Social Security area, $B_{x}^{z}$, is computed for each year, z , of the projection period by applying the age-specific birth rate to the midyear female population aged 14 to 49 as follows:

$$
B_{x}^{z}=b_{x}^{z}\left(\frac{F P_{x}^{z}+F P_{x}^{z+1}}{2}\right)
$$

where,
$B_{x}^{z}=$ number of births to mothers age $x$ in year $z ;$
$b_{x}^{z}=$ birth rate of mothers age $x$ in year $z$; and
$F P_{x}^{z}=$ female population age $x$ at the beginning of year $z$.
The total number of births in a given year is the sum of the number of births to mothers at each age. This total number of births is disaggregated by sex by assuming a gender ratio of 104.8 male births for every 100.0 female births. The total number of births is also disaggregated by population status by assuming $3.0 \%$ of males born are gay and $4.0 \%$ of females born are lesbian.

## Equation 1.8.2 - Deaths

The number of deaths for the Social Security area by age (x), sex (s), and population status (p), $D_{x, s, p}^{z}$, is computed for each projection year by applying the death probabilities for each age and sex, $q_{\chi, s}^{z}$, to the exposed population at the beginning of the year.

$$
D_{\chi, s, p}^{z}=q_{\chi, s}^{z} S_{\chi, s, p}^{z}
$$

## Equation 1.8.5 - Disaggregating the population by marital status

Once the population is projected by single year of age, sex, and population status, it is then disaggregated by population status into the following four marital states; single, married, widowed, and divorced. Estimates of the Social Security area population by single year of age ( $0-99$ and 100+), sex, marital status, and population status as of the starting date of the population projection are obtained from the HISTORICAL POPULATION subprocess. In addition, the HISTORICAL POPULATION subprocess provides the number of married couples by single year of age of husband crossed with single year of age of wife and number of married male/male and female/female marriages, single year of age of spouse 1 crossed with single year of age of spouse 2 , as of the starting date.

All births are assigned to the single category. For a given age, sex, and population status, deaths are assigned by marital status according to the relative differences in death rates by marital status observed for that age and sex during the calendar years 1995 and 1996, as determined in the MORTALITY subprocess. For a given age, sex, and population status immigrants are assigned by marital status according to the beginning of year marital distribution of the Social Security area population for that age and sex.

Once the number of marriages, divorces, and widowings during a year are determined, the population by age, sex, population status, and marital status is updated to represent end of year. The unmarried population at the end of the year is estimated from the population at the beginning of the year by subtracting deaths and marriages and adding new immigrants, widows (or widowers), and divorces during the year. The married population at the end of the year is estimated from the population at the beginning of the year by reducing the population for divorces, widows (or widowers), dissolutions of marriages when both husband and wife dies, and by increasing the population for new immigrants and marriages during the year.

Numbers of new marriages are determined for each projection year. The annual number of opposite-sex marriages occurring at each age of husband crossed with each age of wife is obtained by multiplying the age-ofhusband and-age-of-wife-specific marriage rates with the geometric mean of the midyear unmarried male population and the midyear unmarried female population.

The age-specific midyear unmarried male population ${ }^{20}$ is estimated from the beginning of the year unmarried populations. It is calculated by adjusting the number of unmarried males at the beginning of the year to represent midyear using the relationship between the prior beginning of year and the current beginning of year unmarried male populations. The midyear female unmarried population is approximated similarly.

The numbers of marriages are then distributed by previous marital status (single, widowed, divorced) in the same proportions as would have been produced by applying the previous marital-status-specific marriage rates from the MARRIAGE subprocess to the population by marital status at the beginning of the year.

Numbers of new divorces are determined for each projection year. The number of divorces during a year, occurring at each age of husband crossed with each age of wife, is obtained by multiplying the age-of-husband by age-of-wife divorce rates for that year with the midyear number of married couples in that age crossing.

The number of age-of-husband by age-of-wife midyear married couples is estimated from the beginning of the year married couples. It is calculated by adjusting the number of married couples at the beginning of the year to represent midyear using the relationship between the number of married couples at the beginning of the prior year and the beginning of the current year.

Marriages and divorces for same-sex couples are calculated similarly.
Widowings are computed by applying general population probabilities of death to the marriage prevalence at the beginning of the year. Widowings and deaths by marital status are then reconciled for internal consistency.

## Equation 1.8.6 - Disaggregating the children by parent survival status

Once the population is projected by single year of age, sex, population status, and marital status, the number of children are then categorized by age of father, age of mother, and orphan status. The HISTORICAL POPULATION subprocess provides the historical number of children (ages 0-18), number of women (ages 1449), and the number of married couples by single year of age of husband crossed with single year of age of wife. The projected number of children (ages 0-18), number of women (ages 14-49), and marriage grid age of husband crossed with age of wife is calculated in the projected population.

For women aged 14-49, births are calculated by multiplying the age-specific birth rate, from the FERTILITY subprocess, with the average number of women at the corresponding age. The births are then distributed to the age of husband in the same proportions as the age of husband crossed with age of wife married couples grid.

Each year the number of children is then rolled forward a year to the next age of husband, age of wife, and child age. Parent survival is calculated based on the deaths rates from the MORTALITY subprocess. The number of orphans consists of children with at least one parent deceased. The calculated number of children by age of father and age of mother must match the number of children in the historical or projected population. To accomplish this, the calculated number of children is multiplied by the ratio of the number of children in the historical or projected population to the number of children by age of father and age of mother that was calculated using the fertility rates. For any remaining difference, an adjustment of one is made for each age of husband crossed with age of wife until the total number of children match.

[^12]Once the population is projected by single year of age, sex, population status, marital status, and children, the mean number of children per married couple with children is determined by year and age of householder. The historical mean number of children by year and age of householder in the population program is calculated from the number of children categorized by age of father, age of mother, and the number of married men by age group from the HISTORICAL POPULATION subprocess. Linear regression is used to model the relationship between the mean number of children in the population program to the mean number of children from the U.S. Census Bureau. The regression model is then used to project the mean number of children by age of householder in the population program.

# Process 2: Economics 

## 2. Economic

The Office of the Chief Actuary uses the Economic process to project OASDI employment and earnings-related variables, such as the average wage for indexing and the effective taxable payroll. The Economic process receives input data from the Demography process and provides output data to the Beneficiaries and the Trust Fund Operations \& Actuarial Status processes.

The Economic Process is composed of four subprocesses, U.S. EMPLOYMENT, U.S. EARNINGS, COVERED EMPLOYMENT AND EARNINGS, and TAXABLE PAYROLL. As a rough overview, U.S. EMPLOYMENT and U.S. EARNINGS project U.S. employment and earnings data, respectively, while COVERED EMPLOYMENT AND EARNINGS converts these employment and earnings variables to OASDI covered concepts. TAXABLE PAYROLL, in turn, converts OASDI covered earnings to taxable concepts, which are eventually used to estimate future payroll tax income and future benefit payments.
U.S. EMPLOYMENT and U.S. EARNINGS produce quarterly output, while the output from COVERED EMPLOYMENT AND EARNINGS is annual. TAXABLE PAYROLL produces both.

Two appendices are at the end of this documentation. The first appendix, 2-1, provides details for most of the equations given in the following descriptions of the Economic process. The second appendix, 2-2, provides a listing with explanations of acronyms used in this documentation.

### 2.1. U.S. EMPLOYMENT (USEMP)

## 2.1.a. Overview

The Bureau of Labor Statistics (BLS) publishes historical monthly estimates for civilian U.S. employment-related concepts from the Current Population Survey (CPS). The principal measures include the civilian labor force (LC) and its two components - employment (E) and unemployment ( U ), along with the civilian non-institutional population (N). The BLS also publishes values for the civilian labor force participation rate (LFPR) and the civilian unemployment rate (RU). The LFPR is defined as the ratio of LC to N, while the RU is the ratio of $U$ to LC, expressed to a base of 100 . For many of these concepts, the BLS publishes historical data disaggregated by age, gender, marital status, and presence of children.

For various disaggregated groups ${ }^{1}$, USEMP projects quarterly and annual values for these principal measures of U.S. employment and population. Equations 2.1.1 through 2.1.6 outline the subprocess' overall structure and solution sequence for the total economy. We assume that the military population (M) will remain constant over the first ten years of the projection horizon then grow at the same rate as E . We also assume that the sum of N and M will grow at the same annual rate projected for the Social Security area population (P) (see Demography Process input).

$$
\begin{array}{ll}
\mathrm{M}^{\mathrm{t}}= & \mathrm{M}^{2015} \quad \begin{array}{l}
\text { for } \mathrm{t} \leq 2026 \\
\\
\\
\mathrm{M}^{\mathrm{t}-1} *\left(\mathrm{E}^{\mathrm{t}} / \mathrm{E}^{\mathrm{t}-1}\right) \text { for } \mathrm{t}>2026
\end{array} \\
\mathrm{~N}^{\mathrm{t}}= & {\left[\left(\mathrm{N}^{\mathrm{t}-1}+\mathrm{M}^{\mathrm{t}-1}\right) *\left(\mathrm{P}^{\mathrm{t}} / \mathrm{P}^{\mathrm{t}-1}\right)\right]-\mathrm{M}^{\mathrm{t}}} \\
\mathrm{RU}= & R U(\cdot) \\
\mathrm{LFPR}= & \operatorname{LFPR}(\cdot) \\
\mathrm{LC} & =\operatorname{LFPR} * \mathrm{~N} \\
\mathrm{E} & =\operatorname{LC} *(1-\mathrm{RU} / 100) \tag{2.1.6}
\end{array}
$$

Note: the superscript t represents the projection year.
The Demography Process estimates historical values for the total Social Security area population (P) and an important component, the other immigrant population (OP). OP is further disaggregated into components by visa status: those temporarily authorized to reside or work in the US (OP_A), those who have overstayed their authorization (OP_NA), and those who were never authorized to reside or work in the US (OP_NO). Similarly, USEMP projects annual values for E and employed OP (EO), including its visa-status components (EO_A, EO_NA,

[^13]EO_NO). USEMP also separates EO to those whose earnings are reported and posted to the Master Earnings File (EO_MEF), those whose earnings are reported posted to the Earnings Suspense File (EO_ESF), those in the underground economy (EO_UND). A further subgroup of EO_MEF is also calculated: those who are OASDI covered (EO_MEFC). Equations 2.1.7 through 2.1.14 outline the overall structure of the subprocess used to estimate EO and its sub-components.

$$
\begin{array}{ll}
\text { EO_A } & =E O \_A(\cdot) \\
\text { EO_NA } & =E * O P \_N A / N \\
\text { EO_NO } & =E * O P \_N O / N \\
\text { EO } & =\text { EO_A }+ \text { EO_NA }+ \text { EO_NO } \\
\text { EO_MEF } & =E O \_M E F(\cdot) \\
\text { EO_MEFC } & =E O \_M E F C(\cdot) \\
\text { EO_ESF } & =E O \_E S F(\cdot) \\
\text { EO_UND } & =\text { EO }- \text { EO_MEF }- \text { EO_ESF } \tag{2.1.14}
\end{array}
$$

Finally, for each age/gender group, USEMP projects total "at-any-time" employed other immigrant population (TEO). EO represents the average weekly employment of the other immigrant population during a calendar year. TEO represents the total number of individuals in the other immigrant population who had any employment during the calendar year. (EO can be roughly viewed as the average number of jobs worked by OP during a calendar year, while TEO represents the total number of individuals who worked those jobs.) Effectively, Equations 2.1.15 through 2.1.19 convert every EO age-gender sub-component to an at-any-time TEO age-gender sub-component counterpart.

$$
\begin{array}{ll}
\mathrm{TEO} \text { MEF } & =\text { TEO_MEF }(\cdot) \\
\mathrm{TEO} \text { _MEFC } & =\text { TEO_MEFC }(\cdot) \\
\mathrm{TEO} \text { ESF } & =\text { TEO_ESF }(\cdot) \\
\mathrm{TEO} \text { (UND } & =\text { TEO_UND }(\cdot) \\
\text { TEO } & =\text { TEO_MEF + TEO_ESF + TEO_UND } \tag{2.1.19}
\end{array}
$$

## 2.1.b. Input Data

## Long-Range OASDI Projection Data

These data are updated each year.

## Demography

- Social Security area population as of year-end (1941 - 2099) by age, marital status (single, married, widowed, divorced) and gender (M, F)
- "Other immigrant" population as of year-end (1964 - 2099) by age, gender (M, F), and visa status (OP_A, OP_NA, and OP_NO)
- Number of children by age of child and age of mother (1960-2099)
- Life expectancy by age and gender (1950-2099)
- Exit rates (probability of leaving the "other immigrant" population by other than death) by age and gender.
- Mortality rates by age and gender (1941-2099)


#### Abstract

Trust Fund Operations and Actuarial Status - The Trust Fund Operations and Actuarial Status Process provides no input to the Economic Process sections. However, the LFPRs use input from the Outgo Process from the prior year's Trustees Report. That is, the projected LFPRs for the 2016 Trustees Report use input from the 2015 Trustees Report that includes projections for the disability prevalence rates by age and gender (originally from the Beneficiaries subprocess), and the primary insurance amount (PIA) replacement rates by age and gender. The disability prevalence rate is defined as the ratio of the number of disabled worker beneficiaries to the disability-insured population. The PIA replacement rate is defined as the ratio of a hypothetical medium-scaled worker's PIA to his/her career-average indexed earnings level.


## Trustees’ Assumptions

Each year the Board of Trustees of the OASDI Trust Funds sets the ultimate average annual growth rate values for key economic variables:

- Real wage
- Total economy productivity
- Average hours worked
- Ratio of wages to compensation (RWSD)
- Ratio of compensation to GDP (RWSSY)
- GDP deflator (PGDP)
- Consumer Price Index (CPI)

The Board also sets ultimate values for:

- Annual trust fund real interest rate
- Unemployment rate

These ultimate values are typically reached during the last half of the short-range (first 10 years) of the projection horizon. Earlier projected values are set to provide a smooth transition from the latest actual historical values to the assumed long-range ultimate ones. As
a by-product of this process, values for real GDP, and potential GDP are set. The ratio (RTP) of real to potential GDP is an important summary measure of the economic cycle.

The Trustees also agree on the assumed short-range values for the listed variables.

## Addfactors

Addfactors are adjustments that move an estimate closer to an expected value. They may be used for a variety of reasons associated with data availability, structural changes in the data and/or model, and perceived temporary aberrations in recent historical data. Addfactors were included on male and female LFPRs starting around age 40 to reflect the effects of projected changes in life expectancy.

## Other input data

- U.S. armed forces (EDMIL) by age and gender, estimated by the Department of Defense and published by the Census Bureau on a monthly basis (1948-2000) by single year of age (17 to 64) and gender. These data are no longer produced by Census.
- EDMIL by age and gender, estimated by the Economic Process as the difference in monthly resident plus Armed Forces overseas population and the monthly civilian population. These two populations are available from the Census Bureau on a monthly basis (April 2000 to June 2015) by single year of age (16 to 69) and gender. These data are updated several times a year.
- Data for the mobilized military reservist population, by branch of service (September 2001-September 2015) are reported by the US Department of Defense weekly. This subprocess updates the data several times a year.
- Data from the March Supplement of the Joint BLS/Census Current Population Survey (CPS) by year (1968-2015), for levels of the civilian noninstitutional population, labor force, military, and unemployment. These data are available from the U.S. Census Bureau, via DataFerrett, by single year of age ( 16 to $85+$ ), gender, marital status (never married, married with a spouse present, and married with no spouse present), and presence of children. These data are updated by the U.S. Census Bureau for the BLS annually. This subprocess updates the data every other year (or more often, based on time availability).
- Data from the March Supplement of the CPS by year (1992-2015), for levels of the civilian noninstitutional population. These data are available from the U.S. Census Bureau, via DataFerrett, by single year of age ( 16 to 85+), gender, and educational attainment level. These data are updated by the U.S. Census Bureau for the BLS annually. This subprocess updates the data every year, if time availability allows.
- Data from the CPS (1948-2015) for levels of civilian employment, civilian labor force, civilian unemployment, and civilian noninstitutional population. These data are available from the BLS by age group and gender. These data are updated by the BLS monthly. This
subprocess updates the data several times a year.
- Data from the CPS by year (1994-2015), for the civilian noninstitutional population. These data are available from the BLS by single year of age ( 16 to 85+), gender, marital status, labor force employment status, and (for those not in the labor force) reason for not being in the labor force. These data are updated by the BLS monthly. Monthly data are used to calculate annual averages. This subprocess updates the data every year, if time availability allows.
- Data from the Current Employment Statistics survey (CES) (1964 (varies) to 2015) for establishment employment, average hourly earnings, average weekly earnings, and average weekly hours. These data are available from the BLS by sector. These data are updated by the BLS monthly. This subprocess updates the data several times a year.
- Unpublished data from the CPS (1965 - October 2015) for male and female civilian labor force participation rates for older workers. These data are available from the BLS by single year of age (ages 55-79) and by group ( 75 and over, and 80 and over). These data are updated by the BLS monthly. This subprocess updates the data several times a year.
- The historical resident population is published annually by the Census (1980-2014). These data are available by gender and age group. This subprocess uses data for age groups under 16 and 60 and over. This subprocess updates the data annually.


## 2.1.c. Development of Output

## Equation 2.1.3 - Unemployment Rate (RU)

The RU is disaggregated by age and gender. The age groups include 16-17, 18-19, 20-24, 25-$29,30-34,35-39,40-44,45-49,50-54,55-59,60-64,65-69,70-74$, and 75 and over. Thus, USEMP contains 28 RU equations, 14 for males and 14 for females. Each disaggregated RU is specified using a first-difference model that depends on the distributed lag in the change in the ratio of real to potential GDP (RTP) and an adjustment to ensure that values converge to its estimated trend level. Coefficients are estimated by regression and constrained to an expected aggregate behavior whereby a 2.0 percentage point increase in the RTP elicits a 1.0 percentage point decrease in the RU. Furthermore, projections are constrained to the ultimate age-gender-adjusted RU set by the Trustees. The aggregate age-sex-adjusted RU is dependent on the projected distribution of the labor force by age and gender. See Appendix 2-1 for details on the equations.

## Equation 2.1.4 - Labor Force Participation Rate (LFPR)

The LFPR is disaggregated by age and gender. Age groups include 16 to 17 (i.e., 16-17), 1819, 20-24, 25-29, 30-34, 35-39, 40-44, 45-49, 50-54, 55, 56, ... 99, 100 and over. For age groups between 20 and 54, male and female LFPRs are further disaggregated by marital status, categories of which include never married, ever married with spouse present, and ever married with spouse absent (which includes separated, widowed, and divorced). Female LFPRs disaggregated by age (between 20 and 44) and by marital status are further disaggregated by presence of own child. The groups for presence of own child include females with at least one child under the age of six and females without a child under the age of six. Thus, USEMP contains 153 LFPR equations, 69 for males and 84 for females. See Appendix 2-1 for details on the equations.

Given the level of demographic disaggregation, the aggregate LFPR is dependent on the projected distribution of the population by age, gender, marital status, and presence of own child. Each disaggregated LFPR, however, is dependent on the input variables that are most relevant to the demographic group. For example, only the LFPRs for relevant older workers are dependent on changes to the normal retirement age (NRA). Specific examples of the impact of input data on the disaggregated LFPRs are presented below.

- Disability prevalence ratio (RD) is defined as the ratio of disabled worker beneficiaries to the disability-insured population. An increase in RD lowers the LFPR. RD is adjusted with a multiplicative factor defined as the group's average historical LFPR over a 15-year period. This adjustment implicitly assumes that disability can strike any person in the population with equal probability. For ages 62 to NRA, RDs are not "pure" RDs in that they are subject to the confounding effect of the availability of retirement benefits. For example, at age 62, a marginally disabled individual may opt to begin receiving retirement benefits rather than go through an uncertain disability application/appeals process. For ages NRA and above, RDs are unavailable because at the NRA all disabled-worker beneficiaries become retired-worker beneficiaries. To avoid these problems, RDs for ages

62-74 are set to their cohort RD at age 61. For example, the RD for males age 62 in year ( t ) is set to the RD for males age 61 in year ( $\mathrm{t}-1$ ). For those ages 75 and older, the lagged cohort variable provides information on the influence of disability prevalence rates on labor force participation.

- The unemployment rate (RU) is a measure of the economic cycle. An increase in the lagged and current unemployment rate leads to a decrease in the LFPR. The RU affects most LFPRs.
- The normal retirement age (NRA) is assumed to affect the LFPRs for those age 62 through 69 through an earnings test and replacement rate. The replacement rate is defined as the ratio of a hypothetical worker's PIA to career-average wage level. This value is projected for hypothetical workers with medium-scaled earnings patterns ${ }^{2}$ who retire at ages 62 through 69. The replacement rate is adjusted to include the reduction for early retirement and the delayed retirement credit. An increase in the NRA decreases the adjusted replacement rates, which, in turn, leads to increases in the LFPRs for those between the ages of 62 and 69. The potential earnings test tax rate (POT_ET_TXRT) is used in LFPRs between 62 and 69. It is defined as a tax rate on monthly retirement benefits faced by an individual who opts to collect Social Security benefits before reaching NRA while continuing to work and earn income. An increase in the NRA from 66 to 67 leads to an increase in the potential tax rate for those age 66, which, in turn, leads to a decrease in their LFPR.
- The education distribution of the workforce increases the LFPRs if the level of educational attainment increases.
- The proportion of females with children under age of 6 and their average number of children under age 6 are functions of the ratio of the number of children under age 6 to mothers in each 5 -year age group for ages between 20 and 44 . For females aged 20 to 44 with at least one own child, an increase in the average number of children lowers the LFPR.
- A LFPR increases with its lagged cohort. Lagged cohort variables affect female LFPRs age 55 and over, and male LFPRs age 75 and over.
- The LFPRs for males age 62 through 74 increase with spousal LFPRs.
- For those approximately age 40 and over, an increase in life expectancy leads to an increase in LFPRs .


## Equation 2.1.7 to 2.1.19 - Employed Other Immigrant Population (EO) and At-Any-Time Employed Other Immigrant Population (TEO)

EO is estimated by gender and single-year of age from 16 to 100 based on OP and estimated employment-to-population ratios by visa-status component (OP_A, OP_NA, OP_NO). For this purpose, OP_A is further disaggregated into subgroups by visa type that differ in employment patterns or OASDI coverage status. The other two components are assumed to have equal employment-to-population ratio as the legal permanent resident population of the same age and gender. This portion of USEMP contains 4,250 equations, for 85 ages, 2 genders, and 25 components and subgroups. We separate EO_NO into those who worked in

[^14]2001 and earlier and those who began working in 2002 and later, since we believe that those who worked in 2001 and earlier are more likely to have OASDI covered wages. Each component is then further separated into EO_MEF, EO_MEFC, EO_ESF, and EO_UND.

Every EO sub-component by age, gender, and visa status is converted to its age-gender TEO sub-component counterpart using an age-gender conversion weight. For example, if the sub-component of EO is for never authorized males age 20 to 24, the conversion weight is defined as the ratio of total economy-wide at-any-time employed males age 20 to 24 (TEM2024) to the sum of military and CPS civilian male employment age 20 to 24. For authorized workers and students on temporary visas, conversion weights take into account their partial presence in the year of arrival and the year of departure.

### 2.2. U.S. EARNINGS (USEAR)

## 2.2.a. Overview

In the CPS data, E is separated by class of worker. The broad categories include wage and salary workers (EW), the self-employed (ES), and unpaid family workers (EU). For the nonagricultural sector, a self-employed participation rate (SEPR) is defined as the ratio of ES to E, the proportion of employed persons who are self-employed. For the agricultural sector, the SEPR is defined as the ratio of ES to the civilian noninstitutional population.

USEAR projects quarterly values for these principal classes of employment. Equations 2.2.1 through 2.2.4 outline the subprocess' overall structure and solution sequence.

$$
\begin{array}{ll}
\mathrm{SEPR} & =\operatorname{SEPR}(\cdot) \\
\mathrm{ES} & =\operatorname{SEPR} * \mathrm{E} \\
\mathrm{EU} & =E U(\cdot) \\
\mathrm{EW} & =\mathrm{E}-\mathrm{ES}-\mathrm{EU} \tag{2.2.4}
\end{array}
$$

In the National Income and Product Accounts (NIPA), the Bureau of Economic Analysis (BEA) publishes historical quarterly estimates for gross domestic product (GDP), real GDP, and the GDP price deflator (PGDP). PGDP is equal to the ratio of nominal to real GDP. Potential (or full-employment) GDP is a related concept defined as the level of real GDP that is consistent with a full-employment aggregate RU.

USEAR projects quarterly values for these output measures. Potential GDP is based on the change in full-employment values for: (1) E (including U.S. armed forces), (2) average hours worked per week, and (3) productivity. Full-employment values for E are derived by solving USEMP under full-employment conditions, while the full-employment values for the other variables (average hours worked and productivity) are set by assumption. RTP is the ratio of real GDP to Potential GDP and is set by assumption. RTP reaches 1.0 in the short-range period and remains at 1.0 thereafter. Projected real GDP is set equal to the product of potential GDP and RTP. Nominal GDP is the product of real GDP and PGDP. The growth rate in PGDP is set by assumptions.

The BEA also publishes quarterly values for the principal components of U.S. earnings, including total wage worker compensation (WSS), total wage and salary disbursements (WSD), and total proprietor income (Y). These concepts can be aggregated and rearranged. Total compensation (WSSY) is defined as the sum of WSS and Y. The total compensation ratio (RWSSY) is defined as the ratio of WSSY to the GDP. The income ratio (RY) is defined as the ratio of Y to WSSY. The earnings ratio (RWSD) is defined as the ratio of WSD to WSS.

USEAR projects quarterly values for these principle components of U.S. earnings using Equations 2.2.5 through 2.2.11.

$$
\begin{array}{ll}
\text { RWSSY } & =R W S S Y(\cdot) \\
\text { WSSY } & =\text { RWSSY } * \mathrm{GDP} \\
\mathrm{RY} & =R Y(\cdot) \\
\mathrm{Y} & =\mathrm{RY} * \mathrm{WSSY} \\
\mathrm{WSS} & =\mathrm{WSSY}-\mathrm{Y} \\
\mathrm{RWSD} & =R W S D(\cdot) \\
\mathrm{WSD} & =\mathrm{RWSD} * \mathrm{WSS} \tag{2.2.11}
\end{array}
$$

## 2.2.b. Input Data

Long-Range OASDI Projection Data
Demography- (See Section 2.1.b.)
Economics - Data from Section 2.1 include the total employed (E), E by age and gender, LFPRs by age and gender, the aggregate unemployment rate (RU), and the full-employment concepts for LC, RU, and E.

Trustees Assumptions - (See Section 2.1.b.)

## Addfactors

Addfactors were included on some employment and output variables to smooth the transition between the latest historical data and the projected values. The need for addfactors is reviewed each year and they are implemented if necessary.

## Other input data

- Data from the NIPA (1929 (varies) to 2015) for GDP, income, wages, compensation, personal consumption expenditures, investment, employer contributions for employee pension and insurance funds, and employer contributions for government social insurance. They are published by the BEA quarterly and/or annually. This subprocess updates the data several times a year.
- Ratio of OASDI covered to NIPA wages, and ratio of OASDI taxable to covered wages. NIPA wages by sector are available quarterly and annually from 1947 to 2015. They are published by the BEA and updated several times during the year. OASDI covered and taxable wages (1971 to 2013) are updated annually by the Economic process. Covered and taxable data for more recent historical years are estimated from preliminary tabulations of Form 941 and W-2 data. Projected values for covered ratios are set to the latest historical year for the military, state and local, farm, and private household sectors. The projected value for the federal civilian sector covered ratio is projected to grow to 1.0 by 2030. The projected values for the private non-farm business and the private sector vary with the relative size of the other immigrant population.
- OASDI employee, employer, and self-employed tax rates from 1937 to 2099. These contribution rates are set according to the Social Security Act of 1935 as amended through 2015. The rates are updated when legislation mandates a change.
- The October ratio of the number of teenagers enrolled in school to the civilian noninstitutional population by gender and age group (16-17 and 18-19) for the period 1947 to 2015. An additional new year of data from the Census Bureau is usually available for including in preparation of the next annual Trustees Report. Projected values are set to levels from the latest historical year.
- The historical Consumer Price Index (CPI) is published monthly by the BLS. This subprocess updates the data several times a year.
- The historical CPI for medical services is published monthly by the BLS. Quarterly values are projected based on the projected growth in the aggregate CPI and an additional amount defined as the growth rate differential in the two price measures that was assumed in the latest President's Fiscal Year Budget. The series is updated annually.
- U.S. armed forces (EDMIL) by age and gender were estimated by the Department of Defense and published by the Census Bureau on a monthly basis (1948-2000) by single year of age ( 17 to 64 ) and gender. These data are no longer produced by Census.
- EDMIL by age and gender are estimated by the Economic process as the difference in the monthly resident plus Armed Forces overseas population and the monthly civilian population. These two populations are available from the Census Bureau on a monthly basis (April 2000 to June 2015) by single year of age (16 to 69) and gender. These data are updated several times a year.
- Wages for railroad workers are wages covered by the Railroad Retirement Act. The annual data are for the period 1971 to 2013. An additional year of data from the Railroad Retirement Board is usually available for including in preparation of the next annual Trustees Report.
- Unpublished data from the CPS (1988-2015) on employment by class of worker (i.e., agricultural, non-agricultural, unpaid family, private industry, government, wage and salary, self-employed). These data are available from the BLS by age group and gender. These data are updated by the BLS annually. This subprocess updates the data annually.
- Data from the NIPA (1947-2015) for wages and compensation of households and institutions are published by the BEA quarterly. This subprocess updates the data several times a year.
- Other program-related parameters, including the average indexing wage, the benefit increase, the taxable maximum, and the annual retirement earnings test exempt amounts, are obtained annually from the Short-Range section of OCACT. This subprocess updates the data annually.
- Unpublished data from the CES \& CPS for total hours worked in the economy. These data are available from the BLS. These data are updated by the BLS quarterly (1948-2015) and annually (1948-2014). This subprocess updates the data several times a year.
- The Federal minimum hourly wage is based on the Fair Labor Standards Act from the Department of Labor for 1938 to 2014. The wage is updated when there is legislation mandating a change.
- Time trends (set by Economic process) are used in the agriculture sector for employment, real output, and compensation in the short-range period. These short-range trends are extended for each year's Trustees Report, reflecting a new short-range period.


## 2.2.c. Development of Output

## Equation 2.2.1 - Self-Employed Participation Rate (SEPR)

The SEPR is disaggregated by age, gender, and industry. The age groups include 16-17, 18-$19,20-24,25-34,35-44,45-54,55-64$, and 65 and over. The industry groups include agriculture and non-agriculture.

For the non-agriculture sector, the SEPRs by age and gender are defined as the ratio of the non-agriculture self-employment to total employment. Thus, the aggregate non-agriculture SEPR is dependent on the projected distribution of employment by age and gender. All nonagriculture SEPRs by age and gender are dependent on the RTP. Increases in the RTP lead to decreases in the SEPRs. For female age groups between 20 and 64, the non-agriculture SEPRs are also dependent on the groups LFPRs. Increases in the LFPRs lead to increases in
the SEPRs.
For the agriculture sector, the male SEPRs by age are defined as the ratio of agriculture selfemployment to the civilian noninstitutional population. Thus, the aggregate agriculture SEPR for males is dependent on the projected distribution of the population by age. The agriculture SEPRs for males by age are dependent on the ratio of total agriculture employment (EA) to the total civilian population aged 16 and over. (EA is projected in a farm sub-program. Real farm output is projected to increase with the population, while farm productivity, defined as output per worker, is projected to continue to follow its historical trend. EA is projected as the ratio of farm output to farm productivity.) An increase in the ratio of EA to the total civilian population aged 16 and over leads to an increase in the agriculture SEPRs for males.

The female SEPRs by age for the agriculture sector are defined as the ratio of the female to male agriculture self-employment. Thus, the aggregate agriculture SEPR for females is dependent on the projected distribution of male agriculture employment by age. For female age groups age 18 and over, the SEPRs are dependent on the RTP and the corresponding ratio of total female to male employment. Generally, an increase in the RTP leads to increases in the SEPRs. An increase in the total employment ratio also leads to an increase in the SEPR.

## Equation 2.2.3 - Unpaid Family Workers (EU)

EU is disaggregated by age, gender, and industry. The age groups include 16-17, 18-19, 20-
$24,25-34,35-44,45-54,55-64$, and 65 and over. The industry groups include agriculture and non-agriculture.

From 1970 to 2014, the level of EU fell from about 0.5 to 0.02 million in the agriculture sector and from about 0.5 to 0.06 million in the nonagriculture sector. For projections, the levels of EU by age and gender in the agriculture sector are assumed constant at about five thousand or less. The EUs by age and gender in the nonagriculture sector are projected as a constant ratio to ES.

## Equation 2.2.5-Total Compensation Ratio (RWSSY)

The Trustees set the ultimate annual growth rate for RWSSY. For the short-range period, total WSS, WSD, and Y are aggregated from sector components. Total GDP, WSS, and WSD are divided into the farm and nonfarm sectors. The nonfarm sector is further separated into the government and government enterprises, households, non-profit institutions, and residual (private nonfarm business excluding government enterprises (PBNFXGE)) sectors. Total Y is divided into the farm and residual (i.e., PBNFXGE) sectors.

The methodology used to estimate GDP, WSS, WSD, and Y differs by sector.
Farm - Nominal GDP is the product of real GDP and the farm price deflator. Real farm GDP is projected from estimates of real farm per capita output. EA is projected from estimates of farm productivity. EAW is projected to continue its historical increase relative to EA. Farm
compensation (WSSPF) is the product of estimates for average farm compensation (AWSSPF) and EAW, while farm proprietor income (YF) is the product of estimates of average farm proprietor income (AYF) and EAS. AYF is projected based, in part, on the growth in AWSSPF.

Government and Government Enterprises - This sector is further disaggregated to Federal Civilian, Federal Military, and State and Local. In each sector, WSD is the product of estimates for average wages and employment. WSS is the sum of WSD and estimates for non-wage components of compensation. GDP is the sum of WSS and estimates of consumption of fixed capital.

Household - WSS is the product of estimates for average compensation and employment. WSD is WSS less employer contributions for the OASDHI tax. GDP is the sum of WSS and the gross value added of owner-occupied housing.

Nonprofit Institutions - The Nonprofit Institutions sector is further disaggregated to Health, Education, and Social Services sectors. In each sector, WSS is the product of estimates for average compensation and employment. WSD is WSS less the estimates for non-wage components of compensation. GDP is WSS plus a residual component of output.

Private Nonfarm Business Excluding Government Enterprises (PBNFXGE) - GDP in the PBNFXGE sector is total economy-wide GDP less the sum of the other sector GDPs. WSS is projected as a ratio to GDP less Y. The ratio is projected to be mostly stable, varying only temporarily with changes in RTP. Y is projected to grow with GDP and the ratio of EAS to total employment in the sector.

Thus, total labor compensation (WSSY) is summed from sector components, while the total compensation ratio (RWSSY) is the ratio of total WSSY to total GDP. It is important to note that the pure program-generated estimate for the total RWSSY is adjusted to ensure a smooth transition between the latest historical data and the Trustees' ultimate assumptions.

Equation 2.2.7 - Income Ratio (RY)
Y is disaggregated to the farm and PBNFXGE sectors. (see description for Equation 2.2.5)

## Equation 2.2.10-Earnings Ratio (RWSD)

In the NIPA, the difference between WSS and WSD is defined as employer contributions for employee pension and insurance funds (OLI) and employer contributions for government social insurance (SOC). OLI is mostly health and life insurance, and pension and profit sharing. SOC is composed of employer contributions to Federal and State \& Local government social insurance funds. Federal government funds include OASDI, HI, UI, and other small groups. State and Local government funds mostly include workers’ compensation.

RWSD is defined as the ratio of WSD to WSS. RWSD is projected to mostly decline on a
year-by-year basis over the entire 75-year projection horizon due to projected increases in employer contributions to employee group health insurance premiums (ECEGHIP) and pensions. ECEGHIP is projected by the Center for Medicare and Medicaid Services (CMS) and is consistent with new health care legislation enacted in 2010. Employer contributions to employee pension funds are assumed to increase as life expectancy increases.

### 2.3. OASDI COVERED EMPLOYMENT AND EARNINGS (COV)

## 2.3.a. Overview

Total at-any-time employment (TE) is defined as the sum of total OASDI covered employment (TCE) and total noncovered employment (NCE). TCE can be decomposed to workers who only report OASDI covered self-employed earnings (SEO) and to wage and salary workers who report some OASDI covered wages (WSW). Combination workers (CMB_TOT) are those who have both OASDI covered wages and self-employed income. Workers with some selfemployment income (CSW) are the sum of SEO and CMB_TOT.

COV projects annual values for TE and the principle measures of OASDI covered employment. Equations 2.3.1 through 2.3.9 outline the overall structure and solution sequence used to project these concepts. The combination employment ratio (RCMB) is defined as the ratio of CMB_TOT to WSW.
(Equation 2.3.1 not used in this version.)

$$
\begin{array}{ll}
\mathrm{TE} & =\operatorname{TE}(\cdot) \\
\mathrm{NCE} & =\operatorname{NCE}(\cdot) \\
\mathrm{TCE} & =\mathrm{TE}-\mathrm{NCE} \\
\mathrm{SEO} & =\operatorname{SEO}(\cdot) \\
\mathrm{WSW} & =\mathrm{TCE}-\mathrm{SEO} \\
\mathrm{RCMB} & =R C M B(\cdot) \\
\mathrm{CMB} \text { _TOT } & =\text { RCMB } * \mathrm{WSW}  \tag{2.3.8}\\
\text { CSW } & =\mathrm{SEO}+\mathrm{CMB} \_T O T
\end{array}
$$

Total OASDI covered earnings is defined as the sum of OASDI covered wages (WSC) and total covered self-employed income (CSE_TOT). Both components can be expressed as ratios to their U.S. earnings counterparts. The covered wage ratio (RWSC) is defined as the ratio of WSC to WSD, while the covered self-employed ratio (RCSE) is the ratio of CSE_TOT to Y.

COV projects annual values for the principal measures of OASDI covered earnings using Equations 2.3.10 through 2.3.13.

$$
\begin{array}{ll}
\text { RWSC } & =\operatorname{RWSC}(\cdot) \\
\text { WSC } & =\operatorname{RWSC} * \mathrm{WSD} \\
\text { RCSE } & =\operatorname{RCSE}(\cdot) \\
\text { CSE_TOT } & =\operatorname{RCSE} * \mathrm{Y} \tag{2.3.13}
\end{array}
$$

COV projects various annual measures of average OASDI covered earnings, including the average covered wage (ACW), average covered self-employed income (ACSE), and average covered earnings (ACE).

$$
\begin{array}{ll}
\text { ACW } & =\text { WSC } / \mathrm{WSW} \\
\text { ACSE } & =\text { CSE_TOT } / \mathrm{CSW} \\
\text { ACE } & =(W S C+\text { CSE_TOT }) / \text { TCE } \tag{2.3.16}
\end{array}
$$

The average wage index (AWI) is based on the average wage of all workers with wages reported on Forms W-2 and posted to the Master Earnings File (MEF). By law, it is used to set the OASDI contribution and benefit base (TAXMAX).

COV projects annual values for the AWI and TAXMAX.

$$
\begin{array}{ll}
\text { AWI } & =A W I(\cdot) \\
\text { TAXMAX } & =\text { TAXMAX }(\cdot) \tag{2.3.18}
\end{array}
$$

## 2.3.b. Input Data

## Long-Range OASDI Projection Data

Demography - (See Section 2.1.b.)
Economics- Employment and earnings-related data from Sections 1.1 and 1.2.

## Addfactors

Addfactors were included on some employment variables to smooth the transition from the latest historical data to program estimates. The need for addfactors is reviewed each year and they are implemented if necessary.

## Other input data

- Ratios of OASDI covered to NIPA wages by sector. NIPA wages by sector are available quarterly and annually from 1947 to 2015. They are published by the BEA and updated several times during the year. OASDI covered wages (1971 to 2013) are updated annually by the Economic process. Covered data for the latest historical year are estimated from tabulations of Form 941 and W-2 data.
- U.S. armed forces (EDMIL) by age and gender were estimated by the Department of Defense and published by the Census Bureau on a monthly basis (1948-2000) by single year of age (17 to 64) and gender. These data are no longer produced by Census.
- EDMIL by age and gender are estimated by the Economic process as the difference in the monthly resident plus Armed Forces overseas population and the monthly civilian population. These two populations are available from the Census Bureau on a monthly basis (April 2000 to October 2015) by single year of age (16 to 69) and gender. These data are updated several times a year.
- Railroad employment is covered by the Railroad Retirement Act. The annual historical data are for the period 1971 to 2013. An additional new year of historical data from the Railroad Retirement Board is usually available for inclusion in preparation of the next annual Trustees Report.
- Data obtained from Office of Research, Evaluation, and Statistics (ORES) are tabulations of quarterly Form 941 data. Data currently used are the OASDI, HI, and income taxable wages by sector for the most recent five years. The data represent changes in reported wages since the prior quarterly report. The most recent data are appended to previously reported data. Annual totals are computed and used to derive estimates of OASDI covered wages by sector for the latest historical years.
- Data obtained from the most recently available $1.0 \%$ CWHS active file, maintained on Social Security's mainframe and made available by ORES. The years of data are 1951 to the third year prior to the current Trustees Report year. The data are used for comparison of OASDI covered earnings from other sources.
- Data obtained from extracting information from the 1.0\% Employee-Employer Files, maintained on Social Security's mainframe and made available by ORES. Each year two files are created: a Version 1 file for the third year prior to the current Trustees Report and a Version 3 file for the fifth year prior to the current Trustees Report. Data currently being used are government and farm sector OASDI, HI, and total wages and employment. Data from the latest files are used to estimate OASDI covered wages for the years available on each file.
- Data obtained from quarterly IRS Form 941 files, provided by Office of Systems (OS). Data currently used are the OASDI and HI taxable wages for 1978 to the most recent year available. The data represent changes in reported wages since the prior quarterly report. The most recent data are appended to previously reported data. Annual totals are computed and used to derive estimates of HI taxable wages, which are then used to develop OASDI covered wages for the most recent historical years.
- Data from the Quarterly EPOXY Report, received in hard-copy and, more recently, electronic formats obtained from OS. The data currently used are the number of workers with OASDI taxable earnings, number of workers with HI taxable earnings, distribution of number of HI workers by wage intervals, distribution of number of OASDI workers by wage intervals, number of persons with OASDI taxable wages, number of persons with HI taxable wages, number of persons with OASDI taxable self-employment earnings, and number of persons with HI taxable self-employment earnings.
- Data obtained from the Quarterly Trust Fund Letter, received from Office of Financial Policy and Operations (OFPO). Data currently used are OASDI and HI taxable wages accumulated from all Forms 941 and $\mathrm{W}-2$ to date, and changes in selfemployment earnings and self-reported tips since the prior Letter. These data are for years 1978 to the most recent year available.
- Data obtained from OS on amounts of OASDI taxable wages on the Earnings Suspense File for 1937 through the second year prior to the current Trustees Report year. The data are used in estimating total OASDI covered employment.


## 2.3.c. Development of Output

## Equation 2.3.2-Total Employment (TE)

Based on the CPS, BLS estimates the total number of persons with any work experience (WE) during a calendar year. Average weeks worked (AWW) during a calendar year is defined as AWW = E * $52 /$ WE. Based on a 100 percent count of earnings reports (i.e., Form W-2 and Schedule SE) tabulated by SSA, OCACT estimates the total number of persons employed at any time during a calendar year (TE). Compared to WE, TE is a broader measure of employment. WE is an estimate of the number of workers in the civilian noninstitutional US population age 16 and over. TE is an estimate of employment in the broader

Social Security area population, which includes U.S. territories, the military, and institutions. TE also includes employment of workers who age 15 and younger.

AWW is disaggregated by gender and age and is projected as a function of a time trend and unemployment rate. WE is projected as ( $\mathrm{E} * 52$ ) / AWW. TE is projected as the product of its lagged value and the growth rate for WE. TE is adjusted by two multiplicative factors due to differences between E and TE over the recent historical period. The first factor accounts for the difference in growth between N and P between the last historical value for TE and the last historical value for E . The second factor adjusts the model estimate to the value of our most recent historical year, and our latest estimate based on partial quarterly data for that year.

## Equation 2.3.3 - Non-Covered Employment (NCE)

NCE is disaggregated by age and gender. Age groups include 14-15, 16-17, 18-19, 20-24, $25-29,30-34,35-39,40-44,45-49,50-54,55-59,60-64,65-69$, and 70 and over.
Employment may not be OASDI covered for a variety of reasons mostly related to the type of work. Consequently, NCE is further disaggregated to the type-of-work components listed below.

Federal Civilian Government - All Federal civilian employees are HI (i.e., Medicare) covered. All Federal Civilian employees hired in January 1984 and later are covered under the Federal Employees Retirement System (FERS) and are OASDI covered. Employees hired before January 1984 are covered under the Civil Service Retirement System (CSRS) and are not OASDI covered. This "closed group" of relatively older CSRS employees is projected to fall to near zero by 2030.

State and Local Government - In 1983, about 70 percent of State and Local Government (S\&L) employment and wages were covered under OASDI and HI. Beginning April 1986, all newly hired S\&L employees were covered under HI. Beginning January 1990, all S\&L employees not under an S\&L retirement system were covered under OASDHI.

By 2013, about 28 and 4 percent of S\&L employment (and wages) are still not covered under OASDI and HI respectively. The closed group of relatively older S\&L employees not covered under HI is projected to fall to near zero by 2030. S\&L employment not covered under OASDI is projected to grow at about the same rate as the labor force.

Students at Public Schools - Prior to 2000, students working at S\&L public schools were covered under OASDI and HI if the other school employees were covered. In 2000, legislation offered an "open season" allowing schools to remove their students from coverage. Virtually all major schools opted for removal. Hence, almost no students working at their public schools are covered under OASDI or HI. Students at public schools are projected to grow at about the same rate as the population aged 18 to 24 .

Election Workers - Most S\&L election workers are subject to an earnings test and are not covered under OASDHI. The earnings test was raised from $\$ 100$ to $\$ 1,000$ beginning

January 1995 and indexed beginning in 2000. Election workers are projected to grow at about the same rate as LC.

Private Household - The threshold for coverage of domestic employees' earnings was raised from $\$ 50$ per calendar quarter to $\$ 1,000$ per calendar year (CY) per employee. Domestic workers are no longer covered if under age 18. Private household employment is projected to grow at about the same rate as E and vary with RTP.

Students at Private Schools - All students working in private schools are not covered under OASDHI. Students at private schools are projected to grow at about the same rate as the population aged 18 to 24 .

Railroad - Employers do not submit payments for payroll taxes to the IRS for railroad employees. Railroad employees are projected by the Railroad Retirement Board.

Underground Economy Workers - Set to the at-any-time employed in the other immigrant population who have no reported earnings and therefore are part of the underground economy (i.e., TEO_UND).

## Foreign Students and Exchange Visitors

## Equation 2.3.5 - Self-Employed Only (SEO)

SEO is projected to grow at the same rate as ES.

## Equation 2.3.7 - Ratio of Combination Workers (RCMB)

Total CMB_TOT can be separated into two groups depending on whether they have OASDI covered wages under or over the TAXMAX. CMB_TOT with covered wages under the TAXMAX have taxable wages and self-employed income. CMB_TOT with covered wages over the TAXMAX have taxable wages only. CMB_TOT with covered wages over the TAXMAX would have paid taxes on their self-employed income if the TAXMAX had been eliminated.

Total CMB_TOT is projected as a ratio to WSW. This ratio is dependent on the RTP. If RTP rises, then the CMB_TOT increases.

## Equation 2.3.10-Ratio of Covered Wages (RWSC)

RWSC is disaggregated by the following sectors: Federal Civilian government, Federal Military, S\&L government, and Private.

Federal Civilian government - Total Federal civilian employment and wages are split by retirement system. Those under FERS are OASDI covered, while those under CSRS are not. Hence, the RWSC for the Federal civilian employment is defined as the ratio of wages for
employment under FERS to total Federal civilian wages. Employment and wages are projected for workers under each retirement system. Employment under CSRS is a closed group that is expected to fall to zero by about 2030. Employment under FERS is defined as total Federal employment less employment under CSRS. Total Federal civilian employment is projected to be constant over the short-range period, and about equal to the growth in the LC thereafter. The growth rates in the average wage for those under CSRS and FERS are projected based on, for the first five years, pay raises assumed under the most recent OMB FY Budget and on the growth rate in the CPI.

Federal Military - The RWSC for the Federal military sector is projected to remain constant at its latest actual historical level.

S\&L government - The RWSC for the S\&L government sector is projected to remain constant at its latest actual historical level.

Private - The private sector is separated into sub-sectors including private households, farm, railroad, tips, and a residual private "base". The RWSCs for the private household and farm sub-sectors are projected to remain constant at their latest actual historical levels. By definition, the RWSCs for the railroad and tips sub-sectors are projected to remain constant at 0.0 and 1.0 , respectively. The projected RWSC for the private base sub-sector is dependent on the ratio of EO wage workers in the private base sub-sector who are covered under the OASDI program to all EO wage workers in the private base sub-sector. We assume that all of EO will be wage workers employed in the private residual base sub-sector of the economy and that the proportion of EO that is covered under the OASDI program will decrease. Therefore, we assume that the RWSC for the private residual base sector will also decrease.

## Equation 2.3.12 - Ratio of Covered Self-Employed Earnings (RCSE)

The RCSE is projected to remain constant at its latest actual historical level.

## Equation 2.3.17 - Average Wage Index (AWI)

The growth in the AWI is projected to be equal to the growth in the average wage for employees with any wages (covered and noncovered) posted to the MEF (AWS_MEF). Total wages posted to the MEF (WS_MEF) is equal to WSC less wages posted to the ESF plus any non-OASDI covered wages posted to the MEF. Similarly, the total number of employees with any wages posted to the MEF (WSW_MEF) is equal to WSW less employees posted only to the ESF plus any employees with no OASDI covered wages posted to the MEF.

Equation 2.3.18 - OASDI Taxable Maximum (TAXMAX)
By law, the growth in the AWI is used to increase the TAXMAX.

### 2.4. Effective TAXABLE PAYROLL (TAXPAY)

## 2.4.a. Overview

TAXPAY estimates historical annual taxable earnings data including total employee OASDI taxable wages (WTEE), total employer taxable wages (WTER), and taxable self-employment income (SET). By law, each employee is required to pay the OASDI tax on wages from all covered jobs up to the TAXMAX, while each employer is required to pay the OASDI tax on the wages of each worker up to the TAXMAX. If an employee works more than one covered wage job and the sum of all covered wages exceeds the TAXMAX, the employee but not the employer is due a refund. Hence, WTER is greater than WTEE. The difference (i.e., WTER less WTEE) is defined as multi-employer refund wages (MER).

TAXPAY also estimates the historical annual OASDI effective taxable payroll (ETP). ETP is the amount of earnings in a year which, when multiplied by the combined employee-employer tax rate, yields the total amount of taxes due from wages and self-employment income in the year. ETP is used in estimating OASDI income and in determining income and cost rates and the actuarial balance. ETP is defined as WTER plus SET less one-half of MER.

TAXPAY projects annual values for ETP after first estimating its components. The components in turn are estimated by a collection of ratios. The employee taxable ratio (RWTEE) is defined as the ratio of WTEE to WSC. The multi-employer refund wage ratio (RMER) is defined as the ratio of MER to WSC. The self-employment income taxable ratio (RSET) is defined as the ratio of SET to CSE_TOT. Equations 2.4.1 through 2.4.8 outline the projection methodology.

$$
\begin{align*}
& \text { RWTEE }=\text { RWTEE }(\cdot)  \tag{2.4.1}\\
& \text { WTEE }=\text { RWTEE } * \mathrm{WSC}  \tag{2.4.2}\\
& \text { RMER }=\text { RMER }(\cdot)  \tag{2.4.3}\\
& \text { MER }=\text { RMER } * \mathrm{WSC}  \tag{2.4.4}\\
& \text { WTER }=\text { WTEE }+ \text { MER }  \tag{2.4.5}\\
& \text { RSET }=\text { RSET }(\cdot)  \tag{2.4.6}\\
& \text { SET }=\text { RSET } * \text { CSE_TOT }  \tag{2.4.7}\\
& \text { ETP }=\text { WTER }+ \text { SET }-0.5 * \text { MER } \tag{2.4.8}
\end{align*}
$$

In order to conform to the Trustees’ assumption for an ETP of 0.825 in the final short range year (see below), TAXPAY solves equations 2.4.1 through 2.4.8 iteratively, altering the trend adjustment on RWTEE until the assumed ratio is reached.

Over the short-range projection horizon (i.e., first 10 years), TAXPAY also projects annual OASDI wage tax liabilities (WTL) and self-employment tax liabilities (SEL). In Equation 2.4.9, WTL is the product of the effective taxable wages, defined as WTER less one-half of MER, and the combined OASDI employee-employer tax rate (TRW). In Equation 2.4.10, SEL is the product of SET and the OASDI self-employed tax rate (TRSE).

$$
\begin{align*}
& \text { WTL }=\text { WTER } * \text { TRW }  \tag{2.4.9}\\
& \text { SEL }=\mathrm{SET} * \mathrm{TRSE} \tag{2.4.10}
\end{align*}
$$

Also over the short-range horizon, TAXPAY decomposes WTL into quarterly wage tax liabilities (WTLQ), then to quarterly wage tax collections (WTLQC). TAXPAY also decomposes SEL into quarterly self-employed net income tax collections (SELQC).

$$
\begin{align*}
\text { WTLQ } & =W T L Q(\cdot)  \tag{2.4.11}\\
\text { WTLQC } & =W T L Q C(\cdot)  \tag{2.4.12}\\
\text { SELQC } & =\text { SELQC }(\cdot) \tag{2.4.13}
\end{align*}
$$

Finally, over the first two projected quarters, TAXPAY estimates of WTLQC and SELQC are replaced with ones from the most recent OMB FY Budget.

## 2.4.b. Input Data

## Trustees Assumptions

The Board of Trustees of the OASDI Trust Funds assumes that the ratio of effective OASDI taxable payroll to covered earnings for the final calendar year of the short range period (2025) is 0.825 and remains approximately constant in subsequent years.

Data used to obtain values input directly to model

- Data obtained from ORES by email for the amounts of single and multi-employer refunds for the latest 5 years. Each year, data are updated.
- Data obtained from ORES are tabulations of quarterly Form 941 data. Data currently used are the OASDI, HI, and income taxable wages by sector for the most recent five years. The data represent changes in reported wages since the prior quarterly report. The most recent data are appended to previously reported data. Annual totals are
computed and used to derive estimates of OASDI taxable wages by sector for the latest historical years.
- Data obtained from the most recently available 1.0\% CWHS active file, maintained on Social Security's mainframe and made available by ORES. The years of data are 1951 to the third year prior to the current Trustees Report year. The data are used for comparison of OASDI taxable earnings from other sources.
- Data obtained from quarterly IRS Form 941 files, provided by OS. Data currently used are the OASDI and HI taxable wages for 1978 to the most recent year available. The data represent changes in reported wages since the prior quarterly report. The most recent data are appended to previously reported data. Annual totals are computed and used to derive estimates of OASDI taxable wages for the most recent historical years.
- Data for the most recent ten years from the quarterly EPOXY Report, received in hard-copy and, more recently, electronic formats obtained from OS. The data currently used are the total number of workers with OASDI taxable earnings, total number of workers with OASDI self-employed taxable earnings, distribution of number of HI workers by wage intervals, distribution of number of OASDI workers by wage intervals, number of persons with OASDI taxable wages, number of persons with HI taxable wages, number of persons with OASDI taxable self-employment, number of persons with HI taxable self-employment, number of workers with singleemployer excess wages, and number of workers with multi-employer excess wages. Starting with the EPOXY reports produced during 2015, the data are now tabulated for years 1978 through the most recent year of data available (which was 2014 for the 2015 reports).
- Data obtained from the Quarterly Trust Fund Letter (QTFL), received from OFPO. Data currently used are OASDI and HI taxable wages accumulated from all Forms 941 and W-2 to date, and changes in self-employment earnings and self-reported tips since the prior QTFL. These data are for years 1978 to the most recent year available.


## Long-Range OASDI Projection Data

Historical and projected data from Sections 2.1, 2.2, and 2.3 are used as input. Data for the following variables have final year of 2099. Each variable is shown with the starting year of the data.

| ADDSETROD | Total add factor to OASDI taxable to covered self- <br> employment income ratio, 2015 |
| :--- | :--- |
| ADDWSTREEOD | Total add factor to OASDI taxable to covered wage ratio, <br> 2015 |
| ADDWSTREEODTREND | Component of total add factor to OASDI taxable to <br> covered wage ratio due to trend in ratio, 2015 |
| AIW | Average wage for indexing (\$), 1971 |


| AWSCFM | Average covered wage for farm workers (\$), 1971 |
| :---: | :---: |
| AWSCML | Average covered wage for military (\$), 1971 |
| DMWCHI | Deemed military wage credits for HI (\$ millions), 1983 |
| DMWCOD | Deemed military wage credits for OASDI (\$ millions), 1983 |
| ECFCHO | Number of HI-only covered Federal Civilian workers (millions), 1983 |
| ECFCOD | Number of OASDI covered Federal Civilian workers (millions), 1983 |
| ECHITOT | Number of HI covered workers (millions), 1987 |
| ECSEHI | Number of HI covered self-employed workers (millions), 1988 |
| ECSENOMAX | Number of covered self-employed workers if no taxable maximum (millions), 1988 |
| ECSEO | Number of OASDI covered self-employed only workers (millions), 1981 |
| ECSEOD | Number of OASDI covered self-employed workers (millions), 1981 |
| ECSLNOIS | Number of non-OASDI covered State and Local workers including students (millions), 1983 |
| ECSLP91 | Number of State and Local workers covered under OASDI under pre-1991 law (millions), 1983 |
| ECSLNRP | Number of OASDI covered State and Local workers with no retirement plan (millions), 1983 |
| ECSLOD | Number of OASDI covered State and Local workers (millions), 1983 |
| ECWSHI | Number of HI covered wage workers (millions), 1981 |
| ECWSOD | Number of OASDI covered wage workers (millions), 1981 |
| ECWSOD_MEF | Number of OASDI covered wage workers on the Master Earnings File (MEF) in millions, 1981 |
| ECWSOD(sex, age)_MEF | Number of OASDI covered wage workers on the Master Earnings File (MEF) in millions by sex (M/F) and age group (Under 16, 16-19, 20-24, 25-29, ..., 60-64, 65-69, 70 and over (millions), 1981 |
| ESLCG | Number of State and Local workers not covered under HI (millions), 1983 |
| ESLSTUD | Number of noncovered students at public schools employed by their school (millions), 1983 |
| GAPLAG | Ratio of real to potential GDP (units), 1971 |
| RTP | Ratio of real to potential GDP (units), 1971 |
| RU | Civilian unemployment rate (percent), 1971 |
| SEECCMB | Self-employed earnings of all SE workers who also earned wages in same year (\$ millions), 1991 |
| SEECHI | HI covered self-employed earnings (\$ millions), 1991 |
| SEECNOMAX | Covered self-employed earnings if no taxable maximum (\$ millions), 1991 |


| SEECOD | OASDI covered self-employed earnings (\$ millions), 1991 |
| :---: | :---: |
| SEECOD_OLD | OASDI covered self-employed earnings excluding selfemployed earnings of workers with covered wages greater than or equal to the OASDI taxable maximum (\$ millions), 1971 |
| SEETODCMB | OASDI taxable self-employment earnings of combination workers (\$ billions), 1995 |
| SEETODEXOG | Total OASDI taxable self-employment earnings (\$ millions), 1995 |
| SEETODSEO | OASDI taxable self-employment earnings of selfemployed only workers (\$ billions), 1995 |
| TAXMAXHI | HI taxable maximum (\$)-0 indicates no maximum, 1971 |
| TAXMAXOD | OASDI taxable maximum (\$), 1971 |
| TCFCD | Proportion of annual Federal Civilian wages earned in each quarter (units), 1971 |
| TCMD | Proportion of annual military wages earned in each quarter (units), 1971 |
| TCPD | Proportion of annual private sector wages earned in each quarter (units), 1971 |
| TCSLD | Proportion of annual State and Local wages earned in each quarter (units), 1971 |
| TETODCMB | Total OASDI taxable earnings of combination workers (\$ millions), 1995 |
| WSCCMB | Wages earned in same year by all SE workers with both types of earnings (\$ millions), 1991 |
| WSCFCHO | HI Covered wages of Federal Civilian HI-only workers (\$ millions), 1983 |
| WSCFCOD | OASDI Covered wages of Federal Civilian workers (\$ millions), 1971 |
| WSCFM | Covered wages of farm workers (\$ millions), 1971 |
| WSCHI | HI covered wages (\$ millions), 1971 |
| WSCML | Covered wages of members of the Armed Forces (\$ millions), 1971 |
| WSCOD | OASDI covered wages (\$ millions), 1971 |
| WSCOD_SF | OASDI covered wages on the Suspense File (\$ millions), 1971 |
| WSCPHH | Covered wages of private household workers (\$ millions), 1971 |
| WSCPNF | Covered wages of private nonfarm workers (\$ millions), 1971 |
| WSCSLHI | HI covered State and Local wages (\$ millions), 1971 |
| WSCSLNRP | Covered wages of State and Local workers with no retirement plan (\$ millions), 1991 |
| WSCSLOD | OASDI covered State and Local wages (\$ millions), 1971 |
| WSCSLP91 | Wages of State and Local workers covered under OASDI under pre-1991 law (\$ millions), 1971 |

WSD
WSP
WSS
WSSLCG
WSSLNOIS
WSSLSTUD
WSMEREFODEXOG
WSTEEODEXOG
WSTRRTPHI
WSTTIPSSR
WTWPO

Total NIPA wages (\$ millions), 1971
Total NIPA private sector wages (\$ millions), 1971
Total NIPA compensation (\$ millions), 1971
Wages of State and Local workers not covered under HI (\$ millions), 1983
Wages of non-OASDI covered State and Local workers including students (\$ millions), 1983
Wages of noncovered students at public schools employed by their school (\$ millions), 1983
OASDI multi-employer refund wages (\$ millions), 2014 Total OASDI taxable wages (\$ millions), 2015 Wages of railroad workers taxable under HI (\$ millions), 1971
Taxable tips reported by tip earner instead of employer (\$ millions), 1978
Proportion of annual Postal Service wages earned in each quarter (units), 1971

## Other direct input data

- FICA, SECA, and Federal Employer tax transfers by month from the Department of the Treasury for years 1984 to 2015.
- FICA, SECA, and Federal Employer tax transfers by month from the Department of the Treasury for January 2016.
- FICA and SECA tax transfers by month split by liability period from the Department of the Treasury for 1984 to 2015.
- Historical annual HI taxable self-employment earnings for 1983 to 1993. (Values from 1994 on are equal to HI covered earnings and are obtained from subprocess COV.).
- Historical annual OASDI taxable self-employment earnings for 1971 to 2014.
- Historical annual HI multi-employer refund wages for 1983 to 1993 (Values for 1994 on are zero because of the elimination of the HI taxable maximum.)
- Historical annual OASDI multi-employer refund wages for 1971 to 2014.
- Historical annual HI taxable wages for 1983 to 1993. (Values from 1994 on are equal to HI covered wages and are obtained from subprocess COV.).
- Historical annual OASDI taxable wages for 1971 to 2014.
- Historical annual HI-only taxable Federal Civilian wages for 1987 to 1993 . (Values
from 1994 on are equal to HI -only covered wages and are obtained from subprocess COV.).
- Historical annual OASDI taxable Federal Civilian wages for 1971 to 2014.
- Historical annual HI taxable Federal Civilian wages for 1987 to 1993. (Values from 1994 on are equal to HI covered wages and are obtained from subprocess COV.).
- Historical annual OASDI taxable farm sector wages for 1971 to 2014.
- Historical annual HI taxable farm sector wages for 1991 to 1993. (Values from 1994 on are equal to HI covered wages and are obtained from subprocess COV.).
- Historical annual OASDI taxable military sector wages for 1971 to 2014.
- Historical annual HI taxable military sector wages for 1991 to 1993. (Values from 1994 on are equal to HI covered wages and are obtained from subprocess COV.).
- Historical annual OASDI taxable State and Local government sector wages for 1971 to 2014.
- Historical annual HI taxable State and Local government sector wages for 1987 to 1993. (Values from 1994 on are equal to HI covered wages and are obtained from subprocess COV.).
- Historical annual OASDI taxable tips for employees as reported by employers for 1971 to 2014.
- Historical and projected annual OASDI taxable tips for employers as reported by employers for 1980 to 1987. (No tips were taxable for employers prior to 1980. Employer taxable tips equal employee in 1988 and after.)
- Quarterly OASDI and HI taxable wages for calendar year 2015. (These values do not add up to the annual values as described above. These quarterly distributions are used to split the annual 2015 values.)
- Historical FICA and SECA appropriation adjustments for OASI, DI, and HI by month for 1968 to 2014.
- Estimated FICA and SECA appropriation adjustments for OASI, DI, and HI for March 2016
- Historical FICA revenues for OASI, DI, and HI by quarter for 1984 to 2015.
- Historical SECA revenues for OASI, DI, and HI by quarter for 1984 to 2015.
- Historical Federal Employer revenues for OASI, DI, and HI by quarter for 1984 to 2015.
- Historical Deposits by States for OASI, DI, and HI by quarter for 1984 to 2015. (This is an obsolete type of revenue which has had no valid non-zero amount since 2002.)
- Historical single-employer refunds of excess taxes for OASI, DI, and HI by quarter for 1984 to 2014.
- Historical FICA credits for OASI and DI by quarter for 1984Q1 to 2015.
- Historical SECA credits for OASI, DI, and HI by quarter for 1984Q1 to 2015.
- Historical multi-employer refunds of excess taxes for OASI, DI, and HI by month for 1968 to 2015.
- Miscellaneous historical covered employment and earnings data:
o HI Covered self-employed workers for 1986 to 1993.
o Number of OASDI covered wage workers by age group and gender for 1996.
o HI covered self-employment earnings for 1971 to 1993.
o Covered self-employment earnings if there were no taxable maximum for 1971 to 1993.
o OASDI covered self-employment earnings for 1971 to 1993.
- Miscellaneous historical and fixed projected data:
o Quarterly distribution of annual OASDI taxable farm wages for 1971 to 2024.
o Quarterly OASDI covered private nonfarm sector wages for 1971 to 1977.
o Quarterly OASDI covered State and Local government sector wages for 1971 to 1977.
o Quarterly OASDI covered military sector wages for 1971 to 1977.
o Quarterly OASDI covered Federal Civilian sector wages for 1971 to 1977.
o Quarterly OASDI taxable private nonfarm sector wages for 1971 to 1977.
o Quarterly OASDI taxable State and Local government sector wages for 1971 to 1977.
o Quarterly OASDI taxable military sector wages for 1971 to 1977.
o Quarterly OASDI taxable Federal Civilian sector wages for 1971 to 1977.
o Quarterly OASDI taxable farm sector wages for 1971 to 1977.
o OASDI employee, employer, and self-employment tax rates from 1937 to 2100. These contribution rates are set according to the Social Security Act of 1935 and amendments to the Act through 2015. The rates are updated when legislation mandates a change. The rates were unchanged from 2000 to 2015. The Bipartisan Budget Act of 2015 reallocated the OASI and DI employeeemployer and self-employment tax rates for years 2016 through 2018. The OASDI rates remain the same and the rates revert to the ones in effect for 2000 to 2015 in 2019 and thereafter.
o Annual OASDI employee credit tax rate for 1984.
o Annual OASDI self-employment credit tax rates for 1984 to 1989.
o Annual reductions in OASDI employee and self-employment tax rates due to the payroll tax holiday for 2011 and 2012.
o Annual trend variable for taxable to covered wage ratio calculation for 1971 to 2100 (no longer used)

0 Annual trend variable for taxable to covered self-employment earnings ratio calculation for 1971 to 2100 (no longer used)

- Average OASDI covered wages by age groups and gender for 1996.
- Ratio of OASDI taxable to covered wages by age groups and gender for 1996.
- Corrections to prior FICA appropriation adjustments made in March 2000.
- Projected single-employer refund wages by calendar year for 2015 through 2026.


## 2.4.c. Development of Output

## Equation 2.4.1 - Employee Taxable Ratio (RWTEE)

Over the short-range projection horizon, the projected value for RWTEE is the sum of the model's "raw" estimate and an addfactor consisting of four components. The raw estimate
for RWTEE is dependent on the distribution of workers by wage interval, the RELMAX, RTP, the age-gender distribution of wage workers, a time trend adjustment, and a base-year error adjustment. The projected distribution of workers by wage interval is an average (or amalgam) distribution over the 1993 through 2013 period. Holding other factors constant, a distribution with relatively more workers with wages over the TAXMAX leads to a lower RWTEE. The RELMAX is defined as the ratio of the TAXMAX to the ACW. A higher RELMAX leads to a higher RWTEE.

An increase in the RTP leads to a lower RWTEE. The change in the projected RWTEE due to the change in the age-gender distribution of wage workers is calculated by allowing employment by age and gender to change while holding taxable ratios (and average covered wages) by age and gender constant to levels in 1996. The time trend adjustment reduces the level of RWTEE by about 0.6 percentage point over the short-range projection horizon. The base-year error adjustment starts with the value obtained by subtracting the estimated value of RWTEE for the latest historical (or base) year from the actual value and phases this amount out linearly over the ten years of the short-range projection period.

RWTEE is assumed to remain constant over the long-range projection horizon.
RWTEEs are also projected for various sub-aggregates including Federal Civilian employees under FERS and CSRS, Federal Civilian employees under CSRS only, S\&L employees covered under OASDI, S\&L employees covered under HI only, U.S. armed forces, and agriculture. The RWTEE for each sub-aggregate is dependent only on its sub-aggregate RELMAX, that is, the ratio of the TAXMAX to the sub-aggregate's average covered wage.

## Equation 2.4.2-Employee Taxable Wages (WTEE)

WTEE is computed by multiplying the ratio of taxable employee wages to covered wages by the level of covered wages.

## Equation 2.4.3 - Multi-Employer Refund Wage Ratio (RMER)

The RMER is functionally related to the RWTEE. As RWTEE approaches one, then RMER approaches zero. In between the limit values, RMER is positive. Given the present position of RWTEE and RMER on the function, a projected decline in RWTEE leads to an increase in RMER.

The projected RMER is also dependent on RU. An increase in RU leads to a decrease in RMER.

RMER is assumed to remain constant over the long-range projection horizon.
Equation 2.4.4 - Multi-Employer Refund Wages (MER)
MER is computed by multiplying the ratio of multi-employer refund wages to covered wages
by the level of covered wages.

## Equation 2.4.5-Employer Taxable Wages (WTER)

WTER is computed by adding employer taxable wages to multi-employer refund wages.
Equation 2.4.6 - Self-Employed Net Income Taxable Ratio (RSET)
The RSET is disaggregated by type of self-employed worker, SEO and CMB_TOT.
SEO - The RSET is dependent on the distribution of self-employed workers by income interval and a RELMAX. The projected distribution of self-employed workers by income interval is set to the 2011 distribution. The RELMAX is defined here as the ratio of the TAXMAX to the average income for SEO. A higher RELMAX leads to a higher RSET.

CMB_TOT - Taxable self-employed net income for CMB_TOT is projected in two steps. First, a taxable earnings (wages and self-employment income) ratio for CMB_TOT is projected based on the 2011 distribution and a RELMAX defined as the ratio of the TAXMAX to the average covered earnings. The projected level of taxable earnings for CMB_TOT is the product of the estimated taxable earnings ratio for CMB_TOT and their covered earnings. Second, a taxable wage ratio for CMB_TOT is projected based on a RELMAX defined as the ratio of the TAXMAX to the average covered wage for CMB_TOT. The projected level of taxable wages for CMB_TOT is the product of the estimated taxable wage ratio for CMB_TOT and their covered wages.

Taxable self-employed net income for CMB_TOT is obtained by subtracting taxable wages from taxable earnings for CMB_TOT.

A "combined" RSET is calculated as the ratio of the sum of taxable self-employment income for SEO and CMB_TOT to CSE_TOT. As with the RWTEE, the combined RSET is adjusted over the short-range period due to other factors (i.e., RTP, the age-gender distribution of workers, and a trend). The effect of the other factors are taken from RWTEE and "scaled." That is, RSET is adjusted by a percent effect (as opposed to percentage point) that is equal to the percent change in RWTEE due to changes in these other factors.

It is important to note that while the RWTEE is held constant after the short-range period, the RSETs for self-employed workers are not. After the short-range period, the projected RSETs for SEO and CMB_TOT continue to be dependent on their respective RELMAXs. Since by law the TAXMAX grows at the rate of the AWI and since ACSE is assumed to grow faster than the ACW (since only ACW declines with the growth in fringe benefits), the RELMAXs for self-employed workers decline over the long-range period while the RELMAX for wage workers is approximately constant. Hence, the RSETs for SEO and CMB_TOT are projected to decline over the long-range period while the RWTEE is held constant.

Equation 2.4.7-Taxable Self-Employment Income (SET)

SET is computed by multiplying the self-employment income taxable ratio by the level of covered self-employment income.

## Equation 2.4.8 - Effective Taxable Payroll (ETP)

ETP is computed by adding employer taxable wages and taxable self-employment income and subtracting from that total one-half of multi-employer refund wages. (Only employees can obtain refunds of excess taxes withheld in multi-employer refund wage cases.)

As noted above, in order to meet the Trustees’ assumption that the ratio of ETP to total covered earnings is 0.825 in the last year of the short-range period, equations 2.4.1 through 2.4.8 are solved repeatedly with changes to the time-trend adjustment in equation 2.4.1 until the ratio is obtained.

## Equation 2.4.11 - Quarterly Wage Tax Liabilities (WTLQ)

Total WTLQ is summed from sector components that include Federal Civilian, Federal Military, S\&L, Private Household, Farm, Self-reported Tips, and residual Private Nonfarm. Sector WTLQs are determined by computing ratios of quarterly to annual liabilities for each quarter. These are calculated for the Private Nonfarm, S\&L, Federal Civilian and Military sectors. Each is dependent on the quarterly distribution of WSD and the RWTEE for the relevant sector, and on a payday adjustment that takes into account the actual number of paydays that fall into a particular calendar quarter. WTLQ ratios for the other sectors are assumed to be constants over the projection horizon.

## Equation 2.4.12- Quarterly Wage Tax Collections (WTLQC)

Employers incur tax liabilities when they pay wages to their employees. These liabilities are required to be deposited with the U.S. Treasury by employers based on the amount of total payroll tax liability (income taxes plus Social Security and Medicare taxes withheld) accumulated. Some very large employers must deposit their tax liabilities the next banking day after paying their employees. Other levels of accumulated tax liabilities require depositing within three days, by the middle of the following month, or by the end of the month following the quarter. If employers follow these deposit requirements, the result is that all tax liability for a particular quarter is deposited by the last day of the month following the end of the quarter. Thus, the WTLQC for any particular quarter are the sum of the tax liabilities deposited for wages paid in the same quarter and the liabilities deposited for wages paid in the prior quarter.

WTLQC are summed from sector components that include the Federal Civilian, Federal Military, Farm, S\&L, and residual Private Nonfarm (including Private Household and SelfReported Tips). For the Federal Civilian and Military sectors, the WTLQC are set equal to their respective WTLQ since tax liabilities for the two sectors are considered collected immediately. The WTLQC for Farm is also set equal to its WTLQ, due in part to the fact that farms report tax liabilities annually. For the S\&L and Private Nonfarm sectors, the WTLQC amount is computed by adding the product of WTLQ and the proportion of WTLQ that
should be deposited in the same quarter in which the wages were paid to the WTLQ from the previous quarter which was not deposited in that quarter. Each quarter's proportion is based on the deposit requirements and estimates of accumulated tax liabilities, which in turn are based on firm size (or total wages paid). Separate proportions are estimated for the S\&L and the Private Nonfarm sector because of the large difference in wage distributions between them.

## Equation 2.4.13 - Quarterly Self-Employed Net Income Tax Collections (SELQC)

For wage workers, annual liabilities (WTL) are distributed to quarterly liabilities (WTLQ), which in turn are distributed to quarterly collections (WTLQC). However, for self-employed workers, annual liabilities (SEL) are distributed directly to SELQC, since the SSA only receives self-employed liability amounts on an annual basis (from tabulations of Form 1040 Schedule SE provided by IRS).

SEL for a particular calendar year are distributed as collections to the four quarters of that year and to the first three quarters of the next year. This distribution uses quarterly proportions that are based on an historical pattern of the amount of SEL collected in each month, as estimated by the OTA. The OTA estimates reflect IRS regulations that require selfemployed workers to deposit estimated tax liabilities four times a year (January, April, June, and September). The program computes the collection distribution ratios based on the OTA estimates, which are input to the program. The ratios for projected years are determined by averaging the ratios for all years 1997 through the most recent year for which the OTA estimates are complete.

## Appendix 2-1 <br> Equations

### 2.1 U.S. Employment (USEMP)

## UNEMPLOYMENT RATES, PRELIMINARY

## MALES

```
RM1617_P = RM1617_P.1 + (-36.2076 * DIFF(RTP) - 14.2816 * DIFF(RTP.1) - 26.6756 * DIFF(RTP.2) - 16.9202 * DIFF(RTP.3)) * 50.00/44.72
            RM1617
            Ordinary Least Squares
            QUARTERLY data for 132 periods from 1976Q1 to 2008Q4
                    Date: 23 OCT 2009
                            diff(rm1617)
        = - 36.2076 * diff(rtp) - - 14.2816 * diff(rtp)[-1]
            - 26.6756 * diff(rtp)[-2] - 16.9202 * diff(rtp)[-3]
                (1.63426) (1.06303)
                + 1.64214 * minw - 0.90365 * minw[-1] + 0.06020 * minw[-2]
                (1.76311) (0.70381) (0.04365)
                    - 0.77627 * minw[-3] - 0.12616
                (0.72653) (0.26611)
\begin{tabular}{lrlrlr} 
Sum Sq & 198.967 & Std Err & 1.2719 & LHS Mean & 0.0553 \\
R Sq & 0.1483 & R Bar Sq & 0.0929 & F 8,123 & 2.6769
\end{tabular}
                            D.W.( 1) 2.5142 D.W.( 4) 2.3291
```

RM1819_P = RM1819_P. $1+(-48.4227$ * DIFF(RTP) - 25.8766 * DIFF(RTP.1) - 21.7466 * DIFF(RTP.2) + 1.1551 * DIFF(RTP.3)) * 50.00/44.72

RM1819
Ordinary Least Squares
QUARTERLY data for 132 periods from 1976Q1 to 2008Q4
Date: 23 OCT 2009
diff(rm1819)
$=\underset{(3.45103)}{-48.4227}$ * $\operatorname{diff(rtp)} \underset{(1.77685)}{25.8766}$ * $\operatorname{diff(rtp)[-1]}$
-21.7466 * $\operatorname{diff}(r t p)[-2]+\underset{(0.08165)}{1.15512}$ * $\operatorname{diff}(r t p)[-3]$ (1.49900) (0.08165)

+ 0.62723 * minw - 0.48738 * minw[-1] + 0.67739 * minw[-2] (0.75770) (0.42710) (0.55270) - 0.79385 * minw[-3] - 0.11294 (0.83595) (0.26803)
$\begin{array}{lrlrll}\text { Sum Sq } & 157.172 & \text { Std Err } & 1.1304 & \text { LHS Mean } & 0.0235\end{array}$
R Sq 0.1890 R Bar Sq 0.1362 F 8,123 3.5828
D.W.( 1) 2.6908 D.W. ( 4) 2.4565

RM2024_P = RM2024_P. $1+(-51.6518 * \operatorname{DIFF}(R T P)-16.6465 * \operatorname{DIFF}(R T P .1)-13.1350 * \operatorname{DIFF}(R T P .2)-10.9309 * \operatorname{DIFF}(R T P .3)) *$ 50.00/44.72

```
    RM2024
    Ordinary Least Squares
    QUARTERLY data for 132 periods from 1976Q1 to 2008Q4
    Date: 23 OCT 2009
```

```
            diff(rm2024)
                        = - 51.6518 * diff(rtp) - 16.6465 * diff(rtp)[-1]
                (7.75482) (2.35721)
                - 13.1350 * diff(rtp)[-2] - 10.9309 * diff(rtp)[-3] + 0.00093
                (1.86731)
                    (1.59404)
                                    (0.01922)
\begin{tabular}{lrlllll} 
Sum Sq & 38.7297 & Std Err & 0.5522 & LHS Mean & -0.0048 \\
R Sq & 0.4356 & R Bar Sq & 0.4178 & F 4,127 & 24.5002 \\
D.W. ( 1) & 2.4679 & D.W.( 4) & 2.2856 & & &
\end{tabular}
RM2529_P = RM2529_P.1 + (-37.9533 * DIFF(RTP) - 17.3941 * DIFF(RTP.1) - 14.9170 * DIFF(RTP.2) - 7.0513 * DIFF(RTP.3)) * 50.00/44.72
\begin{tabular}{|c|c|c|c|c|c|}
\hline \multicolumn{6}{|l|}{RM2529} \\
\hline \multicolumn{6}{|l|}{Ordinary Least Squares} \\
\hline \multicolumn{6}{|l|}{QUARTERLY data for 129 periods from 1976Q4 to 2008Q4} \\
\hline \multicolumn{6}{|l|}{Date: 23 OCT 2009} \\
\hline \multicolumn{6}{|l|}{diff(rm2529)} \\
\hline \multicolumn{6}{|l|}{\[
=-\underset{(7.06307)}{-37.9533} * \operatorname{diff(rtp)}-\underset{(3.05222)}{-17.3941} * \operatorname{diff(rtp)[-1]}
\]} \\
\hline \multicolumn{6}{|r|}{} \\
\hline Sum Sq & 23.9433 & Std Err & 0.4394 & LHS Mean & 0.0095 \\
\hline R Sq & 0.4417 & R Bar Sq & 0.4237 & F 4,124 & 24.5278 \\
\hline D.W.( 1) & 2.1341 & D.W. ( 4) & 2.4166 & & \\
\hline
\end{tabular}
RM3034_P = RM3034_P.1 + (-23.6417 * DIFF(RTP) - 14.1284 * DIFF(RTP.1) - 7.5008 * DIFF(RTP.2) - 9.7232 * DIFF(RTP.3)) * 50.00/44.72 RM3034
Ordinary Least Squares QUARTERLY data for 129 periods from 1976Q4 to 2008Q4 Date: 23 OCT 2009
diff(rm3034)
```

```
\[
\begin{aligned}
& =-23.6417 \text { * diff(rtp) - } 14.1286 \text { * diff(rtp)[-1] } \\
& \text { (6.21241) (3.50067) } \\
& \text { - } 7.50079 \text { * diff(rtp)[-2] - } 9.7232 \text { * diff(rtp)[-3] + } 0.01058 \\
& \text { (1.85832) (2.50593) (0.38580) } \\
& \begin{array}{lrlrlrr}
\text { Sum Sq } & 12.0091 & \text { Std Err } & 0.3112 & \text { LHS Mean } & 0.0119 \\
\text { R Sq } & 0.4221 & \text { R Bar Sq } & 0.4034 & \text { F } 4,124 & 22.6397
\end{array} \\
& \text { D.W. (1) } 2.1876 \text { D.W. (4) } 1.8816
\end{aligned}
\]
RM3539_P = RM3539_P. \(1+(-27.6828\) * DIFF(RTP) -5.4850 * DIFF(RTP.1) - 10.8974 * DIFF(RTP.2) - 9.8932 * DIFF(RTP.3)) * 50.00/44.72 RM3539
Ordinary Least Squares
QUARTERLY data for 129 periods from 1976Q4 to 2008Q4
Date: 23 OCT 2009
diff(rm3539)
\(=\underset{(6.57840)}{-27.6828}\) * \(\operatorname{diff(rtp)}-\underset{(1.22901)}{5.48498}\) * \(\operatorname{diff(rtp)[-1]}\)
\begin{tabular}{|c|c|c|c|c|c|}
\hline Sum Sq & 14.6843 & Std Err & 0.3441 & LHS Mean & 0.0130 \\
\hline R Sq & 0.3795 & R Bar Sq & 0.3595 & F 4,124 & 18.9589 \\
\hline D.W.( 1) & 2.3381 & D.W. ( 4) & 1.9092 & & \\
\hline
\end{tabular}
RM4044_P = RM4044_P. \(1+(-14.6558 * \operatorname{DIFF}(R T P)-14.9735 * \operatorname{DIFF}(\) RTP.1 \()-8.2594 * \operatorname{DIFF}(R T P .2)-5.5023 * \operatorname{DIFF}(R T P .3)) * 50.00 / 44.72\) RM4044
Ordinary Least Squares
QUARTERLY data for 129 periods from 1976Q4 to 2008Q4
```

Date: 23 OCT 2009
diff(rm4044)


RM4549_P = RM4549_P. $1+(-20.7806 * \operatorname{DIFF}(R T P)-11.5121 * \operatorname{DIFF}(R T P .1)-9.9409 * \operatorname{DIFF}(R T P .2)+1.5480 * \operatorname{DIFF}($ RTP.3 $)$ * 50.00/44.72

```
RM4549
    Ordinary Least Squares
    QUARTERLY data for }129\mathrm{ periods from 1976Q4 to 2008Q4
    Date: 23 OCT 2009
    diff(rm4549)
        = - 20.7806 * diff(rtp) - - 11.5121 * * * diff(rtp)[-1]
            - 9.9409 * diff(rtp)[-2] + 1.54797 * diff(rtp)[-3] + 0.00874
            (2.39795)
                        (0.38844)
                                    (0.31046)
    Sum Sq 12.6680 Std Err 0.3196 LHS Mean 0.0114
    R Sq 0.3249 R Bar Sq 0.3031 F 4,124 14.9185
```

RM5054_P = RM5054_P.1 + (-19.3341 * DIFF(RTP) $-9.5336 * \operatorname{DIFF}($ RTP.1) $-8.8784 * \operatorname{DIFF}($ RTP.2 $)-7.6218 * \operatorname{DIFF}($ RTP.3) $) * 50.00 / 44.72$
RM5054
Ordinary Least Squares
QUARTERLY data for 129 periods from 1976Q4 to 2008Q4
Date: 23 OCT 2009
diff(rm5054)
$=\underset{(4.72314)}{-19.3341} * \operatorname{diff}(\mathrm{rtp}) \underset{(2.19601)}{-}{ }_{(2.5336}^{*} \operatorname{diff}(\mathrm{rtp})[-1]$
-8.87840
$(2.04491)$$* \operatorname{diff(rtp)[-2]} \begin{gathered}-\underset{(1.82617)}{7.62180} *\end{gathered}$ diff(rtp)[-3]$+\underset{(0.36742)}{0.01083}$
Sum Sq 13.8950 Std Err 0.3347 LHS Mean 0.0118
$\begin{array}{lllllll}R & \text { Sq } 0.2957 & \text { R Bar Sq } 0.2730 \quad F & 4,124 & 13.0163\end{array}$
D.W.( 1) 2.1290 D.W.( 4) 1.7836
RM5559_P = RM5559_P.1 + (-25.9031 * DIFF(RTP) - 11.4442 * DIFF(RTP.1) - 4.5421 * DIFF(RTP.2) + 0.55815 * DIFF(RTP.3)) *
50.00/44.72
RM5559
Ordinary Least Squares
QUARTERLY data for 129 periods from 1976 Q4 to 2008Q4
Date: 23 OCT 2009
diff(rm5559)
$=-25.9031$ * diff(rtp) - 11.4442 * diff(rtp)[-1]
(5.21572) (2.17280)

Sum Sq 20.4526 Std Err 0.4061 LHS Mean 0.0068
$\begin{array}{lllllll}\text { R Sq } & 0.2605 & \text { R Bar Sq } 0.2366 & \text { F } & 424 & 10.9177\end{array}$
$\begin{array}{llllll}\text { D.W.( 1) } 2.2469 & \text { D.W. ( 4) } & 1.7904 & & \end{array}$
RM6064_P = RM6064_P. $1+(1.3133 * \operatorname{DIFF}($ RTP $)-12.9625 * \operatorname{DIFF}($ RTP.1) $-2.4816 * \operatorname{DIFF}($ RTP. 2$)-14.4797 * \operatorname{DIFF}($ RTP.3 $)$ * 50.00/44.72

```
            RM6064
            Ordinary Least Squares
                    QUARTERLY data for }129\mathrm{ periods from 1976Q4 to 2008Q4
                    Date: 23 OCT 2009
                    diff(rm6064)
                    = (1.31332 * * diff(rtp) - - 12.9625 * diff(rtp)[-1]
                - 2.48164 * diff(rtp)[-2] - 14.4797 * diff(rtp)[-3] + 0.00491
                    (0.44909) (2.72583) (0.13091)
Sum Sq 22.5085 Std Err 0.4261 LHS Mean 0.0021
R Sq 0.1187 R Bar Sq 0.0903 F 4,124 4.1768
D.W.(1) 2.3485 D.W.( 4) 1.9007
RM6569_P = RM6569_P.1 + (-19.5151 * DIFF(RTP) + 4.9785 * DIFF(RTP.1) -13.3449 * DIFF(RTP.2) + 2.4706 * DIFF(RTP.3)) *
50.00/44.72
    RM6569
    Ordinary Least Squares
    QUARTERLY data for }129\mathrm{ periods from 1976Q4 to 2008Q4
    Date: 23 OCT 2009
    diff(rm6569)
        = - 19.5151 * diff(rtp) + 4.97852 * diff(rtp)[-1]
                (2.18595) (0.52582)
                - 13.3449 * diff(rtp)[-2] + 2.47056 * diff(rtp)[-3] - 0.01208
                (1.40935) (0.27142) (0.18783)
\begin{tabular}{lrlrlr} 
Sum Sq & 66.0895 & Std Err & 0.7301 & LHS Mean & -0.0096 \\
R Sq & 0.0551 & R Bar Sq & 0.0246 & F 4,124 & 1.8065
\end{tabular}
    D.W.( 1) 2.6235 D.W.( 4) 1.5080
RM7074_P = RM7074_P.1 + (4.1938 * DIFF(RTP) - 5.9012 * DIFF(RTP.1) - 27.0406 * DIFF(RTP.2) + 7.0400 * DIFF(RTP.3)) * 50.00/44.72
    RM7074
    Ordinary Least Squares
    QUARTERLY data for 111 periods from 1981Q2 to 2008Q4
    Date: 23 OCT 2009
    diff(rm7074)
        = 4.19384 * diff(rtp) - 5. 5.90117 * * diff(rtp)[-1]
                - 27.0406 * diff(rtp)[-2] + 7.03995 * diff(rtp)[-3] + 0.02434
                (1.59539) (0.43075) (0.24587)
\begin{tabular}{lrlrlll} 
Sum Sq & 114.781 & Std Err & 1.0406 & LHS Mean & 0.0185 \\
R Sq & 0.0282 & R Bar Sq & -0.0085 & F 4,106 & 0.7684
\end{tabular}
    (1) 2.8303
RM75O_P = RM75O_P.1 + (-12.1042 * DIFF(RTP) - 15.6142 * DIFF(RTP.1) + 7.06185 * DIFF(RTP.2) - 2.5738 * DIFF(RTP.3)) * 50.00/44.72
    RM750
    Ordinary Least Squares
    QUARTERLY data for }111\mathrm{ periods from 1981Q2 to 2008Q4
    Date: 23 OCT 2009
    diff(rm75o)
        = - 12.1042 * diff(rtp) - 15.6142 * diff(rtp)[-1]
            (0.80507) (0.98509)
                + 7.06185 * diff(rtp)[-2] - 2.57381 * diff(rtp)[-3] + 0.00860
                (0.45088) (0.17042) (0.09395)
    Sum Sq }r\mathrm{ 98.0128 
    R Sq 0.0212 R Bar Sq -0.0157 F 4,106 0.5749
    D.W.( 1) 2.6726 D.W.( 4) 1.8788
```


## FEMALES

```
RF1617_P = RF1617_P.1 + (-27.3243 * DIFF(RTP) + 13.4173 * DIFF(RTP.1) - 50.4583 * DIFF(RTP.2) - 0.3678 * DIFF(RTP.3))*50.00/44.72
    RF1617
        Ordinary Least Squares
        QUARTERLY data for 132 periods from 1976Q1 to 2008Q4
    Date: 23 OCT 2009
    diff(rf1617)
        = - 27.3243 * diff(rtp) + + (0.81297) 13.4173 * diff(rtp)[-1]
            - 50.4583 * diff(rtp)[-2] - 0.36782 * diff(rtp)[-3]
                (3.23806) (0.02421)
                + 0.33050 * minw + 0.19356 * minw[-1] + 0.18090 * minw[-2]
                (0.37169) (0.15791) (0.13742)
                - 0.68394 * minw[-3] - 0.13675
                (0.67051) (0.30213)
    Sum Sq 181.339 Std Err 1.2142 LHS Mean -0.0136
    R Sq 0.1227 R Bar Sq 0.0656 F 8,123 2.1501
    D.W.( 1) 2.9150 D.W.( 4) 2.4862 
```

RF1819_P = RF1819_P. $1+(-42.6358$ * DIFF(RTP) -13.6261 * DIFF(RTP.1) + 9.5650 * DIFF(RTP.2) - 31.4798 * DIFF(RTP.3) $)$ 50.00/44.72
RF1819
Ordinary Least Squares
QUARTERLY data for 132 periods from 1976Q1 to 2008Q4
Date: 23 OCT 2009
diff(rf1819)

```
        = - 42.6358 * diff(rtp) - 13.6261 * diff(rtp)[-1]
            (3.54124) (1.09043)
            + 9.5650 * diff(rtp)[-2] - 31.4798 * diff(rtp)[-3]
                (0.76838) (2.59333)
                + 0.27394 * minw - 0.95221 * minw[-1] + 1.01588 * minw[-2]
                (0.38566) (0.97247) (0.96600)
                - 0.32609 * minw[-3] - 0.05888
                (0.40019) (0.16286)
    Sum Sq rllll
    D.W.( 1) 2.7048 D.W.( 4) 2.3991
```

RF2024_P = RF2024_P. $1+(-16.9400 * \operatorname{DIFF}(R T P)-13.2669 * \operatorname{DIFF}(R T P .1)-7.8323 * \operatorname{DIFF}(R T P .2)-8.6887 * \operatorname{DIFF}(R T P .3) * 50.00 / 44.72$
RF2024
Ordinary Least Squares
QUARTERLY data for 132 periods from 1976 Q1 to 2008Q4
Date: 23 OCT 2009
diff(rf2024)
$=\underset{(3.09139)}{-16.9400}$ * $\operatorname{diff(rtp)}-\underset{(2.28348)}{13.2669}$ * $\operatorname{diff(rtp)[-1]}$

Sum Sq 26.2142 Std Err 0.4543 LHS Mean -0.0275
$\begin{array}{lllllll}R & \text { Sq } & 0.1917 & R & \text { Bar Sq } & 0.1663 & F \\ 4.127 & 7.5310\end{array}$
$\begin{array}{lll}\text { D.W.( 1) } & 2.5252 & \text { D.W. ( 4) } \\ 2.1988\end{array}$
RF2529_P = RF2529_P. $1+(-15.5798 * \operatorname{DIFF}(R T P)-11.9097 * \operatorname{DIFF}(R T P .1)-9.8424 * \operatorname{DIFF}(R T P .2)-2.7555 * \operatorname{DIFF}(R T P .3)$ ) $50.00 / 44.72$

```
RF2529
    Ordinary Least Squares
    QUARTERLY data for }129\mathrm{ periods from 1976Q4 to 2008Q4
    Date: 23 OCT 2009
    diff(rf2529)
        = - 15.5798 * diff(rtp) - 11.9097 * diff(rtp)[-1]
            (3.32423) (2.39607)
                - 9.8424 * diff(rtp)[-2] - 2.75548 * diff(rtp)[-3] - 0.01837
                (1.97999) (0.57663) (0.54398)
    Sum Sq 18.2145 Std Err 0.3833 LHS Mean -0.0172
    R Sq 0.2094 R Bar Sq 0.1839 F 4,124 8.2130
    D.W.( 1) 2.3764 D.W.( 4) 2.0455
RF3034_P = RF3034_P.1 + (-12.5396 * DIFF(RTP) - 1.6601 * DIFF(RTP.1) - 21.0289 * DIFF(RTP.2) + 0.0881 * DIFF(RTP.3))*50.00/44.72
    RF3034
    Ordinary Least Squares
    QUARTERLY data for }129\mathrm{ periods from 1976Q4 to 2008Q4
    Date: 23 OCT 2009
    diff(rf3034)
        = - 12.5396 * diff(rtp) - 1.66005 * diff(rtp)[-1]
            (2.58233) (0.32234)
            - 21.0289 * diff(rtp)[-2] + + 0.08813 * diff(rtp)[-3] - - 0.01780) (0.01851 
    Sum Sq 19.5533 Std Err 0.3971 LHS Mean -0.0179
    R Sq 0.1929 R Bar Sq 0.1669 F 4,124 7.4100
    D.W.( 1) 2.5217 D.W.( 4) 2.0509
RF3539_P = RF3539_P.1 + (-21.9314 * DIFF(RTP) - 3.0139 * DIFF(RTP.1) - 7.8723 * DIFF(RTP.2) - 6.4785 * DIFF(RTP.3))*50.00/44.72;
    RF3539
    Ordinary Least Squares
    QUARTERLY data for 129 periods from 1976Q4 to 2008Q4
    Date: 23 OCT 2009
    diff(rf3539)
        = - 21.9314 * diff(rtp) - - 3.01391 * diff(rtp)[-1]
            - 7.87232 * diff(rtp)[-2] - % (1.47846 * diff(rtp)[-3] - 0.0.00217 (0.0694)
\begin{tabular}{lrlrll} 
Sum Sq & 15.6401 & Std Err & 0.3551 & LHS Mean & -0.0005 \\
R Sq & 0.2453 & R Bar Sq & 0.2210 & F 4,124 & 10.0778 \\
D.W.( 1) & 2.3835 & D.W.( 4) & 1.7499 & & \\
\end{tabular}
RF4044_P = RF4044_P.1 + (-7.7893 * DIFF(RTP) - 7.7152 * DIFF(RTP.1) - 5.7849 * DIFF(RTP.2) - 2.7298 * DIFF(RTP.3))*50.00/44.72
    RF4044
    Ordinary Least Squares
    QUARTERLY data for }129\mathrm{ periods from 1976Q4 to 2008Q4
    Date: 23 OCT 2009
    diff(rf4044)
        = - 7.78933 * diff(rtp) - 7.71518 * diff(rtp)[-1]
            (1.61580) (1.50905)
                - 5.78494 * diff(rtp)[-2] - 2.72977 * diff(rtp)[-3] - 0.00986
                (1.13141) (0.55538) (0.28389)
    Sum Sq 19.2707 Std Err 0.3942 LHS Mean -0.0095
    R Sq 0.0780 R Bar Sq 0.0483 F 4,124 2.6232
    D.W.( 1) 2.2897 D.W.( 4) 2.2106
```

```
RF4549_P = RF4549_P.1 + (-7.8747 * DIFF(RTP) -12.5212 * DIFF(RTP.1) + 3.56675 * DIFF(RTP.2) - 5.4812 * DIFF(RTP.3))*50.00/44.72
    RF4549
        Ordinary Least Squares
        QUARTERLY data for 129 periods from 1976Q4 to 2008Q4
        Date: 23 OCT 2009
    diff(rf4549)
        = - 7.87468 * diff(rtp) - - 12.5212 * diff(rtp)[-1]
                (1.73362) (2.59919)
                + 3.56675 * diff(rtp)[-2] - 5.48119 * diff(rtp)[-3] - 0.00968
                (0.74033) (1.18351) (0.29587)
    Sum Sq 17.1092 Std Err 0.3715 LHS Mean -0.0093
    R Sq 0.1055 R Bar Sq 0.0767 F 4,124 3.6575
    D.W.( 1) 2.5032 D.W.( 4) 2.1339
RF5054_P = RF5054_P.1 + (-9.7818 * DIFF(RTP) - 3.1242 * DIFF(RTP.1) - 14.0327 * DIFF(RTP.2) - 4.0364 * DIFF(RTP.3)**0.00/44.72
    RF5054
    Ordinary Least Squares
    QUARTERLY data for 129 periods from 1976Q4 to 2008Q4
    Date: 23 OCT 2009
    diff(rf5054)
        = - 9.7818 * diff(rtp) - - 3.12420 * * diff(rtp)[-1]
                - 14.0327 * diff(rtp)[-2] - 4.03638 * diff(rtp)[-3] - - 0.00195 (0.91457) (0.06247)
    Sum Sq 15.5373 Std Err 0.3540 LHS Mean -0.0021
    R Sq 0.1556 R Bar Sq 0.1283 F 4,124 5.7111
    D.W.( 1) 2.3211 D.W.( 4) 1.9891
```

RF5559_P = RF5559_P. $1+(-2.4665$ * DIFF(RTP) -4.8191 * DIFF(RTP.1) $-11.4418 * \operatorname{DIFF}(R T P .2)-3.5854 * \operatorname{DIFF}(R T P .3) * 50.00 / 44.72$ RF5559
Ordinary Least Squares
QUARTERLY data for 129 periods from 1976Q4 to 2008Q4
Date: 23 OCT 2009
diff(rf5559)
$=-2.46651 * \operatorname{diff}(r t p)-4.81906 * \operatorname{diff}(r t p)[-1]$
(0.39811) (0.73342)
- 11.4418 * diff(rtp)[-2] - 3.58538 * diff(rtp)[-3] - 0.00784
(1.74120) (0.56759) (0.17561)

| Sum Sq | 31.8294 | Std Err | 0.5066 | LHS Mean | -0.0088 |  |
| :--- | ---: | :--- | ---: | :--- | :--- | :--- |
| R Sq | 0.0474 | R Bar Sq | 0.0166 | F | 4,124 | 1.5416 |

RF6064_P = RF6064_P. $1+(-22.1139$ * DIFF(RTP) +4.5539 * DIFF(RTP.1) - 6.2406 * DIFF(RTP.2) -7.0337 * DIFF(RTP.3) $) * 50.00 / 44.72$ RF6064
Ordinary Least Squares
QUARTERLY data for 129 periods from 1976Q4 to 2008Q4
Date: 23 OCT 2009
diff(rf6064)


Sum Sq 53.3640 Std Err 0.6560 LHS Mean -0.0008

```
    R Sq 0.0732 R Bar Sq 0.0433 F 4,124 2.4497
    D.W.( 1) 3.0445 D.W.( 4) 1.3297
RF6569_P = RF6569_P.1 + (9.2541 * DIFF(RTP) + 7.6281 * DIFF(RTP.1) - 22.5230 * DIFF(RTP.2) + 0.2738 * DIFF(RTP.3))*50.00/44.72
    RF6569
    Ordinary Least Squares
    QUARTERLY data for 129 periods from 1976Q4 to 2008Q4
    Date: 23 OCT 2009
    diff(rf6569)
    = (0.25410 * diff(rtp) + (0.62811 (0.62521) * diff(rtp)[-1]
                - 22.5230 * diff(rtp)[-2] + 0.27380 * diff(rtp)[-3] - 0.01724
                (1.84584) (0.02334) (0.20808)
    Sum Sq 109.749 Std Err 0.9408 LHS Mean -0.0199
    R Sq 0.0329 R Bar Sq 0.0017 F 4,124 1.0547
    D.W.( 1) 3.0050 D.W.( 4) 1.0241
RF7074_P = RF7074_P.1 + (24.2237 * DIFF(RTP) + 8.3386 * DIFF(RTP.1) - 13.5317 * DIFF(RTP.2) - 8.1855 * DIFF(RTP.3))*50.00/44.72
    RF7074
    Ordinary Least Squares
    QUARTERLY data for 111 periods from 1981Q2 to 2008Q4
    Date: 23 OCT 2009
    diff(rf7074)
        = 24.2237 * diff(rtp) + + 8.33858 * diff(rtp)[-1]
            (1.28864) (0.42077)
                - (0.69102) * diff(rtp)[-2] - 8.18546 * (0.43349) diff(rtp)[-3] + + 0.01891 (0.16535)
    Sum Sq 153.213 Std Err 1.2023 LHS Mean 0.0075
    R Sq 0.0218 R Bar Sq -0.0151 F 4,106 0.5903
    D.W.( 1) 2.6506 D.W.( 4) 1.6645
RF75O_P = RF75O_P.1 + (-28.8294 * DIFF(RTP) + 55.5911 * DIFF(RTP.1) - 31.0676 * DIFF(RTP.2) - 15.8580 * DIFF(RTP.3)*50.00/44.72
    RF750
    Ordinary Least Squares
        QUARTERLY data for 111 periods from 1981Q2 to 2008Q4
        Date: 23 OCT 2009
    diff(rf75o)
        = - 28.8294 * diff(rtp) + 55.5911 * diff(rtp)[-1]
            (1.09028) (1.99419)
                - 31.0676 * diff(rtp)[-2] - 15.8580 * diff(rtp)[-3] + 0.03090
                (1.12786) (0.59703) (0.19206)
\begin{tabular}{lrlrll} 
Sum Sq & 303.162 & Std Err & 1.6912 & LHS Mean & 0.0324 \\
R Sq & 0.0508 & R Bar Sq & 0.0149 & F 4,106 & 1.4170
\end{tabular}
    1.4170
```


## UNEMPLOYMENT RATES, AGE-GENDER ADJUSTED, PRELIMINARY

```
RUM_ASA_P = (RM1617_P * LM1617_BY + RM1819_P * LM1819_BY + RM2024_P * LM2024_BY + RM2529_P * LM2529_BY +
    RM3034_P * LM3034_BY + RM3539_P * LM3539_BY + RM4044_P * LM4044_BY + RM4549_P * LM4549_BY +
    RM5054_P * LM5054_BY + RM5559_P * LM5559_BY + RM6064_P * LM6064_BY + RM6569_P * LM6569_BY +
    RM7074_P * LM7074_BY + RM75O_P * LM75O_BY)/ LCM_BY
```

```
RUF_ASA_P = (RF1617_P * LF1617_BY + RF1819_P * LF1819_BY + RF2024_P * LF2024_BY + RF2529_P * LF2529_BY +
    RF3034_P * LF3034_BY + RF3539_P * LF3539_BY + RF4044_P * LF4044_BY + RF4549_P * LF4549_BY +
    RF5054_P * LF5054_BY + RF5559_P * LF5559_BY + RF6064_P * LF6064_BY + RF6569_P * LF6569_BY +
    RF7074_P * LF7074_BY + RF75O_P * LF75O_BY)/ LCF_BY
RU_ASA_P = (RUM_ASA_P * LCM_BY + RUF_ASA_P * LCF_BY) / LC_BY
```


## UNEMPLOYMENT RATES

## MALES

```
RM1617= RM1617_P * (1 + RU_ASA_ADJ / RU_ASA_P)
RM1819= RM1819_P * (1 + RU_ASA_ADJ / RU_ASA_P)
RM2024= RM2024_P * (1 + RU_ASA_ADJ / RU_ASA_P)
RM2529= RM2529_P * (1 + RU_ASA_ADJ / RU_ASA_P)
RM3034= RM3034_P * (1 + RU_ASA_ADJ / RU_ASA_P)
RM3539= RM3539_P * (1 + RU_ASA_ADJ / RU_ASA_P)
RM4044= RM4044_P * (1 + RU_ASA_ADJ / RU_ASA_P)
RM4549= RM4549_P * (1 + RU_ASA_ADJ / RU_ASA_P)
RM5054= RM5054_P * (1 + RU_ASA_ADJ / RU_ASA_P)
RM5559= RM5559_P * (1 + RU_ASA_ADJ / RU_ASA_P)
RM6064= RM6064_P * (1 + RU_ASA_ADJ / RU_ASA_P)
RM6569= RM6569_P * (1 + RU_ASA_ADJ / RU_ASA_P)
RM7074= RM7074_P * (1 + RU_ASA_ADJ / RU_ASA_P)
RM75O = RM75O_P * (1 + RU_ASA_ADJ / RU_ASA_P)
```


## FEMALES

RF1617 $=$ RF1617_P * (1 + RU_ASA_ADJ / RU_ASA_P)
RF1819 $=$ RF1819_P * (1 + RU_ASA_ADJ / RU_ASA_P)
RF2024 $=$ RF2024_P * (1 + RU_ASA_ADJ / RU_ASA_P)
RF2529 $=$ RF2529_P * (1 + RU_ASA_ADJ / RU_ASA_P)
RF3034 $=$ RF3034_P * (1 + RU_ASA_ADJ / RU_ASA_P)
RF3539 $=$ RF3539_P * $(1+$ RU_ASA_ADJ / RU_ASA_P $)$
RF4044 $=$ RF4044_P * (1 + RU_ASA_ADJ / RU_ASA_P)
RF4549 $=$ RF4549_P * $(1+$ RU_ASA_ADJ / RU_ASA_P)
RF5054 $=$ RF5054_P * (1 + RU_ASA_ADJ / RU_ASA_P)
RF5559= RF5559_P * (1 + RU_ASA_ADJ / RU_ASA_P)
RF6064 $=$ RF6064_P * $\left(1+\mathrm{RU} \_A S A \_A D J / R U \_A S A \_P\right)$
RF6569 $=$ RF6569_P * $(1+$ RU_ASA_ADJ / RU_ASA_P $)$
RF7074 $=$ RF7074_P * (1 + RU_ASA_ADJ / RU_ASA_P)
RF75O = RF75O_P * (1 + RU_ASA_ADJ / RU_ASA_P)

## UNEMPLOYMENT RATES, AGE-GENDER ADJUSTED

```
RUM_ASA = (RM1617 * LM1617_BY+ RM1819 * LM1819_BY + RM2024 * LM2024_BY + RM2529 * LM2529_BY +
    RM3034 * LM3034_BY + RM3539 * LM3539_BY + RM4044 * LM4044_BY + RM4549 * LM4549_BY +
    RM5054 * LM5054_BY + RM5559 * LM5559_BY + RM6064 * LM6064_BY + RM6569 * LM6569_BY +
    RM7074 * LM7074_BY + RM75O * LM75O_BY)/ LCM_BY
RUF_ASA = (RF1617 * LF1617_BY + RF1819 * LF1819_BY + RF2024 * LF2024_BY + RF2529 * LF2529_BY +
    RF3034 * LF3034_BY + RF3539 * LF3539_BY + RF4044 * LF4044_BY + RF4549 * LF4549_BY +
    RF5054 * LF5054_BY + RF5559 * LF5559_BY + RF6064 * LF6064_BY + RF6569 * LF6569_BY +
    RF7074 * LF7074_BY + RF75O * LF75O_BY)/ LCF_BY
RU_ASA = (RUM_ASA * LCM_BY + RUF_ASA * LCF_BY) / LC_BY
```


## UNEMPLOYMENT RATES, FULL EMPLOYMENT DIFFERENTIALS

## MALES



## FEMALES

DRF1617_FE $=(-27.3243 *(1-\mathrm{RTP})+13.4173 *(1-\mathrm{RTP} .1)-50.4583 *(1-\mathrm{RTP} .2)-0.3678 *(1-\mathrm{RTP} .3)) * 50.00 / 44.72$ DRF1819_FE $=(-42.6358 *(1-$ RTP $)-13.6261 *(1-$ RTP.1 $)+9.5650 *(1-$ RTP. 2$)-31.4798 *(1-$ RTP. $)$ ) $) * 50.00 / 44.72$ DRF2024_FE $=(-16.9400 *(1-$ RTP $)-13.2669 *(1-R T P .1)-7.8323 *(1-$ RTP. $)-8.6887 *(1-$ RTP.3 $) * 50.00 / 44.72$ DRF2529_FE $=(-15.5798 *(1-$ RTP $)-11.9097 *(1-$ RTP. 1$)-9.8424 *(1-$ RTP. 2$)-2.7555 *(1-$ RTP.3 $)) * 50.00 / 44.72$ DRF3034_FE $=(-12.5396$ * $(1-$ RTP $)-1.6601 *(1-$ RTP.1 $)-21.0289 *(1-R T P .2)+0.0881 *(1-R T P .3)) * 50.00 / 44.72$ DRF3539_FE $=(-21.9314 *(1-$ RTP $)-3.0139 *(1-$ RTP.1 $)-7.8723 *(1-$ RTP.2 $)-6.4785 *(1-$ RTP.3 $) * 50.00 / 44.72$ DRF4044_FE $=(-7.7893$ * $(1-$ RTP $)-7.7152 *(1-R T P .1)-5.7849 *(1-R T P .2)-2.7298 *(1-R T P .3)) * 50.00 / 44.72$ DRF4549_FE $=(-7.8747 *(1-$ RTP $)-12.5212 *(1-$ RTP. 1$)+3.56675 *(1-$ RTP.2 $)-5.4812 *(1-$ RTP.3 $) * 50.00 / 44.72$ DRF5054_FE $=(-9.7818 *(1-$ RTP $)-3.1242 *(1-$ RTP.1 $)-14.0327 *(1-$ RTP. 2$)-4.0364 *(1-$ RTP.3 $)) * 50.00 / 44.72$ DRF5559_FE $=(-2.4665 *(1-$ RTP $)-4.8191 *(1-$ RTP.1 $)-11.4418 *(1-$ RTP. 2$)-3.5854 *(1-$ RTP.3 $)) * 50.00 / 44.72$ DRF6064_FE $=(-22.1139 *(1-\mathrm{RTP})+4.5539 *(1-\mathrm{RTP} .1)-6.2406$ * ( $1-\mathrm{RTP} .2)-7.0337$ * ( $1-\mathrm{RTP} .3$ ) $) * 50.00 / 44.72$ DRF6569_FE $=(9.2541 *(1-$ RTP $)+7.6281 *(1-$ RTP.1 $)-22.5230 *(1-$ RTP. 2$)+0.2738 *(1-$ RTP.3 $) * 50.00 / 44.72$ DRF7074_FE $=(24.2237 *(1-$ RTP $)+8.3386 *(1-$ RTP. 1$)-13.5317 *(1-$ RTP. 2$)-8.1855 *(1-$ RTP. 3$)) * 50.00 / 44.72$ DRF75O_FE $=(-28.8294 *(1-$ RTP $)+55.5911 *(1-R T P .1)-31.0676 *(1-$ RTP. 2$)-15.8580 *(1-$ RTP.3 $)) * 50.00 / 44.72$

UNEMPLOYMENT RATES, FULL EMPLOYMENT DIFFERENTIALS

TOTALS

MALES

RM1617_FE = RM1617 + DRM1617_FE RM1819_FE $=$ RM1819 + DRM1819_FE RM2024_FE = RM2024 + DRM2024_FE RM2529_FE $=$ RM2529 + DRM2529_FE RM3034_FE = RM3034 + DRM3034_FE RM3539_FE $=$ RM3539 + DRM3539_FE RM4044_FE = RM4044 + DRM4044_FE RM4549_FE = RM4549 + DRM4549_FE RM5054_FE $=$ RM5054 + DRM5054_FE RM5559_FE $=$ RM5559 + DRM5559_FE RM6064_FE = RM6064 + DRM6064_FE RM6569_FE $=$ RM6569 + DRM6569_FE RM7074_FE = RM7074 + DRM7074_FE RM75O_FE = RM75O + DRM75O_FE

RF1617_FE = RF1617 + DRF1617_FE RF1819_FE $=$ RF1819 + DRF1819_FE RF2024_FE $=$ RF2024 + DRF2024_FE RF2529_FE $=$ RF2529 + DRF2529_FE

```
RF3034_FE = RF3034 + DRF3034_FE
```

RF3539_FE $=$ RF3539 + DRF3539_FE
RF4044_FE $=$ RF4044 + DRF4044_FE
RF4549 FE $=$ RF4549 + DRF4549 FE
RF5054_FE $=$ RF5054 + DRF5054_FE
RF5559 FE $=$ RF5559 + DRF5559 FE
RF6064_FE $=$ RF6064 + DRF6064_FE
RF6569_FE $=$ RF6569 + DRF6569_FE
RF7074_FE $=$ RF7074 + DRF7074_FE
RF75O_FE = RF75O + DRF75O_FE

## LABOR FORCE PARTICIPATION RATES (LFPR)

## MALE LFPR EQUATIONS

AGE 16 TO 19

```
PM1617_P = 0.98298-0.000026
    -0.368046 * RM1617DI
    -(- 0.78720 + 0.01330 * TR_PM1617 + 0.00158 * RM1617 + 0.00180 * RM1617.1 + 0.00115 * RM1617.2 + 0.00014 * RM1617.3 -
    0.00072 * RM1617.4 - 0.00094 * RM1617.5);
PM1819 P = 0.97979-0.000070
    -0.614469 * RM1819DI
    - (- 0.50476 + 0.00764 * TR_PM1819 + 0.00626 * MOVAVG(5,RM1819));
PM1617 = PM1617 P;
PM1819 = PM1819_P;
```

AGE 20 TO 54
NEVER MARRIED
PM2024NM_P $=1.04005-0.00523-0.000143-0.00225 *$ TR_PM2024
-0.00063 * RM2024-0.00077 * RM2024.1-0.00059 * RM2024.2-0.00027 * RM2024.3 + $0.00005 *$ RM2024.4 + 0.00020 *
RM2024.5
- 0.788895 * RM2024DI;
PM2529NM_P = 0.97919-0.00809 + 0.000424-0.00070 * TR_PM2529
-0.00028 * RM2529-0.00044 * RM2529.1-0.00050 * RM2529.2-0.00047 * RM2529.3-0.00037 * RM2529.4-0.00021 *
RM2529.5
- 0.887239 * RM2529DI;
PM3034NM_P $=0.90427+0.000567$
- 0.00046 * RM3034-0.00061 * RM3034.1-0.00054 * RM3034.2-0.00036 * RM3034.3-0.00014 * RM3034.4 + 0.00001 *
RM3034.5
- 0.883076 * RM3034DI;
PM3539NM_P $=$ 0.86825-0.000105
-0.00004 * RM3539-0.00010 * RM3539.1-0.00016 * RM3539.2-0.00021 * RM3539.3-0.00021 * RM3539.4-0.00015 *
RM3539.5
- 0.846107 * RM3539DI;

PM4044NM_P $=$ 0.83977-0.001481
$-0.00057 *$ RM4044-0.00066 * RM4044.1-0.00044 * RM4044.2-0.00009 * RM4044.3 + 0.00022 * RM4044.4 + $0.00031 *$ RM4044.5
-0.807518 * RM4044DI;
PM4549NM_P $=0.81116-0.004005$

- 0.00002 * RM4549-0.00016 * RM4549.1-0.00034 * RM4549.2-0.00049 * RM4549.3-0.00054 * RM4549.4-0.00040 * RM4549.5
- 0.765856 * RM4549DI;

PM5054NM_P $=0.77540-0.010316$
+0.00112 * RM5054 + 0.00103 * RM5054.1 + 0.00023 * RM5054.2 - 0.00078 * RM5054.3-0.00149 * RM5054.4-0.00139 * RM5054.5

- 0.715161 * RM5054DI;

AGE 20 TO 54

MARRIED, SPOUSE PRESENT

```
PM2024MS_P = 1.03184-0.00733 + 0.000543-0.00069 * TR_PM2024
    - 0.00063 * RM2024 - 0.00077 * RM2024.1 - 0.00059 * RM2024.2 - 0.00027 * RM2024.3 + 0.00005 * RM2024.4 + 0.00020 *
    RM2024.5
    -0.937270 * RM2024DI;
PM2529MS_P = 1.00110-0.00498 + 0.001239-0.00025 * TR_PM2529
    -0.00028 * RM2529-0.00044 * RM2529.1-0.00050 * RM2529.2 - 0.00047 * RM2529.3 - 0.00037 * RM2529.4 - 0.00021 *
    RM2529.5
    -0.956095 * RM2529DI;
PM3034MS_P = 0.97120 + 0.001964 + 0.16457 * 1/(TR_PM3034-85)
    -0.00046 * RM3034-0.00061 * RM3034.1-0.00054 * RM3034.2-0.00036 * RM3034.3-0.00014 * RM3034.4 + 0.00001 *
    RM3034.5
    -0.961197 * RM3034DI;
PM3539MS_P = 0.98068 + 0.002566
    - 0.00004 * RM3539-0.00010 * RM3539.1-0.00016 * RM3539.2 - 0.00021 * RM3539.3 - 0.00021 * RM3539.4 - 0.00015 *
    RM3539.5
    -0.958567 * RM3539DI;
PM4044MS_P = 0.98250 + 0.003434
    -0.00057 * RM4044-0.00066 * RM4044.1-0.00044 * RM4044.2-0.00009 * RM4044.3 + 0.00022 * RM4044.4 + 0.00031 *
    RM4044.5
    - 0.950227 * RM4044DI;
PM4549MS_P = 0.98115 + 0.004829
    -0.00002 * RM4549-0.00016 * RM4549.1-0.00034 * RM4549.2-0.00049 * RM4549.3-0.00054 * RM4549.4-0.00040 *
    RM4549.5
    -0.935691 * RM4549DI;
PM5054MS_P = 0.94484 + 0.004732
    + 0.00112 * RM5054 + 0.00103 * RM5054.1 + 0.00023 * RM5054.2 - 0.00078 * RM5054.3 - 0.00149 * RM5054.4 - 0.00139 *
    RM5054.5
    + 0.09796 * (RF5054CU6+RF5054C6O) - 0.901430 * RM5054DI;
```


## AGE 20 TO 54

MARRIED, SPOUSE ABSENT

PM2024MA_P $=1.14087-0.01412+0.000270-0.00232 *$ TR_PM2024
-0.00063 * RM2024-0.00077 * RM2024.1-0.00059 * RM2024.2-0.00027 * RM2024.3 + 0.00005 * RM2024.4 + 0.00020 * RM2024.5

- 0.878382 * RM2024DI;

PM2529MA_P $=$ 0.98602-0.00788 $+0.000733-0.00051 *$ TR_PM2529
-0.00028 * RM2529-0.00044 * RM2529.1-0.00050 * RM2529.2-0.00047 * RM2529.3-0.00037 * RM2529.4-0.00021 * RM2529.5

- 0.913332 * RM2529DI;

PM3034MA_P $=0.93933+0.001197$
-0.00046 * RM3034-0.00061 * RM3034.1-0.00054 * RM3034.2-0.00036 * RM3034.3-0.00014 * RM3034.4 + 0.00001 * RM3034.5

- 0.918267 * RM3034DI;

```
PM3539MA_P = 0.92354 + 0.001209
    -0.00004 * RM3539-0.00010 * RM3539.1-0.00016 * RM3539.2-0.00021 * RM3539.3-0.00021 * RM3539.4-0.00015 *
    RM3539.5
    -0.901443 * RM3539DI;
PM4044MA_P = 0.91512 + 0.001111
    - 0.00057 * RM4044-0.00066 * RM4044.1-0.00044 * RM4044.2 - 0.00009 * RM4044.3 + 0.00022 * RM4044.4 + 0.00031 *
        RM4044.5
        - 0.882777 * RM4044DI;
PM4549MA_P = 0.89473 + 0.000334
        -0.00002 * RM4549-0.00016 * RM4549.1-0.00034 * RM4549.2-0.00049 * RM4549.3-0.00054 * RM4549.4-0.00040 *
        RM4549.5
        -0.849283 * RM4549DI;
PM5054MA_P = 0.84912-0.004345
    + 0.00112 * RM5054 + 0.00103 * RM5054.1 + 0.00023 * RM5054.2-0.00078 * RM5054.3 - 0.00149 * RM5054.4 - 0.00139 *
    RM5054.5
    -0.789075 * RM5054DI;
```

AGE 20 TO 54

```
PM2024_P = (PM2024NM_P * NM2024NM + PM2024MS_P * NM2024MS + PM2024MA_P * NM2024MA) / NM2024;
PM2529_P = (PM2529NM_P * NM2529NM + PM2529MS_P * NM2529MS + PM2529MA_P * NM2529MA) / NM2529;
PM3034_P = (PM3034NM_P * NM3034NM + PM3034MS_P * NM3034MS + PM3034MA_P * NM3034MA) / NM3034;
PM3539_P = (PM3539NM_P * NM3539NM + PM3539MS_P * NM3539MS + PM3539MA_P * NM3539MA) / NM3539;
PM4044_P = (PM4044NM_P * NM4044NM + PM4044MS_P * NM4044MS + PM4044MA_P * NM4044MA) / NM4044;
PM4549_P = (PM4549NM_P * NM4549NM + PM4549MS_P * NM4549MS + PM4549MA_P * NM4549MA) / NM4549;
PM5054_P = (PM5054NM_P * NM5054NM + PM5054MS_P * NM5054MS + PM5054MA_P * NM5054MA) / NM5054;
PM2024 = PM2024_P;
PM2529 = PM2529_P;
PM3034 = PM3034_P;
PM3539 = PM3539_P;
PM4044 = PM4044_P;
PM4549 = PM4549_P;
PM5054 = PM5054_P;
PM2024NM = PM2024NM_P * PM2024 / PM2024_P;
PM2529NM = PM2529NM_P * PM2529 / PM2529_P;
PM3034NM = PM3034NM_P * PM3034 / PM3034_P;
PM3539NM = PM3539NM_P * PM3539 / PM3539_P;
PM4044NM = PM4044NM_P * PM4044 / PM4044_P;
PM4549NM = PM4549NM_P * PM4549 / PM4549_P;
PM5054NM = PM5054NM_P * PM5054 / PM5054_P;
PM2024MS = PM2024MS_P * PM2024 / PM2024_P;
PM2529MS = PM2529MS_P * PM2529 / PM2529_P;
PM3034MS = PM3034MS_P * PM3034 / PM3034_P;
PM3539MS = PM3539MS_P * PM3539 / PM3539_P;
PM4044MS = PM4044MS_P *PM4044 / PM4044_P;
PM4549MS = PM4549MS_P * PM4549 / PM4549_P;
PM5054MS = PM5054MS_P * PM5054 / PM5054_P;
PM2024MA = PM2024MA_P * PM2024 / PM2024_P;
PM2529MA = PM2529MA_P *PM2529 / PM2529_P;
PM3034MA = PM3034MA_P * PM3034/PM3034_P;
PM3539MA = PM3539MA_P * PM3539 / PM3539_P;
PM4044MA = PM4044MA_P * PM4044 / PM4044_P;
PM4549MA = PM4549MA_P * PM4549 / PM4549_P;
PM5054MA = PM5054MA_P * PM5054/PM5054_P;
```

```
PM55_P = - 0.76902 + 0.000142-0.818694 * RM55DI + PM55E_DE + PM55_DM
    + 0.00062 * RM5559 + 0.00041 * RM5559.1-0.00026 * RM5559.2-0.00101 * RM5559.3 - 0.00147 * RM5559.4 - 0.00126 *
    RM5559.5;
PM56_P = - 0.76098-0.002143-0.798138*RM56DI + PM56E_DE + PM56_DM
    + 0.00062 * RM5559 + 0.00041 * RM5559.1-0.00026 * RM5559.2-0.00101 * RM5559.3-0.00147 * RM5559.4 - 0.00126 *
    RM5559.5;
PM57_P = - 0.71065-0.004979-0.776220 * RM57DI + PM57E_DE + PM57_DM
    + 0.00062 * RM5559 + 0.00041 * RM5559.1-0.00026 * RM5559.2-0.00101 * RM5559.3-0.00147 * RM5559.4-0.00126 *
    RM5559.5;
PM58_P = - 0.69412-0.007970-0.755977 * RM58DI + PM58E_DE + PM58_DM
    + 0.00062 * RM5559 + 0.00041 * RM5559.1-0.00026 * RM5559.2-0.00101 * RM5559.3-0.00147 * RM5559.4-0.00126 *
    RM5559.5;
PM59_P = - 0.60153-0.012123-0.730720 * RM59DI + PM59E_DE + PM59_DM
    + 0.00062 * RM5559 + 0.00041 * RM5559.1-0.00026 * RM5559.2-0.00101 * RM5559.3-0.00147 * RM5559.4-0.00126 *
    RM5559.5;
PM60_P = - 0.58858-0.011072 + 0.000362-0.682175 * RM60DI + PM60E_DE + PM60_DM
    + 0.00203 * RM6064 + 0.00160 * RM6064.1-0.00021 * RM6064.2 - 0.00235 * RM6064.3-0.00374 * RM6064.4-0.00331 *
    RM6064.5
    + 0.01566 * PF58;
PM61_P = - 0.55646-0.019672 + 0.002517-0.647667 * RM61DI + PM61E_DE + PM61_DM
    +0.00203 * RM6064 + 0.00160 * RM6064.1-0.00021 * RM6064.2-0.00235 * RM6064.3 - 0.00374 * RM6064.4 - 0.00331 *
    RM6064.5
        + 0.08544 * PF59;
```

AGE 62 TO 74

PM62_P $=0.26329 *$ PF60 - 0.29161-0.544582 * RM61DI. $4+$ PM62E_DE + PM62_DM - $0.056836+0.008814$
+0.00203 * RM6064 + 0.00160 * RM6064.1-0.00021 * RM6064.2-0.00235 * RM6064.3-0.00374 * RM6064.4-0.00331 * RM6064.5

- 0.60 * RRADJ_M62 - 0.02 * POT_ET_TXRT_62;

PM63_P $=0.40940 *$ PF61-0.29495-0.480592 * RM61DI. $8+$ PM63E_DE + PM63_DM - $0.085295+0.017121$
+0.00203 * RM6064 + 0.00160 * RM6064.1-0.00021 * RM6064.2-0.00235 * RM6064.3-0.00374 * RM6064.4-0.00331 * RM6064.5

- 0.55 * RRADJ_M63-0.02 * POT_ET_TXRT_63;

PM64_P $=0.47933 *$ PF62-0.22665-0.435189 * RM61DI. 12 + PM64E_DE + PM64_DM - $0.113466+0.030849$
+0.00203 * RM6064 + 0.00160 * RM6064.1-0.00021 * RM6064.2-0.00235 * RM6064.3-0.00374 * RM6064.4-0.00331 * RM6064.5

- 0.50 * RRADJ_M64-0.02 * POT_ET_TXRT_64;

PM65_P $=0.72722 *$ PF63-0.35819-0.368105 * RM61DI. $16+$ PM65E_DE + PM65_DM - $0.016932+0.059764$
+0.00067 * RM6569 + 0.00040 * RM6569.1-0.00040 * RM6569.2-0.00127 * RM6569.3-0.00178 * RM6569.4-0.00151 * RM6569.5

- 0.45 * RRADJ_M65-0.02 * POT_ET_TXRT_65;

PM66_P $=0.38684 *$ PF64-0.20883-0.330352 * RM61DI. $20+$ PM66E_DE + PM66_DM + 0.000194 + 0.036539
+0.00067 * RM6569 + 0.00040 * RM6569.1-0.00040 * RM6569.2-0.00127 * RM6569.3-0.00178 * RM6569.4-0.00151 * RM6569.5

- 0.40 * RRADJ_M66-0.02 * POT_ET_TXRT_66;

PM67_P $=0.35012 *$ PF65-0.15975-0.294277 * RM61DI. $24+$ PM67E_DE + PM67_DM + 0.042362-0.012475
+0.00067 * RM6569 + 0.00040 * RM6569.1-0.00040 * RM6569.2-0.00127 * RM6569.3-0.00178 * RM6569.4-0.00151 * RM6569.5

- 0.35 * RRADJ_M67-0.02 * POT_ET_TXRT_67;

PM68_P $=0.95984 *$ PF66-0.26305-0.269452 * RM61DI. 28 + PM68E_DE + PM68_DM + 0.057835-0.027492
+0.00067 * RM6569 + 0.00040 * RM6569.1-0.00040 * RM6569.2-0.00127 * RM6569.3-0.00178 * RM6569.4-0.00151 * RM6569.5

- 0.30 * RRADJ_M68-0.02 * POT_ET_TXRT_68;

```
PM69_P = 0.74113 * PF67-0.22589-0.246975 * RM61DI.32 + PM69E_DE + PM69_DM + 0.047260-0.018867
    + 0.00067 * RM6569 + 0.00040 * RM6569.1-0.00040 * RM6569.2-0.00127 * RM6569.3-0.00178 * RM6569.4 - 0.00151 *
    RM6569.5
    - 0.30 * RRADJ_M69 - 0.02 * POT_ET_TXRT_69;
PM70_P = 0.46445 * PF68-0.23451-0.220464 * RM61DI.36 + PM70E_DE + PM70_DM + 0.037323-0.010528
    -0.00013 * RM7074-0.00016 * RM7074.1-0.00013 * RM7074.2 - 0.00006 * RM7074.3 + 0.00000 * RM7074.4 + 0.00003 *
    RM7074.5;
PM71_P = 0.27684 * PF69-0.20679-0.202537 * RM61DI. 40 + PM71E_DE + PM71_DM + 0.030867-0.005653
    -0.00013 * RM7074-0.00016 * RM7074.1-0.00013 * RM7074.2 - 0.00006 * RM7074.3 + 0.00000 * RM7074.4 + 0.00003 *
    RM7074.5;
PM72_P = 0.77240 * PF70-0.25289-0.186071 * RM61DI.44 + PM72E_DE + PM72_DM + 0.040290-0.012330
    -0.00013 * RM7074-0.00016 * RM7074.1-0.00013 * RM7074.2 - 0.00006 * RM7074.3 + 0.00000 * RM7074.4 + 0.00003 *
    RM7074.5;
PM73_P = 0.65971 * PF71-0.19394-0.167545 * RM61DI.48 + PM73E_DE + PM73_DM + 0.036021-0.009631
    -0.00013 * RM7074-0.00016 * RM7074.1-0.00013 * RM7074.2-0.00006 * RM7074.3 + 0.00000 * RM7074.4 + 0.00003 *
    RM7074.5;
PM74_P = 0.78464 * PF72-0.17649-0.147134 * RM61DI.52 + PM74E_DE + PM74_DM + 0.033451-0.010839
    -0.00013 * RM7074-0.00016 * RM7074.1 - 0.00013 * RM7074.2-0.00006 * RM7074.3 + 0.00000 * RM7074.4 + 0.00003 *
    RM7074.5;
```

AGE 75 TO 79

PM75_P $=$ PM74.4 * $0.920+$ DPM75O_FE;
PM76_P $=$ PM75.4 * 0.920 + DPM75O_FE;
PM77_P $=$ PM76.4 * $0.920+$ DPM75O_FE;
PM78_P $=$ PM77.4 * 0.920 + DPM75O_FE;
PM79_P $=$ PM78.4 * 0.920 + DPM75O_FE;

AGE 55 TO 79

PM55 = PM55_P;
PM56 = PM56_P;
PM57 = PM57_P;
PM58 = PM58_P;
PM59 = PM59_P;
PM60 = PM60_P;
PM61 = PM61_P;
PM62 = PM62_P;
PM63 = PM63_P;
PM64 = PM64_P;
PM65 = PM65_P;
PM66 = PM66_P;
PM67 = PM67_P;
PM68 = PM68_P;
PM69 = PM69_P;
PM70 = PM70_P;
PM71 = PM71_P;
PM72 = PM72_P;
PM73 = PM73_P;
PM74 = PM74_P;
PM75 = PM75_P;
PM76 = PM76_P;
PM77 = PM77_P;
PM78 = PM78_P;
PM79 = PM79_P;

```
PM80_P = PM79.4 *0.965**(1) + DPM75O_FE;
PM81_P = PM79.8 *0.965**( 2) + DPM75O_FE;
PM82_P = PM79.12 *0.965**(3) + DPM75O_FE;
PM83_P = PM79.16 *0.965**(4) + DPM75O_FE;
PM84_P = PM79.20 *0.965**(5) + DPM75O_FE;
PM85_P = MOVAVG(8,PM79.24) *0.965**( 6) + DPM75O_FE;
PM86_P = MOVAVG(8,PM79.28) * 0.965**(7) + DPM75O_FE;
PM87_P = MOVAVG(8,PM79.32) * 0.965**(8) + DPM75O_FE;
PM88_P = MOVAVG(8,PM79.36) * 0.965**(9) + DPM75O_FE;
PM89_P = MOVAVG(8,PM79.40) * 0.965**(10) + DPM75O_FE;
PM90_P = MOVAVG(8,PM79.44) * 0.965**(11) + DPM75O_FE;
PM91_P = MOVAVG(8,PM79.48) * 0.965**(12) + DPM75O_FE;
PM92_P = MOVAVG(8,PM79.52) *0.965**(13) + DPM75O_FE;
PM93_P = MOVAVG(8,PM79.56) * 0.965**(14) + DPM75O_FE;
PM94_P = MOVAVG(8,PM79.60) * 0.965**(15) + DPM75O_FE;
PM95_P = PM94_P * 0.965 + DPM75O_FE;
PM96_P = PM95_P * 0.965 + DPM75O_FE;
PM97_P = PM96_P * 0.965 + DPM75O_FE;
PM98_P = PM97_P * 0.965 + DPM75O_FE;
PM99_P = PM98_P * 0.965 + DPM75O_FE;
PM100_P = PM99_P * 0.965 + DPM75O_FE;
PM80O_P = (PM80_P*NM80 + PM81_P*NM81 + PM82_P*NM82 + PM83_P*NM83 + PM84_P*NM84 + PM85_P*NM85 +
        PM86_P*NM86 + PM87_P*NM87 + PM88_P*NM88 + PM89_P*NM89 +
    PM90_P*NM90 + PM91_P*NM91 + PM92_P*NM92 + PM93_P*NM93 + PM94_P*NM94 + PM95_P*NM95 + PM96_P*NM96 +
        PM97_P*NM97 + PM98_P*NM98 + PM99_P*NM99 +
    PM100_P*NM100 ) / NM80O;
PM80O = PM80O_P;
PM80 = PM80_P * PM80O / PM800_P;
PM81 = PM81_P * PM80O / PM800_P;
PM82 = PM82_P * PM80O/PM800_P;
PM83 = PM83_P * PM80O / PM800_P;
PM84 = PM84_P * PM80O / PM80O_P;
PM85 = PM85_P * PM80O / PM800_P;
PM86 = PM86_P * PM80O / PM80O_P;
PM87 = PM87_P * PM80O / PM800_P;
PM88 = PM88_P * PM80O / PM800_P;
PM89 = PM89_P * PM80O / PM800_P;
PM90 = PM90_P * PM80O / PM800_P;
PM91 = PM91_P * PM80O / PM800_P;
PM92 = PM92_P * PM80O / PM800_P;
PM93 = PM93_P * PM80O / PM800_P;
PM94 = PM94_P * PM80O / PM800_P;
PM95 = PM95_P * PM80O / PM800_P;
PM96 = PM96_P * PM80O / PM800_P;
PM97 = PM97_P * PM80O / PM800_P;
PM98 = PM98_P * PM80O / PM800_P;
PM99 = PM99_P * PM80O / PM80O_P;
PM100 = PM100_P * PM80O / PM80O_P;
```


## FEMALE LFPR EQUATIONS

AGE 16 TO 19

PF1617_P $=0.98681-0.000047-(0.377204 *$ RF1617DI $)$
$-(-0.00741+0.23393$ * RF1617CU6 + 0.00051 * MOVAVG(5,RF1617))
$-(-0.69608+0.01166 *$ TR_PF1617 $+0.00616 * \operatorname{MOVAVG}(5, R F 1617))$;
PF1819_P $=0.98200-0.000069$

```
    - (0.584930 * RF1819DI)
    -(-0.00080 + 0.22814 * RF1819CU6 + 0.00318 * MOVAVG(5,RF1819))
    -(-0.53433+0.00764 * TR_PF1819 + 0.00667 * MOVAVG(5,RF1819));
PF1617 = PF1617_P;
PF1819 = PF1819_P;
AGE 20 TO 44
FEMALES - NEVER MARRIED WITH AT LEAST 1 OWN CHILD UNDER AGE 6
```

```
PF2024NMC6U_P = ( 0.70868-0.000477
```

PF2024NMC6U_P = ( 0.70868-0.000477
-0.00087 * RF2024-0.00099 * RF2024.1-0.00063 * RF2024.2 - 0.00007 * RF2024.3 + 0.00041 * RF2024.4 + 0.00052 *
-0.00087 * RF2024-0.00099 * RF2024.1-0.00063 * RF2024.2 - 0.00007 * RF2024.3 + 0.00041 * RF2024.4 + 0.00052 *
RF2024.5
RF2024.5
-0.661872 * RF2024DI)
-0.661872 * RF2024DI)
* 1.0160;
* 1.0160;
PF2529NMC6U_P = ( 0.74861-0.000966
PF2529NMC6U_P = ( 0.74861-0.000966
-0.00056 * RF2529-0.00070 * RF2529.1-0.00057 * RF2529.2-0.00029 * RF2529.3-0.00002 * RF2529.4 + 0.00013 *
-0.00056 * RF2529-0.00070 * RF2529.1-0.00057 * RF2529.2-0.00029 * RF2529.3-0.00002 * RF2529.4 + 0.00013 *
RF2529.5
RF2529.5
-0.700572 * RF2529DI)
-0.700572 * RF2529DI)
* 0.9981;
* 0.9981;
PF3034NMC6U_P = ( 0.73944-0.001998
PF3034NMC6U_P = ( 0.73944-0.001998
-0.00081 * RF3034-0.00065 * RF3034.1 + 0.00005 * RF3034.2 + 0.00089 * RF3034.3 + 0.00143 * RF3034.4 + 0.00128 *
-0.00081 * RF3034-0.00065 * RF3034.1 + 0.00005 * RF3034.2 + 0.00089 * RF3034.3 + 0.00143 * RF3034.4 + 0.00128 *
RF3034.5
RF3034.5
-0.690741 * RF3034DI)
-0.690741 * RF3034DI)
* 0.9980;
* 0.9980;
PF3539NMC6U_P = ( 0.75363-0.003230
PF3539NMC6U_P = ( 0.75363-0.003230
-0.00195 * RF3539-0.00216 * RF3539.1-0.00128 * RF3539.2 + 0.00002 * RF3539.3 + 0.00111 * RF3539.4 + 0.00132 *
-0.00195 * RF3539-0.00216 * RF3539.1-0.00128 * RF3539.2 + 0.00002 * RF3539.3 + 0.00111 * RF3539.4 + 0.00132 *
RF3539.5
RF3539.5
-0.701148 * RF3539DI)
-0.701148 * RF3539DI)
* 0.9989;
* 0.9989;
PF4044NMC6U_P = ( 0.73920-0.005589
PF4044NMC6U_P = ( 0.73920-0.005589
-0.00026 * RF4044-0.00050 * RF4044.1-0.00068 * RF4044.2-0.00076 * RF4044.3-0.00070 * RF4044.4-0.00046 *
-0.00026 * RF4044-0.00050 * RF4044.1-0.00068 * RF4044.2-0.00076 * RF4044.3-0.00070 * RF4044.4-0.00046 *
RF4044.5
RF4044.5
-0.677955 * RF4044DI)
-0.677955 * RF4044DI)
* 0.9989;

```
            * 0.9989;
```


## FEMALES - NEVER MARRIED NO OWN CHILDREN UNDER AGE 6

```
PF2024NMNC6_P = (1.13654-0.000206-0.00366 * TR_PF2024NMNC6
```

            -0.00087 * RF2024-0.00099 * RF2024.1-0.00063 * RF2024.2-0.00007 * RF2024.3 + 0.00041 * RF2024.4 + 0.00052 *
        RF2024.5
            - 0.742995 * RF2024DI)
            * 1.0160;
    PF2529NMNC6_P $=(0.98148+0.000457-0.00111 *$ TR_PF2529NMNC6
-0.00056 * RF2529-0.00070 * RF2529.1-0.00057 * RF2529.2-0.00029 * RF2529.3-0.00002 * RF2529.4 + 0.00013 *
RF2529.5
- 0.853721 * RF2529DI)
*0.9981;
PF3034NMNC6_P $=(0.84901+0.000598$
-0.00081 * RF3034-0.00065 * RF3034.1 + 0.00005 * RF3034.2 + $0.00089 *$ RF3034.3 + 0.00143 * RF3034.4 + 0.00128 *
RF3034.5
- 0.838634 * RF3034DI)
* 0.9980;
PF3539NMNC6_P $=(0.84953+0.000015$
$-0.00195 *$ RF3539-0.00216 * RF3539.1-0.00128 * RF3539.2 + 0.00002 * RF3539.3 + $0.00111 *$ RF3539.4 + 0.00132 *
RF3539.5

```
    -0.822121 * RF3539DI)
    * 0.9989;
PF4044NMNC6_P = (0.83790-0.000878
    -0.00026 * RF4044-0.00050 * RF4044.1-0.00068 * RF4044.2-0.00076 * RF4044.3-0.00070 * RF4044.4-0.00046 *
    RF4044.5
    -0.799021 * RF4044DI)
    * 0.9989;
```

FEMALES - MARRIED SPOUSE PRESENT WITH AT LEAST 1 OWN CHILD UNDER AGE 6

```
PF2024MSC6U_P = ( 0.69043-0.000935 + 0.000015
    -0.00087 * RF2024-0.00099 * RF2024.1-0.00063 * RF2024.2 - 0.00007 * RF2024.3 + 0.00041 * RF2024.4 + 0.00052 *
        RF2024.5
            -0.524765 * RF2024DI - 0.1*IF2024MSC6U)
            * 1.0160;
PF2529MSC6U_P = ( 0.76218-0.001821-0.000004
            -0.00056 * RF2529-0.00070 * RF2529.1-0.00057 * RF2529.2-0.00029 * RF2529.3-0.00002 * RF2529.4 + 0.00013 *
        RF2529.5
        -0.608428*RF2529DI - 0.1*IF2529MSC6U)
        * 0.9981;
PF3034MSC6U_P = ( 0.78186-0.002891-0.000006
        -0.00081 * RF3034-0.00065 * RF3034.1 + 0.00005 * RF3034.2 + 0.00089 * RF3034.3 + 0.00143 * RF3034.4 + 0.00128 *
        RF3034.5
            -0.639819 * RF3034DI - 0.1*IF3034MSC6U)
            * 0.9980;
PF3539MSC6U_P = ( 0.79072-0.004987-0.000005
        -0.00195 * RF3539-0.00216 * RF3539.1-0.00128 * RF3539.2 + 0.00002 * RF3539.3 + 0.00111 * RF3539.4 + 0.00132 *
        RF3539.5
            -0.635646 * RF3539DI - 0.1*IF3539MSC6U)
            * 0.9989;
PF4044MSC6U_P = ( 0.79356-0.007041-0.000008
    -0.00026 * RF4044-0.00050 * RF4044.1-0.00068 * RF4044.2 - 0.00076 * RF4044.3-0.00070 * RF4044.4 - 0.00046 *
        RF4044.5
    - 0.640644 * RF4044DI - 0.1*IF4044MSC6U)
    * 0.9989;
```

FEMALES - MARRIED SPOUSE PRESENT NO OWN CHILDREN UNDER AGE 6

```
PF2024MSNC6_P = (0.79421-0.000075
        -0.00087 * RF2024-0.00099 * RF2024.1 - 0.00063 * RF2024.2 - 0.00007 * RF2024.3 + 0.00041 * RF2024.4 + 0.00052 *
        RF2024.5
            -0.781956 * RF2024DI)
            * 1.0160;
PF2529MSNC6_P = ( 0.83502 + 0.000186
        -0.00056 * RF2529-0.00070 * RF2529.1-0.00057 * RF2529.2 - 0.00029 * RF2529.3 - 0.00002 * RF2529.4 + 0.00013 *
        RF2529.5
            -0.824520 * RF2529DI)
            * 0.9981;
PF3034MSNC6_P = ( 0.80379-0.000056
        -0.00081 * RF3034-0.00065 * RF3034.1 + 0.00005 * RF3034.2 + 0.00089 * RF3034.3 + 0.00143 * RF3034.4 + 0.00128 *
        RF3034.5
            -0.801357 * RF3034DI)
            * 0.9980;
PF3539MSNC6_P = ( 0.80906-0.000984
        - 0.00195 * RF3539-0.00216 * RF3539.1 - 0.00128 * RF3539.2 + 0.00002 * RF3539.3 + 0.00111 * RF3539.4 + 0.00132 *
        RF3539.5
```

```
    -0.784868 * RF3539DI)
    * 0.9989;
PF4044MSNC6_P = ( 0.82602-0.001249
    -0.00026 * RF4044-0.00050 * RF4044.1-0.00068 * RF4044.2-0.00076 * RF4044.3-0.00070 * RF4044.4-0.00046 *
    RF4044.5
    - 0.789470 * RF4044DI)
    * 0.9989;
```

FEMALES - MARRIED SPOUSE ABSENT WITH AT LEAST 1 OWN CHILD UNDER AGE 6

```
PF2024MAC6U_P = ( 0.95787-0.000403 + 0.000007
    0.00087 * RF2024-0.00099 * RF2024.1 - 0.00063 * RF2024.2 - 0.00007 * RF2024.3 + 0.00041 * RF2024.4 + 0.00052 *
        RF2024.5
    -0.683999 * RF2024DI - 0.16722 * IF2024MAC6U)
    * 1.0160;
PF2529MAC6U_P = ( 0.90653-0.000676
    -0.00056 * RF2529-0.00070 * RF2529.1-0.00057 * RF2529.2 - 0.00029 * RF2529.3-0.00002 * RF2529.4 + 0.00013 *
        RF2529.5
    - 0.731798 * RF2529DI - 0.10000 * IF2529MAC6U)
    * 0.9981;
PF3034MAC6U_P = ( 0.88071-0.001042
    -0.00081 * RF3034-0.00065 * RF3034.1 + 0.00005 * RF3034.2 + 0.00089 * RF3034.3 + 0.00143 * RF3034.4 + 0.00128 *
    RF3034.5
    -0.745172 * RF3034DI - 0.10000 * IF3034MAC6U)
    * 0.9980;
PF3539MAC6U_P = ( 0.90258-0.002058
    -0.00195 * RF3539-0.00216 * RF3539.1-0.00128 * RF3539.2 + 0.00002 * RF3539.3 + 0.00111 * RF3539.4 + 0.00132 *
        RF3539.5
            - 0.744846 * RF3539DI - 0.10000 * IF3539MAC6U)
            * 0.9989;
PF4044MAC6U_P = 0.89876-0.003975-0.000005
    - 0.00026 * RF4044-0.00050 * RF4044.1-0.00068 * RF4044.2 - 0.00076 * RF4044.3-0.00070 * RF4044.4 - 0.00046 *
    RF4044.5
    - 0.719419 * RF4044DI - 0.10000 * IF4044MAC6U)
    * 0.9989;
```

FEMALES - MARRIED SPOUSE ABSENT NO OWN CHILDREN UNDER AGE 6

PF2024MANC6_P $=(0.75174-0.000258+0.000005$
-0.00087 * RF2024-0.00099 * RF2024.1-0.00063 * RF2024.2-0.00007 * RF2024.3 + 0.00041 * RF2024.4 + 0.00052 * RF2024.5
-0.727427 * RF2024DI)

* 1.0160;

PF2529MANC6_P $=($ 0.82060 - 0.000015
-0.00056 * RF2529-0.00070 * RF2529.1-0.00057 * RF2529.2-0.00029 * RF2529.3-0.00002 * RF2529.4 + 0.00013 * RF2529.5

- 0.802962 * RF2529DI)
* 0.9981;

PF3034MANC6_P $=(0.83806+0.000442$
-0.00081 * RF3034-0.00065 * RF3034.1 + $0.00005 *$ RF3034.2 $+0.00089 *$ RF3034.3 + 0.00143 * RF3034.4 + 0.00128 * RF3034.5

- 0.829717 * RF3034DI)
* 0.9980;

PF3539MANC6_P $=(0.86613+0.000458$
-0.00195 * RF3539-0.00216 * RF3539.1-0.00128 * RF3539.2 + 0.00002 * RF3539.3 + 0.00111 * RF3539.4 + 0.00132 * RF3539.5

```
-0.838637 * RF3539DI)
* 0.9989;
PF4044MANC6_P = ( 0.85937+0.000112
    -0.00026 * RF4044-0.00050 * RF4044.1-0.00068 * RF4044.2-0.00076 * RF4044.3-0.00070 * RF4044.4-0.00046 *
    RF4044.5
    - 0.824438 * RF4044DI)
    * 0.9989;
```

PF2024NM_P = (PF2024NMC6U_P * NF2024NMC6U + PF2024NMNC6_P * NF2024NMNC6) / NF2024NM;
PF2024MS_P = (PF2024MSC6U_P * NF2024MSC6U + PF2024MSNC6_P * NF2024MSNC6) / NF2024MS;
PF2024MA_P = (PF2024MAC6U_P * NF2024MAC6U + PF2024MANC6_P * NF2024MANC6) / NF2024MA;
PF2529NM_P = (PF2529NMC6U_P * NF2529NMC6U + PF2529NMNC6_P * NF2529NMNC6) / NF2529NM;
PF2529MS_P = (PF2529MSC6U_P * NF2529MSC6U + PF2529MSNC6_P * NF2529MSNC6) / NF2529MS;
PF2529MA_P $=\left(\mathrm{PF} 2529 \mathrm{MAC6U} \mathrm{\_P} *\right.$ NF2529MAC6U + PF2529MANC6_P $*$ NF2529MANC6 $) /$ NF2529MA;
PF3034NM_P = (PF3034NMC6U_P * NF3034NMC6U + PF3034NMNC6_P * NF3034NMNC6) / NF3034NM;
PF3034MS_P = (PF3034MSC6U_P * NF3034MSC6U + PF3034MSNC6_P * NF3034MSNC6) / NF3034MS;
PF3034MA_P $=($ PF3034MAC6U_P $*$ NF3034MAC6U + PF3034MANC6_P $*$ NF3034MANC6) $/$ NF3034MA;
PF3539NM_P $=($ PF3539NMC6U_P $*$ NF3539NMC6U + PF3539NMNC6_P $*$ NF3539NMNC6) / NF3539NM;
PF3539MS_P = (PF3539MSC6U_P * NF3539MSC6U + PF3539MSNC6_P * NF3539MSNC6) / NF3539MS;
PF3539MA_P $=($ PF3539MAC6U_P $*$ NF3539MAC6U + PF3539MANC6_P $*$ NF3539MANC6) $/$ NF3539MA;
PF4044NM_P = (PF4044NMC6U_P * NF4044NMC6U + PF4044NMNC6_P * NF4044NMNC6) / NF4044NM;
PF4044MS_P = (PF4044MSC6U_P * NF4044MSC6U + PF4044MSNC6_P * NF4044MSNC6) / NF4044MS;
PF4044MA_P $=($ PF4044MAC6U_P $*$ NF4044MAC6U + PF4044MANC6_P $*$ NF4044MANC6) $/$ NF4044MA;

AGE 45 TO 54

## FEMALES - NEVER MARRIED 45 TO 54

```
PF4549NM_P = 0.03650-0.003349 + PF4549E_DE
    -0.00076 * RF4549-0.00070 * RF4549.1 - 0.00018 * RF4549.2 + 0.00049 * RF4549.3 + 0.00096 * RF4549.4 + 0.00091 * RF4549.5
    - 0.774819 * RF4549DI;
PF5054NM_P = 0.05788-0.008250 + PF5054E_DE
    + 0.00003 * RF5054-0.00011 * RF5054.1-0.00032 * RF5054.2 - 0.00051 * RF5054.3-0.00059 * RF5054.4 - 0.00045 * RF5054.5
    -0.732696 * RF5054DI;
```

FEMALES - MARRIED SPOUSE PRESENT 45 TO 54

```
PF4549MS_P = 0.03842-0.003545 + PF4549E_DE
    -0.00076 * RF4549-0.00070 * RF4549.1-0.00018 * RF4549.2 + 0.00049 * RF4549.3 + 0.00096 * RF4549.4 + 0.00091 *
    RF4549.5
    - 0.771341 * RF4549DI - 0.15 * RF4549MSCU6;
PF5054MS_P = - 0.40180-0.008735 + PF5054E_DE
    + 0.00003 * RF5054-0.00011 * RF5054.1-0.00032 * RF5054.2-0.00051 * RF5054.3-0.00059 * RF5054.4-0.00045 * RF5054.5
    + 0.00454 * TR_PF5054MS - 0.726715 * RF5054DI - 0.12 * RF5054MSCU6;
```

FEMALES - MARRIED SPOUSE ABSENT 45 TO 54

PF4549MA_P $=0.06830-0.001541+$ PF4549E_DE
-0.00076 * RF4549-0.00070 * RF4549.1-0.00018 * RF4549.2 + 0.00049 * RF4549.3 + 0.00096 * RF4549.4 + 0.00091 * RF4549.5

- 0.806961 * RF4549DI - 0.1 * RF4549MACU6;

PF5054MA_P $=0.08983-0.006021+$ PF5054E_DE
+0.00003 * RF5054-0.00011 * RF5054.1-0.00032 * RF5054.2-0.00051 * RF5054.3-0.00059 * RF5054.4-0.00045 * RF5054.5 -0.760157 * RF5054DI - 0.2 * RF5054MACU6;

AGE 20 TO 54

```
PF2024_P = (PF2024NM_P * NF2024NM + PF2024MS_P * NF2024MS + PF2024MA_P * NF2024MA) / NF2024;
PF2529_P = (PF2529NM_P * NF2529NM + PF2529MS_P * NF2529MS + PF2529MA_P * NF2529MA) / NF2529;
PF3034_P = (PF3034NM_P * NF3034NM + PF3034MS_P * NF3034MS + PF3034MA_P * NF3034MA) / NF3034;
PF3539_P = (PF3539NM_P * NF3539NM + PF3539MS_P * NF3539MS + PF3539MA_P * NF3539MA) / NF3539;
PF4044_P = (PF4044NM_P * NF4044NM + PF4044MS_P * NF4044MS + PF4044MA_P * NF4044MA) / NF4044;
PF4549_P = (PF4549NM_P * NF4549NM + PF4549MS_P * NF4549MS + PF4549MA_P * NF4549MA) / NF4549;
PF5054_P = (PF5054NM_P * NF5054NM + PF5054MS_P * NF5054MS + PF5054MA_P * NF5054MA) / NF5054;
PF2024 = PF2024_P;
PF2529 = PF2529_P;
PF3034 = PF3034_P;
PF3539 = PF3539_P;
PF4044 = PF4044_P;
PF4549 = PF4549_P;
PF5054 = PF5054_P;
```

PF2024NM = PF2024NM_P * PF2024 / PF2024_P;
PF2529NM $=$ PF2529NM_P *PF2529 / PF2529_P;
PF3034NM $=$ PF3034NM_P * PF3034 / PF3034_P;
PF3539NM $=$ PF3539NM_P * PF3539 / PF3539_P;
PF4044NM $=$ PF4044NM_P *PF4044 / PF4044_P;
PF4549NM $=$ PF4549NM_P *PF4549 / PF4549_P;
PF5054NM $=$ PF5054NM_P *PF5054 / PF5054_P;
PF2024MS = PF2024MS_P * PF2024 / PF2024_P;
PF2529MS = PF2529MS_P * PF2529 / PF2529_P;
PF3034MS = PF3034MS_P * PF3034 / PF3034_P;
PF3539MS = PF3539MS_P * PF3539 / PF3539_P;
PF4044MS $=$ PF4044MS_P *PF4044 / PF4044_P;
PF4549MS = PF4549MS_P *PF4549 / PF4549_P;
PF5054MS = PF5054MS_P * PF5054/PF5054_P;
PF2024MA = PF2024MA_P * PF2024 / PF2024_P;
PF2529MA $=$ PF2529MA_P *PF2529 / PF2529_P;
PF3034MA $=$ PF3034MA_P * PF3034 / PF3034_P;
PF3539MA $=$ PF3539MA_P *PF3539 / PF3539_P;
PF4044MA $=$ PF4044MA_P *PF4044 / PF4044_P;
PF4549MA $=$ PF4549MA_P *PF4549 / PF4549_P;
PF5054MA $=$ PF5054MA_P * PF5054 / PF5054_P;
PF2024NMC6U = PF2024NMC6U_P * PF2024 / PF2024_P;
PF2529NMC6U $=$ PF2529NMC6U_P * PF2529 / PF2529_P;
PF3034NMC6U = PF3034NMC6U_P * PF3034 / PF3034_P;
PF3539NMC6U $=$ PF3539NMC6U_P * PF3539 / PF3539_P;
PF4044NMC6U = PF4044NMC6U_P * PF4044 / PF4044_P;
PF2024NMNC6 = PF2024NMNC6_P * PF2024 / PF2024_P;
PF2529NMNC6 $=$ PF2529NMNC6_P * PF2529 / PF2529_P;
PF3034NMNC6 $=$ PF3034NMNC6_P * PF3034 / PF3034_P;
PF3539NMNC6 $=$ PF3539NMNC6_P * PF3539 / PF3539_P;
PF4044NMNC6 = PF4044NMNC6_P * PF4044 / PF4044_P;
PF2024MSC6U = PF2024MSC6U_P * PF2024 / PF2024_P;
PF2529MSC6U = PF2529MSC6U_P * PF2529 / PF2529_P;
PF3034MSC6U $=$ PF3034MSC6U_P * PF3034 / PF3034_P;
PF3539MSC6U = PF3539MSC6U_P * PF3539 / PF3539_P;
PF4044MSC6U $=$ PF4044MSC6U_P $*$ PF4044 / PF4044_P;
PF2024MSNC6 $=$ PF2024MSNC6_P * PF2024 / PF2024_P;
PF2529MSNC6 = PF2529MSNC6_P * PF2529 / PF2529_P;
PF3034MSNC6 $=$ PF3034MSNC6_P $*$ PF3034 / PF3034_P;

```
PF3539MSNC6 = PF3539MSNC6_P * PF3539 / PF3539_P;
PF4044MSNC6 = PF4044MSNC6_P * PF4044 / PF4044_P;
PF2024MAC6U = PF2024MAC6U_P * PF2024 / PF2024_P;
PF2529MAC6U = PF2529MAC6U_P * PF2529 / PF2529_P
PF3034MAC6U = PF3034MAC6U_P * PF3034 / PF3034_P
PF3539MAC6U = PF3539MAC6U_P * PF3539 / PF3539_P;
PF4044MAC6U = PF4044MAC6U_P * PF4044 / PF4044_P
PF2024MANC6 = PF2024MANC6_P * PF2024 / PF2024_P
PF2529MANC6 = PF2529MANC6_P * PF2529 / PF2529_P;
PF3034MANC6 = PF3034MANC6_P * PF3034 / PF3034_P
PF3539MANC6 = PF3539MANC6_P * PF3539 / PF3539_P
PF4044MANC6 = PF4044MANC6_P * PF4044 / PF4044_P;
```

AGE 55 TO 61

```
PF55_P = - 0.678427 * RF55DI-0.011198
    + 0.00064 * RF5559 + 0.00041 * RF5559.1-0.00029 * RF5559.2-0.00107 * RF5559.3-0.00155 * RF5559.4 - 0.00132 * RF5559.5
    + PF55E_DE + PF55_DM + 0.00368 * PF55COH48-0.90941;
PF56_P = - 0.651951 * RF56DI - 0.014942
    + 0.00064 * RF5559 + 0.00041 * RF5559.1-0.00029 * RF5559.2 - 0.00107 * RF5559.3-0.00155 * RF5559.4 - 0.00132 * RF5559.5
    + PF56E_DE + PF56_DM + 0.00486 * PF56COH48-0.96865;
PF57_P = - 0.634496 * RF57DI - 0.018113
    + 0.00064 * RF5559 + 0.00041 * RF5559.1-0.00029 * RF5559.2 - 0.00107 * RF5559.3-0.00155 * RF5559.4 - 0.00132 * RF5559.5
    + PF57E_DE + PF57_DM + 0.00344 * PF57COH48-0.85033;
PF58_P = - 0.604503* RF58DI - 0.023096
    + 0.00064 * RF5559 + 0.00041 * RF5559.1-0.00029 * RF5559.2-0.00107 * RF5559.3-0.00155 * RF5559.4 - 0.00132 * RF5559.5
    + PF58E_DE + PF58_DM + 0.00362 * PF58COH48-0.83081;
PF59_P = - 0.571801*RF59DI - 0.029463
    + 0.00064 * RF5559 + 0.00041 * RF5559.1-0.00029 * RF5559.2-0.00107 * RF5559.3-0.00155 * RF5559.4 - 0.00132 * RF5559.5
    + PF59E_DE + PF59_DM + 0.00470 * PF59COH48 - 0.85665;
PF60_P = - 0.526008*RF60DI - 0.033474
    + 0.00141 * RF6064 + 0.00166 * RF6064.1 + 0.00116 * RF6064.2 + 0.00033 * RF6064.3 - 0.00041 * RF6064.4 - 0.00066 * RF6064.5
    + PF60E_DE + PF60_DM + 0.00819 * PF60COH48-1.18744;
PF61_P = - 0.487750 * RF61DI - 0.041036
    + 0.00141 * RF6064 + 0.00166 * RF6064.1 + 0.00116 * RF6064.2 + 0.00033 * RF6064.3 - 0.00041 * RF6064.4 - 0.00066 * RF6064.5
    + PF61E_DE + PF61_DM + 0.00520 * PF61COH48-0.87850;
```

AGE 62 TO 74

```
PF62_P = - 0.412411 * RF61DI.4-0.063463
```

    +0.00141 * RF6064 + 0.00166 * RF6064.1 + 0.00116 * RF6064.2 + 0.00033 * RF6064.3-0.00041 * RF6064.4-0.00066 * RF6064.5
    + PF62E_DE + PF62_DM + 0.00523 * PF62COH48-0.56287-0.5100 * RRADJ_F62-0.02 * POT_ET_TXRT_62;
    ```
PF63_P = - 0.359185 * RF61DI.8-0.082104
```

    +0.00141 * RF6064 + 0.00166 * RF6064.1 + 0.00116 * RF6064.2 + 0.00033 * RF6064.3-0.00041 * RF6064.4-0.00066 * RF6064.5
    + PF63E_DE + PF63_DM + 0.00659 * PF63COH48-0.68289-0.4675 * RRADJ_F63-0.02 * POT_ET_TXRT_63;
    PF64_P $=-0.323866 *$ RF61DI. $12-0.094477$
+0.00141 * RF6064 + 0.00166 * RF6064.1 + 0.00116 * RF6064.2 + 0.00033 * RF6064.3-0.00041 * RF6064.4-0.00066 * RF6064.5
+ PF64E_DE + PF64_DM + 0.00745 * PF64COH48-0.72813-0.4250 * RRADJ_F64-0.02 * POT_ET_TXRT_64;
PF65_P $=-0.256938 *$ RF61DI $16+0.035143$
+0.00029 * RF6569 + 0.00014 * RF6569.1-0.00023 * RF6569.2-0.00063 * RF6569.3-0.00086 * RF6569.4-0.00072 * RF6569.5
+ PF65E_DE + PF65_DM + 0.00348 * PF65COH48-0.37490-0.3825 * RRADJ_F65-0.02 * POT_ET_TXRT_65;

PF66_P $=-0.226911 *$ RF61DI $20+0.028402$

```
    + 0.00029 * RF6569 + 0.00014 * RF6569.1-0.00023 * RF6569.2 - 0.00063 * RF6569.3-0.00086 * RF6569.4-0.00072 * RF6569.5
    + PF66E_DE + PF66_DM + 0.00512 * PF66COH48-0.52851-0.3400 * RRADJ_F66 - 0.02 * POT_ET_TXRT_66;
PF67_P = 0.00029 * RF6569 + 0.00014 * RF6569.1-0.00023 * RF6569.2-0.00063 * RF6569.3-0.00086 * RF6569.4-0.00072 * RF6569.5
    - 0.206835 * RF61DI. 24 + 0.025203 + PF67E_DE + PF67_DM + 0.00518 * PF67COH48-0.53735 - 0.2975 * RRADJ_F67 - 0.02 *
        POT_ET_TXRT_67;
PF68_P = 0.00029 * RF6569 + 0.00014 * RF6569.1-0.00023 * RF6569.2 - 0.00063 * RF6569.3-0.00086 * RF6569.4 - 0.00072 * RF6569.5
    - 0.180608 * RF61DI. 28 + 0.020367 + PF68E_DE + PF68_DM + 0.00400 * PF68COH48-0.41046 - 0.2550 * RRADJ_F68 - 0.02 *
        POT_ET_TXRT_68;
PF69_P = 0.00029 * RF6569 + 0.00014 * RF6569.1-0.00023 * RF6569.2 - 0.00063 * RF6569.3-0.00086 * RF6569.4 - 0.00072 * RF6569.5
    - 0.162386 * RF61DI. 32 + 0.017869 + PF69E_DE + PF69_DM + 0.00400 * PF69COH48 - 0.36706 - 0.2550 * RRADJ_F69 - 0.02 *
        POT_ET_TXRT_69;
PF70_P = - 0.00009 * RF7074-0.00028 * RF7074.1-0.00048 * RF7074.2-0.00063 * RF7074.3-0.00064 * RF7074.4 - 0.00046 * RF7074.5
    - 0.136965 * RF61DI.36 + 0.015491 + PF70E_DE + PF70_DM + 0.00424 * PF70COH48 - 0.49282;
PF71_P = - 0.00009 * RF7074-0.00028 * RF7074.1-0.00048 * RF7074.2-0.00063 * RF7074.3-0.00064 * RF7074.4 - 0.00046 * RF7074.5
    - 0.123245 * RF61DI. 40 + 0.014381 + PF71E_DE + PF71_DM + 0.00300 * PF71COH48 - 0.37390;
PF72_P = - 0.00009 * RF7074-0.00028 * RF7074.1-0.00048 * RF7074.2 - 0.00063 * RF7074.3 - 0.00064 * RF7074.4 - 0.00046 * RF7074.5
    - 0.107978 * RF61DI.44 + 0.013644 + PF72E_DE + PF72_DM + 0.00286 * PF72COH48 - 0.34670;
PF73_P = - 0.00009 * RF7074-0.00028 * RF7074.1-0.00048 * RF7074.2-0.00063 * RF7074.3-0.00064 * RF7074.4 - 0.00046 * RF7074.5
        - 0.094426 * RF61DI. 48 + 0.012196 + PF73E_DE + PF73_DM + 0.00370 * PF73COH48 - 0.41058;
PF74_P = - 0.00009 * RF7074-0.00028 * RF7074.1-0.00048 * RF7074.2 - 0.00063 * RF7074.3-0.00064 * RF7074.4 - 0.00046 * RF7074.5
    - 0.082404 * RF61DI. }2\mathrm{ + 0.010762 + PF74E_DE + PF74_DM + 0.00304 * PF74COH48 - 0.35113;
```

AGE 75 TO 79

```
PF75_P = PF74.4 * 0.900 + DPF75O_FE;
PF76_P = PF75.4 * 0.900 + DPF75O_FE;
PF77_P = PF76.4 * 0.900 + DPF75O_FE;
PF78_P = PF77.4 * 0.900 + DPF75O_FE;
PF79_P = PF78.4 * 0.900 + DPF75O_FE;
```

AGE 55 TO 79

```
PF55 = PF55_P;
PF56 = PF56_P;
PF57 = PF57_P;
PF58 = PF58_P;
PF59 = PF59_P;
PF60 = PF60_P;
PF61 = PF61_P;
PF62 = PF62_P;
PF63 = PF63_P;
PF64 = PF64_P;
PF65 = PF65_P;
PF66 = PF66_P;
PF67 = PF67_P;
PF68 = PF68_P;
PF69 = PF69_P;
PF70 = PF70_P;
PF71 = PF71_P;
PF72 = PF72_P;
PF73 = PF73_P;
PF74 = PF74_P;
PF75 = PF75_P;
PF76 = PF76_P;
PF77 = PF77_P;
PF78 = PF78_P;
```

```
PF79 = PF79_P;
```

AGE 80 AND OVER

| PF80_P $=$ PF79.4 | $* 0.965^{* *}(1)+$ DPF75O_FE; |
| :--- | :--- |
| PF81_P $=$ PF79.8 | $* 0.965^{* *}(2)+$ DPF75O_FE; |
| PF82_P $=$ PF79.12 | $* 0.965^{* *}(3)+$ DPF75O_FE; |
| PF83_P $=$ PF79.16 | $* 0.965^{* *}(4)+$ DPF75O_FE; |
| PF84_P $=$ PF79.20 | $* 0.965^{* *}(5)+$ DPF75O_FE; |

PF85_P $=$ MOVAVG(8,PF79.24) ${ }^{*} 0.965^{* *}(6)+$ DPF75O_FE;
PF86_P $=$ MOVAVG(8,PF79.28) ${ }^{*} 0.965^{* *}(7)+$ DPF75O_FE;
PF87_P $=$ MOVAVG(8,PF79.32) ${ }^{*} 0.965^{* *}(8)+$ DPF75O_FE;
PF88_P $=$ MOVAVG(8,PF79.36) ${ }^{*} 0.965^{* *}(9)+$ DPF75O_FE;
PF89_P $=$ MOVAVG(8,PF79.40) ${ }^{*} 0.965^{* *}(10)+$ DPF75O_FE;
PF90_P $=$ MOVAVG(8,PF79.44) ${ }^{*} 0.965^{* *}(11)+$ DPF75O_FE;
PF91_P $=$ MOVAVG(8,PF79.48) $* 0.965 * *(12)+$ DPF75O_FE;
PF92_P $=$ MOVAVG(8,PF79.52) ${ }^{*} 0.965^{* *}(13)+$ DPF75O_FE;
PF93_P $=$ MOVAVG(8,PF79.56) ${ }^{*} 0.965^{* *}(14)+$ DPF75O_FE;
PF94_P $=$ MOVAVG(8,PF79.60) ${ }^{*} 0.965^{* *}(15)+$ DPF75O_FE;
PF95_P $=$ PF94_P * $0.965+$ DPF75O_FE;
PF96_P $=$ PF95_P $* 0.965+$ DPF75O_FE;
PF97_P $=$ PF96_P * $0.965+$ DPF75O_FE;
PF98_P $=$ PF97_P $* 0.965+$ DPF75O_FE;
PF99_P $=$ PF98_P $* 0.965+$ DPF75O_FE;
PF100_P $=$ PF99_P $* 0.965+$ DPF75O_FE;
PF80O_P $=($ PF80_P $*$ NF80 + PF81_P $* N F 81+$ PF82_P $* N F 82+$ PF83_P*NF83 + PF84_P $*$ NF84 + PF85_P $* N F 85+$ PF86_P $* N F 86+$
PF87_P*NF87 + PF88_P*NF88 + PF89_P*NF89 +
PF90_P*NF90 + PF91_P*NF91 + PF92_P*NF92 + PF93_P*NF93 + PF94_P*NF94 + PF95_P*NF95 + PF96_P*NF96 +
PF97_P*NF97 + PF98_P*NF98 + PF99_P*NF99 +
PF100_P*NF100 ) / NF80O;
PF80O = PF80O_P;
PF80 = PF80_P * PF800 / PF80O_P;
PF81 = PF81_P * PF800 / PF800_P;
PF82 $=$ PF82_P $*$ PF800 / PF800_P;
PF83 = PF83_P * PF800 / PF80O_P;
PF84 = PF84_P * PF800 / PF800_P;
PF85 = PF85_P * PF800 / PF80O_P;
PF86 = PF86_P * PF800 / PF800_P;
PF87 = PF87_P * PF800 / PF80O_P;
PF88 $=$ PF88_P $*$ PF800 / PF80O_P;
PF89 = PF89_P * PF800 / PF800_P;
PF90 $=$ PF90_P $*$ PF800 $/$ PF80O_P;
PF91 = PF91_P *PF800 / PF800_P;
PF92 = PF92_P * PF800 / PF800_P;
PF93 = PF93_P * PF800 / PF800_P;
PF94 $=$ PF94_P $*$ PF800 / PF80O_P;
PF95 = PF95_P * PF800 / PF800_P;
PF96 $=$ PF96_P * PF800 / PF80O_P;
PF97 = PF97_P * PF800 / PF800_P;
PF98 = PF98_P * PF800 / PF80O_P;
PF99 $=$ PF99_P $*$ PF800 $/$ PF800_P;
PF100 = PF100_P * PF800 / PF80O_P;

## LFPR EQUATIONS, AGE 16 AND OVER

PM16O_P=(
PM1617_P * NM1617 +
PM1819_P * NM1819 +
PM2024_P * NM2024 +
PM2529_P * NM2529 +

```
PM3034_P * NM3034 +
PM3539_P * NM3539 +
PM4044_P * NM4044 +
PM4549_P * NM4549 +
PM5054_P * NM5054 +
PM55_P * NM55 +
PM56_P * NM56 +
PM57_P * NM57 +
PM58_P * NM58 +
PM59_P * NM59 +
PM60_P * NM60 +
PM61_P * NM61 +
PM62_P * NM62 +
PM63 P * NM63 +
PM64_P * NM64 +
PM65_P * NM65 +
PM66_P * NM66 +
PM67_P * NM67 +
PM68_P * NM68 +
PM69_P * NM69 +
PM70_P * NM70 +
PM71_P * NM71 +
PM72 P * NM72 +
PM73_P * NM73 +
PM74_P * NM74 +
PM75_P * NM75 +
PM76_P * NM76 +
PM77_P * NM77 +
PM78_P * NM78 +
PM79_P * NM79 +
PM80O_P * NM80O ) / NM16O;
PF16O_P= (
PF1617_P * NF1617 +
PF1819_P * NF1819 +
PF2024_P * NF2024 +
PF2529_P * NF2529 +
PF3034_P * NF3034 +
PF3539_P * NF3539 +
PF4044_P * NF4044 +
PF4549_P * NF4549 +
PF5054_P * NF5054 +
PF55_P * NF55 +
PF56_P * NF56 +
PF57_P * NF57 +
PF58_P * NF58 +
PF59_P * NF59 +
PF60_P * NF60 +
PF61_P * NF61 +
PF62_P * NF62 +
PF63_P * NF63 +
PF64 P * NF64 +
PF65_P * NF65 +
PF66_P * NF66 +
PF67_P * NF67 +
PF68_P * NF68 +
PF69_P * NF69 +
PF70_P * NF70 +
PF71 P * NF71 +
PF72_P * NF72 +
PF73 P * NF73 +
PF74_P * NF74 +
PF75_P * NF75 +
PF76_P * NF76 +
PF77_P * NF77 +
PF78_P * NF78 +
PF79_P * NF79 +
PF80O_P * NF80O ) / NF16O;
P16O_P = (PM16O_P * NM16O + PF16O_P * NF16O) / (NM16O + NF16O);
```

LCM_P = PM16O_P * NM16O;
LCF_P $=$ PF16O_P $*$ NF16O;
LC_P = LCM_P + LCF_P;

## LABOR FORCE PARTICIPATION RATES, FULL EMPLOYMENT

## MALE LFPR EQUATIONS, FULL EMPLOYMENT

```
DPM1617_FE = (- 0.00158 * RM1617_FE - 0.00180 * RM1617_FE.1 - 0.00115 * RM1617_FE.2 - 0.00014 * RM1617_FE. 3 + 0.00072 *
RM1617_FE.4 + 0.00094 * RM1617_FE.5) - (- 0.00158 * RM1617-0.00180 * RM1617.1-0.00115 * RM1617.2 - 0.00014 * RM1617.3 +
0.00072 * RM1617.4 + 0.00094 * RM1617.5)
    Restricted Ordinary Least Squares
        QUARTERLY data for 148 periods from 1971Q1 to 2007Q4
        Date: 2 SEP 2009
    pm1617_dpk
        = - 0.00158 * rm1617_dpk - 0.00180 * rm1617_dpk[-1]
                (5.56270) (6.13183)
                - 0.00115 * rm1617_dpk[-2] - 0.00014 * rm1617_dpk[-3]
                (7.49261) (0.94429)
                + 0.00072 * rm1617_dpk[-4] + 0.00094 * rm1617_dpk[-5]
                        (2.47378) (3.30450)
    Polynomial lags:
                rm1617_dpk
                from 0 to 5 degree 3 near far
```

    Sum Sq 0.0343 Std Err 0.0149 LHS Mean -0.0011 Res Mean 0.0037
    \(\begin{array}{llllllll}\text { R Sq } & 0.2826 & \text { R Bar Sq } 0.2777 & \text { F } 2,146 & 28.7598 & \text { \%RMSE } & 46265.0\end{array}\)
    \(\begin{array}{llll}\text { D.W. ( 1) } & 0.3942 & \text { D.W. ( 4) } & 0.9128\end{array}\)
    DPM1819_FE $=(-0.00108$ * RM1819_FE - 0.00127 * RM1819_FE. $1-0.00088$ * RM1819_FE. $2-0.00023$ * RM1819_FE. 3 + 0.00034 * RM1819_FE. $4+0.00053$ * RM1819_FE.5) - (-0.00108 * RM1819-0.00127 * RM1819.1-0.00088 * RM1819.2 - 0.00023 * RM1819.3 + 0.00034 * RM1819.4 + 0.00053 * RM1819.5)

```
        Restricted Ordinary Least Squares
        QUARTERLY data for 148 periods from 1971Q1 to 2007Q4
        Date: 2 SEP 2009
        pm1819_dpk
        = - 0.00108 * rm1819_dpk - 0.00127 * rm1819_dpk[-1]
            (5.22267) (5.94236)
                        - 0.00088 * rm1819_dpk[-2] - 0.00023 * rm1819_dpk[-3]
                (7.96365) (2.16526)
                        + 0.00034 * rm1819_dpk[-4] + 0.00053 * rm1819_dpk[-5]
                        (1.61951) (2.56002)
        Polynomial lags:
            rm1819_dpk
            from 0 to 5 degree 3 near far
```

| Sum Sq | 0.0192 | Std Err | 0.0114 | LHS Mean | -0.0040 | Res Mean | 0.0008 |
| :--- | ---: | :--- | :--- | :--- | :--- | :--- | ---: |
| R Sq | 0.2729 | R Bar Sq | 0.2679 | F 2,146 | 27.3976 | \%RMSE | 17596.7 |
| D.W.( 1) | 0.7981 | D.W.( 4) | 1.2825 |  |  |  |  |

DPM2024_FE $=(-0.00063$ * RM2024_FE - 0.00077 * RM2024_FE. $1-0.00059$ * RM2024_FE. $2-0.00027$ * RM2024_FE. 3 + 0.00005 * RM2024_FE. $4+0.00020$ * RM2024_FE.5) - (- 0.00063 * RM2024-0.00077 * RM2024.1-0.00059 * RM2024.2-0.00027 * RM2024.3 + 0.00005 * RM2024.4 + 0.00020 * RM2024.5)

```
Restricted Ordinary Least Squares
    QUARTERLY data for }148\mathrm{ periods from 1971Q1 to 2007Q4
    Date: 2 SEP 2009
    pm2024_dpk
        = - - 0.00063 (4.32837) * rm2024_dpk - - 0.00077 (5.14659)
        - 0.00059 * rm2024_dpk[-2] - 0.00027 * rm2024_dpk[-3]
                (7.47597) (3.39844)
        + 0.00005 * rm2024_dpk[-4] + 0.00020 * rm2024_dpk[-5]
                (0.35674) (1.36468)
```

Polynomial lags:

```
rm2024_dpk
from 0 to 5 degree 3 near far
```

| Sum Sq | 0.0067 | Std Err | 0.0068 | LHS Mean | -0.0028 | Res Mean | -0.0002 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| R Sq | 0.2336 | R Bar Sq | 0.2284 | F 2,146 | 22.2560 | \%RMSE | 32980.5 |
| D.W.( 1) | 0.5915 | D.W. ( 4) | 1.0423 |  |  |  |  |

DPM2529_FE $=(-0.00028 *$ RM2529_FE - 0.00044 * RM2529_FE. $1-0.00050$ * RM2529_FE. $2-0.00047$ * RM2529_FE. 3 - 0.00037 * RM2529_FE. $4-0.00021$ * RM2529_FE.5) - (- 0.00028 * RM2529-0.00044 * RM2529.1-0.00050 * RM2529.2-0.00047 * RM2529.3 0.00037 * RM2529.4-0.00021 * RM2529.5)

```
Restricted Ordinary Least Squares
    QUARTERLY data for 101 periods from 1982Q4 to 2007Q4
    Date: 2 SEP 2009
```

pm2529_dpk

```
= - 0.00028 * rm2529_dpk - 0.00044 * rm2529_dpk[-1]
                    (1.40609) (2.14232)
                    -0.00050 * rm2529_dpk[-2] - 0.00047 * rm2529_dpk[-3]
                    (4.60340) (4.70117)
                    - 0.00037 * rm2529_dpk[-4] - 0.00021 * rm2529_dpk[-5]
                    (1.87119) (1.06920)
```

Polynomial lags:

```
                rm2529_dpk
```

                from 0 to 5 degree 3 near far
    | Sum Sq | 0.0026 | Std Err | 0.0049 | LHS Mean | -0.0033 | Res Mean | -0.0016 |
| :--- | :---: | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| R Sq | 0.0492 | R Bar Sq | 0.0396 | F 2, 99 | 2.5598 | \%RMSE | 48872.7 |
| D.W.( 1) | 0.7362 | D.W. ( 4) | 1.0116 |  |  |  |  |

DPM3034_FE $=(-0.00046 *$ RM3034_FE - 0.00061 * RM3034_FE.1-0.00054 * RM3034_FE. $2-0.00036$ * RM3034_FE. $3-0.00014$ * RM3034_FE. 4 + 0.00001 * RM3034_FE.5) - (- 0.00046 * RM3034-0.00061 * RM3034.1-0.00054 * RM3034.2 - 0.00036 * RM3034.30.00014 * RM3034.4 + 0.00001 * RM3034.5)

```
    Restricted Ordinary Least Squares
    QUARTERLY data for 101 periods from 1982Q4 to 2007Q4
    Date: 2 SEP 2009
    pm3034_dpk
```

```
= - 0.00046 * rm3034_dpk - 0.00061 * rm3034_dpk[-1]
                (1.81713) (2.34240)
    - 0.00054 * rm3034_dpk[-2] - 0.00036 * rm3034_dpk[-3]
        (4.05252) (2.76527)
    - 0.00014 * rm3034_dpk[-4] + 0.00001 * rm3034_dpk[-5]
        (0.56353)
                        (0.02760)
```

Polynomial lags:

```
rm3034_dpk
from 0 to 5 degree 3 near far
```

| Sum Sq | 0.0024 | Std Err | 0.0049 | LHS Mean | -0.0015 | Res Mean | -0.0001 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| R Sq | 0.1497 | R Bar Sq | 0.1411 | F 2, 99 | 8.7134 | \%RMSE | 50015.8 |
| D.W.( 1) | 0.6227 | D.W.( 4) | 0.8678 |  |  |  |  |

```
DPM3539_FE = (-0.00004 * RM3539_FE - 0.00010 * RM3539_FE.1-0.00016 * RM3539_FE.2 - 0.00021 * RM3539_FE.3-0.00021 *
RM3539_FE.4-0.00015 * RM3539_FE.5) - (- 0.00004 * RM3539 - 0.00010 * RM3539.1-0.00016 * RM3539.2-0.00021 * RM3539.3-
0.00021 * RM3539.4-0.00015 * RM3539.5)
    Restricted Ordinary Least Squares
    QUARTERLY data for 101 periods from 1982Q4 to 2007Q4
    Date: 2 SEP 2009
    pm3539_dpk
        = - 0.00004 * rm3539_dpk - 0.00010 * rm3539_dpk[-1]
            (0.12650) (0.32986)
            - 0.00016 * rm3539_dpk[-2] - 0.00021 * rm3539_dpk[-3]
                (1.10007) (1.48792)
                - 0.00021 * rm3539_dpk[-4] - 0.00015 * rm3539_dpk[-5]
                (0.72412) (0.51868)
    Polynomial lags:
                rm3539_dpk
                from 0 to 5 degree 3 near far
\begin{tabular}{lrlrlrll} 
Sum Sq & 0.0028 & Std Err & 0.0052 & LHS Mean & -0.0017 & Res Mean & -0.0010 \\
R Sq & -0.0664 & R Bar Sq & -0.0772 & F 2, 99 & NC & \%RMSE & 1817.43
\end{tabular}
```

DPM4044_FE $=(-0.00057 *$ RM4044_FE - 0.00066 * RM4044_FE. $1-0.00044 *$ RM4044_FE. $2-0.00009 *$ RM4044_FE. $3+0.00022$ *
RM4044_FE. 4 + 0.00031 * RM4044_FE.5) - (- 0.00057 * RM4044-0.00066 * RM4044.1-0.00044 * RM4044.2-0.00009 * RM4044.3 +
0.00022 * RM4044.4 +0.00031 * RM4044.5)
Restricted Ordinary Least Squares
QUARTERLY data for 101 periods from 1982Q4 to 2007Q4
Date: 2 SEP 2009
pm4044_dpk

```
        = - 0.00057 * rm4044_dpk - 0.00066 * rm4044_dpk[-1]
                (1.75853) (1.97180)
                - 0.00044 * rm4044_dpk[-2] - 0.00009 * rm4044_dpk[-3]
                (2.56550) (0.54915)
                + 0.00022 * rm4044_dpk[-4] + 0.00031 * rm4044_dpk[-5]
                (0.66426) (0.95499)
```

    Polynomial lags:
            rm4044_dpk
    from 0 to 5 degree 3 near far

| Sum Sq | 0.0029 | Std Err | 0.0054 | LHS Mean | -0.0012 | Res Mean | -0.0004 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| R Sq | 0.0268 | R Bar Sq | 0.0170 | F 2, 99 | 1.3647 | \%RMSE | 310.335 |
| D.W.( 1) | 0.5028 | D.W. ( 4) | 0.7333 |  |  |  |  |

DPM4549_FE $=(-0.00002 *$ RM4549_FE - $0.00016 *$ RM4549_FE. $1-0.00034 *$ RM4549_FE. $2-0.00049 *$ RM4549_FE. $3-0.00054$ * RM4549_FE. $4-0.00040$ * RM4549_FE.5) - (- 0.00002 * RM4549-0.00016 * RM4549.1-0.00034 * RM4549.2-0.00049 * RM4549.3 0.00054 * RM4549.4-0.00040 * RM4549.5)

```
Restricted Ordinary Least Squares
QUARTERLY data for 101 periods from 1982Q4 to 2007Q4
Date: 2 SEP 2009
```

pm4549_dpk

$$
\begin{aligned}
& =-0.00002 \text { * rm4549_dpk - } 0.00016 \text { * rm4549_dpk[-1] } \\
& \text { (0.06650) (0.55897) } \\
& \text { - } 0.00034 \text { * rm4549_dpk[-2] - } 0.00049 \text { * rm4549_dpk[-3] } \\
& \text { (2.31789) (3.39192) } \\
& \text { - } 0.00054 \text { * rm4549_dpk[-4] - 0.00040 * rm4549_dpk[-5] } \\
& \text { (1.92566) (1.46813) }
\end{aligned}
$$

Polynomial lags:

```
rm4549_dpk
from 0 to 5 degree 3 near far
```

| Sum Sq | 0.0015 | Std Err | 0.0039 | LHS Mean | -0.0013 | Res Mean | -0.0003 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| R Sq | 0.0719 | R Bar Sq | 0.0625 | F 2, 99 | 3.8350 | \%RMSE | 4821.60 |
| D.W.( 1) | 0.7819 | D.W.( 4) | 1.0968 |  |  |  |  |

DPM5054_FE $=(0.00112 *$ RM5054_FE $+0.00103 *$ RM5054_FE. $1+0.00023 *$ RM5054_FE. $2-0.00078 *$ RM5054_FE. $3-0.00149 *$ RM5054_FE. $4-0.00139$ * RM5054_FE.5) - ( 0.00112 * RM5054 + 0.00103 * RM5054.1 + 0.00023 * RM5054.2 - 0.00078 * RM5054.3 0.00149 * RM5054.4-0.00139 * RM5054.5)

```
Restricted Ordinary Least Squares
    QUARTERLY data for 101 periods from 1982Q4 to 2007Q4
    Date: 2 SEP 2009
pm5054_dpk
    = 0.00112 * rm5054_dpk + 0.00103 * rm5054_dpk[-1]
        (3.16181) (2.84569)
        + 0.00023 * rm5054_dpk[-2] - 0.00078 * rm5054_dpk[-3]
                (1.30991)
                (4.52791)
                - 0.00149 * rm5054_dpk[-4] - 0.00139 * rm5054_dpk[-5]
                (4.15166) (3.96094)
```

Polynomial lags:
rm5054_dpk
from 0 to 5 degree 3 near far

| Sum Sq | 0.0037 | Std Err | 0.0061 | LHS Mean | -0.0012 | Res Mean | 0.0001 |  |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | ---: | ---: |
| R Sq | 0.1417 | R Bar Sq | 0.1330 | F 2, 99 | 8.1732 | \%RMSE | 42597.0 |  |
| D.W.( 1) | 1.0445 | D.W. ( 4) | 1.5173 |  |  |  |  |  |

DPM5559_FE $=(0.00062$ * RM5559_FE + 0.00041 * RM5559_FE. $1-0.00026$ * RM5559_FE. $2-0.00101$ * RM5559_FE. 3 - 0.00147 * RM5559_FE.4-0.00126 * RM5559_FE.5) - ( 0.00062 * RM5559 + 0.00041 * RM5559.1-0.00026 * RM5559.2-0.00101 * RM5559.3 0.00147 * RM5559.4-0.00126 * RM5559.5)

[^15]```
Date: 2 SEP 2009
pm5559_dpk
    = 0.00062 * rm5559_dpk + 0.00041 * rm5559_dpk[-1]
    (1.23508) (0.79452)
    - 0.00026 * rm5559_dpk[-2] - 0.00101 * rm5559_dpk[-3]
        (1.02810) (4.16786)
        - 0.00147 * rm5559_dpk[-4] - 0.00126 * rm5559_dpk[-5]
        (2.91432) (2.53301)
```

Polynomial lags:
rm5559_dpk
from 0 to 5 degree 3 near far

| Sum Sq | 0.0066 | Std Err | 0.0080 | LHS Mean | -0.0037 | Res Mean | -0.0014 |
| :--- | ---: | :--- | :--- | :--- | :--- | :--- | ---: |
| R Sq | 0.0164 | R Bar Sq | 0.0065 | F 2, 99 | 0.8277 | \%RMSE | 256947 |
| D.W.( 1) | 0.8793 | D.W. ( 4) | 1.2992 |  |  |  |  |

DPM6064_FE $=(0.00203 *$ RM6064_FE $+0.00160 *$ RM6064_FE. $1-0.00021 *$ RM6064_FE. $2-0.00235 *$ RM6064_FE. 3 - 0.00374 * RM6064_FE. $4-0.00331$ * RM6064_FE.5) - ( 0.00203 * RM6064 + 0.00160 * RM6064.1-0.00021 * RM6064.2 - 0.00235 * RM6064.3 0.00374 * RM6064.4-0.00331 * RM6064.5)

```
Restricted Ordinary Least Squares
    QUARTERLY data for 101 periods from 1982Q4 to 2007Q4
    Date: 2 SEP 2009
    pm6064_dpk
    = 0.00203 * rm6064_dpk + 0.00160 * rm6064_dpk[-1]
        (2.66008) (2.03611)
        - 0.00021 * rm6064_dpk[-2] - 0.00235 * rm6064_dpk[-3]
                (0.51861) (5.71949)
        - 0.00374 * rm6064_dpk[-4] - 0.00331 * rm6064_dpk[-5]
        (4.77362) (4.35298)
```

    Polynomial lags:
        rm6064_dpk
        from 0 to 5 degree 3 near far
    | Sum Sq | 0.0165 | Std Err | 0.0110 | LHS Mean | -0.0095 | Res Mean | -0.0067 |
| :--- | ---: | :--- | ---: | :--- | ---: | ---: | ---: | ---: |
| R Sq | -0.2740 | R Bar Sq | -0.2869 | F 2, 99 | NC | \%RMSE | 21316.0 |

DPM6569_FE $=(0.00067 *$ RM6569_FE $+0.00040 *$ RM6569_FE. $1-0.00040 *$ RM6569_FE. $2-0.00127$ * RM6569_FE. 3 - 0.00178 * RM6569_FE. $4-0.00151$ * RM6569_FE.5) - ( 0.00067 * RM6569 + 0.00040 * RM6569.1-0.00040 * RM6569.2-0.00127 * RM6569.3 0.00178 * RM6569.4-0.00151 * RM6569.5)

Polynomial lags:
rm6569_dpk
from 0 to 5 degree 3 near far

| Sum Sq | 0.0210 | Std Err | 0.0138 | LHS Mean | -0.0059 | Res Mean | -0.0047 |  |
| :--- | ---: | :--- | ---: | :--- | ---: | ---: | ---: | ---: |
| R Sq | -0.0832 | R Bar Sq | -0.0942 | F 2, 99 | NC | \%RMSE | 36465.0 |  |
| D.W.( 1) | 0.5122 | D.W. ( 4) | 0.7034 |  |  |  |  |  |

DPM7074_FE $=(-0.00013 *$ RM7074_FE - 0.00016 * RM7074_FE. $1-0.00013 *$ RM7074_FE. $2-0.00006 *$ RM7074_FE. 3 + 0.00000 * RM7074_FE. 4 + 0.00003 * RM7074_FE.5) - (- 0.00013 * RM7074-0.00016 * RM7074.1-0.00013 * RM7074.2 - 0.00006 * RM7074.3 + 0.00000 * RM7074.4 + 0.00003 * RM7074.5)

```
    Restricted Ordinary Least Squares
        QUARTERLY data for 101 periods from 1982Q4 to 2007Q4
        Date: 2 SEP 2009
    pm7074_dpk
        = - 0.00013 * rm7074_dpk - 0.00016 * rm7074_dpk[-1]
            (0.30689) (0.35702)
                - 0.00013 * rm7074_dpk[-2] - 0.00006 * rm7074_dpk[-3]
                (0.44247) (0.21289)
                + 0.00000 * rm7074_dpk[-4] + 0.00003 * rm7074_dpk[-5]
                (0.00874) (0.08371)
```

    Polynomial lags:
            rm7074_dpk
            from 0 to 5 degree 3 near far
    | Sum Sq | 0.0114 | Std Err | 0.0080 | LHS Mean | -0.0073 | Res Mean | -0.0071 |  |
| :--- | ---: | :--- | ---: | :--- | :--- | :--- | :--- | :--- |
| R Sq | -0.8795 | R Bar Sq | -0.8984 | F 2, 99 | NC | \%RMSE | 1540.73 |  |
| D.W.( 1) | 0.5068 | D.W. ( 4) | 0.7160 |  |  |  |  |  |

DPM75O_FE $=(-0.00043$ * RM75O_FE - 0.00051 * RM75O_FE. $1-0.00036$ * RM75O_FE. $2-0.00010$ * RM75O_FE. 3 + 0.00013 * RM75O_FE. $4+0.00021$ * RM75O_FE.5) - (- 0.00043 * RM75O-0.00051 * RM75O.1-0.00036 * RM75O. 2 - 0.00010 * RM75O.3 + 0.00013 * RM75O. 4 + 0.00021 * RM75O.5)

Restricted Ordinary Least Squares QUARTERLY data for 101 periods from 1982Q4 to 2007Q4 Date: 2 SEP 2009
pm750_dpk

| (1.82379) | (2.03996) |
| :---: | :---: |



Polynomial lags:
rm750_dpk from 0 to 5 degree 3 near far

| Sum Sq | 0.0039 | Std Err | 0.0058 | LHS Mean | -0.0031 | Res Mean | -0.0025 |  |
| :--- | ---: | :--- | ---: | :--- | :--- | :--- | :--- | :--- |
| R Sq | -0.2424 | R Bar Sq | -0.2549 | F 2, 99 | NC | \%RMSE | 19877.0 |  |
| D.W.( 1) | 0.8382 | D.W.( 4) | 1.0542 |  |  |  |  |  |

## FEMALE LFPR EQUATIONS, FULL EMPLOYMENT DIFFERENTIALS

```
DPF1617_FE = (- 0.00224 * RF1617_FE - 0.00239 * RF1617_FE.1 - 0.00126 * RF1617_FE.2 + 0.00035 * RF1617_FE. 3 + 0.00163 *
RF1617_FE.4 + 0.00178 * RF1617_FE.5) - (- 0.00224 * RF1617 - 0.00239 * RF1617.1-0.00126 * RF1617.2 + 0.00035 * RF1617.3 + 0.00163
* RF1617.4 + 0.00178 * RF1617.5)
    Restricted Ordinary Least Squares
        QUARTERLY data for 148 periods from 1971Q1 to 2007Q4
        Date: 2 SEP 2009
    pf1617_dpk
            = -0.00224 * rf1617_dpk - 0.00239 * rf1617_dpk[-1]
                    (5.90261) (6.13505)
                - 0.00126 * rf1617_dpk[-2] + 0.00035 * rf1617_dpk[-3]
                (6.26010) (1.72973)
                + 0.00163 * rf1617_dpk[-4] + 0.00178 * rf1617_dpk[-5]
                (4.17464) (4.69274)
    Polynomial lags:
                rf1617_dpk
                from 0 to 5 degree 3 near far
\begin{tabular}{lrlllrlr} 
Sum Sq & 0.0354 & Std Err & 0.0153 & LHS Mean & 0.0006 & Res Mean & 0.0030 \\
R Sq & 0.2161 & R Bar Sq & 0.2107 & F 2,146 & 20.1227 & \%RMSE & 28883.0 \\
D.W.( 1) & 0.4141 & D.W. ( 4) & 0.9322 & & & &
\end{tabular}
DPF1819_FE = (-0.00124 * RF1819_FE - 0.00147 * RF1819_FE.1-0.00106 * RF1819_FE.2-0.00035 * RF1819_FE. 3 + 0.00030 *
RF1819_FE.4 + 0.00053 * RF1819_FE.5) - (- 0.00124 * RF1819-0.00147 * RF1819.1-0.00106 * RF1819.2 - 0.00035 * RF1819.3 + 0.00030 *
RF1819.4 + 0.00053 * RF1819.5)
    Restricted Ordinary Least Squares
    QUARTERLY data for 148 periods from 1971Q1 to 2007Q4
    Date: 2 SEP 2009
    pf1819_dpk
        = - 0.00124 * rf1819_dpk - 0.00147 * rf1819_dpk[-1]
                (4.35316) (5.03839)
                    -0.00106 * rf1819_dpk[-2] - 0.00035 * rf1819_dpk[-3]
                (7.08785) (2.38461)
                + 0.00030 * rf1819_dpk[-4] + 0.00053 * rf1819_dpk[-5]
                (1.03892) (1.89331)
    Polynomial lags:
                rf1819_dpk
                from 0 to 5 degree 3 near far
\begin{tabular}{llllllllr} 
Sum Sq & 0.0217 & Std Err & 0.0120 & LHS Mean & -0.0031 & Res Mean & 0.0021 \\
R Sq & 0.2645 & R Bar Sq & 0.2594 & F 2,146 & 26.2493 & \%RMSE & 53411.9
\end{tabular}
            D.W.( 1) 0.8392 D.W.( 4) 1.3288
DPF2024_FE = (-0.00087 * RF2024_FE - 0.00099 * RF2024_FE.1-0.00063 * RF2024_FE.2-0.00007 * RF2024_FE. 3 + 0.00041 *
RF2024_FE.4 + 0.00052 * RF2024_FE.5) - (- 0.00087 * RF2024-0.00099 * RF2024.1-0.00063 * RF2024.2 - 0.00007 * RF2024.3 + 0.00041
* RF2024.4 + 0.00052 * RF2024.5)
    Restricted Ordinary Least Squares
    QUARTERLY data for 148 periods from 1971Q1 to 2007Q4
    Date: 2 SEP 2009
```

```
pf2024_dpk
    = - 0.00087 * rf2024_dpk - 0.00099 * rf2024_dpk[-1]
            (1.95040)
                (2.16369)
            - 0.00063 * rf2024_dpk[-2] - 0.00007 * rf2024_dpk[-3]
        (2.80524) (0.30336)
        + 0.00041 * rf2024_dpk[-4] + 0.00052 * rf2024_dpk[-5]
        (0.90125) (1.17782)
```

Polynomial lags:
rf2024_dpk
from 0 to 5 degree 3 near far

| Sum Sq | 0.0212 | Std Err | 0.0120 | LHS Mean | -0.0017 | Res Mean | 0.0002 |
| :--- | ---: | :--- | :--- | :--- | :--- | :--- | ---: |
| R Sq | 0.0408 | R Bar Sq | 0.0343 | F 2,146 | 3.1076 | \%RMSE | 51970.8 |
| D.W.( 1) | 0.2424 | D.W. ( 4) | 0.5144 |  |  |  |  |

DPF2529_FE $=(-0.00056 *$ RF2529_FE $-0.00070 *$ RF2529_FE. $1-0.00057 *$ RF2529_FE. $2-0.00029 *$ RF2529_FE. $3-0.00002 *$
RF2529_FE. $4+0.00013$ * RF2529_FE.5) - (- 0.00056 * RF2529-0.00070 * RF2529.1-0.00057 * RF2529.2-0.00029 * RF2529.3-0.00002 *
RF2529.4 + 0.00013 * RF2529.5)
Restricted Ordinary Least Squares
QUARTERLY data for 101 periods from 1982 Q4 to 2007 Q4
Date: 2 SEP 2009
pf2529_dpk

Polynomial lags:
rf2529_dpk
from 0 to 5 degree 3 near far

| Sum Sq | 0.0129 | Std Err | 0.0110 | LHS Mean | 0.0014 | Res Mean | 0.0030 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | ---: |
| R Sq | 0.0265 | R Bar Sq | 0.0167 | F 2, 99 | 1.3467 | \%RMSE | 1908.57 |
| D.W.( 1) | 0.2611 | D.W. ( 4) | 0.6593 |  |  |  |  |

DPF3034_FE $=(-0.00081 *$ RF3034_FE $-0.00065 *$ RF3034_FE. $1+0.00005 *$ RF3034_FE. $2+0.00089 *$ RF3034_FE. $3+0.00143$ *
RF3034_-FE. $4+0.00128$ * RF3034_FE.5) $-(-0.00081$ * RF3034-0.00065 * RF3034.1 + 0.00005 * RF3034.2 + 0.00089 * RF3034.3 + 0.00143

* RF3034.4 + 0.00128 * RF3034.5)
Restricted Ordinary Least Squares
QUARTERLY data for 101 periods from 1982Q4 to 2007Q4
Date: 2 SEP 2009
pf3034_dpk
$=-0.00081$ * rf3034_dpk - 0.00065 * rf3034_dpk[-1]
(1.30466) (1.02194)
+0.00005 * rf3034_dpk[-2] + 0.00089 * rf3034_dpk[-3]
(0.16121) (2.80112)
+0.00143 * rf3034_dpk[-4] + 0.00128 * rf3034_dpk[-5]
(2.29055) (2.08640)
Polynomial lags:

```
    rf3034_dpk
    from 0 to 5 degree 3 near far
\begin{tabular}{lrlrlrlrr} 
Sum Sq & 0.0083 & Std Err & 0.0088 & LHS Mean & 0.0040 & Res Mean & 0.0026 \\
R Sq & -0.1326 & R Bar Sq & -0.1440 & F 2, 99 & NC & \%RMSE & 29049.2 \\
D.W.( 1) & 0.3221 & D.W. ( 4) & 0.8675 & & & & &
\end{tabular}
DPF3539_FE = (-0.00195 * RF3539_FE - 0.00216 * RF3539_FE.1-0.00128 * RF3539_FE.2 + 0.00002 * RF3539_FE. 3 + 0.00111 *
RF3539_FE.4 + 0.00132 * RF3539_FE.5) - (- 0.00195 * RF3539 - 0.00216 * RF3539.1-0.00128 * RF3539.2 + 0.00002 * RF3539.3 + 0.00111
* RF3539.4 + 0.00132 * RF3539.5)
    Restricted Ordinary Least Squares
        QUARTERLY data for 101 periods from 1982Q4 to 2007Q4
        Date: 2 SEP 2009
        pf3539_dpk
            = - 0.00195 * rf3539_dpk - 0.00216 * rf3539_dpk[-1]
                (2.96355) (3.19201)
                - 0.00128 * rf3539_dpk[-2] + 0.00002 * rf3539_dpk[-3]
                (3.74105) (0.06643)
                + 0.00111 * rf3539_dpk[-4] + 0.00132 * rf3539_dpk[-5]
                    (1.67998) (2.03681)
        Polynomial lags:
                rf3539_dpk
                from 0 to 5 degree 3 near far
\begin{tabular}{llllllll} 
Sum Sq & 0.0079 & Std Err & 0.0089 & LHS Mean & -0.0021 & Res Mean & -0.0007 \\
R Sq & 0.0801 & R Bar Sq & 0.0708 & F 2, 99 & 4.3116 & \%RMSE & 6918.68 \\
D.W.( 1) & 0.3914 & D.W. ( 4) & 0.5620 & & & &
\end{tabular}
DPF4044_FE \(=(-0.00026 *\) RF4044_FE \(-0.00050 *\) RF4044_FE. \(1-0.00068 *\) RF4044_FE. \(2-0.00076 *\) RF4044_FE. \(3-0.00070 *\) RF4044_FE. \(4-0.00046\) * RF4044_FE.5) - (- 0.00026 * RF4044-0.00050 * RF4044.1-0.00068 * RF4044.2-0.00076 * RF4044.3-0.00070 * RF4044.4-0.00046 * RF4044.5)
Restricted Ordinary Least Squares
QUARTERLY data for 101 periods from 1982Q4 to 2007Q4 Date: 2 SEP 2009
pf4044_dpk
```

```
\[
\begin{aligned}
& =-0.00026 \text { * rf4044_dpk - 0.00050 * rf4044_dpk[-1] } \\
& \text { (0.53052) (0.99291) } \\
& \text { - } 0.00068 \text { * rf4044_dpk[-2] - } 0.00076 \text { * rf4044_dpk[-3] } \\
& \text { (2.55091) (2.90760) } \\
& \text { - } 0.00070 \text { * rf4044_dpk[-4] - 0.00046 * rf4044_dpk[-5] } \\
& \text { (1.41817) (0.96194) }
\end{aligned}
\]
Polynomial lags:
rf4044_dpk
from 0 to 5 degree 3 near far
\begin{tabular}{lllllllr} 
Sum Sq & 0.0041 & Std Err & 0.0061 & LHS Mean & 0.0005 & Res Mean & 0.0021 \\
R Sq & 0.1306 & R Bar Sq & 0.1218 & F 2, 99 & 7.4378 & \%RMSE & 4872.89 \\
D.W.( 1) & 0.5681 & D.W. ( 4) & 0.8422 & & & &
\end{tabular}
DPF4549_FE \(=(-0.00076 *\) RF4549_FE \(-0.00070 *\) RF4549_FE. \(1-0.00018 *\) RF4549_FE. \(2+0.00049 *\) RF4549_FE. \(3+0.00096 *\) RF4549_FE. \(4+0.00091\) * RF4549_FE.5) - ( 0.00076 * RF4549-0.00070 * RF4549.1-0.00018 * RF4549.2 + 0.00049 * RF4549.3 + 0.00096 * RF4549.4 + 0.00091 * RF4549.5)
```

```
Restricted Ordinary Least Squares
    QUARTERLY data for 101 periods from 1982Q4 to 2007Q4
    Date: 2 SEP 2009
    pf4549_dpk
        = - 0.00076 * rf4549_dpk - 0.00070 * rf4549_dpk[-1]
                (1.20345) (1.08117)
                - 0.00018 * rf4549_dpk[-2] + 0.00049 * rf4549_dpk[-3]
                (0.50034) (1.36204)
                + 0.00096 * rf4549_dpk[-4] + 0.00091 * rf4549_dpk[-5]
                    (1.47572) (1.44905)
    Polynomial lags:
                rf4549_dpk
                from 0 to 5 degree 3 near far
\begin{tabular}{lrlrlrlrl} 
Sum Sq & 0.0046 & Std Err & 0.0063 & LHS Mean & 0.0028 & Res Mean & 0.0027 \\
R Sq & -0.1835 & R Bar Sq & -0.1954 & F 2, 99 & NC & \%RMSE & 316691
\end{tabular}
    D.W.( 1) 0.7137 D.W.( 4) 0.9879
DPF5054_FE = ( 0.00003 * RF5054_FE - 0.00011 * RF5054_FE.1-0.00032 * RF5054_FE.2-0.00051 * RF5054_FE.3 - 0.00059 *
RF5054_FE.4-0.00045 * RF5054_FE.5) - ( 0.00003 * RF5054-0.00011 * RF5054.1 - 0.00032 * RF5054.2 - 0.00051 * RF5054.3 - 0.00059 *
RF5054.4-0.00045 * RF5054.5)
Restricted Ordinary Least Squares
    QUARTERLY data for 101 periods from 1982Q4 to 2007Q4
    Date: 2 SEP 2009
    pf5054_dpk
```

```
        = 0.00003 * rf5054_dpk - 0.00011 * rf5054_dpk[-1]
```

        = 0.00003 * rf5054_dpk - 0.00011 * rf5054_dpk[-1]
                (0.05929) (0.18046)
                (0.05929) (0.18046)
                - 0.00032 * rf5054_dpk[-2] - 0.00051 * rf5054_dpk[-3]
                - 0.00032 * rf5054_dpk[-2] - 0.00051 * rf5054_dpk[-3]
                (1.13804) (1.96823)
                (1.13804) (1.96823)
                - 0.00059 * rf5054_dpk[-4] - 0.00045 * rf5054_dpk[-5]
                - 0.00059 * rf5054_dpk[-4] - 0.00045 * rf5054_dpk[-5]
                (1.04732) (0.80816)
                (1.04732) (0.80816)
    Polynomial lags:
        rf5054_dpk
        from 0 to 5 degree 3 near far
    | Sum Sq | 0.0054 | Std Err | 0.0073 | LHS Mean | -0.0004 | Res Mean | 0.0010 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | ---: |
| R Sq | 0.0704 | R Bar Sq | 0.0611 | F 2, 99 | 3.7511 | \%RMSE | 66278.7 |

D.W.( 1) 0.8578 D.W.( 4) 1.1467
DPF5559_FE = ( 0.00064 * RF5559_FE + 0.00041 * RF5559_FE.1-0.00029 * RF5559_FE.2 - 0.00107 * RF5559_FE.3 - 0.00155 *
RF5559_FE.4-0.00132 * RF5559_FE.5) - (0.00064 * RF5559 + 0.00041 * RF5559.1-0.00029 * RF5559.2 - 0.00107 * RF5559.3 - 0.00155 *
RF5559.4-0.00132 * RF5559.5)
Restricted Ordinary Least Squares
QUARTERLY data for 101 periods from 1982Q4 to 2007Q4
Date: 2 SEP 2009
pf5559_dpk
= 0.00064 * rf5559_dpk + 0.00041 * rf5559_dpk[-1]
(0.89335) (0.55546)
- 0.00029 * rf5559_dpk[-2] - 0.00107 * rf5559_dpk[-3]
(0.75840) (2.90227)
- 0.00155 * rf5559_dpk[-4] - 0.00132 * rf5559_dpk[-5]

```

Polynomial lags:
rf5559_dpk
from 0 to 5 degree 3 near far
\begin{tabular}{lrlllllr} 
Sum Sq & 0.0102 & Std Err & 0.0102 & LHS Mean & -0.0011 & Res Mean & 0.0005 \\
R Sq & 0.0769 & R Bar Sq & 0.0676 & F 2, 99 & 4.1224 & \%RMSE & 4599.06 \\
D.W.( 1) & 0.8791 & D.W. ( 4) & 0.8944 & & & &
\end{tabular}

DPF6064_FE \(=(0.00141 *\) RF6064_FE \(+0.00166 *\) RF6064_FE. \(1+0.00116 *\) RF6064_FE. \(2+0.00033 *\) RF6064_FE. \(3-0.00041 *\) RF6064_FE. \(4-0.00066\) * RF6064_FE.5) - ( 0.00141 * RF6064 + 0.00166 * RF6064.1 + 0.00116 * RF6064.2 + 0.00033 * RF6064.3 - 0.00041 * RF6064.4-0.00066 * RF6064.5)

Restricted Ordinary Least Squares
QUARTERLY data for 101 periods from 1982Q4 to 2007Q4
Date: 2 SEP 2009
pf6064_dpk
\(=0.00141\) * rf6064_dpk + 0.00166 * rf6064_dpk[-1] (1.84124) (2.11172) +0.00116 * rf6064_dpk[-2] + 0.00033 * rf6064_dpk[-3] (2.99557) (0.93319)
- 0.00041 * rf6064_dpk[-4] - 0.00066 * rf6064_dpk[-5] (0.54220) (0.88278)

Polynomial lags:
rf6064_dpk
from 0 to 5 degree 3 near far
\begin{tabular}{llllllll} 
Sum Sq & 0.0113 & Std Err & 0.0105 & LHS Mean & 0.0007 & Res Mean & -0.0017 \\
R Sq & 0.1074 & R Bar Sq & 0.0984 & F 2, 99 & 5.9541 & \%RMSE & 8386.58 \\
D.W.( 1) & 0.6636 & D.W. ( 4) & 0.8453 & & & &
\end{tabular}

DPF6569_FE \(=(0.00029 *\) RF6569_FE \(+0.00014 *\) RF6569_FE. \(1-0.00023 *\) RF6569_FE. \(2-0.00063 *\) RF6569_FE. \(3-0.00086 *\) RF6569_FE. \(4-0.00072\) * RF6569_FE.5) - ( 0.00029 * RF6569 + 0.00014 * RF6569.1-0.00023 * RF6569.2-0.00063 * RF6569.3-0.00086 * RF6569.4-0.00072 * RF6569.5)
\begin{tabular}{|c|c|c|c|c|c|c|c|}
\hline \multicolumn{8}{|l|}{Restricted Ordinary Least Squares} \\
\hline \multicolumn{8}{|l|}{QUARTERLY data for 101 periods from 1982Q4 to 2007Q4} \\
\hline \multicolumn{8}{|l|}{Date: 2 SEP 2009} \\
\hline \multicolumn{8}{|l|}{pf6569_dpk} \\
\hline \multicolumn{8}{|l|}{\[
\begin{gathered}
=\begin{array}{c}
0.00029 \\
(0.39369)
\end{array}{ }^{*} \text { rf6569_dpk }+\underset{(0.00014}{(0.18572)}{ }^{*} \text { rf6569_dpk[-1] }
\end{gathered}
\]} \\
\hline \multicolumn{8}{|c|}{\[
\begin{gathered}
-0.00023 \text { * rf6569_dpk[-2] }-\underset{(1.29792)}{-0.00063} \text { * rf6569_dpk[-3] }
\end{gathered}
\]} \\
\hline \multicolumn{8}{|c|}{\[
\begin{gathered}
-0.00086 \text { * rf6569_dpk[-4] }-\underset{(0.00072}{(1.10980)} \text { * rf6569_dpk[-5] }
\end{gathered}
\]} \\
\hline \multicolumn{8}{|l|}{Polynomial lags:} \\
\hline \multicolumn{8}{|c|}{rf6569_dpk} \\
\hline Sum Sq & 0.0161 & Std Err & 0.0097 & LHS Mean & -0.0088 & Res Mean & -0.0083 \\
\hline R Sq & -0.8729 & R Bar Sq & -0.8918 & F 2, 99 & NC & \%RMSE & 5935.52 \\
\hline D.W.( 1) & 0.3864 & D.W.( 4) & 0.3951 & & & & \\
\hline
\end{tabular}

DPF7074_FE \(=(-0.00009 *\) RF7074_FE - 0.00028 * RF7074_FE.1-0.00048 * RF7074_FE. \(2-0.00063\) * RF7074_FE. \(3-0.00064\) * RF7074_FE. \(4-0.00046\) * RF7074_FE.5) - (- 0.00009 * RF7074-0.00028 * RF7074.1-0.00048 * RF7074.2-0.00063 * RF7074.3-0.00064 * RF7074.4-0.00046 *RF7074.5)
 Polynomial lags:
rf7074_dpk
from 0 to 5 degree 3 near far
\begin{tabular}{lrlrlrrrr} 
Sum Sq & 0.0038 & Std Err & 0.0059 & LHS Mean & -0.0043 & Res Mean & -0.0018 \\
R Sq & -0.1504 & R Bar Sq & -0.1620 & F 2, 99 & NC & \%RMSE & 14566.3 \\
D.W.( 1) & 0.7462 & D.W. ( 4) & 1.0245 & & & & &
\end{tabular}

DPF75O_FE \(=(-0.00007 *\) RF75O_FE \(-0.00012 *\) RF75O_FE. \(1-0.00015 *\) RF75O_FE. \(2-0.00014\) * RF75O_FE. \(3-0.00012\) *RF75O_FE. \(4-\) 0.00007 * RF75O_FE.5) - (- 0.00007 * RF75O-0.00012 * RF75O.1-0.00015 * RF75O. \(2-0.00014\) * RF75O. 3-0.00012 * RF75O.4-0.00007 * RF75O.5)
```

Restricted Ordinary Least Squares
QUARTERLY data for 101 periods from 1982Q4 to 2007Q4
Date: 2 SEP 2009
pf750_dpk
= - 0.00007 * rf75o_dpk - 0.00012 * rf75o_dpk[-1]
(0.83244) (1.28137)
- 0.00015 * rf750_dpk[-2] - 0.00014 * rf75o_dpk[-3]
(2.41595) (2.48652)
- 0.00012 * rf75o_dpk[-4] - 0.00007 * rf75o_dpk[-5]
(1.27553) (0.79882)
Polynomial lags:
rf75o_dpk
from 0 to 5 degree 3 near far

| Sum Sq | 0.0008 | Std Err | 0.0028 | LHS Mean | 0.0001 | Res Mean | 0.0004 |
| :--- | ---: | :--- | :--- | :--- | :--- | :--- | :--- |
| R Sq | 0.0818 | R Bar Sq | 0.0725 | F 2, 99 | 4.4079 | \%RMSE | 43067.0 |
| D.W.( 1) | 0.7531 | D.W. ( 4) | 1.3841 |  |  |  |  |

```

\subsection*{2.2 U.S. Earnings (USEAR)}

Annual Employment Equations
NonagriculturalWage Workers, Private Household Workers
\begin{tabular}{|c|c|}
\hline EF1617NAWPH = & MAX (0.001, -0.20802 * RTP - 0.40988 * RTP. \(1+0.01015+61.2465\) * 1/YEAR - 0.00965 * MINW/CPIW_U + 0.01561 * NU10/NF1617-0.13398) * EF1617 \\
\hline EF1819NAWPH = & MAX ( \(0.001,-0.03363\) * RTP - 0.12989 * RTP. \(1-0.00661+8.44701\) * 1/YEAR - 0.00539 * MINW/CPIW_U +0.00345 * NU10/NF1819 + 0.07597) * EF1819 \\
\hline EF2024NAWPH = & MAX ( \(0.001,-0.18707\) * MOVAVG (5, RTP) \(-0.00223+2.12060\) * 1/YEAR + 0.00820 * NU10/NF2024 + 0.14537) * EF2024 \\
\hline EF2534NAWPH = & \[
\begin{aligned}
& \text { MAX }(0.001,0.01874 * \text { RTP - } 0.04167 * \text { MOVAVG (5, RTP.1) }-0.00090+1.55167 * 1 / \text { YEAR }+0.01021 * \\
& \text { NU10/NF2534 - } 0.00170) * \text { EF2534 }
\end{aligned}
\] \\
\hline EF3544NAWPH = & \[
\begin{aligned}
& (0.00622 * \text { RTP }-0.06062 * \text { MOVAVG }(5, \text { RTP. } 1)+0.00008+0.29372 * \text { MOVAVG }(3, \text { EF2534NAWPH.9/EF2534.9 })+ \\
& 0.06187) * \text { EF3544 }
\end{aligned}
\] \\
\hline EF4554NAWPH = & \[
\begin{aligned}
& (0.02788 * \text { RTP }-0.10996 * \text { MOVAVG }(5, \text { RTP.1) }-0.00349+0.53068 * \text { MOVAVG }(3, \text { EF3544NAWPH.9/EF3544.9) }+ \\
& 0.08883) * \text { EF4554 }
\end{aligned}
\] \\
\hline EF5564NAWPH = & \((0.05939\) * RTP -0.10618 * MOVAVG (2, RTP.1) - \(0.00579+0.66195\) * MOVAVG (3, EF4554NAWPH.9/EF4554.9) +
\(0.05966) ~ * ~ E F 5564 ~\) \\
\hline EF65ONAWPH & \((0.22642\) * RTP - \(0.02069+0.33505\) * MOVAVG (3, EF5564NAWPH.9) - 0.19707) \\
\hline EM1617NAWPH = & MAX ( \(0.001,-0.05284\) * RTP - 0.17833 * RTP. \(1-0.00768+9.19738\) * 1/YEAR - 0.00588 * MINW/CPIW_U + 0.16862) * EM1617 \\
\hline EM1819NAWPH = & MAX ( \(0.001,-0.07122\) * RTP - 0.03737 * RTP. \(1-0.00282+3.76796\) * 1/YEAR - 0.00499 * MINW/CPIW_U + 0.08727) * EM1819 \\
\hline EM2024NAWPH = & MAX (0.001, -0.00450 * RTP - 0.02345 * RTP.1-0.00113-0.00057 * MINW/CPIW_U + 0.03265) * EM2024 \\
\hline EM2534NAWPH = & MAX (0.001, -0.00490 * RTP.1-0.00054-0.00051 * MINW/CPIW_U + 0.00789) * EM2534 \\
\hline EM3544NAWPH = & (-0.00446 * RTP. \(1-0.00041-0.00053\) * MINW/CPIW_U + 0.00726) * EM3544 \\
\hline EM4554NAWPH = & \((-0.00039+0.00129) *\) EM4554 \\
\hline EM5564NAWPH = & \((-0.00015+0.00200) *\) EM5564 \\
\hline EM65ONAWPH = & \((-0.00679+0.64405\) * MOVAVG (3, EM5564NAWPH.9) + 0.00231) \\
\hline
\end{tabular}

Nonagricultural Self-employed workers
EF1617NAS \(=(0.12015 *\) RTP.1-0.10551) * EF1617
Ordinary Least Squares
ANNUAL data for 5 periods from 2000 to 2004
Date: 9 NOV 2005
\begin{tabular}{|c|c|c|c|}
\hline ef1617nas/ef1617 & 17 = & \[
\begin{aligned}
& 0.12015 \text { * rtp. } 1 \text { - } \\
& (1.96868)
\end{aligned}
\] & \[
\begin{aligned}
& 0.10551 \\
& (1.73441)
\end{aligned}
\] \\
\hline Sum Sq 0 & 0.0000 & & \\
\hline Std Error 0 & 0.0030 & & \\
\hline LHS Mean 0 & 0.0142 & & \\
\hline R-Squared 0 & 0.5637 & & \\
\hline R Bar Squared 0 & 0.4182 & & \\
\hline F-stat 1, 3 & 3.8757 & & \\
\hline D.W. (1) 1.5 & 1.5620 & & \\
\hline D.W. (2) 2 & 2.3626 & & \\
\hline \multicolumn{4}{|l|}{EF1819NAS \(=(0.11184 *\) RTP. \(1-0.10241) *\) EF1819} \\
\hline \multicolumn{4}{|l|}{Ordinary Least Squares} \\
\hline \multicolumn{4}{|l|}{ANNUAL data for 5 periods from 2000 to 2004} \\
\hline \multicolumn{4}{|l|}{Date: 9 NOV 2005} \\
\hline ef1819nas/ef1819 & 19 = & \[
\begin{aligned}
& 0.11184 \text { * rtp. } 1- \\
& (2.99537)
\end{aligned}
\] & \[
\begin{aligned}
& 0.10241 \\
& (2.75170)
\end{aligned}
\] \\
\hline Sum Sq 0 & 0.0000 & & \\
\hline Std Error 0 & 0.0018 & & \\
\hline LHS Mean 0. & 0.0090 & & \\
\hline R-Squared 0 & 0.7494 & & \\
\hline R Bar Squared 0 & 0.6659 & & \\
\hline F-STAT 1, 3 & 8.9722 & & \\
\hline D.W. (1) 3 & 3.2586 & & \\
\hline D.W. (2) 0 & 0.9766 & & \\
\hline EF2024NAS = & (0.0890 & *TP.1-0.07176) & F2024 \\
\hline
\end{tabular}
\begin{tabular}{|c|c|c|c|}
\hline \multicolumn{4}{|l|}{Ordinary Least Squares} \\
\hline \multicolumn{4}{|l|}{ANNUAL data for 5 periods from 2000 to 2004} \\
\hline \multicolumn{4}{|l|}{Date: 9 NOV 2005} \\
\hline ef2024nas/ef2024 & \(024=\) & \[
\begin{aligned}
& 0.08908 * \text { rtp. } 1- \\
& (2.54605)
\end{aligned}
\] & \[
\begin{aligned}
& 0.07176 \\
& (2.05763)
\end{aligned}
\] \\
\hline \multicolumn{4}{|l|}{Sum Sq 0.0000} \\
\hline \multicolumn{4}{|l|}{Std Error 0.0017} \\
\hline \multicolumn{4}{|l|}{LHS Mean 0.0170} \\
\hline \multicolumn{4}{|l|}{R-Squared 0.6836} \\
\hline \multicolumn{4}{|l|}{R Bar Squared 0.5782} \\
\hline \multicolumn{4}{|l|}{F-STAT 1, 36.4824} \\
\hline \multicolumn{4}{|l|}{D.W. (1) 2.6600} \\
\hline \multicolumn{4}{|l|}{D.W. (2) 1.5247} \\
\hline \multicolumn{4}{|l|}{EF2534NAS \(=(0.00906 *\) RTP. \(1+0.03539) *\) EF2534} \\
\hline \multicolumn{4}{|l|}{Ordinary Least Squares} \\
\hline \multicolumn{4}{|l|}{ANNUAL data for 5 periods from 2000 to 2004} \\
\hline \multicolumn{4}{|l|}{Date: 9 NOV 2005} \\
\hline ef2534nas/ef2534 & \(534=\) & \[
\begin{aligned}
& 0.00906 \text { * rtp. } 1+ \\
& (0.34277)
\end{aligned}
\] & \[
\begin{aligned}
& 0.03539 \\
& (1.34366)
\end{aligned}
\] \\
\hline \multicolumn{4}{|l|}{Sum Sq 0.0000} \\
\hline \multicolumn{4}{|l|}{Std Error 0.0013} \\
\hline \multicolumn{4}{|l|}{LHS Mean 0.0444} \\
\hline \multicolumn{4}{|l|}{R-Squared 0.0377} \\
\hline \multicolumn{4}{|l|}{R Bar Squared 0.2831} \\
\hline \multicolumn{4}{|l|}{F-STAT 1, 30.1175} \\
\hline \multicolumn{4}{|l|}{D.W. (1) 3.0818} \\
\hline \multicolumn{4}{|l|}{D.W. (2) 1.1094} \\
\hline \multicolumn{4}{|l|}{EF3544NAS \(=(-0.01869 *\) RTP. \(1+0.08087) *\) EF3544} \\
\hline \multicolumn{4}{|l|}{Ordinary Least Squares} \\
\hline \multicolumn{4}{|l|}{ANNUAL data for 5 periods from 2000 to 2004} \\
\hline \multicolumn{4}{|l|}{Date: 9 NOV 2005} \\
\hline \multicolumn{2}{|l|}{ef3544nas/ef3544 =} & \[
\begin{aligned}
& -0.01869 * \text { rtp. } 1+ \\
& (0.70565)
\end{aligned}
\] & \[
\begin{aligned}
& 0.08087 \\
& (3.06320)
\end{aligned}
\] \\
\hline \multicolumn{4}{|l|}{Sum Sq 0.0000} \\
\hline \multicolumn{4}{|l|}{Std Error 0.0013} \\
\hline \multicolumn{4}{|l|}{LHS Mean 0.0622} \\
\hline \multicolumn{4}{|l|}{R-Squared 0.1424} \\
\hline \multicolumn{4}{|l|}{R Bar Squared 0.1435} \\
\hline \multicolumn{4}{|l|}{F-STAT 1, 30.4979} \\
\hline \multicolumn{4}{|l|}{D.W. (1) 2.2440} \\
\hline \multicolumn{4}{|l|}{D.W. (2) 2.1852} \\
\hline \multicolumn{4}{|l|}{EF4554NAS \(=(0.07232 *\) RTP. \(1-0.00701) *\) EF4554} \\
\hline \multicolumn{4}{|l|}{Ordinary Least Squares} \\
\hline \multicolumn{4}{|l|}{ANNUAL data for 5 periods from 2000 to 2004} \\
\hline \multicolumn{4}{|l|}{Date: 9 NOV 2005} \\
\hline ef4554nas/ef4554 & \(554=\) & \[
\begin{aligned}
& 0.07232 * \text { rtp. } 1- \\
& (2.86756)
\end{aligned}
\] & \[
\begin{aligned}
& 0.00701 \\
& (0.27876)
\end{aligned}
\] \\
\hline \multicolumn{4}{|l|}{Sum Sq 0.0000} \\
\hline \multicolumn{4}{|l|}{Std Error 0.0012} \\
\hline \multicolumn{4}{|l|}{LHS Mean 0.0651} \\
\hline \multicolumn{4}{|l|}{R-Squared 0.7327} \\
\hline \multicolumn{4}{|l|}{R Bar Squared 0.6436} \\
\hline \multicolumn{4}{|l|}{F-STAT 1, 38.2229} \\
\hline \multicolumn{4}{|l|}{D.W. (1) 1.7821} \\
\hline \multicolumn{4}{|l|}{D.W. (2) 2.7029} \\
\hline \multicolumn{4}{|l|}{\multirow[t]{3}{*}{\begin{tabular}{l}
EF5564NAS \(=(0.07872 *\) RTP. \(1+0.00466) *\) EF5564 Ordinary Least Squares \\
ANNUAL data for 5 periods from 2000 to 2004
\end{tabular}}} \\
\hline & & & \\
\hline & & & \\
\hline
\end{tabular}

Date: 9 NOV 2005
\(\left.\begin{array}{lll}\text { ef5564nas/ef5564 } & \begin{array}{l}0.07872 \\ (1.38159)\end{array} \\ \\ \text { rtp. } 1\end{array}+\begin{array}{l}0.00466 \\ (0.08196)\end{array}\right)\)

Ordinary Least Squares
ANNUAL data for 5 periods from 2000 to 2004
Date: 9 NOV 2005
\begin{tabular}{rl} 
ef7074nas/ef7074 \(=\quad\) & 0.12265 \\
\((16.4939)\)
\end{tabular}
\begin{tabular}{ll} 
Sum Sq & 0.0011 \\
Std Error & 0.0166 \\
LHS Mean & 0.1226 \\
R-Squared & 0.0000 \\
R Bar Squared & 0.0000 \\
F 0, 4 & NC \\
D.W. (1) & 1.0289 \\
D.W. (2) & 1.7188
\end{tabular}

Ordinary Least Squares
ANNUAL data for 5 periods from 2000 to 2004
Date: 9 NOV 2005

\begin{tabular}{ll} 
Sum Sq & 0.0000 \\
Std Error & 0.0022 \\
LHS Mean & 0.0203 \\
R-Squared & 0.8961 \\
R Bar Squared & 0.8614 \\
F-STAT 1, & 25.8611 \\
D.W. (1) & 2.4658 \\
D.W. (2) & 1.6839
\end{tabular}

EM1819NAS \(=(-0.05782 *\) RTP. \(1+0.07265) *\) EM1819
Ordinary Least Squares
ANNUAL data for 5 periods from 2000 to 2004
Date: 9 NOV 2005
\begin{tabular}{lrl} 
em1819nas/em1819 \(=\) & \begin{tabular}{l}
-0.05782 \\
\((3.43044)\)
\end{tabular} & rtp. \(1+\)\begin{tabular}{l}
0.07265 \\
\((4.32458)\)
\end{tabular} \\
& & \\
Sum Sq & 0.0000 & \\
Std Error & 0.0008 & \\
LHS Mean & 0.0150 & \\
R-Squared & 0.7969 & \\
R Bar Squared & 0.7291 & \\
F-STAT 1, 3 & 11.7679 & \\
D.W. (1) & 3.3262 & \\
D.W. (2) & 1.0399 &
\end{tabular}

EM2024NAS \(=(-0.09206 *\) RTP. \(1+0.11567) *\) EM2024
Ordinary Least Squares
ANNUAL data for 5 periods from 2000 to 2004
Date: 9 NOV 2005
\begin{tabular}{|c|c|c|c|}
\hline em2024nas/em & 2024 & \[
\begin{aligned}
& -0.09206 \text { * rtp. } 1+ \\
& (2.44839)
\end{aligned}
\] & \[
\begin{aligned}
& 0.11567 \\
& (3.08618)
\end{aligned}
\] \\
\hline Sum Sq & 0.0000 & & \\
\hline Std Error & 0.0018 & & \\
\hline LHS Mean & 0.0239 & & \\
\hline R-Squared & 0.6665 & & \\
\hline R Bar Squared & 0.5553 & & \\
\hline F-STAT 1, 3 & 5.9946 & & \\
\hline D.W. (1) & 2.1493 & & \\
\hline D.W. (2) & 1.7046 & & \\
\hline \multicolumn{4}{|l|}{\multirow[t]{4}{*}{\begin{tabular}{l}
EM2534NAS \(=(-0.09661 *\) RTP. \(1+0.14843) *\) EM \\
Ordinary Least Squares \\
ANNUAL data for 5 periods from 2000 to 2004 \\
Date: 9 NOV 2005
\end{tabular}}} \\
\hline & & & \\
\hline & & & \\
\hline & & & \\
\hline \multicolumn{2}{|l|}{em2534nas/em2534} & \[
\begin{aligned}
& -0.09661 \text { * rtp. } 1+ \\
& (2.81478)
\end{aligned}
\] & \[
\begin{aligned}
& 0.14843 \\
& (4.33847)
\end{aligned}
\] \\
\hline \multicolumn{4}{|l|}{Sum Sq 0.0000} \\
\hline \multicolumn{4}{|l|}{Std Error 0.0017} \\
\hline \multicolumn{4}{|l|}{LHS Mean 0.0522} \\
\hline \multicolumn{4}{|l|}{R-Squared 0.7254} \\
\hline \multicolumn{4}{|l|}{R Bar Squared 0.6338} \\
\hline \multicolumn{4}{|l|}{F-STAT 1, 37.9230} \\
\hline \multicolumn{4}{|l|}{D.W. (1) 1.8300} \\
\hline \multicolumn{4}{|l|}{D.W. (2) 2.9632} \\
\hline \multicolumn{4}{|l|}{EM3544NAS \(=(0.02739 *\) RTP. \(1+0.05236) *\) EM3544} \\
\hline \multicolumn{4}{|l|}{\multirow[t]{2}{*}{\begin{tabular}{l}
Ordinary Least Squares \\
ANNUAL data for 5 periods from 2000 to 2004
\end{tabular}}} \\
\hline & & & \\
\hline \multicolumn{4}{|l|}{Date: 9 NOV 2005} \\
\hline \multicolumn{2}{|l|}{em3544nas/em3544} & \[
\begin{aligned}
& 0.02739 \text { * rtp. } 1+ \\
& (0.61129)
\end{aligned}
\] & \[
\begin{aligned}
& 0.05236 \\
& (1.17241)
\end{aligned}
\] \\
\hline
\end{tabular}

Sum Sq 0.0000
\begin{tabular}{lr} 
Std Error & 0.0022 \\
LHS Mean & 0.0797 \\
R-Squared & 0.1108 \\
R Bar Squared & 0.1857 \\
F-STAT 1, 3 & 0.3737 \\
D.W. (1) & 2.5508 \\
D.W. (2) & 2.2676
\end{tabular}

EM4554NAS \(=(0.06217 *\) RTP. \(1+0.03411) *\) EM4554
Ordinary Least Squares
ANNUAL data for 5 periods from 2000 to 2004
Date: 9 NOV 2005
\begin{tabular}{|c|c|c|c|}
\hline em4554nas/em & 4554 & \[
\begin{aligned}
& 0.06217 * \text { rtp. } 1+ \\
& (1.91738)
\end{aligned}
\] & \[
\begin{aligned}
& 0.03411 \\
& (1.05544)
\end{aligned}
\] \\
\hline Sum Sq & 0.0000 & & \\
\hline Std Error & 0.0016 & & \\
\hline LHS Mean & 0.0961 & & \\
\hline R-Squared & 0.5507 & & \\
\hline R Bar Squared & 0.4009 & & \\
\hline F-STAT 1, 3 & 3.6764 & & \\
\hline D.W. (1) & 2.5497 & & \\
\hline D.W. (2) & 1.5554 & & \\
\hline \multicolumn{4}{|l|}{EM5564NAS \(=(-0.04776 *\) RTP. \(1+0.16626) *\) EM5564} \\
\hline \multicolumn{4}{|l|}{Ordinary Least Squares} \\
\hline \multicolumn{4}{|l|}{ANNUAL data for 5 periods from 2000 to 2004} \\
\hline \multicolumn{4}{|l|}{Date: 9 NOV 2005} \\
\hline \multicolumn{2}{|l|}{em5564nas/em5564} & \[
\begin{aligned}
& -0.04776 * \text { rtp. } 1+ \\
& (0.60480)
\end{aligned}
\] & \[
\begin{aligned}
& 0.16626 \\
& (2.11226)
\end{aligned}
\] \\
\hline Sum Sq & 0.0000 & & \\
\hline Std Error & 0.0039 & & \\
\hline LHS Mean & 0.1187 & & \\
\hline R-Squared & 0.1087 & & \\
\hline R Bar Squared & 0.1884 & & \\
\hline F-STAT 1, 3 & 0.3658 & & \\
\hline D.W. (1) & 2.9234 & & \\
\hline D.W. (2) & 1.9432 & & \\
\hline
\end{tabular}

EM65ONAS \(=(0.16527 *\) EM6569 \(+0.17798 *\) EM7074 \(+0.19058 *\) EM75O \()\)
Ordinary Least Squares
ANNUAL data for 5 periods from 2000 to 2004
Date: 9 NOV 2005
\begin{tabular}{lrl} 
em6569nas/em6569 & \(=\)\begin{tabular}{l}
0.16527 \\
\((53.9126)\)
\end{tabular} \\
Sum Sq & 0.0002 & \\
Std Error & 0.0069 & \\
LHS Mean & 0.1653 & \\
R-Squared & 0.0000 & \\
R Bar Squared & 0.0000 & \\
F-stat 0, 4 & NC & \\
D.W. (1) & 1.7716 & \\
D.W. (2) & 2.9645 &
\end{tabular}

Ordinary Least Squares
ANNUAL data for 5 periods from 2000 to 2004
Date: 9 NOV 2005
em7074nas/em7074 \(=0.17798\)
(22.9265)

Sum Sq 0.0012
Std Error 0.0174
LHS Mean 0.1780
R-Squared 0.0000
R Bar Squared 0.0000
\begin{tabular}{ll} 
F-stat 0, 4 & NC \\
D.W. (1) & 1.7116 \\
D.W. (2) & 2.1991
\end{tabular}

Ordinary Least Squares
ANNUAL data for 5 periods from 2000 to 2004
Date: 9 NOV 2005
\begin{tabular}{lll} 
em75onas/em75o & \(=\) & \begin{tabular}{l}
0.19058 \\
\((20.1892)\)
\end{tabular} \\
& & \\
Sum Sq & 0.0018 & \\
Std Error & 0.0211 & \\
LHS Mean & 0.1906 & \\
R-Squared & 0.0000 & \\
R Bar Squared & 0.0000 \\
F-stat 0, 4 & NC & \\
D.W. (1) & 2.7330 & \\
D.W. (2) & 0.9992 &
\end{tabular}

Nonagricultural Unpaid Family Workers
EF1617NAU \(=0.00012 *\) ENAS
EF1819NAU \(=0.00025 *\) ENAS
EF2024NAU \(=0.00024 *\) ENAS
EF2534NAU \(=0.00117\) * ENAS
EF3544NAU \(=0.00218 *\) ENAS
EF4554NAU \(=0.00226 *\) ENAS
EF5564NAU \(=0.00083 *\) ENAS
EF65ONAU \(=(0.00027+0.00021+0.00008) *\) ENAS
EM1617NAU \(=0.00028 *\) ENAS
EM1819NAU \(=0.00033 *\) ENAS
EM2024NAU \(=0.00050 *\) ENAS
EM2534NAU \(=0.00044 *\) ENAS
EM3544NAU \(=0.00043 *\) ENAS
EM4554NAU \(=0.00052 *\) ENAS
EM5564NAU \(=0.00037 *\) ENAS
EM65ONAU \(=(0.00023+0.00010+0.00011) *\) ENAS

Agricultural Wage Workers
```

EM1617AW = EAW * (-0.00594-0.09353* RTP + 5.28754*EM1617/E + 0.08116)
EM1819AW = EAW * (-0.00131-0.18120* RTP + 3.87151*EM1819/E + 0.16636)
EM2024AW = EAW * (-0.00664 + 0.10493 * RTP + 2.00153 * EM2024/E - 0.08191)
EM2534AW = EAW * (-0.02065 + 0.38358* RTP - 0.98380*EM2534/E + 0.00751)
EM3544AW = EAW * (0.00402-0.15663 * RTP + 1.72119 * EM3544/E + 0.05679)
EM4554AW = EAW * (0.00834 + 0.03746 * RTP + 0.46522 * EM4554/E + 0.00144)
EM5564AW = EAW * (-0.00655 + 0.03521*RTP + 0.46852 * EM5564/E - 0.00037)
EM65OAW = EAW * (-0.00114 + 0.07640 * RTP + 3.25911* EM65O/E - 0.10058)
EF1617AW = EAW * (-0.00055-0.05470 * RTP + 1.41760 * EF1617/E + 0.04979)
EF1819AW = EAW * (0.00102-0.07375 * RTP + 0.78394*EF1819/E + 0.07226)
EF2024AW = EAW * (0.00112-0.05971 * RTP + 0.57256 * EF2024/E + 0.05907)
EF2534AW = EAW * (0.00623 + 0.08868*RTP + 1.00897 * EF2534/E - 0.15142)
EF3544AW = EAW * (0.00687-0.00259 * RTP + 0.51319 * EF3544/E - 0.00937)
EF4554AW = EAW * (0.00185 + 0.08747 * RTP + 0.28022 * EF4554/E - 0.08053)
EF5564AW = EAW * (-0.00140-0.03001*RTP - 0.59383 * EF5564/E + 0.07088)
EF65OAW }=\mathrm{ EAW * (0.00096 + 0.06768*RTP + 1.04213 * EF65O/E - 0.07359)

```

Agricultural Self-employed Workers
```

EM1617AS = NM1617 * (0.00528 + 0.00404)
EM1819AS = NM1819 * (0.00309 + 0.28448 * EA / (NM16O+ NF16O) - 0.00165)
EM2024AS }=\mathrm{ NM2024* (-0.00181 + 0.97958*EA / (NM16O+ NF16O) - 0.01093)
EM2534AS = NM2534* (-0.00263 + 1.23186 * EA / (NM16O+ NF16O) - 0.01021)
EM3544AS = NM3544* (-0.00151 + 1.66765*EA / (NM16O+NF16O) - 0.01450)

```
\begin{tabular}{|c|c|}
\hline EM4554AS = & NM4554 * (-0.00381 + 2.86654 * EA / (NM16O+ NF16O) - 0.03175) \\
\hline EM5564AS = & NM5564 * (-0.00460 + 2.78817 * EA / (NM16O+ NF16O) - 0.02398) \\
\hline EM650AS = & NM65O * (0.00079 + 1.76904 * EA / (NM16O+ NF16O) - 0.01437) \\
\hline EF1617AS & NF1617 * (0.00181 + 0.00030) \\
\hline EF1819AS = & EM1819AS * (-0.02393 + 0.63672 * RTP + 0.98791 * EF1819/EM1819-1.43926) \\
\hline EF2024AS = & EM2024AS * (0.07353-0.40207* RTP + 0.57572 * EF2024/EM2024-0.01117) \\
\hline EF2534AS = & EM2534AS * (0.16575 + 0.16967 * RTP + 0.55503 * EF2534/EM2534-0.43412) \\
\hline EF3544AS = & EM3544AS * (0.15848 + 0.37839 * RTP + 0.37764 * EF3544/EM3544-0.45362) \\
\hline EF4554AS = & EM4554AS * (0.21947 + 0.29497* RTP + 0.58974 * EF4554/EM4554-0.51966) \\
\hline EF5564AS = & EM5564AS * \(0.20892+0.36294 *\) RTP +0.65320 * EF5564/EM5564-0.66626) \\
\hline EF65OAS = & EM65OAS * (0.16242 + 0.54916 * RTP + 0.06199 * EF65O/EM65O-0.47556) \\
\hline
\end{tabular}

\(E N A=E-E A\)
Nonagricultural Wage Workers, Private Household Workers:
\begin{tabular}{ll} 
EF1617NAWPH_R \(=\) & MAX \((0.001,-0.20802 *\) MOVAVG (4, RTP.1) \(-0.40988 *\) MOVAVG \((4\), RTP. \()+0.01015+61.2465 * 1 / Y E A R-\) \\
EF1819NAWPH_R \(=\) & \(0.00965 *\) MINW/CPIW_U + 0.01561 * NU10/NF1617-0.13398) * EF1617 + EF1617NAWPH.ADJ \\
MAX \((0.001,-0.03363 *\) MOVAVG (4, RTP.1) \(-0.12989 *\) MOVAVG (4, RTP.5) \(-0.00661+8.44701 * 1 / Y E A R-\)
\end{tabular}
    0.00539 * MINW/CPIW_U + 0.00345 * NU10/NF1819 + 0.07597) * EF1819 + EF1819NAWPH.ADJ
EF2024NAWPH_R \(=\quad\) MAX \((0.001,-0.18707 *\) MOVAVG (20, RTP.1) \(-0.00223+2.12060 * 1 / Y E A R+0.00820 *\) NU10/NF2024 +
    0.14537 ) * EF2024 + EF2024NAWPH.ADJ
EF2534NAWPH_R \(=\quad\) MAX ( \(0.001,0.01874 *\) MOVAVG (4, RTP.1) \(-0.04167 *\) MOVAVG (20, RTP.5) \(-0.00090+1.55167 * 1 / Y E A R+\)
    0.01021 * NU10/NF2534-0.00170) * EF2534 + EF2534NAWPH.ADJ
EF3544NAWPH_R \(=(0.00622 *\) MOVAVG (4, RTP.1) \(-0.06062 *\) MOVAVG (20, RTP.5) \(+0.00008+0.29372 *\) MOVAVG (12,
    EF2534NAWPH.36/EF2534.36) + 0.06187) * EF3544 + EF3544NAWPH.ADJ
EF4554NAWPH_R \(=(0.02788 *\) MOVAVG (4, RTP.1) \(-0.10996 *\) MOVAVG (20, RTP.5) \(-0.00349+0.53068 *\) MOVAVG (12,
    EF3544NAWPH.36/EF3544.36) + 0.08883) * EF4554 + EF4554NAWPH.ADJ
EF5564NAWPH_R \(=\quad(0.05939 *\) MOVAVG (4, RTP.1) \(-0.10618 *\) MOVAVG (8, RTP.5) \(-0.00579+0.66195 *\) MOVAVG (12,
EF65ONAWPH_R \(=\quad(0.22642 *\) MOVAVG (4, RTP.1) \(-0.02069+0.33505 *\) MOVAVG (12, EF5564NAWPH.36) -0.19707\()+\)
        EF65ONAWPH.ADJ
EM1617NAWPH_R \(=\quad\) MAX \((0.001,-0.05284 * \operatorname{MOVAVG}(4\), RTP.1) \(-0.17833 * \operatorname{MOVAVG}(4\), RTP. \()-0.00768+9.19738 * 1 / Y E A R-\)
        0.00588 * MINW/CPIW_U + 0.16862) * EM1617+ EM1617NAWPH.ADJ
EM1819NAWPH_R \(=\quad\) MAX (0.001, \(-0.07122 *\) MOVAVG (4, RTP.1) - 0.03737 * MOVAVG (4, RTP. 5\()-0.00282+3.76796\) * 1/YEAR -
        0.00499 * MINW/CPIW_U + 0.08727) * EM1819+ EM1819NAWPH.ADJ
EM2024NAWPH_R \(=\quad\) MAX (0.001, -0.00450 * MOVAVG (4, RTP.1) - 0.02345 * MOVAVG (4, RTP.5) - 0.00113-0.00057 *
        MINW/CPIW_U + 0.03265) * EM2024 + EM2024NAWPH.ADJ
EM2534NAWPH_R = MAX (0.001, - 0.00490 * MOVAVG (4, RTP.5) - 0.00054-0.00051 * MINW/CPIW_U + 0.00789) * EM2534 +
        EM2534NAWPH.ADJ
EM3544NAWPH_R \(=(-0.00446 *\) MOVAVG (4, RTP.5) \(-0.00041-0.00053 *\) MINW/CPIW_U + 0.00726) \(*\) EM3544 +
        EM3544NAWPH.ADJ
EM4554NAWPH_R \(=(-0.00039+0.00129) *\) EM4554 + EM4554NAWPH.ADJ
EM5564NAWPH_R \(=(-0.00015+0.00200) *\) EM5564 + EM5564NAWPH.ADJ
EM65ONAWPH_R \(=(-0.00679+0.64405 *\) MOVAVG \((12\), EM5564NAWPH.36 \()+0.00231)+\) EM65ONAWPH.ADJ
ENAWPH_R = EF1617NAWPH_R + EF1819NAWPH_R + EF2024NAWPH_R + EF2534NAWPH_R + EF3544NAWPH_R +
    EF4554NAWPH_R + EF5564NAWPH_R + EF65ONAWPH_R + EM1617NAWPH_R + EM1819NAWPH_R +


Nonagricultural Self-employed Workers:
EF1617NAS_R \(=(0.12015 *\) RTP. \(1-0.10551) *\) EF1617 + EF1617NAS.ADJ
EF1819NAS_R \(=(0.11184 *\) RTP. \(1-0.10241) *\) EF1819 + EF1819NAS.ADJ
EF2024NAS_R \(=(0.08908 *\) RTP. \(1-0.07176) *\) EF2024 + EF2024NAS.ADJ
EF2534NAS_R \(=(0.00906 *\) RTP. \(1+0.03539) *\) EF2534 + EF2534NAS.ADJ
EF3544NAS_R \(=(-0.01869 *\) RTP. \(1+0.08087) *\) EF3544 + EF3544NAS.ADJ
EF4554NAS_R \(=(0.07232 *\) RTP. \(1-0.00701) *\) EF4554 + EF4554NAS.ADJ
EF5564NAS_R \(=(0.07872 *\) RTP. \(1+0.00466) *\) EF5564 + EF5564NAS.ADJ
EF65ONAS_R \(=(0.10940 *\) EF6569 \(+0.12265 *\) EF7074 \(+0.14137 *\) EF75O \()+\) EF65ONAS.ADJ
EM1617NAS_R \(=(-0.23035 *\) RTP. \(1+0.24985) *\) EM1617 + EM1617NAS.ADJ
EM1819NAS_R \(=(-0.05782 *\) RTP. \(1+0.07265) *\) EM1819 + EM1819NAS.ADJ
EM2024NAS_R \(=(-0.09206 *\) RTP. \(1+0.11567) *\) EM2024 + EM2024NAS.ADJ
EM2534NAS_R \(=(-0.09661 *\) RTP. \(1+0.14843) *\) EM2534 + EM2534NAS.ADJ
EM3544NAS_R \(=(0.02739 *\) RTP. \(1+0.05236) *\) EM3544 + EM3544NAS.ADJ
EM4554NAS_R \(=(0.06217 *\) RTP. \(1+0.03411) *\) EM4554 + EM4554NAS.ADJ
EM5564NAS_R \(=(-0.04776 *\) RTP. \(1+0.16626) *\) EM5564 + EM5564NAS.ADJ
EM65ONAS_R \(=(0.16527 *\) EM6569 \(+0.17798 *\) EM7074 \(+0.19058 *\) EM75O \()+\) EM65ONAS.ADJ

ENAS_R = EF1617NAS_R + EF1819NAS_R + EF2024NAS_R + EF2534NAS_R + EF3544NAS_R + EF4554NAS_R + EF5564NAS_R + EF65ONAS_R + EM1617NAS_R + EM1819NAS_R + EM2024NAS_R + EM2534NAS_R + EM3544NAS_R + EM4554NAS_R + EM5564NAS_R + EM65ONAS_R

ENAS \(=\) IF LONGRANGE \(=0\)
THEN ENAS_R
ELSE ENA * (ENAS.1/ENA.1)
\begin{tabular}{ll} 
EM1617NAS \(=\) & EM1617NAS_R \(*(\) ENAS/ENAS_R) \\
EM1819NAS \(=\) & EM1819NAS_R \(*(\) ENAS/ENAS_R) \\
EM2024NAS \(=\) & EM2024NAS_R \(*(\) ENAS/ENAS_R) \\
EM2534NAS \(=\) & EM2534NAS_R \(*(\) ENAS/ENAS_R) \\
EM3544NAS \(=\) & EM3544NAS_R \(*(\) ENAS/ENAS_R) \\
EM4554NAS \(=\) & EM4554NAS_R \(*(\) ENAS/ENAS_R) \\
EM5564NAS \(=\) & EM5564NAS_R \(*(\) ENAS/ENAS_R) \\
EM65ONAS \(=\) & EM65ONAS_R \(*(\) ENAS/ENAS_R) \\
& \\
EF1617NAS \(=\) & EF1617NAS_R \(*(\) ENAS/ENAS_R)
\end{tabular}
\begin{tabular}{ll} 
EF1819NAS \(=\) & EF1819NAS_R \(*(\) ENAS/ENAS_R \()\) \\
EF2024NAS \(=\) & EF2024NAS_R \(*(\) ENAS/ENAS_R \()\) \\
EF2534NAS \(=\) & EF2534NAS_R \(*(\) ENAS/ENAS_R \()\) \\
EF3544NAS \(=\) & EF3544NAS_R \(*(\) ENAS/ENAS_R \()\) \\
EF4554NAS \(=\) & EF4554NAS_R \(*(\) ENAS/ENAS_R \()\) \\
EF5564NAS \(=\) & EF5564NAS_R \(*(\) ENAS/ENAS_R \()\) \\
EF65ONAS \(=\) & EF65ONAS_R \(*(\) ENAS/ENAS_R \()\) \\
& \\
EFNAS \(=\) & EF1617NAS + EF1819NAS + EF2024NAS + EF2534NAS + EF3544NAS + EF4554NAS + EF5564NAS + EF65ONAS \\
EMNAS \(=\) & EM1617NAS + EM1819NAS + EM2024NAS + EM2534NAS + EM3544NAS + EM4554NAS + EM5564NAS + \\
& EM65ONAS
\end{tabular}

Nonagricultural Unpaid Family Workers
EF1617NAU_R = 0.00012 * ENAS + EF1617NAU.ADJ
EF1819NAU_R \(=0.00025 *\) ENAS + EF1819NAU.ADJ
EF2024NAU_R \(=0.00024\) * ENAS + EF2024NAU.ADJ
EF2534NAU_R \(=0.00117\) * ENAS + EF2534NAU.ADJ
EF3544NAU_R \(=0.00218 *\) ENAS + EF3544NAU.ADJ
EF4554NAU_R \(=0.00226\) * ENAS + EF4554NAU.ADJ
EF5564NAU_R \(=0.00083\) * ENAS + EF5564NAU.ADJ
EF65ONAU_R \(=(0.00027+0.00021+0.00008) *\) ENAS + EF65ONAU.ADJ
EM1617NAU_R \(=0.00028\) * ENAS + EM1617NAU.ADJ
EM1819NAU_R \(=0.00033 *\) ENAS + EM1819NAU.ADJ
EM2024NAU_R \(=0.00050\) * ENAS + EM2024NAU.ADJ
EM2534NAU_R \(=0.00044 *\) ENAS + EM2534NAU.ADJ
EM3544NAU_R \(=0.00043\) * ENAS + EM3544NAU.ADJ
EM4554NAU_R \(=0.00052 *\) ENAS + EM4554NAU.ADJ
EM5564NAU_R = 0.00037 * ENAS + EM5564NAU.ADJ
EM65ONAU_R \(=(0.00023+0.00010+0.00011) *\) ENAS + EM65ONAU.ADJ
ENAU_R \(=\) EF1617NAU_R + EF1819NAU_R + EF2024NAU_R + EF2534NAU_R + EF3544NAU_R + EF4554NAU_R + EF5564NAU_R + EF65ONAU_R + EM1617NAU_R + EM1819NAU_R + EM2024NAU_R + EM2534NAU_R + EM3544NAU_R + EM4554NAU_R + EM5564NAU_R + EM65ONAU_R

ENAU \(=\quad\) ENAU_R
```

EM1617NAU = EM1617NAU_R * (ENAU/ENAU_R)
EM1819NAU = EM1819NAU_R * (ENAU/ENAU_R)
EM2024NAU = EM2024NAU_R * (ENAU/ENAU_R)
EM2534NAU = EM2534NAU_R * (ENAU/ENAU_R)
EM3544NAU = EM3544NAU_R * (ENAU/ENAU_R)
EM4554NAU = EM4554NAU_R * (ENAU/ENAU_R)
EM5564NAU = EM5564NAU_R * (ENAU/ENAU_R)
EM65ONAU = EM65ONAU_R * (ENAU/ENAU_R)
EF1617NAU = EF1617NAU_R * (ENAU/ENAU_R)
EF1819NAU = EF1819NAU_R * (ENAU/ENAU_R)
EF2024NAU = EF2024NAU_R * (ENAU/ENAU_R)
EF2534NAU = EF2534NAU_R * (ENAU/ENAU_R)
EF3544NAU = EF3544NAU_R * (ENAU/ENAU_R)
EF4554NAU = EF4554NAU_R * (ENAU/ENAU_R)
EF5564NAU = EF5564NAU_R *(ENAU/ENAU_R)
EF65ONAU = EF65ONAU_R * (ENAU/ENAU_R)
EFNAU = EF1617NAU + EF1819NAU + EF2024NAU + EF2534NAU + EF3544NAU + EF4554NAU + EF5564NAU + EF65ONAU
EMNAU = EM1617NAU + EM1819NAU + EM2024NAU + EM2534NAU + EM3544NAU + EM4554NAU + EM5564NAU +
EM65ONAU

```

Agricultural Workers
\(\mathrm{EA}=\mathrm{IF}\) LONGRANGE \(=0\)
THEN GDPPF09 / (1.125 * 1.138 * EXP ( -0.20541 + 0.03254 * YEAR - 0.07829 + 0.37854\()\) )
ELSE E * EA.1/E. 1

Agricultural Wage Workers
\(\mathrm{EAW}=\quad \mathrm{IF}\) LONGRANGE \(=0\)
```

THEN EA * (0.00893 * YEAR + 0.33159 * RTP - 0.67943)

```
ELSE EA * (EAW.1/EA.1)

EM1617AW_R EM1819AW_R = EM2024AW_R = EM2534AW_R = EM3544AW_R = EM4554AW_R = EM5564AW_R = EM650AW_R =

EF1617AW_R = EF1819AW_R = EF2024AW_R = EF2534AW_R = EF3544AW_R = EF4554AW_R = EF5564AW_R = EF650AW_R =

MAX ( 0 , EAW * ( \(-0.00594-0.09353\) * MOVAVG (2, RTP.1) +5.28754 * EM1617/E + 0.08116) + EM1617AW.ADJ \()\) MAX (0, EAW * \((-0.00131-0.18120 *\) MOVAVG (2, RTP.1) \(+3.87151 *\) EM1819/E +0.16636\()+\) EM1819AW.ADJ) MAX (0, EAW * \((-0.00664+0.10493\) * MOVAVG (2, RTP.1) +2.00153 * EM2024/E - 0.08191) + EM2024AW.ADJ) MAX (0, EAW * \((-0.02065+0.38358 *\) MOVAVG (2, RTP.1) \(-0.98380 *\) EM2534/E + 0.00751) + EM2534AW.ADJ) MAX (0, EAW * (0.00402-0.15663 * MOVAVG (2, RTP.1) + 1.72119 * EM3544/E + 0.05679) + EM3544AW.ADJ) MAX ( 0 , EAW * \((0.00834+0.03746\) * MOVAVG \((2\), RTP.1 \()+0.46522 *\) EM4554/E + 0.00144) + EM4554AW.ADJ \()\) MAX (0, EAW * ( \(-0.00655+0.03521\) * MOVAVG (2, RTP.1) +0.46852 * EM5564/E - 0.00037) + EM5564AW.ADJ) MAX (0, EAW * \((-0.00114+0.07640\) * MOVAVG (2, RTP.1) +3.25911 * EM65O/E - 0.10058) + EM65OAW.ADJ)

MAX ( 0, EAW \(^{*}(-0.00055-0.05470 *\) MOVAVG (2, RTP.1) \(+1.41760 *\) EF1617/E +0.04979\()+\) EF1617AW.ADJ \()\) MAX (0, EAW * (0.00102-0.07375 * MOVAVG (2, RTP.1) + 0.78394 * EF1819/E + 0.07226) + EF1819AW.ADJ) MAX ( 0 , EAW * (0.00112-0.05971 * MOVAVG (2, RTP.1) +0.57256 * EF2024/E + 0.05907) + EF2024AW.ADJ) MAX ( 0, EAW \(^{*}(0.00623+0.08868\) * MOVAVG (2, RTP.1) +1.00897 * EF2534/E - 0.15142) + EF2534AW.ADJ) MAX ( 0 , EAW * ( \(0.00687-0.00259\) * MOVAVG (2, RTP.1) +0.51319 * EF3544/E - 0.00937) + EF3544AW.ADJ) MAX ( 0 , EAW * ( \(0.00185+0.08747\) * MOVAVG (2, RTP.1) +0.28022 * EF4554/E - 0.08053) + EF4554AW.ADJ) MAX ( 0 , EAW * \((-0.00140-0.03001 *\) MOVAVG (2, RTP.1) \(-0.59383 *\) EF5564/E +0.07088\()+\) EF5564AW.ADJ \()\) MAX (0, EAW * \((0.00096+0.06768 *\) MOVAVG (2, RTP.1) +1.04213 * EF65O/E - 0.07359) + EF65OAW.ADJ \()\)

EAW_R = EF1617AW_R + EF1819AW_R + EF2024AW_R + EF2534AW_R + EF3544AW_R + EF4554AW_R + EF5564AW_R + EF65OAW_R + EM1617AW_R + EM1819AW_R+ EM2024AW_R + EM2534AW_R + EM3544AW_R + EM4554AW_R + EM5564AW_R + EM65OAW_R
\begin{tabular}{|c|c|}
\hline & EM1617AW_R * (EAW/EAW_R) \\
\hline & EM1819AW_R * (EAW/EAW_R) \\
\hline EM2024 & EM2024AW_R * (EAW/EAW_R) \\
\hline 12534 & EM2534AW_R * (EAW/EAW_R) \\
\hline EM3544AW = & EM3544AW_R * (EAW/EAW_R) \\
\hline , & EM4554AW_R * (EAW/EAW_R) \\
\hline EM5564AW & EM5564AW_R * (EAW/EAW_R) \\
\hline EM650AW & EM650AW_R * (EAW/EAW_R) \\
\hline EF1617AW = & EF1617AW_R * (EAW/EAW_R) \\
\hline 1819AW & EF1819AW_R * (EAW/EAW_R) \\
\hline 2024AW & EF2024AW_R * (EAW/EAW_R) \\
\hline 2534AW & EF2534AW_R * (EAW/EAW_R) \\
\hline EF3544AW = & EF3544AW_R * (EAW/EAW_R) \\
\hline EF4554AW = & EF4554AW_R * (EAW/EAW_R) \\
\hline EF5564AW = & EF5564AW_R * (EAW/EAW_R) \\
\hline EF65OAW = & EF650AW_R * (EAW/EAW_R) \\
\hline
\end{tabular}

EFAW \(=\) EF1617AW+ EF1819AW+ EF2024AW+ EF2534AW+ EF3544AW+ EF4554AW+ EF5564AW+ EF65OAW \(\mathrm{EMAW}=\mathrm{EM} 1617 \mathrm{AW}+\mathrm{EM} 1819 \mathrm{AW}+\mathrm{EM} 2024 \mathrm{AW}+\mathrm{EM} 2534 A W+\) EM3544AW + EM4554AW+ EM5564AW + EM65OAW
\begin{tabular}{|c|c|}
\hline \multicolumn{2}{|l|}{Self-employed Agricultural Workers} \\
\hline EAS \(=\) EA & - EAW \\
\hline EM1617AS_R = & MAX (0, NM1617 * (0.00528 + 0.00404) + EM1617AS.ADJ) \\
\hline EM1819AS_R = & MAX (0, NM1819 * (0.00309 + 0.28448* EA / (NM16O+ NF16O) - 0.00165) + EM1819AS.ADJ) \\
\hline EM2024AS_R = & MAX (0, NM2024 * (-0.00181 + 0.97958*EA / (NM16O+ NF16O) - 0.01093) + EM2024AS.ADJ) \\
\hline EM2534AS_R = & MAX (0, NM2534 * (-0.00263 + 1.23186 * EA / (NM16O+ NF16O) - 0.01021) + EM2534AS.ADJ) \\
\hline EM3544AS_R = & MAX (0, NM3544 * ( \(-0.00151+1.66765 *\) EA / (NM16O+ NF16O) - 0.01450) + EM3544AS.ADJ \()\) \\
\hline EM4554AS_R = & MAX (0, NM4554 * (-0.00381 + 2.86654*EA / (NM16O+ NF16O) - 0.03175) + EM4554AS.ADJ) \\
\hline EM5564AS_R = & MAX (0, NM5564 * (-0.00460 + 2.78817 * EA / (NM16O+ NF16O) - 0.02398) + EM5564AS.ADJ) \\
\hline EM65OAS_R = & MAX (0, NM65O * (0.00079 + 1.76904 * EA / (NM16O+ NF16O) - 0.01437) + EM65OAS.ADJ) \\
\hline EF1617AS_R = & MAX (0, NF1617 * (0.00181 + 0.00030) + EF1617AS.ADJ) \\
\hline EF1819AS_R = & MAX (0, EM1819AS. \(1 *(-0.02393+0.63672 *\) MOVAVG \((2\), RTP. 1\()+0.98791 *\) EF1819/EM1819 - 4.43926 \()+\) EF1819AS.ADJ) \\
\hline EF2024AS_R = & MAX (0, EM2024AS. 1 * (0.07353-0.40207 * MOVAVG (2, RTP.1) + 0.57572 * EF2024/EM2024-0.01117) + EF2024AS.ADJ) \\
\hline EF2534AS_R = & MAX (0, EM2534AS. 1 * \((0.16575+0.16967 * \operatorname{MOVAVG}(2\), RTP.1) \(+0.55503 *\) EF2534/EM2534-0.43412 \()+\) EF2534AS.ADJ) \\
\hline EF3544AS_R = & MAX (0, EM3544AS. \(1 *(0.15848+0.37839 * \operatorname{MOVAVG}(2\), RTP.1 \()+0.37764 *\) EF3544/EM3544-0.45362 \()+\) EF3544AS.ADJ) \\
\hline EF4554AS_R = & MAX (0, EM4554AS. 1 * ( \(0.21947+0.29497\) * MOVAVG (2, RTP.1) +0.58974 * EF4554/EM4554-0.51966) + EF4554AS.ADJ) \\
\hline EF5564AS_R = & MAX (0, EM5564AS. \(1 *(0.20892+0.36294 *\) MOVAVG (2, RTP.1) +0.65320 * EF5564/EM5564-0.66626) + \\
\hline
\end{tabular}

```

EF2534AU = EF2534AU_R *(EAU/EAU_R)
EF3544AU = EF3544AU_R *(EAU/EAU_R)
EF4554AU = EF4554AU_R *(EAU/EAU_R)
EF5564AU = EF5564AU_R * (EAU/EAU_R)
EF65OAU = EF65OAU_R * (EAU/EAU_R)
EFAU = EF1617AU + EF1819AU + EF2024AU + EF2534AU + EF3544AU + EF4554AU + EF5564AU + EF65OAU
EMAU = EM1617AU + EM1819AU + EM2024AU + EM2534AU + EM3544AU + EM4554AU + EM5564AU + EM65OAU

```

\section*{ANNUAL "AVERAGE HOURS WORKED" EQUATIONS}

Total Nonagricultural Wage Workers
Males
\begin{tabular}{|c|c|}
\hline AHWM1617NAW = & 34.4953 * RTP + 1.03247 * MINW/CPIW_U - 31.4229 * RNM1617S- 0.12369 * \(1+12.1981\) \\
\hline AHWM1819NAW = & 32.4361 * RTP + 12.8742 * RTP. 1 + 1.28624 * MINW/CPIW_U- 16.0989 * RNM1819S - 0.27834 * 1-8.58664 \\
\hline AHWM2024NAW = &  \\
\hline AHWM2534NAW = & 17.0559 * RTP - 0.29076 * \(1+0.04542\) *YEAR + 22.5121 \\
\hline AHWM3544NAW = & \(18.3314 *\) RTP -0.30475 * \(1+0.04409\) *YEAR +22.3275 \\
\hline AHWM4554NAW = & 16.0678 * RTP -0.12289 * \(1+0.07446\) *YEAR + 21.4366 \\
\hline AHWM5564NAW = & 10.8277 * RTP + 0.23715 * \(1+30.8975\) \\
\hline AHWM65ONAW = & \(-0.28038 * 1+5.40682\) * RTP + 25.4797 \\
\hline \multicolumn{2}{|l|}{Females} \\
\hline AHWF1617NAW = & \(17.2598 *\) RTP \(+7.31262 *\) RTP. \(1-24.9241\) * RNF1617S +16.9218 \\
\hline AHWF1819NAW = & 18.7922 * RTP + 13.1066 * RTP. \(1-18.9417\) * RNF1819S-16.2182 * PF1819 + 0.81920 * \(1+16.7826\) \\
\hline AHWF2024NAW = & \(-219.154 *\) PF2024 + 12.4226 * RTP - 0.98470 * \(1+158.354 *(\text { PF2024 })^{2}+98.6034\) \\
\hline AHWF2534NAW = & -39.2904 * PF2534 + 39.6515 * (PF2534) \({ }^{2}+9.10839\) * RTP- \(0.33503 * 1+35.8525\) \\
\hline AHWF3544NAW = & -39.2904 * PF3544 + 39.6515 * (PF3544) \({ }^{2}+7.40115\) * RTP- \(0.38214 * 1+36.9133\) \\
\hline AHWF4554NAW = & 8.64511 * RTP - 63.8042 * PF4554 + 59.9107 * (PF4554) \({ }^{2}-0.58355 * 1+44.0723\) \\
\hline AHWF5564NAW = & 6.61506 * RTP - 0.28969 * \(1+8.36882\) * PF5564 + 24.8288 \\
\hline AHWF65ONAW = & \(-0.43916 * 1+26.5228+(8.36882 / 2) *(\) PF65O- 0.086) \\
\hline \multicolumn{2}{|l|}{Nonagricultural Self-employed Workers} \\
\hline AHWNAS = & 24.9592 * RTP - 17.2194 * EFNAS/ENAS + 22.0120 \\
\hline \multicolumn{2}{|l|}{Nonagricultural Unpaid Family Workers} \\
\hline AHWNAU = & 25.5124 * EF2534NAU + EF3544NAU + EF4554NAU)/ENAU + 19.2730 \\
\hline \multicolumn{2}{|l|}{Agricultural Wage Workers} \\
\hline AHWAW = & 6.58073 * RTP + 14.9130 * RTP. \(1+19.7800\) \\
\hline \multicolumn{2}{|l|}{Agricultural Self-employed Workers} \\
\hline AHWAS = & 2.45830 * PGDPAF/PGDP - 3.61107 * \(1+44.5318\) \\
\hline \multicolumn{2}{|l|}{Agricultural Unpaid Family Workers} \\
\hline AHWAU = & 39.3563 \\
\hline
\end{tabular}

\section*{QUARTERLY "AVERAGE HOURS WORKED" EQUATIONS}

Total
AHW =
\[
(\mathrm{AHWNA} * \mathrm{ENA}+\mathrm{AHWA} * \mathrm{EA}) / \mathrm{E}
\]

Nonagricultural sector
AHWNA \(=\quad(\) AHWNAW \(*(\) E-ENAS-ENAU-EA \()+\) AHWNAS \(*\) ENAS + AHWNAU \(*\) ENAU \() /\) ENA
Nonagricultural Wage Workers
AHWNAW \(=\quad(A H W M N A W ~ * ~(E M ~-~ E M N A S ~-~ E M N A U ~-~ E M A W ~-~ E M A S ~-~ E M A U) ~+~ A H W F N A W ~ * ~(E F ~-~ E F N A S ~-~ E F N A U ~-~\)
EFAW - EFAS - EFAU))/ (E - ENAS - ENAU - EA)

Males
\begin{tabular}{|c|c|}
\hline AHWMNAW1 = & AHWM1617NAW * (EM1617-EM1617NAS - EM1617NAU - EM1617AW - EM1617AS - EM1617AU) \\
\hline AHWMNAW2 = & AHWM1819NAW * (EM1819-EM1819NAS - EM1819NAU - EM1819AW - EM1819AS - EM1819AU) \\
\hline AHWMNAW3 = & AHWM2024NAW * (EM2024-EM2024NAS - EM2024NAU - EM2024AW - EM2024AS - EM2024AU) \\
\hline AHWMNAW4 = & AHWM2534NAW * (EM2534-EM2534NAS - EM2534NAU - EM2534AW - EM2534AS - EM2534AU) \\
\hline AHWMNAW5 = & AHWM3544NAW * (EM3544-EM3544NAS - EM3544NAU - EM3544AW - EM3544AS - EM3544AU) \\
\hline AHWMNAW6 = & AHWM4554NAW * (EM4554-EM4554NAS - EM4554NAU - EM4554AW - EM4554AS - EM4554AU) \\
\hline AHWMNAW7 = & AHWM5564NAW * (EM5564-EM5564NAS - EM5564NAU - EM5564AW - EM5564AS - EM5564AU) \\
\hline AHWMNAW8 & AHWM65ONAW * (EM65O - EM65ONAS - EM65ONAU - EM65OAW - EM65OAS - EM65OAU) \\
\hline AHWMNAW = & \begin{tabular}{l}
(AHWMNAW1 + AHWMNAW2 + AHWMNAW3 + AHWMNAW4 + AHWMNAW5 + AHWMNAW6 + \\
AHWMNAW7 + AHWMNAW8) / (EM - EMNAS - EMNAU- EMAW - EMAS - EMAU)
\end{tabular} \\
\hline AHWM1617NA & 34.4953 * MOVAVG (2, RTP.1) + 1.03247 * MINW/CPIW_U - 31.4229 * RNM1617S - \(0.12369 * 1+12.1981+\) AHWM1617NAW.ADJ \\
\hline AHWM1819NAW \(=\) & \[
\begin{aligned}
& 32.4361 * \text { MOVAVG (2, RTP.1) + 12.8742 * MOVAVG (2, RTP.5) + } 1.28624 * \text { MINW/CPIW_U - } 16.0989 * \\
& \text { RNM1819S - 0.27834 * } 1-8.58664 \text { + AHWM1819NAW.ADJ }
\end{aligned}
\] \\
\hline AHWM2024NAW = & 20.1161 * MOVAVG (2, RTP.1) - 10.2292 * RNM2024S - 2.28628 * (PM2024NM * NM2024NM/ (PM2024NM * NM2024NM + PM2024MS * NM2024MS + PM2024MA * NM2024MA)) + 23.2252 + AHWM2024NAW.ADJ \\
\hline AHWM2534NAW = & MIN (45, 17.0559 * MOVAVG (2, RTP.1) - 0.29076 * \(1+0.04542\) *YEAR + 22.5121) + AHWM2534NAW.ADJ \\
\hline AHWM3544NAW = & MIN (45, 18.3314 * MOVAVG (2, RTP.1) - 0.30475 * \(1+0.04409\) *YEAR + 22.3275) + AHWM3544NAW.ADJ \\
\hline AHWM4554NAW = & MIN (45, 16.0678 * MOVAVG (2, RTP.1) - 0.12289 * \(1+0.07446\) *YEAR + 21.4366) + AHWM4554NAW.ADJ \\
\hline AHWM5564NAW = & 10.8277 * MOVAVG (2, RTP.1) + \(0.23715 * 1+30.8975\) + AHWM5564NAW.ADJ \\
\hline AHWM650NAW = & \(-0.28038 * 1+5.40682\) * MOVAVG (2, RTP. \()\) + 25.4797 + AHWM65ONAW.ADJ \\
\hline
\end{tabular}

Females

AHWFNAW1 = AHWF1617NAW* (EF1617 -EF1617NAS-EF1617NAU-EF1617AW-EF1617AS-EF1617AU) AHWFNAW2 \(=\quad\) AHWF1819NAW* \((E F 1819-E F 1819 N A S-E F 1819 N A U-E F 1819 A W-E F 1819 A S-E F 1819 A U)\) AHWFNAW3 = AHWFNAW4 = AHWFNAW5 = AHWFNAW6 = AHWFNAW7 = AHWFNAW8 = AHWFNAW = AHWF2024NAW* (EF2024 -EF2024NAS-EF2024NAU-EF2024AW-EF2024AS-EF2024AU) AHWF2534NAW* (EF2534 -EF2534NAS-EF2534NAU-EF2534AW-EF2534AS-EF2534AU) AHWF3544NAW* (EF3544 -EF3544NAS-EF3544NAU-EF3544AW-EF3544AS-EF3544AU) AHWF4554NAW* (EF4554 -EF4554NAS-EF4554NAU-EF4554AW-EF4554AS-EF4554AU) AHWF5564NAW* (EF5564 -EF5564NAS-EF5564NAU-EF5564AW-EF5564AS-EF5564AU) AHWF65ONAW * (EF65O -EF65ONAS-EF65ONAU-EF65OAW-EF65OAS-EF65OAU) (AHWFNAW1 + AHWFNAW2 + AHWFNAW3 + AHWFNAW4 + AHWFNAW5 + AHWFNAW6 + AHWFNAW7 + AHWFNAW8) / (EF - EFNAS - EFNAU - EFAW - EFAS - EFAU)
AHWF1617NAW \(=17.2598 *\) MOVAVG (2, RTP.1) \(+7.31262 *\) MOVAVG (2, RTP.5) \(-24.9241 *\) RNF1617S \(+16.9218+\) AHWF1617NAW.ADJ
AHWF1819NAW \(=18.7922 *\) MOVAVG (2, RTP.1) \(+13.1066 *\) MOVAVG (2, RTP.5) \(-18.9417 *\) RNF1819S \(-16.2182 *\) PF1819 + 0.81920 * 1 + 16.7826 + AHWF1819NAW.ADJ

AHWF2024NAW \(=-219.154 *\) PF2024 \(+12.4226 *\) MOVAVG (2, RTP.1) \(-0.98470 * 1+158.354 *(\text { PF2024 })^{2}+98.6034+\) AHWF2024NAW.ADJ
AHWF2534NAW \(=-39.2904 *\) PF2534 \(+39.6515 *(\text { PF2534 })^{2}+9.10839 *\) MOVAVG (2, RTP.1) \(-0.33503 * 1+35.8525+\) AHWF2534NAW.ADJ
AHWF3544NAW \(=-39.2904 *\) PF3544 \(+39.6515 *(\text { PF3544 })^{2}+7.40115 *\) MOVAVG \((2\), RTP. \()-0.38214 * 1+36.9133+\) AHWF3544NAW.ADJ
AHWF4554NAW \(=8.64511 *\) MOVAVG (2, RTP.1) \(-63.8042 *\) PF4554 \(+59.9107 *(\text { PF4554 })^{2}-0.58355 * 1+44.0723+\) AHWF4554NAW.ADJ
AHWF5564NAW \(=6.61506 *\) MOVAVG (2, RTP.1) \(-0.28969 * 1+8.36882 *\) PF5564 \(+24.8288+\) AHWF5564NAW.ADJ AHWF65ONAW \(=-0.43916 * 1+26.5228+(8.36882 / 2) *(\) PF65O- 0.086 \()+\) AHWF65ONAW.ADJ

Nonagricultural Self-employed
AHWNAS \(=\quad 24.9592 *\) MOVAVG (2, RTP.1) \(-17.2194 *\) EFNAS/ENAS \(+22.0120+\) AHWNAS.ADJ
Nonagricultural Unpaid Family Workers
AHWNAU \(=\quad 25.5124^{*}(\) EF2534NAU + EF3544NAU + EF4554NAU \() /\) ENAU \(+19.2730+\) AHWNAU.ADJ
Agricultural sector
AHWA \(=(\) AHWAW \(*\) EAW + AHWAS \(*\) EAS + AHWAU \(*\) EAU \() /\) EA
Agricultural Wage Workers
AHWAW \(=66.58073 * \operatorname{MOVAVG}(2\), RTP.1 \()+14.9130 * \operatorname{MOVAVG}(2\), RTP. \()+19.7800+\) AHWAW.ADJ
Agricultural Self-employed
AHWAS \(=\quad 2.45830 *\) PGDPAF/PGDP \(-3.61107 * 1+44.5318+\) AHWAS.ADJ

Agricultural Unpaid Family Workers
AHWAU \(=39.3563+\) AHWAU.ADJ

OTHER EMPLOYMENT MEASURES
\begin{tabular}{|c|c|}
\hline \multicolumn{2}{|l|}{Federal Civilian Government and Government Enterprises} \\
\hline EGFC = & IF LONGRANGE \(=0\) \\
\hline & THEN (EGFC. 1 * 1.0094 \({ }^{0.25}\) ) \\
\hline & ELSE (EGFC. 1 * (E_FE/E_FE.1)) \\
\hline \multirow[t]{3}{*}{EGEFCPS =} & IF LONGRANGE \(=0\) \\
\hline & THEN (EGEFCPS. 1 * \(1.0075{ }^{0.25}\) ) \\
\hline & ELSE (EGEFCPS. 1 * (E_FE/E_FE.1)) \\
\hline EGGEFC = & EGFC + EGEFCPS \\
\hline \multicolumn{2}{|l|}{State and Local Government and Government Enterprises} \\
\hline \multirow[t]{3}{*}{EGGESL =} & IF LONGRANGE \(=0\) \\
\hline & THEN EGGESL. 1 * (LC_FE/LC_FE.4) \({ }^{0.25}\) \\
\hline & ELSE EGGESL. 1 * (E_FE/E_FE.1) \\
\hline \multicolumn{2}{|l|}{Military} \\
\hline \multirow[t]{3}{*}{DNEDMIL =} & IF (EDMIL-EDMIL.4) < 0 \\
\hline & THEN (EDMIL-EDMIL.4) \\
\hline & ELSE 0 \\
\hline EP = & E-EGGESL - EGGEFC - EAS - ENAS \\
\hline
\end{tabular}

Compensation and Output Sectors

Price Deflator for Medical Services
CPIWMS \(=\quad\) CPIWMS. \(1 *\left(1+\left((\text { CPIW_U/CPIW_U.4 })^{0.25}-1\right) *\right.\) CPIWMSWT \()\)
\begin{tabular}{|c|c|}
\hline \multicolumn{2}{|l|}{Unemployment Insurance and Workers Compensation Effective Tax Rates} \\
\hline TMAXUI_SL = & TMAXUI_SL. 1 *AWSUI.1/AWSUI. 2 \\
\hline RELMAX_UI = & TMAXUI_SL/AWSUI.1/1000 \\
\hline CR_UI = & 0.775 \\
\hline TRATIO_UI = & 0.96996 * RELMAX_UI-0.13744 * MOVAVG (4, RTP.1) + 0.10368 * MOVAVG (4, RTP.5) + 0.04887 \\
\hline TRATE_UI = & 0.00143 * MOVAVG (4, RU.5) + 0.00128 * MOVAVG (4, RU.9) + 0.00057 * MOVAVG (4, RU. 13\()+0.00356\) \\
\hline RUIWS1 = & CR_UI * TRATIO_UI * TRATE_UI \\
\hline RUIWS2 = & 0.32476 * MOVAVG (4, RUIWS1.8 * (WSP. 8 -WSPRRB. 8 +WSGGESL.8)) / (WSP. 1 -WSPRR2. 1 +WSGGESL.1) \\
\hline
\end{tabular}

Workers' Compensation
RWCWS = RWCWS. \(1-(\) RWCWS. \(1-0.0144) / 12\)
State and Local Government and Government Enterprises TAXMAXQ \(=\quad\) IF QTR \(=1\) THEN 300.* ROUND (MOVAVG (4, AWSE.5)/MOVAVG (4, AWSE.9) * 1000.* TAXMAX.1/300 +0.5)/1000 ELSE TAXMAX. 1

Wages
AWSPL \(=\) MOVAVG (8, AWSP.1)
AWSSPL \(=\quad\) MOVAVG (8, AWSSP.1)
AWSGGESL \(=\) IF LONGRANGE \(=0\)
THEN AWSGGESL. 1 * AWSPL/AWSPL. 1 ELSE AWSGGESL. 1 * AVG_GDP/AVG_GDP. 1 * \((1+\text { WS_TO_WSS_D/100 })^{0.25}\)
WSGGESL \(=\) AWSGGESL * EGGESL

Employer Contribution for Government Social Insurance
OASDISL_L \(=(\) EMPTROASI + EMPTRDI \() * 0.978 *\) CSLA \(*\) WSGGESL
HISL_L \(=\) EMPTRHI \(* 1.0\) * CSLHI * WSGGESL
SOC_UISL \(=(-0.02821 *\) MOVAVG (4, RTP.2) +0.03145\() *\) WSGGESL
RSOCSL_WC = RSOCSL_WC. \(1-(\) RSOCSL_WC.1-0.176)/12
SOC_WCSL = RSOCSL_WC * RWCWS * WSGGESL
SOC_SL \(=\) (OASDISL_L + HISL_L + SOC_UISL + SOC_WCSL)

Employer Contributions for Employee Pension and Insurance funds
```

Workers' Compensation - employees and annuitants
OLI_WCSL = (1-RSOCSL_WC) * RWCWS * WSGGESL
Pensions
OLI_RETSL = WSGGESL * (OLI_RETSL.1/WSGGESL.1)
Life Insurance - employees and annuitants
OLI_GLI_SL = 2.0 * EGGESL * ((WSGGESL/EGGESL) +2.0) * 0.075 * 26/1000
Health Insurance - employees and annuitants
OLI_GHI_SL = (OLI_GHI_SL.1 / EGGESL.1) * CPIWMS/CPIWMS.1 * EGGESL * RGR_GHI
Total
OLI_SL = (OLI_GLI_SL + OLI_GHI_SL + OLI_WCSL + OLI_RETSL)
RCWSSL = (1 + (SOC_SL + OLI_SL)/WSGGESL)
WSSGGESL = IF LONGRANGE = 0
THEN RCWSSL*WSGGESL
ELSE (WSSGGESL.1/EGGESL.1) * AVG_GDP/AVG_GDP.1 * EGGESL
WSSGSL = WSSGGESL * WSSGSL.1/WSSGGESL. }
WSSGESL = WSSGGESL - WSSGSL
CFCGSL = IF LONGRANGE = 0
THEN WSSGSL * RCFCGSL
ELSE CFCGSL.1 * WSSGGESL/WSSGGESL. }
GDPGSL = WSSGSL + CFCGSL
CFCGESL = IF LONGRANGE =0
THEN WSSGESL * RCFCGESL
ELSE CFCGESL. 1 * WSSGGESL/WSSGGESL. }
GDPGESL = WSSGESL + CFCGESL
GDPGGESL = GDPGSL + GDPGESL

```

Federal Civilian General Government and Government Enterprises

Wages
General Government and Government Enterprises
Civilian pay raise
CRAZ1 \(=\quad\) IF LONGRANGE \(=0\)
    THEN ((IF QTR = 1 THEN (0.82429 * (AWSP.6/AWSP.10-1) -0.005) ELSE 0))
    ELSE (IF QTR = 1 THEN (AWSP.6/AWSP. 10 -1) ELSE 0)
Military pay raise
MRAZ \(=\quad\) IF LONGRANGE \(=0\)
    THEN ((IF QTR = 1 THEN ( 0.82429 * (AWSP.6/AWSP.10-1) -0.005) ELSE 0))
    ELSE (IF QTR = 1 THEN (AWSP.6/AWSP.10-1) ELSE 0)
AWSGGEFC \(=\) IF LONGRANGE \(=0\)
    THEN (AWSGGEFC. 1 * (1 +1.0 *CRAZ1 +0.0015)
    ELSE AWSGGEFC. 1 * AVG_GDP/AVG_GDP. 1 * \(\left(1+W S \_T O \_W S S \_D / 100\right)^{0.25}\)
WSGGEFC \(=\) AWSGGEFC * EGGEFC
CSRS workers
AWEFC_N \(=\) IF LONGRANGE \(=0\)
    THEN (AWEFC_N. 1 * ( \(1+1.0\) *CRAZ1 + 0.00082) \()\)
    ELSE AWEFC_N. 1 * AVG_GDP/AVG_GDP. 1 * \(\left(1+W S \_T O \_W S S \_D / 100\right)^{0.25}\)
WEFC_N \(=\) AWEFC_N \(*\) TEFC_N
```

Government Enterprises (Mostly U.S. Postal Service)
AWSGEFC = IF LONGRANGE = 0
THEN (AWSGEFC. 1 * (1 +1.0 *CRAZ1 +0.0015))
ELSE AWSGEFC.1 * AVG_GDP/AVG_GDP.1 * (1 +WS_TO_WSS_D/100) 0.25
WSGEFC = AWSGEFC * EGEFCPS

```

General Government
WSGFC \(=\quad\) WSGGEFC -WSGEFC
AWSGFC \(=\quad\) WSGFC/EGFC

Employer Contribution for Government Social Insurance
General Government and Government Enterprises
```

OASDIFC_L = (EMPTROASI + EMPTRDI) * 1.04 * (WSGGEFC - WEFC_N) * ADJ_FSA_FC
HIFC_L = EMPTRHI* 1.055 * WSGGEFC * ADJ_FSA_FC
SOCF_UIFC }=(-0.05934*\mathrm{ RTP + 0.06165) * WSGGEFC
SOCF_WC = 0.0159 * WSGGEFC
SOC_\overline{FC = (SOCF_UIFC + SOCF_WC + OASDIFC_L + HIFC_L })=()

```

Employer Contributions for Employee Pension and Insurance funds

General Government and Government Enterprises

\section*{Pensions}

OLI_CSRS1 \(=\quad((0.174 *\) WSGEFC \(+0.07 *\) WSGFC \() /\) WSGGEFC \() *\) WEFC_N
OLI_FERS1 \(=0.107 *\left(W S G G E F C * 0.9-W E F C \_N\right)\)
OLI_FERSFC = 0.048 * (WSGGEFC * 0.9 -WEFC_N)

OLI_RETFC \(=\) OLI_CSRS1 + OLI_FERS1 + OLI_FERSFC + OLIF_RETFCO
Life Insurance - employees and annuitants
OLI_GLI_FC \(=2.0\) * EGGEFC \(*((W S G G E F C / E G G E F C) ~+~ 2.0) ~ * ~ 0.075 ~ * ~ 26 / 1000 ~\)
Health Insurance - employees and annuitants
OLI_GHI_FC \(=\) (OLI_GHI_FC. \(1 /\) EGGEFC. 1\() *\) CPIWMS/CPIWMS. \(1 *\) EGGEFC \(*\) RGR_GHI
OLI_FC \(=\) (OLI_GHI_FC + OLI_GLI_FC + OLI_RETFC)

Compensation
General Government and Government Enterprises
RCWSF \(=\quad(1+(\) SOC_FC + OLI_FC \() / W S G G E F C)\)
WSSGGEFC \(=\) IF LONGRANGE \(=0\)
THEN RCWSF * WSGGEFC
ELSE (WSSGGEFC.1/EGGEFC.1) * AVG_GDP/AVG_GDP. 1 * EGGEFC
WSSGFC \(=\quad\) IF LONGRANGE \(=0\)
THEN RCWSF * WSGFC
ELSE (WSSGFC. 1 / (EGGEFC. 1 - EGEFCPS.1)) * AVG_GDP/AVG_GDP. \(1 *\) (EGGEFC - EGEFCPS)
WSSGEFC \(=\) IF LONGRANGE \(=0\)
THEN RCWSF * WSGEFC
ELSE (WSSGEFC.1/EGEFCPS.1) * AVG_GDP/AVG_GDP. 1 * EGEFCPS

Consumption of Fixed Capital
General Government and Government Enterprises
CFCGFC \(=\quad\) IF LONGRANGE \(=0\)
THEN WSSGFC * RCFCGFC
ELSE CFCGFC. 1 * WSSGGEFC / WSSGGEFC. 1
CFCGEFC \(=\) IF LONGRANGE \(=0\)
THEN WSSGEFC * RCFCGEFC
ELSE CFCGEFC. 1 * WSSGGEFC / WSSGGEFC. 1

Gross Domestic Product
```

General Government and Government Enterprises
GDPGFC = WSSGFC + CFCGFC
GDPGEFC = WSSGEFC + CFCGEFC
GDPGGEFC = GDPGFC + GDPGEFC

```

Federal Government Military
Wages
AWSGFM \(=\quad\) IF LONGRANGE \(=0\)
```

        THEN (AWSGFM.1 * (1.0027 +1.0 * MRAZ))
        ELSE AWSGFM.1 * AVG_GDP/AVG_GDP. 1 * (1 + WS_TO_WSS_D/100) 0.25
    WSGFM =
AWSGFM * (EDMIL+ EDMIL_R)
Employer Contribution for Government Social Insurance
OASDIFM_L = (EMPTROASI + EMPTRDI) * 0.9975 * CML * WSGFM
HIFM_L = EMPTRHI* 1.0 *CML * WSGFM
SOCF_UIFM = MAX (0.001, (-0.05263 * DIFF (EDMIL + EDMIL_R) - 0.03079 * RTP + 0.03310)) * WSGFM
SOCF_MIFM = 0.30 * CPIWMS * (EDMIL + EDMIL_R)
SOC_FM = (SOCF_UIFM + SOCF_MIFM + OASDIFM_L + HIFM_L)
Employer Contributions for Employee Pension and Insurance funds
OLI_RETFM = (OLI_RETFM.1/WSGFM.1 - (OLI_RETFM.1/WSGFM.1-0.472)/12) * WSGFM
Compensation
RCWSM = (1 + (OLI_RETFM + SOC_FM)/WSGFM)
WSSGFM = IF LONGRANGE = 0
THEN RCWSM * WSGFM
ELSE (WSSGFM.1/EDMIL.1) * AVG_GDP/AVG_GDP.1 * EDMIL

```
Consumption of Fixed Capital
CFCGFM \(=\quad\) IF LONGRANGE \(=0\)
    THEN WSSGFM * RCFCGFM
    ELSE CFCGFM. 1 * WSSGFM/WSSGFM. 1
Gross Domestic Product
GDPGFM \(=\) WSSGFM + CFCGFM
GDPGF \(=\quad\) GDPGFC + GDPGFM
GDPGGE \(=\) GDPGGEFC + GDPGGESL + GDPGFM

Total (Civilian and Military) Federal General Government and Government Enterprises
WSSGF = WSSGFC + WSSGFM
WSSGE = WSSGEFC + WSSGESL
WSSG = WSSGF + WSSGSL
GDPGE = GDPGEFC + GDPGESL
GDPG \(=\) GDPGF + GDPGSL
NIPA Farm Output and Earnings
GDPPF09 = IF LONGRANGE \(=0\)
    THEN EXP (- 3.52340 + 0.02055 * YEAR) * N_SSA * 1.125 * 1.138
    ELSE GDPPF09.1 * GDP09/GDP09.1
PGDPAF \(=\quad\) IF LONGRANGE \(=0\)
    THEN PGDPAF. 1 * ((PGDP/PGDP.1) \(\left.{ }^{4}-0.01\right)^{0.25}\)
    ELSE PGDPAF. 1 * ((PGDP/PGDP. 1\(\left.)^{4}\right)^{0.25}\)
GDPPF \(=\quad\) GDPPF09 \(*\) PGDPAF
AYF_K = ((YF.1 /EAS.1) / (WSSPF.1/EAW.1) -5.0)*. \(8+5.0\)
WSSPF \(=\quad\) IF LONGRANGE \(=0\)
    THEN EAW * MOVAVG (4, WSSP.2/EP.2) * (3.15749 / (YEAR-65) - 0.43419 * RTP + 0.68725)
    ELSE (WSSPF.1/EAW.1) * AVG_GDP/AVG_GDP. 1 * EAW
WSPF \(=\quad\) IF LONGRANGE \(=0\)
    THEN WSSPF * (MOVAVG (12, (WSP.1/WSSP.1)) + 0.015)
    ELSE (WSPF. 1 /WSSPF.1) * (WSP.1/WSSP.1) / (WSP.2/WSSP.2) * WSSPF
AWSPF = WSPF/EAW
\(\mathrm{YF}=\quad\) AYF_K \(*(W S S P F / E A W) *\) EAS

GDP, WSS and WS, Private Households \& Nonprofit Institutions

Private Households
\begin{tabular}{ll} 
Compensation \& Wages \\
\begin{tabular}{ll} 
WSSPH \(=\) & IF LONGRANGE \(=0\) \\
& THEN \((((\) WSSPH.1/ENAWPH.1)/MOVAVG \((4\), WSSP.3/EP.3) -0.41\() * 0.875+0.41) *\) MOVAVG (4, WSSP.2/EP.2) *
\end{tabular} \\
ENAWPH &
\end{tabular}
```

        ELSE (AVG_GDP/AVG_GDP.1) * ENAWPH * (WSSPH.1/ENAWPH.1)
    WSPH = IF LONGRANGE = 0
THEN WSSPH / (1 +CPH* 1 * (EMPTROASI + EMPTRDI + EMPTRHI))
ELSE (AWSPH.1 * ENAWPH.1/WSSPH.1) * (1 + WS_TO_WSS_D/100) 0.25 * WSSPH
AWSPH = WSPH / ENAWPH

```
Owner Occupied Housing
\(\mathrm{OOH}=\mathrm{OOH} .1 *(\mathrm{KGDP} 09 * \mathrm{PGDP}) /(\mathrm{KGDP} 09.1 *\) PGDP.1 \()\)
Gross Value Added
GDPPH \(=\) IF LONGRANGE \(=0\)
    THEN WSSPH + OOH
    ELSE (AVG_GDP/AVG_GDP.1) * ENAWPH * (GDPPH.1/ENAWPH.1)

Nonprofit Institutions

Health Services
EPHS_EST \(=\) IF LONGRANGE \(=0\)
    THEN EPHS_EST. 1 + 0.275/4
    ELSE EPHS_EST. 1 * (E_FE/E_FE.1)
AWSSPHS \(=\) IF LONGRANGE \(=0\)
    THEN AWSSPHS. 1 * AWSSPL/AWSSPL. 1
    ELSE AWSSPHS. 1 * AVG_GDP/AVG_GDP. 1
WSSPHS = AWSSPHS* EPHS_EST
Educational Services
EPES_EST \(=\) IF LONGRANGE \(=0\)
    THEN EPES_EST. 1 + 0.075/4
    ELSE EPES_EST. 1 * (E_FE/E_FE.1)
AWSSPES \(=\) IF LONGRANGE \(=0\)
    THEN AWSSPES. 1 * AWSSPL/AWSSPL. 1
    ELSE AWSSPES. 1 * AVG_GDP/AVG_GDP. 1
WSSPES \(=\) AWSSPES* EPES EST
Social Services
EPSS_EST \(=\) IF LONGRANGE \(=0\)
    THEN EPSS EST. \(1+0.075 / 4\)
    ELSE EPSS_EST. 1 * (E_FE/E_FE.1)
AWSSPSS \(=\quad\) IF LONGRANGE \(=0\)
    THEN AWSSPSS. 1 * AWSSPL/AWSSPL. 1
    ELSE AWSSPSS. 1 * AVG_GDP/AVG_GDP. 1
WSSPSS \(=\) AWSSPSS*EPSS_EST

Gross Value Added
WSSPNI \(=\quad\) WSSPNI. \(1 *(\) WSSPHS + WSSPES + WSSPSS \() /(\) WSSPHS. \(1+\) WSSPES. \(1+\) WSSPSS. 1\()\)
WSPNI \(=\quad\) IF LONGRANGE \(=0\)
    THEN WSSPNI * (WSPNI.1/WSSPNI.1) * ((WSP.1/WSSP.1) / (WSP.9/WSSP.9) \()^{(1 / 8)}\)
    ELSE WSSPNI * (WSPNI.1/WSSPNI.1) * (1 +WS_TO_WSS_D/100 \()^{0.25}\)
GDPPNI \(=\quad\) IF LONGRANGE \(=0\)
    THEN WSSPNI / ((WSSPNI.1/GDPPNI. \(1-0.866)\) * \(0.8+0.866)\)
    ELSE WSSPNI /0.866

Private Output and Compensation
```

ROASDIP_L = (EMPTROASI + EMPTRDI) * TXRP * CP
RHIP L = EMPTRHI * 1.0 * CP
RSOC_UIP = 0.00109 * MOVAVG (4, RU.2) + 0.00045 * MOVAVG (4, RU.10) + 0.00048 * MOVAVG (4, RU.18) - 0.00331
RSOC_WCP = RWCWS * RSOCSL_WC
RSOCF_PBG = 0.00022
OLI
ROLI_WCP = RWCWS * (1 - RSOCSL_WC)
ROLI_SU = 0.0005
OLI_GLI_P = 0.0025 * EP * AWSP. }
OLI_GHI_P = (OLI_GHI_P.1/EP.1) * CPIWMS/CPIWMS.1 * EP * RGR_GHI

```
```

ROLI_PPPS = MAX (ROLI_PPPS.1, 0.00031 * YEAR + 0.00866)

```

Employee Compensation and Non-Farm Proprietor Income (WSS and YNF)
\begin{tabular}{|c|c|}
\hline WSSGGE & (WSSGGESL + WSSGGEFC + WSSGFM) \\
\hline GDPPBNFXGE & \(=(\mathrm{GDP}-\mathrm{GDPGGE}-\mathrm{GDPPF}-\mathrm{GDPPH}-\mathrm{GDPPNI})\) \\
\hline RWSSPBNFXGE & 0.30026 * rtp. \(1+0.31936+(0.971 * 0.621-(0.30026 * 1.0+0.31936))+\) rwsspbnfxge_adj \\
\hline ENAW & ENA-ENAS-ENAU \\
\hline ENAWPBXGE & ```
= ENAW - (ENAWPH + EGGEFC + EGGESL + WSSPNI / (WSSPHS + WSSPES + WSSPSS) * (EPHS_EST +
EPES_EST + EPSS_EST))
``` \\
\hline ENAWSPBXGE & \(=\) ENAWPBXGE + ENAS \\
\hline AYNF_K & \(=((\) YNF.1/ENAS. \(1 /(\) WSSPBNFXGE.1/ENAWPBXGE.1) \()-1.65) * 0.9+1.65\) \\
\hline AYF & YF/EAS \\
\hline AWSSPF & WSSPF/EAW \\
\hline AYNF & YNF/ENAS \\
\hline AWSSPBNFXGE & \(=\) WSSPBNFXGE/ENAWPBXGE \\
\hline YNF & \[
\begin{aligned}
& =\text { YNF. } 1 *(\text { GDPPBNFXGE / GDPPBNFXGE. } 1) *(\text { ENAS } /(\text { ENAS }+ \text { ENAWPBXGE })) /(\text { ENAS. } 1 /(\text { ENAS. } 1+ \\
& \text { ENAWPBXGE.1) })
\end{aligned}
\] \\
\hline WSSPBNFXGE & \(=\quad\) rwsspbnfxge * (gdppbnfxge - ynf) \\
\hline WSSP & WSSPBNFXGE+WSSPF+WSSPH + WSSPNI \\
\hline RCWSP & ```
= WSSP / (WSSP - SOCF_RETRR - OLI_GLI_P - OLI_GHI_P) * (1 + ROASDIP_L + RHIP_L + RSOC_UIP +
RSOC_WCP + RSOCF_PBG + ROLI_WCP + ROLI_SU+ ROLI_PPPS)
``` \\
\hline WS & \(=\quad\) IF WS_TO_WSS_DYR \(=0\) \\
\hline & THEN (WSGGESL + WSGGEFC + WSGFM + WSSP/RCWSP) ELSE WSS * WS.1/WSS. 1 * ( 1 + WS TO WSS D/100 \()^{0.25}\) \\
\hline WSP & \(=\) (WS - WSGGESL - WSGGEFC - WSGFM) \\
\hline AWSP & \(=\) WSP/EP \\
\hline AWSSP & \(=\mathrm{WSSP} / \mathrm{EP}\) \\
\hline
\end{tabular}

Other Variables
\begin{tabular}{|c|c|}
\hline WSD & WS \\
\hline WSDP & (WSD -WSGGESL - WSGGEFC - WSGFM) \\
\hline AWSE & WS / (E + EDMIL - EAS - ENAS) \\
\hline AWSUI & (WS -WSGGEFC -WSGFM) / (E - EGGEFC - EAS - ENAS) \\
\hline WSS & (WSSP + WSSGGE) \\
\hline OLI & OLI_GGE + OLI_P \\
\hline SOC & SOC_GGE + SOC_P \\
\hline OLI_GGE & OLI_FC + OLI_SL + OLI_RETFM \\
\hline SOC_GGE & SOC_FC + SOC_FM + SOC_SL \\
\hline SOC_UIP & RSOC_UIP * WSP \\
\hline SOC_WCP & RSOC_WCP * WSP \\
\hline OASDIP_L & ROASDIP_L * WSP \\
\hline HIP_L & RHIP_L * WSP \\
\hline SOCF_PBG & RSOCF_PBG * WSP \\
\hline SOCF_RETRR & 0.20 * WSPRRB \\
\hline SOC_P & SOC_UIP + SOC_WCP + OASDIP_L + HIP_L + SOCF_PBG + SOCF_RETRR \\
\hline OLI_WCP & ROLI_WCP * WSP \\
\hline OLI_SU & ROLI_SU * WSP \\
\hline OLI_PPPS & ROLI_PPPS * WSP \\
\hline OLI_P & OLI_WCP + OLI_SU + OLI_GHI_P + OLI_GLI_P + OLI_PPPS \\
\hline OLI_PPS & OLI_PPPS + OLI_RETFC + OLI_RETFM + OLI_RETSL \\
\hline OLI_GHI & OLI_GHI_P + OLI_GHI_FC + OLI_GHI_SL \\
\hline OLI_GLI & OLI_GLI_P + OLI_GLI_FC + OLI_GLI_SL \\
\hline OLI_WC & OLI_WCP + OLI_WCSL \\
\hline SOCSL_WC & SOC_WCSL + SOC_WCP \\
\hline SOCF_UIFED & SOCF_UIFC + SOCF_UIFM \\
\hline SOCF_UIS & (SOC_UIP + SOC_UISL) * RUIWS1 / (RUIWS1 + RUIWS2) \\
\hline SOCF_UIF & (SOC_UIP + SOC_UISL) - SOCF_UIS \\
\hline SOCF_OASDI & OASDIP_L + OASDISL_L + OASDIFC_L + OASDIFM_L \\
\hline SOCF_HI & HIP_L + HISL_L + HIFC_L + HIFM_L \\
\hline
\end{tabular}

\subsection*{2.3 OASDI Covered Employment and Earnings (COV)}

Total At-Any-Time Employment
Ages 0 through 15 , where \(s=\) sex;a=age \(0,1,2,3, \ldots 15\); \(i=\) calendar year)

> he_m_sy(s,a,i) = (he_m_sy(s,a,histend) / nsy_a(s,a,histend) \&
> + he_m_sy(s,a,histend-1) / nsy_a(s,a,histend-1) \&
> \(\quad+\) he_m_sy(s,a,histend-2) / nsy_a(s,a,histend-2) ) / 3 \&
> \(\quad *\) nsy_a(s,a,i)

HI covered workers age groups 10-13 and 14-15, by sex and calendar year
\[
\begin{aligned}
& \text { he_m_1013(s,i) }=\text { sum(he_m_sy }(\mathrm{s}, 10: 13, \mathrm{i})) \\
& \text { he_m_1415(s,i) }=\text { sum(he_m_sy(s,14:15,i) }
\end{aligned}
\]

HI covered workers age group 15u, by sex and calendar year
he_m_15u(s,i) = sum(he_m_sy(s,0:9,i)) + he_m_1013(s,i) + he_m_1415(s,i)
OASDI covered = HI covered, by sex, single year of age, and calendar year
ce_m_sy(s,a,i)=he_m(s,a,i)

OASDI covered workers age groups 10-13 and 14-15, by sex and calendar year
```

ce_m_1013(s,i) = sum(ce_m_sy(s,10:13,i))
ce_m_1415(s,i) = sum(ce_m_sy(s,14:15,i)

```

OASDI covered workers age group 15u, by sex and calendar year
```

ce_m_15u(s,i) = sum(ce_m_sy(s,0:9,i)) +ce_m_1013(s,i) + ce_m_1415(s,i)

```

Male Disaggregates Aged 16 and Over
Preliminary
Average Weeks Worked
```

AWWM1617_P = 0.35417 * TREND_TE - 0.17568 * RM1617 + 8.80229;
AWWM1819_P = 0.28021 * TREND_TE - 0.16633 * RM1819 + 15.4522;
AWWM2024_P = 0.17338 * TREND_TE - 0.32843 * RM2024 + 31.4779;
AWWM2529_P = 0.04353 * TREND_TE - 0.32898 * RM2529 + 46.8304;
AWWM3034_P = 0.04353 * TREND_TE - 0.32898 * RM3034 + 46.8304;
AWWM3539_P = 0.03256 * TREND_TE - 0.26573 * RM3539 + 47.9512;
AWWM4044_P = 0.03256 * TREND_TE - 0.26573 * RM4044 + 47.9512;
AWWM4549_P = 0.03923 * TREND_TE - 0.23146 * RM4549 + 46.8025;
AWWM5054_P = 0.03923 * TREND_TE - 0.23146 * RM5054 + 46.8025;
AWWM5559_P = 0.04481 * TREND_TE - 0.28537 * RM5559 + 45.0196;
AWWM6064_P = 0.14488 * TREND_TE + 0.01910 * RM6064 + 29.5797;
AWWM6569_P = 0.21483 * TREND_TE - 0.23366 * RM6569 + 21.1343;
AWWM70O_P = 0.09069 * TREND_TE + 0.43206 * RM70O + 28.6903;

```
AWWM1617_PL \(=0.35417\) * TREND_TE. \(1-0.17568 *\) RM1617.1 + 8.80229;
AWWM1819_PL \(=0.28021\) * TREND_TE. \(1-0.16633\) * RM1819.1 + 15.4522;
AWWM2024_PL \(=0.17338 *\) TREND_TE.1-0.32843 * RM2024.1 + 31.4779;
AWWM2529_PL \(=0.04353\) * TREND_TE. \(1-0.32898 *\) RM2529.1 +46.8304 ;
AWWM3034_PL \(=0.04353\) * TREND_TE.1-0.32898 * RM3034.1 + 46.8304;
AWWM3539_PL \(=0.03256\) * TREND_TE. \(1-0.26573\) * RM3539.1 + 47.9512;
AWWM4044_PL \(=0.03256\) * TREND_TE. \(1-0.26573 *\) RM4044.1 + 47.9512;
AWWM4549_PL \(=0.03923\) * TREND_TE. \(1-0.23146 *\) RM4549.1 + 46.8025;
AWWM5054_PL \(=0.03923\) * TREND_TE. \(1-0.23146 *\) RM5054.1 + 46.8025;
AWWM5559_PL \(=0.04481\) * TREND_TE.1-0.28537 * RM5559.1 + 45.0196;
```

AWWM6064_PL = 0.14488 * TREND_TE.1 + 0.01910 * RM6064.1 + 29.5797
AWWM6569_PL = 0.21483 * TREND_TE.1 - 0.23366 * RM6569.1 + 21.1343;
AWWM70O_PL = 0.09069 * TREND_TE.1 + 0.43206 * RM70O.1 + 28.6903;

```

Work Experience
```

WEM1617_3_P = EM1617 * 52 / AWWM1617_P;
WEM1819_3_P = EM1819 * 52 / AWWM1819_P;
WEM2024_3_P = EM2024 * 52 / AWWM2024_P;
WEM2529_3_P = EM2529 * 52 / AWWM2529_P;
WEM3034_3_P = EM3034 * 52 / AWWM3034_P;
WEM3539_3_P = EM3539 * 52 / AWWM3539_P;
WEM4044_3_P = EM4044 * 52 / AWWM4044_P;
WEM4549_3_P = EM4549 * 52 / AWWM4549_P;
WEM5054_3_P = EM5054 * 52 / AWWM5054_P;
WEM5559_3_P = EM5559 * 52 / AWWM5559_P;
WEM6064_3_P = EM6064 * 52 / AWWM6064_P;
WEM6569_3_P = EM6569 * 52 / AWWM6569_P;
WEM70O_3_P = EM700 * 52 / AWWM70O_P

```
WEM1617_3_PL = EM1617.1 * 52 / AWWM1617_PL;
WEM1819_3_PL = EM1819.1 * \(52 /\) AWWM1819_PL;
WEM2024_3_PL = EM2024.1 * 52 / AWWM2024_PL;
WEM2529_3_PL = EM2529.1 * 52 / AWWM2529_PL;
WEM3034_3_PL = EM3034.1 * 52 / AWWM3034_PL;
WEM3539_3_PL = EM3539.1 * 52 / AWWM3539_PL;
WEM4044_3_PL = EM4044.1 * 52 / AWWM4044_PL;
WEM4549_3_PL = EM4549.1 * 52 / AWWM4549_PL;
WEM5054_3_PL = EM5054.1 * 52 / AWWM5054_PL;
WEM5559_3_PL = EM5559.1 * 52 / AWWM5559_PL;
WEM6064_3_PL = EM6064.1 * 52 / AWWM6064_PL;
WEM6569_3_PL = EM6569.1 * 52 / AWWM6569_PL;
WEM70O_3_PL = EM700.1 * 52 / AWWM70O_PL ;

Total Employed
TEM1617_P = ((WEM1617_3_P / WEM1617_3_PL) * (TEM1617.1 - NM1617M.1) + NM1617M) * MULT1_TEM1617 * MULT2_TEM1617; TEM1819_P \(=((\) WEM1819_3_P \(/\) WEM1819_3_PL \() *(\) TEM1819.1 - NM1819M.1 \()+\) NM1819M \() *\) MULT1_TEM1819 * MULT2_TEM1819; TEM2024_P = ((WEM2024_3_P / WEM2024_3_PL) * (TEM2024.1 - NM2024M.1) + NM2024M) * MULT1_TEM2024 * MULT2_TEM2024;

TEM2529_P = ((WEM2529_3_P / WEM2529_3_PL) * (TEM2529.1 - NM2529M.1) + NM2529M) * MULT1_TEM2529 * MULT2_TEM2529;
TEM3034_P = ((WEM3034_3_P / WEM3034_3_PL) * (TEM3034.1 - NM3034M.1) + NM3034M) * MULT1_TEM3034 * MULT2_TEM3034;

TEM4044_P = ((WEM4044_3_P / WEM4044_3_PL) * (TEM4044.1 - NM4044M.1) + NM4044M) * MULT1_TEM4044 * MULT2_TEM4044;
TEM4549_P = ((WEM4549_3_P / WEM4549_3_PL) * (TEM4549.1 - NM4549M.1) + NM4549M) * MULT1_TEM4549 * MULT2_TEM4549;
TEM5054_P = ((WEM5054_3_P / WEM5054_3_PL) * (TEM5054.1 - NM5054M.1) + NM5054M) * MULT1_TEM5054 * MULT2_TEM5054;
TEM5559_P = ((WEM5559_3_P / WEM5559_3_PL) * (TEM5559.1 - NM5559M.1 \()\) + NM5559M \()\) * MULT1_TEM5559 * MULT2_TEM5559; TEM6064_P = ((WEM6064_3_P / WEM6064_3_PL) * (TEM6064.1)) * MULT1_TEM6064 * MULT2_TEM6064;



WEM16O_3_P = WEM1617_3_P + WEM1819_3_P + WEM2024_3_P + WEM2529_3_P + WEM3034_3_P + WEM3539_3_P + WEM4044_3_P + WEM4549_3_P
+ WEM5054_3_P + WEM5559_3_P + WEM6064_3_P + WEM6569_3_P + WEM70O_3_P;
AWWM16O_P = EM16O * \(52 /\) WEM16O_3_P;

TEM16O_P = TEM1617_P + TEM1819_P + TEM2024_P + TEM2529_P + TEM3034_P + TEM3539_P + TEM4044_P + TEM4549_P + TEM5054_P + TEM5559_P + TEM6064_P + TEM6569_P + TEM70O_P;

Final (Pre-TE.ADD)
Average Weeks Worked
AWWM1617 = AWWM1617_P;
AWWM1819 = AWWM1819_P;
AWWM2024 = AWWM2024_P;
AWWM2529 = AWWM2529_P;
AWWM3034 = AWWM3034_P;
AWWM3539 = AWWM3539_P;
AWWM4044 = AWWM4044_P;
AWWM4549 = AWWM4549_P;
AWWM5054 = AWWM5054_P;

AWWM5559 = AWWM5559_P;
AWWM6064 = AWWM6064_P;
AWWM6569 = AWWM6569_P;
AWWM700 = AWWM70O_P;

Work Experience
WEM1617_3 = WEM1617_3_P;
WEM1819_3 = WEM1819_3_P;
WEM2024_3 = WEM2024_3_P;
WEM2529_3 = WEM2529_3_P; WEM3034_3 = WEM3034_3_P;
WEM3539_3 = WEM3539_3_P;
WEM4044_3 = WEM4044_3_P;
WEM4549_3 = WEM4549_3_P;
WEM5054_3 = WEM5054_3_P;
WEM5559_3 = WEM5559_3_P;
WEM6064_3 = WEM6064_3_P;
WEM6569_3 = WEM6569_3_P;
WEM70O_3 = WEM70O_3_P;

Total Employed
TEM1617 = TEM1617_P;
TEM1819 = TEM1819_P;
TEM2024 = TEM2024_P;
TEM2529 = TEM2529_P;
TEM3034 = TEM3034_P;
TEM3539 = TEM3539_P;
TEM4044 = TEM4044_P;
TEM4549 = TEM4549_P;
TEM5054 = TEM5054_P;

TEM5559 = TEM5559_P;
TEM6064 = TEM6064_P;
TEM6569 = TEM6569_P
TEM700 = TEM700_P ;

WEM16O_3 = WEM16O_3_P;
AWWM16O = AWWM16O_P;
TEM16O = TEM16O_P ;

Female Disaggregates Aged 16 and Over
Preliminary
Average Weeks Worked
```

AWWF1617_P = 0.44829 * TREND_TE - 0.06786 * RF1617 + 0.37331;
AWWF1819_P = 0.24262 * TREND_TE - 0.00950 * RF1819 + 17.0791;
AWWF2024_P = 0.18291 * TREND_TE - 0.06453 * RF2024 + 26.6823;
AWWF2529_P = 0.07674 * TREND_TE - 0.10131 * RF2529 + 40.0210;
AWWF3034_P = 0.07674 * TREND_TE - 0.10131 * RF3034 + 40.0210;
AWWF3539_P = 0.05004 * TREND_TE - 0.10322 * RF3539 + 43.9974;
AWWF4044_P = 0.05004 * TREND_TE - 0.10322 * RF4044 + 43.9974;
AWWF4549_P = 0.08209 * TREND_TE - 0.17307 * RF4549 + 41.1649;
AWWF5054_P = 0.08209 * TREND_TE - 0.17307 * RF5054 + 41.1649;
AWWF5559_P = 0.04868 * TREND_TE + 0.17072 * RF5559 + 41.9580;
AWWF6064_P = 0.14339 * TREND_TE + 0.01918 * RF6064 + 29.1567;
AWWF6569_P = 0.01857 * TREND_TE + 0.64199 * RF6569 + 36.1193;
AWWF70O_P = 0.20193 * TREND_TE + 0.92866 * RF70O + 14.2412;

```

AWWF1617_PL \(=0.44829 *\) TREND_TE. \(1-0.06786 *\) RF1617.1 + 0.37331; AWWF1819_PL \(=0.24262\) * TREND_TE. \(1-0.00950 *\) RF1819.1 + 17.0791; AWWF2024_PL \(=0.18291\) * TREND_TE. \(1-0.06453\) * RF2024.1 + 26.6823;

AWWF2529_PL \(=0.07674\) * TREND_TE. \(1-0.10131 *\) RF2529.1 + 40.0210; AWWF3034_PL \(=0.07674 *\) TREND_TE. \(1-0.10131 *\) RF3034.1 +40.0210 ; AWWF3539_PL \(=0.05004\) * TREND_TE. \(1-0.10322\) * RF3539.1 + 43.9974; AWWF4044_PL \(=0.05004\) * TREND_TE. \(1-0.10322\) * RF4044.1 + 43.9974; AWWF4549_PL \(=0.08209\) * TREND_TE. \(1-0.17307\) * RF4549.1 + 41.1649; AWWF5054_PL \(=0.08209\) * TREND_TE. \(1-0.17307\) * RF5054.1 + 41.1649;

AWWF5559_PL \(=0.04868 *\) TREND_TE. \(1+0.17072 *\) RF5559.1 + 41.9580; AWWF6064_PL \(=0.14339\) * TREND_TE. \(1+0.01918\) * RF6064.1 + 29.1567; AWWF6569_PL \(=0.01857\) * TREND_TE. \(1+0.64199 *\) RF6569.1 +36.1193 ; AWWF70O_PL \(=0.20193\) * TREND_TE. \(1+0.92866\) * RF70O. 1 + 14.2412;

Work Experience
```

WEF1617_3_P = EF1617 * 52 / AWWF1617_P;
WEF1819_3_P = EF1819 * 52 / AWWF1819_P;
WEF2024_3_P = EF2024 * 52 / AWWF2024_P;
WEF2529_3_P = EF2529 * 52 / AWWF2529_P;
WEF3034_3_P = EF3034 * 52 / AWWF3034_P;
WEF3539_3_P = EF3539 * 52 / AWWF3539_P;
WEF4044_3_P = EF4044 * 52 / AWWF4044_P;
WEF4549_3_P = EF4549 * 52 / AWWF4549_P;
WEF5054_3_P = EF5054 * 52 / AWWF5054_P;
WEF5559_3_P = EF5559 * 52 / AWWF5559_P;
WEF6064_3_P = EF6064 * 52 / AWWF6064_P;
WEF6569_3_P = EF6569 * 52 / AWWF6569_P;
WEF70O_3_P = EF70O * 52 / AWWF70O_P;

```
WEF1617_3_PL = EF1617.1 * 52 / AWWF1617_PL;
WEF1819_3_PL = EF1819.1 * 52 / AWWF1819_PL;
WEF2024_3_PL = EF2024.1 * 52 / AWWF2024_PL;
WEF2529_3_PL = EF2529.1 * 52 / AWWF2529_PL;
WEF3034_3_PL = EF3034.1 * 52 / AWWF3034_PL;
WEF3539_3_PL = EF3539.1 * 52 / AWWF3539_PL;
WEF4044_3_PL = EF4044.1 * 52 / AWWF4044_PL;
WEF4549_3_PL = EF4549.1 * 52 / AWWF4549_PL;
WEF5054_3_PL = EF5054.1 * \(52 /\) AWWF5054_PL;
WEF5559_3_PL = EF5559.1 * 52 / AWWF5559_PL;
WEF6064_3_PL = EF6064.1 * 52 / AWWF6064_PL;
WEF6569_3_PL = EF6569.1 * 52 / AWWF6569_PL;
WEF70O_3_PL = EF700.1 * 52 / AWWF700_PL ;

Total Employed
```

TEF1617_P = ((WEF1617_3_P / WEF1617_3_PL) * (TEF1617.1 - NF1617M.1) + NF1617M) * MULT1_TEF1617 * MULT2_TEF1617;
TEF1819_P = ((WEF1819_3_P / WEF1819_3_PL) * (TEF1819.1 - NF1819M.1) + NF1819M) * MULT1_TEF1819 * MULT2_TEF1819;
TEF2024_P = ((WEF2024_3_P / WEF2024_3_PL) * (TEF2024.1 - NF2024M.1) + NF2024M) * MULT1_TEF2024 * MULT2_TEF2024;
TEF2529_P = ((WEF2529_3_P / WEF2529_3_PL) * (TEF2529.1 - NF2529M.1) + NF2529M) * MULT1_TEF2529 * MULT2_TEF2529;
TEF3034_P = ((WEF3034_3_P / WEF3034_3_PL) * (TEF3034.1 - NF3034M.1) + NF3034M) * MULT1_TEF3034 * MULT2_TEF3034;
TEF3539_P = ((WEF3539_3_P / WEF3539_3_PL) * (TEF3539.1 - NF3539M.1) + NF3539M) * MULT1_TEF3539 * MULT2_TEF3539;
TEF4044_P = ((WEF4044_3_P / WEF4044_3_PL) * (TEF4044.1 - NF4044M.1) + NF4044M) * MULT1_TEF4044 * MULT2_TEF4044;
TEF4549_P = ((WEF4549_3_P / WEF4549_3_PL) * (TEF4549.1 - NF4549M.1) + NF4549M) * MULT1_TEF4549 * MULT2_TEF4549;
TEF5054_P = ((WEF5054_3_P / WEF5054_3_PL) * (TEF5054.1 - NF5054M.1) + NF5054M) * MULT1_TEF5054 * MULT2_TEF5054;
TEF5559_P = ((WEF5559_3_P / WEF5559_3_PL) * (TEF5559.1 - NF5559M.1) + NF5559M) * MULT1_TEF5559 * MULT2_TEF5559;
TEF6064_P = ((WEF6064_3_P / WEF6064_3_PL) * (TEF6064.1)) * MULT1_TEF6064 * MULT2_TEF6064;
TEF6569_P = ((WEF6569_3_P / WEF6569_3_PL) * (TEF6569.1)) * MULT1_TEF6569 * MULT2_TEF6569;
TEF70O_P = ((WEF70O_3_P / WEF70O_3_PL ) * (TEF70O.1 )) * MULT1_TEF70O * MULT2_TEF70O;
WEF16O_3_P = WEF1617_3_P + WEF1819_3_P + WEF2024_3_P + WEF2529_3_P + WEF3034_3_P + WEF3539_3_P + WEF4044_3_P +
WEF4549_3_P
+ WEF5054_3_P + WEF5559_3_P + WEF6064_3_P + WEF6569_3_P + WEF70O_3_P;
AWWF16O_P = EF16O * 52 / WEF16O_3_P;
TEF16O_P = TEF1617_P + TEF1819_P + TEF2024_P + TEF2529_P + TEF3034_P + TEF3539_P + TEF4044_P + TEF4549_P + TEF5054_P
+ TEF5559_P}+ TEF6064_P + TEF6569_P + TEF70O_P;

```

Final (Pre-TE.ADD)
Average Weeks Worked
AWWF1617 = AWWF1617_P; AWWF1819 = AWWF1819_P; AWWF2024 = AWWF2024_P;

AWWF2529 = AWWF2529_P; AWWF3034 = AWWF3034_P; AWWF3539 = AWWF3539_P; AWWF4044 = AWWF4044_P; AWWF4549 = AWWF4549_P; AWWF5054 = AWWF5054_P;

AWWF5559 = AWWF5559_P; AWWF6064 = AWWF6064_P; AWWF6569 = AWWF6569_P; AWWF70O = AWWF70O_P;

Work Experience
WEF1617_3 = WEF1617_3_P; WEF1819_3 = WEF1819_3_P; WEF2024_3 = WEF2024_3_P;

WEF2529_3 = WEF2529_3_P; WEF3034_3 = WEF3034_3_P; WEF3539_3 = WEF3539_3_P; WEF4044_3 = WEF4044_3_P; WEF4549_3 = WEF4549_3_P; WEF5054_3 = WEF5054_3_P;

WEF5559_3 = WEF5559_3_P; WEF6064_3 = WEF6064_3_P; WEF6569_3 = WEF6569_3_P; WEF70O_3 = WEF70O_3_P;

Total Employed
```

TEF1617 = TEF1617_P;
TEF1819 = TEF1819_P;
TEF2024 = TEF2024_P;
TEF2529 = TEF2529_P;
TEF3034 = TEF3034_P;
TEF3539 = TEF3539_P;
TEF4044 = TEF4044_P;
TEF4549 = TEF4549_P;
TEF5054 = TEF5054_P;
TEF5559 = TEF5559_P;
TEF6064 = TEF6064_P;
TEF6569 = TEF6569_P;
TEF70O = TEF70O_P;

```
WEF16O_3 = WEF16O_3_P;
AWWF16O = AWWF16O_P;
TEF16O = TEF16O_P ;

Combined, Age 16 and Over
WE16O_3_P = WEM16O_3_P + WEF16O_3_P;
AWW16O_P \(=\) E16O * \(52 /\) WE16O_3_P;
WE16O_3 = WE16O_3_P ;
AWW16O = AWW16O_P ;
Self-Employed Only
SEOCMB \(=0.039 *(\) TEFC_N_N + TESL_N_N + EPRRB \()\)
SEOCMBL1 \(=0.039 *(\) TEFC_N_N. \(1+\) TESL_N_N. \(1+\) EPRRB. 1\()\)
SEO \(=(\) SEO \(1 *(\) EAS + ENAS \() /(\) EAS. \(1+\) ENAS. 1\()+(\) SEOCMB - SEOCMBL1 \()) *\) MULTSEO
Combination Workers
CMB_TOT \(=((-0.01468+0.06227 *\) RTP.1-0.0008 \() *\) WSWA - SEOCMB \() *\) MULTCMB
CSW_TOT \(=\) SEO + CMB_TOT
AW_CMBTOT \(=1.4953 *\) ACWA
W_CMBTOT = AW_CMBTOT * CMB_TOT
CMB_WRELMAX = TAXMAX/AW_CMBTOT
CMB Wage Andover Curve
CMB_WAO1 = IF (CMB_WRELMAX < 0.0543009)
    THEN 1-0.722659 * CMB_WRELMAX \({ }^{0.65}-0.461913 *\) CMB_WRELMAX \(^{0.8}\)
    ELSE IF (CMB_WRELMAX \(<0.1086018\) )
        THEN \(-1.02884 *\) CMB_WRELMAX \({ }^{0.6}+0.324761 *\) CMB_WRELMAX \(^{1.6}+1.02015\)
        ELSE IF (CMB_WRELMAX < 0.1629027)
                        THEN -0.906607 * CMB_WRELMAX \({ }^{0.7}+0.947662\)
                    ELSE IF (CMB_WRELMAX < 0.2172037)
                            THEN -0.813951 * CMB_WRELMAX \({ }^{0.55}+0.991722\)
                            ELSE IF (CMB_WRELMAX < 0.3258055)
                            THEN -0.755135 * CMB_WRELMAX \({ }^{0.55}+0.964593\)
                            ELSE 0
CMB_WAO2 = IF (CMB_WRELMAX \(<0.5430091\) )
        THEN -0.649755 * CMB_WRELMAX \({ }^{0.6}+0.886467\)
        ELSE IF (CMB_WRELMAX < 0.7059119)
        THEN -0.573205 * CMB_WRELMAX \({ }^{0.7}+0.810122\)
            ELSE IF (CMB_WRELMAX < 0.9231155)
            THEN - 5.22264 * CMB WRELMAX \({ }^{0.06}+5.47514\)
            ELSE IF (CMB_WRELMAX < 1.0860183)
            THEN - 2.02619 * CMB_WRELMAX \({ }^{0.15}+2.27963\)
                ELSE IF (CMB_WRELMAX < 1.5204256)
```

                    THEN 0.605192 * EXP (-0.2 * CMB_WRELMAX) - 0.827158 * EXP (-0.8 * CMB_WRELMAX)
    ```
                    +1.52918 * EXP (-1.5 * CMB_WRELMAX) - 0.212269
                    ELSE 0
```

CMB_WAO3 = IF (CMB_WRELMAX $<1.8462311$ )
THEN 0.19139 * EXP ( -0.6 * CMB_WRELMAX $)+0.764408$ * EXP ( 1.8 * CMB_WRELMAX $)+0.0194903$
ELSE IF (CMB_WRELMAX < 2.3077888)
THEN 0.12964 * EXP ( -0.5 * CMB_WRELMAX) + 0.644861 * EXP (- 1.5 * CMB_WRELMAX) + 0.0183343
ELSE IF (CMB_WRELMAX < 2.9865502)
THEN 0.361318 * EXP ( -0.8 * CMB_WRELMAX) +0.0219491
ELSE IF (CMB_WRELMAX < 4.3440731)
THEN 0.193202 * EXP ( -0.45 * CMB_WRELMAX) + 0.00425171
ELSE IF (CMB_WRELMAX<5.4300913)
THEN 0.0560412 * EXP ( -0.25 * CMB_WRELMAX) + 0.311286 * EXP ( -0.8 * CMB_WRELMAX $)+$
0.00297316
ELSE 0
CMB_WAO4 = IF (CMB_WRELMAX $<13.5752283$ )
THEN 0.0995677 * EXP ( -0.32 * CMB_WRELMAX $)+0.00355234$
ELSE IF (CMB_WRELMAX < 21.7203653)
THEN 0.041159 * EXP ( -0.19 * CMB_WRELMAX $)+0.00156765$
ELSE IF (CMB_WRELMAX < 678.7614168)
THEN 0.265022 * CMB_WRELMAX ${ }^{(-1.555)}$
ELSE 0
CMB_WAO $=$ IF $\left(C M B \_W R E L M A X ~<0.3258055\right)$
THEN CMB_WAO1
ELSE IF (CMB_WRELMAX < 1.5204256)
THEN CMB_WAO2
ELSE IF (CMB_WRELMAX < 5.4300913)
THEN CMB_WAO3
ELSE CMB_WAO4
CMB $\quad=\left(1-\left(\mathrm{CMB} \_W A O-0.019\right)\right) *$ CMB_TOT
CSW $=\mathrm{SEO}+\mathrm{CMB}$
SEOCMB_HI $=0.039 *\left(T E F C \_N \_N+T E S L \_N \_N \_H I\right)$
SEO_HI $=$ SEO - SEOCMB_HI
CMB_HI $=$ CMB_TOT + SEOCMB_HI
CSW_HI $=$ SEO_HI + CMB_HI

```

NIPA Wages
Private Residual Sector
\begin{tabular}{lll} 
PSDPB & \(=\) & WSDP - WSPH - WSPF - WSPRRB - TIPS_SR \\
WSIPS_SR & \(=\) & \((0.000508328 *\) RTP -0.000481700\() *\) GDP * 1.26393
\end{tabular}

OASDI Wages
\begin{tabular}{l} 
Covered Employment and Wages - Federal Civilian Government \\
TEFC
\end{tabular}\(=\quad(\) TEFC. \(1 /\) EGGEFC.1) \(*\) EGGEFC
\begin{tabular}{lll} 
Covered Employment and Wages - State and Local Govt. \\
TESL & \(=\) & (TESL.1/EGGESLMAX.1) * EGGESLMAX \\
TESL_O & \(=\) & (TESL_O.1/TESL.1) * TESL \\
TESL_N & \(=\) & (TESL-TESL_O) \\
TESL_N_O & \(=\) & (TESL_N_O_HI + TESL_N_O_NHI) \\
TESL_N_O_HI & \(=\) & (TESL_N - TESL_N_O_NHI - TESL_N_N_NHI) * CER_MQGE_O \\
TESL_N_O_NHI & \(=\) & (TESL_N_O_NHI_S + TESL_N_O_NHI_E + TESL_N_O_NHI_NS)
\end{tabular}
```

```
= TESL_N_S.1 * (NF1819 + NF2024 + NM1819 + NM2024) / (NF1819.1 + NF2024.1 + NM1819.1 +
```

```
= TESL_N_S.1 * (NF1819 + NF2024 + NM1819 + NM2024) / (NF1819.1 + NF2024.1 + NM1819.1 +
NM2024.1)
NM2024.1)
= TESL_N_E. 1 * (TESL / TESL.1)
```

= TESL_N_E. 1 * (TESL / TESL.1)

```
```

= TESL_N_S * (TESL_N_O_NHI_S.1/TESL_N_S.1)

```
= TESL_N_S * (TESL_N_O_NHI_S.1/TESL_N_S.1)
= TESL_N_E * 0.6
= TESL_N_E * 0.6
= TESL_N_O_NHI_NS.1 * ESR_NS
= TESL_N_O_NHI_NS.1 * ESR_NS
= (TESL_N-TESL_N_O)
= (TESL_N-TESL_N_O)
= (TESL_N_N - TESL_N_N_NHI)
= (TESL_N_N - TESL_N_N_NHI)
= (TESL_N_N_NHI_S + TESL_N_N_NHI_E + TESL_N_N_NHI_NS)
= (TESL_N_N_NHI_S + TESL_N_N_NHI_E + TESL_N_N_NHI_NS)
= (TESL_N_S - TESL_N_O_NHI_S)
= (TESL_N_S - TESL_N_O_NHI_S)
= (TESL_N_E - TESL_N_O_NHI_E)
= (TESL_N_E - TESL_N_O_NHI_E)
= TESL_N_N_NHI_NS.1 * ESR_NS
= TESL_N_N_NHI_NS.1 * ESR_NS
= (WESL.1/WSGGESL.1) * WSGGESL
= (WESL.1/WSGGESL.1) * WSGGESL
= (WESL_O.1/WSGGESL.1) * WSGGESL
= (WESL_O.1/WSGGESL.1) * WSGGESL
= (WESL - WESL_O)
= (WESL - WESL_O)
= (WESL_N - WESL_N_NHI)
= (WESL_N - WESL_N_NHI)
= (WESL_N_NHI_S + WESL_N_NHI_E + WESL_N_NHI_NS)
= (WESL_N_NHI_S + WESL_N_NHI_E + WESL_N_NHI_NS)
= WESL_N_NHI_S.1 * (TESL_N_S/TESL_N_S.1) * (AWSGGESL/AWSGGESL.1)
= WESL_N_NHI_S.1 * (TESL_N_S/TESL_N_S.1) * (AWSGGESL/AWSGGESL.1)
= WESL_N_NHI_E. }1*\mathrm{ * (TESL_N_E/TESL_N_E.1) * (AWSGGESL/AWSGGESL.1)
= WESL_N_NHI_E. }1*\mathrm{ * (TESL_N_E/TESL_N_E.1) * (AWSGGESL/AWSGGESL.1)
= IF (AWR_NS = 0) THEN 0 ELSE AWR_NS/AWR_NS. }
= IF (AWR_NS = 0) THEN 0 ELSE AWR_NS/AWR_NS. }
= IF (ESR_NS = 0) THEN 0 ELSE WESL_N_NHI_NS.1 * (TESL_N_O_NHI_NS + TESL_N_N_NHI_NS)/
= IF (ESR_NS = 0) THEN 0 ELSE WESL_N_NHI_NS.1 * (TESL_N_O_NHI_NS + TESL_N_N_NHI_NS)/
(TESL_N_O_NHI_NS.1 + TESL_N_N_NHI_NS.1) *(AWSGGESL/AWSGGESL.1) * RAWR_NS
```

(TESL_N_O_NHI_NS.1 + TESL_N_N_NHI_NS.1) *(AWSGGESL/AWSGGESL.1) * RAWR_NS

```
TESL_N_O_NHI_S
TESL_N_O_NHI_E
TESL_N_O_NHI_NS
TESL_N_N
TESL_N_N_HI
TESL_N_N_NHI
TESL_N_N_NHI_S
TESL_N_N_NHI_E
TESL_N_N_NHI_NS
WESL
WESL_O
WESL_N
WESL_N_HI
WESL_N_NHI
WESL_N_NHI_S
WESL_N_NHI_E
RAWR NS
WESL_N_NHI_NS

Self-Employed Earnings Sector
\begin{tabular}{|c|c|}
\hline Covered SENE & \\
\hline CSE_TOT & \(=(\mathrm{YF}+\mathrm{YNF}) /(\mathrm{YF} .1+\mathrm{YNF} .1) *\) CSE_TOT. 1 \\
\hline CSE_CMB_N &  \\
\hline CSE & = CSE_TOT - CSE_CMB_N \\
\hline ACSE_SEO & \(=\quad(\mathrm{CSE}\) _TOT \(/(\mathrm{SEO}+0.416488 *\) CMB_TOT \()\) ) \\
\hline ACSE_CMB_TOT & \(=0.416488 *\) ACSE_SEO \\
\hline CSE_SEO & \(=\) ACSE_SEO * SEO \\
\hline CSE_CMB_TOT & \(=\) ACSE_CMB_TOT \({ }^{*}\) CMB_TOT \\
\hline CSE_CMB & \(=\) CSE_CMB_TOT - CSE_CMB_N \\
\hline ACSE_CMB & \(=\) CSE_CMB/CMB \\
\hline
\end{tabular}
\begin{tabular}{ll} 
Present Law OASDI and HI Covered Wages and Earnings \\
WSGMLC & \(=\) CML * WSGFM \\
WSGFCA & \(=\) WEFC_O \\
CFCA & \(=\) WSGFCA/WSGGEFC \\
CSLHI & \(=\) (WESL_O+WESL_N_HI)/WSGGESL \\
WSGSLCA & \(=\) WESL_O \\
WSPH_O & \(=\) CPH *WSPH \\
WSPF_O & \(=\) WSPF_O.1 * WSPF/WSPF.1 \\
CPF & \(=\) WSPF_O/WSPF \\
WSPRR_O & \(=\) CPRR *WSPRRB \\
WSPC & \(=\) WSPH_O + WSPF_O + WSPRR_O + TIPS_SR + WSPB_O \\
CP & \(=\) WSPC/WSDP \\
WSCA & \(=\) WSPC + WSGSLCA + WSGFCA + WSGMLC) \\
COVERNA & \(=\) WSCA + CSE) \\
ACWA & \(=\) WSCA/WSWA \\
ASE & \(=\) CSE/CSW \\
ASEHI & \(=\) CSE_TOT/CSW_HI \\
ACEA & \(=\) COVERNA/TCEA \\
ACSLW & \(=\) WESL_O/TESL_O * MULTACSLW \\
ACMW & \(=\) ACMW.1 * AWSGFM/AWSGFM.1 \(*\) CML/CML.1 \\
ACFCW & \(=\) WEFC_O/TEFC_O \\
ACFMW & \(=\) ACFMW.1 * (AIW.1/AIW.3) \({ }^{0.5}\) \\
TEPH_N & \(=\) ENAWPH * (1-CPH) \\
TEP_N_N_S & \(=\) TEP_N_N_S.1 * (NF1819 + NF2024 + NM1819 + NM2024) / (NF1819.1 + NF2024.1 + NM1819.1 + NM2024.1) \\
WSWA & \(=\) (TCEA-SEO)
\end{tabular}
Present Law HI Covered Wages and Earnings
\begin{tabular}{ll} 
WSCAHI_ADD & \(=\) \\
WSCA * WSCAHI_ADD.1/WSCA. 1 \\
TCEAHI & \(=\) \\
WSWAHI & \(=\) \\
TCEA + TEFC_N_N + TESL_N_N_HI)
\end{tabular}
```

WSCAHI = WSCA + WEFC_N + WESL_N_HI + WSCAHI_ADD
ACWAHI = WSCAHI/WSWAHI
COVERNHI = WSCAHI + CSE_TOT
ACEAHI = COVERNHI/TCEAHI

```
```

Complete Coverage concepts
WSWC = (WSWAHI + TEPH_N + EPRRB + TEP_N_N_S + TEPO_N + TESL_N_N_NHI) + LOST_MF
ACWC = WSD/WSWC
AIW = IF AIW_GR_YR = 0
THEN AIW.1 *ACWC/ACWC. 1 * MULTAIW
ELSE AIW.1 * (1 +AIW_GR/100)

```
\begin{tabular}{|c|c|c|}
\hline \multicolumn{3}{|l|}{Taxable Maximums} \\
\hline RAIW & \(=\) & AIW.2/AIWBASE \\
\hline TAXMAXB1 & = & RAIW * TMAXBASE * 1000/300 \\
\hline TAXMAXB2 & = & IF TAXMAXB1 - ROUND (TAXMAXB1) >= 0.5 THEN ROUND (TAXMAXB1) +1 \\
\hline & & ELSE ROUND (TAXMAXB1) \\
\hline \multirow[t]{3}{*}{TAXMAXB3} & \multirow[t]{3}{*}{\(=\)} & IF TAXMAXB2 < TAXMAX. 1 \\
\hline & & THEN TAXMAX. 1 * 1000/300 \\
\hline & & ELSE TAXMAXB2 \\
\hline \multirow[t]{3}{*}{TAXMAX} & \multirow[t]{3}{*}{\(=\)} & IF BENINC. \(1<=0.001\) \\
\hline & & THEN TAXMAX. 1 \\
\hline & & ELSE 300 * TAXMAXB3/1000 \\
\hline
\end{tabular}
Deemed Military Wage Credits
EDMILAF = EDMIL * 1.1
EDMILT \(=(2.00303-50.7517 / Y E A R) *\) EDMILAF
EDMILR = EDMILT - EDMILAF
MWC_ED_O \(=1.2 *\) EDMILAF * 0.997
MWC_ED_HI \(=1.2 *\) EDMILAF
AMWC_GO2 \(=\operatorname{MIN}(1.2, \operatorname{AWSGFM} *(2 / 52) *(1 / 3))\)
MWC_EDR_O = AMWC_GO2 * EDMILR * (1-0.017)
MWC_EDR_HI= MWC_EDR_O + ((1.2 +AMWC_GO2) * 0.5) * EDMILR * 0.017
MWC_O = MWC_ED_O + MWC_EDR_O
MWC_HI \(=\) MWC_ED_HI+ MWC_EDR_HI

\section*{2-4 Effective Taxable Payroll (TAXPAY)}

\section*{2-4.1 Ratio of taxable employee to total covered OASDI wages (RWTEE)}
```

if (relmax > 10.46164264d0) then
RWTEE = -(0.19381d0 / 0.75d0) * relmax**(-0.75d0) + \&
1.000181751d0
else if (relmax > 3.24372838d0) then
RWTEE = -(0.00896d0 / 0.15d0) * exp(-0.15d0 * relmax) - \&
(0.05161d0 / 0.40d0) * exp(-0.40d0 * relmax) - \&
(0.79023d0 / 1.25d0) * exp(-1.25d0 * relmax) + \&
0.00057d0 * relmax + 0.964197529d0
else if (relmax > 1.95516643d0) then
RWTEE =-(0.076346d0 / 0.25d0)* exp(-0.25d0 * relmax) - \&
(1.34462d0 / 1.50d0) * exp(-1.50d0 * relmax) - \&
0.01046d0 * relmax + 1.059665027d0
else if (relmax > 1.03872350d0) then
RWTEE = (0.17160d0 / 0.25d0) * exp(-0.25d0 * relmax) - \&
(1.27296d0 / 1.2d0) * exp(-1.2d0 * relmax) + \&
0.09125d0 * relmax + 0.30631048d0
else if (relmax > 0.46210354d0) then
RWTEE = (0.50949d0 / 1.5d00) * relmax**1.50d0 - \&
(1.95934d0 / 0.69d0) * exp(-0.69d0 * relmax) - \&
1.15006d0 * relmax + 2.847263595d0
else
RWTEE = relmax - \&
(0.31285d0 / 1.5d0) * relmax**1.5d0 - \&
(0.35995d0 / 2.00d0) * relmax**2d0
end if
Where
relmaxc = OASDI taxable maximum / average covered OASDI wage
RWTEE = Ratio of OASDI taxable employee to covered wages
2-4.2 Taxable employee OASDI wages (WTEE)
WTEE = RWTEE * WSC
Where
RWTEE = Ratio of OASDI taxable employee to covered wages
WSC = OASDI total covered wages
WTEE = OASDI taxable employee wages
2-4.3 Ratio of multi-employer refund wages to total OASDI covered wages (RMER)
RMER = ( MER(-1) / WSC(-1) ) - 0.03217 * ( RWTEE - RWTEE(-1) ) - 0.00024*(RU - RU(-1) )
Where
MER(-1) = Multi-employer refund wages in prior year
RMER = Ratio of multi-employer refund wages to total OASDI covered wages
RU = Annual average civilian unemployment rate
RWTEE = Ratio of OASDI taxable employee to covered wages
WSC(-1) = OASDI total covered wages in prior year
2-4.4 Multi-employer refund wages (MER)
MER = RMER * WSC
Where
MER = OASDI multi-employer refund wages
RMER = Ratio of multi-employer refund wages to total OASDI covered wages
WSC = OASDI total covered wages
2-4.5 Taxable employer OASDI wages (WTER)
WTER = WTEE + MER

```

Where
\begin{tabular}{ll} 
MER & \(=\) OASDI multi-employer refund wages \\
WTEE & \(=\) OASDI taxable employee wages \\
WTER & \(=\) OASDI taxable employer wages
\end{tabular}

\section*{2-4.6 Ratio of taxable to covered self-employment earnings (RSET)}

\section*{Preliminary}
BASECT \(=47831.98\)
BASECW \(=36831.79\)
BASEO \(=23448.56\)

\section*{Self-employed only}

SECSEO = CSE - SECCMB
ASESEO = SECSEO / SEO
ASEO96 = ASESEO(2011)
ASESEO = ASESEO * BASEO / ASEO2011
\(\mathrm{O}=\) TAXMAX / ASESEO
if ( \(\mathrm{O}<0.030138750 \mathrm{~d} 0\) ) then \(\mathrm{OTR}=\mathrm{O}-(8.07753 \mathrm{~d} 0 / 2.5 \mathrm{~d} 0) * \mathrm{O}^{* *} 2.5 \mathrm{~d} 0\)
else if ( \(\mathrm{O}<0.120555001 \mathrm{~d} 0\) ) then OTR \(=-(1.11653 \mathrm{~d} 0 / 1.5 \mathrm{~d} 0) * \mathrm{O}^{* *} 1.5 \mathrm{~d} 0+(1.07133 \mathrm{~d} 0 / 2.15 \mathrm{~d} 0) * \mathrm{O}^{* *} 2.15 \mathrm{~d} 0-\&\) (5.49830d0 / 3.3d0) * O**3.3d0 + 1.12622d0 * O -0.000670705 d 0
else if ( \(\mathrm{O}<0.2411100020 \mathrm{~d} 0\) ) then
OTR \(=-(4.16459 \mathrm{~d} 0 / 1.3 \mathrm{~d} 0){ }^{*} \mathrm{O}^{* * 1.3 d 0 ~+~(4.56181 \mathrm{~d} 0 / 1.75 \mathrm{~d} 0) ~ * ~} \mathrm{O}^{* *} 1.75 \mathrm{~d} 0-\&\) (3.42097d0 / 2.4d0) * O**2.4d0 + 2.24266d0 * O -0.013377428 d 0
else if ( \(\mathrm{O}<0.542497505 \mathrm{~d} 0\) ) then
\(\mathrm{OTR}=(1.50592 \mathrm{~d} 0 / 1.25 \mathrm{~d} 0) * \mathrm{O}^{* *} 1.25 \mathrm{~d} 0+(0.47570 \mathrm{~d} 0 / 1.6 \mathrm{~d} 0) * \mathrm{O}^{* *} 1.6 \mathrm{~d} 0-\&\) (2.3111d0 / 2.4d0) * dexp(-2.4d0 * O) - 1.92373d0 * O + 0.96224334d0
else if ( \(\mathrm{O}<1.175411260 \mathrm{~d} 0\) ) then
OTR \(=-(0.98661 \mathrm{~d} 0 / 0.85 \mathrm{~d} 0) * \operatorname{dexp}(-0.85 \mathrm{~d} 0 * \mathrm{O})+(1.04020 \mathrm{~d} 0 / 1.55 \mathrm{~d} 0) * \&\) \(\operatorname{dexp}(-1.55 \mathrm{~d} 0\) * O) - (1.18513d0 / 3.1d0) \(* \operatorname{dexp}(-3.1 \mathrm{~d} 0 * \mathrm{O})-\&\) 0.06690 d 0 * \(\mathrm{O}+0.879238152 \mathrm{~d} 0\)
else if ( \(\mathrm{O}<1.506937513 \mathrm{~d} 0\) ) then
OTR \(=-(0.13975 \mathrm{~d} 0 / 0.55 \mathrm{~d} 0) * \operatorname{dexp}(-0.55 * \mathrm{O})-(0.45736 \mathrm{~d} 0 / 1.75 \mathrm{~d} 0) * \&\)
\(\operatorname{dexp}(-1.75 \mathrm{~d} 0 * O)+(0.03592 \mathrm{~d} 0 / 2.6 \mathrm{~d} 0) * \operatorname{dexp}(-2.6 \mathrm{~d} 0 * \mathrm{O})+\&\) \(0.02918 \mathrm{~d} 0 * \mathrm{O}+0.603323298 \mathrm{~d} 0\)
else if ( \(\mathrm{O}<1.989157517 \mathrm{~d} 0\) ) then OTR \(=(0.11412 \mathrm{~d} 0 / 0.3 \mathrm{~d} 0) * \operatorname{dexp}(-0.3 \mathrm{~d} 0 * O)-(0.61682 \mathrm{~d} 0 / 1.15 \mathrm{~d} 0) * \&\) \(\operatorname{dexp}(-1.15 \mathrm{~d} 0 * O)+(0.26454 \mathrm{~d} 0 / 2.5 \mathrm{~d} 0) * \operatorname{dexp}(-2.5 \mathrm{~d} 0 * \mathrm{O})+\&\) \(0.09195 \mathrm{~d} 0 * \mathrm{O}+0.229687651 \mathrm{~d} 0\)
else if ( \(\mathrm{O}<2.501516271 \mathrm{~d} 0\) ) then
OTR \(=-(0.28669 \mathrm{~d} 0 / 0.55 \mathrm{~d} 0) * \operatorname{dexp}(-0.55 \mathrm{~d} 0 * O)+(1.25448 \mathrm{~d} 0 / 1.85 \mathrm{~d} 0) * \&\) \(\operatorname{dexp}(-1.85 \mathrm{~d} 0 * O)-(3.74662 \mathrm{~d} 0 / 2.6 \mathrm{~d} 0) * \operatorname{dexp}(-2.6 \mathrm{~d} 0 * \mathrm{O})+\&\) \(0.00414 \mathrm{~d} 0 * \mathrm{O}+0.725889276 \mathrm{~d} 0\)
else if ( \(\mathrm{O}<3.218848666 \mathrm{~d} 0\) ) then
OTR \(=-(0.09665 \mathrm{~d} 0 / 0.2 \mathrm{~d} 0) * \operatorname{dexp}(-0.2 \mathrm{~d} 0 * \mathrm{O})-(2.54208 \mathrm{~d} 0 / 1.9 \mathrm{~d} 0) * \&\) dexp(-1.9d0 * O) + (7.90355d0 / 2.75d0) * dexp(-2.75d0 * O) - \& \(0.00240 \mathrm{~d} 0 * \mathrm{O}+0.916500569 \mathrm{~d} 0\)
else if ( \(\mathrm{O}<30.138750251 \mathrm{~d} 0\) ) then
OTR \(=-(0.02699 \mathrm{~d} 0 / 0.1 \mathrm{~d} 0) * \operatorname{dexp}(-0.1 \mathrm{~d} 0 * \mathrm{O})-(0.11714 \mathrm{~d} 0 / 0.4 \mathrm{~d} 0) * \&\) \(\operatorname{dexp}(-0.4 \mathrm{~d} 0 * O)+0.00058 \mathrm{~d} 0\) * O + 0.926940904d0
else if ( \(\mathrm{O}<399 \mathrm{~d} 0\) ) then OTR \(=-(0.96152 \mathrm{~d} 0 / 0.8 \mathrm{~d} 0){ }^{*} \mathrm{O}^{* *}(-0.8 \mathrm{~d} 0)+1.009974913 \mathrm{~d} 0\)
else
OTR = 1d0
end if
SETSEO=OTR*SECSEO
OASDI taxable wages of workers with both wages and self-employment earnings
AWSCMB=WSCCMB/CMBNT
AWSCMB96=AWSCMB(1996)
AWSCMB=AWSCMB*BASECW/AWSCMB96
CW=TAXMAX/AWSCMB
if ( CW < 0.019370968d0 ) then
CWTR \(=\) CW - (1.34906d0 /1.71d0) * CW**1.71d0
else if ( \(\mathrm{CW}<0.067798388 \mathrm{~d} 0\) ) then
```

    CWTR = (0.37771d0 / 1.45d0) * CW**1.45d0 - (1.56008d0 / 1.6d0) * CW**1.6d0 + &
        1.00003d0 * CW - 0.00001304d0
    else if (CW < 0.184024197d0 ) then
CWTR = -(1.81225d0 / 1.8d0) * CW**1.8d0 + (1.81896d0 / 2.65d0) * CW**2.65d0 + \&
0.99117d0 * CW + 0.00007587d0
else if (CW < 0.271193553d0 ) then
CWTR = -(1.39762d0 / 1.3d0) * CW**1.3d0 + (0.37353d0 / 2.2d0)* CW**2.2d0 + \&
1.42720d0 * CW - 0.00529457d0
else if (CW < 0.445532266d0 ) then
CWTR = -(1.23848d0 / 1.2d0) * CW**1.2d0 - (0.29723d0 / 6.5d0) * CW**6.5d0 + \&
1.51516d0 * CW - 0.00103355d0
else if (CW < 0.619870979d0 ) then
CWTR = -(12.1962d0 / 1.2d0) * dexp(-1.2d0*CW) + (5.29477d0 / 1.55d0) * \&
dexp(-1.55d0 * CW) + 13.4878d0 * CW - 3.5869199d0
else if (CW < 0.987919372d0 ) then
CWTR = -(0.30024d0 / 0.35d0) * dexp(-0.35d0 * CW) - (0.59052d0 / 1.35d0) * \&
dexp(-1.35d0 * CW) - 0.11893d0 * CW + 1.308982d0
else if (CW < 1.239741958d0 ) then
CWTR = -(0.58675d0 / 0.4d0) * dexp(-0.4d0 * CW) - (0.73173d0 / 2.9d0) * \&
dexp(-2.9d0 * CW) - 0.18767d0 * CW + 1.65698379d0
else if (CW < 1.801500032d0) then
CWTR }=(3.05611d0 / 0.25d0)* dexp(-0.25d0 *CW) - (2.98387d0 / 0.55d0) * \&
dexp(-0.55d0 * CW) + (0.30583d0 / 1.3d0) * dexp(-1.3d0 * CW) + \&
0.98343d0 * CW - 6.96526531d0
else if (CW < 3.680483937d0 ) then
CWTR = (-0.09559d0 / 0.3d0) * dexp(-0.3d0 * CW) - (0.67837d0 / 1.4d0) * \&
dexp(-1.4d0 * CW) + 0.00384d0 * CW + 0.82409021d0
else if ( CW < 38.741936176d0 ) then
CWTR = -(0.03095d0 / 0.15d0) * dexp(-0.15d0 *CW) - (0.21195d0 / 0.63d0) * \&
dexp(-0.63d0 * CW) + 0.00069d0 * CW + 0.87915846d0
else if ( CW < 1d3) then
CWTR = -(0.19828d0 / 0.43d0) * CW***(-0.43d0) + 1.00096844d0
else
CWTR = 1d0
end if
WSTCMB=CWTR*WSCCMB
OASDI taxable earnings of workers with both wages and self-employment earnings
TECCMB=SECCMB+WSCCMB
ATECMB=TECCMB/CMBNT
ATECMB96=ATECMB(1996)
ATECMB=ATECMB*BASECT/ATECMB96
CT=TAXMAX/ATECMB
end if
TETCMB=CTTR*TECCMB
SETCMB=TETCMB-WSTCMB

```

\section*{Ratio OASDI taxable to covered self-employment earnings}

\begin{tabular}{ll} 
CT & \(=\)\begin{tabular}{l} 
Ratio OASDI taxable maximum to average earnings of workers with both self-employment earnings and OASDI \\
taxable wages
\end{tabular} \\
CW & \(=\)\begin{tabular}{l} 
Ratio OASDI taxable maximum to average self-employment earnings of workers with both self-employment earnings \\
and OASDI taxable wages
\end{tabular} \\
CTTR & \(=\)\begin{tabular}{l} 
Ratio of OASDI taxable to covered earnings for workers with both wages and self-employment earnings \\
CWTR
\end{tabular} \\
O & \(=\) Ratio OASDI taxable maximum to average self-employment earnings of workers with no OASDI taxable wages \\
OTR & \(=\) Ratio of OASDI taxable self-employment to covered earnings for workers with no OASDI taxable wages \\
SECCMB & \(=\) OASDI covered self-employment earnings of workers with both self-employment earnings and OASDI taxable \\
SECSEO & \(=\) OASDI covered self-employment earnings of workers with no OASDI taxable wages \\
SEO & \(=\) Number or workers with OASDI covered self-employment earnings and no OASDI taxable wages \\
SETCMB & \(=\) OASDI taxable self-employment earnings of workers with both OASDI taxable wages and self-employment earnings \\
SETSEO & \(=\) OASDI taxable self-employment earnings of workers with no OASDI taxable wages \\
TAXMAX & \(=\) OASDI taxable maximum \\
TECCMB & \(=\) OASDI covered earnings of workers with both wages and self-employed earnings \\
TETCMB & \(=\) OASDI taxable earnings of workers with both wages and self-employed earnings \\
WSCCMB & \(=\) OASDI covered wages of workers with both wages and self-employed earnings \\
WSTCMB & \(=\) OASDI taxable wages of workers with both wages and self-employed earnings
\end{tabular}

2-4.7 OASDI taxable self-employment earnings (SET)
SET \(=\) SETR * CSE
Where
CSE = OASDI covered self-employment earnings
SET \(=\) OASDI taxable self-employment earnings
SETR \(=\) Ratio of OASDI taxable to covered self-employment earnings

\section*{2-4.8 OASDI effective taxable payroll (ETP)}

ETP=WTER + SET- \(0.5^{*}\) MER
Where
\begin{tabular}{ll} 
ETP & \(=\) OASDI effective taxable payroll \\
MER & \(=\) OASDI multi-employer refund wages \\
SET & \(=\) OASDI taxable self-employment earnings \\
WTER & \(=\) Annual OASDI taxable employer wages
\end{tabular}

\section*{2-4.9 OASDI taxable wage liability (WTL)}
```

WTL = WTER * TRW

```

Where
TRW = OASDI combined employee-employer tax rate
WTL = Annual OASDI taxable wage liabilities
WTER \(=\) Annual OASDI taxable employer wages

\section*{2-4.10 OASDI taxable self-employment liability (SEL)}
```

SEL = SET * TRSE

```

Where
\begin{tabular}{ll} 
SEL & \(=\) OASDI taxable self-employment earnings liabilities \\
SET & \(=\) OASDI taxable self-employment earnings \\
TRSE & \(=\) OASDI self-employment tax rate
\end{tabular}

\section*{2-4.11 OASDI quarterly taxable wage liability (WTLQ)}

\section*{Federal Civilian}

\section*{Annual total wages (OASDI + MQGE)}

BAFCW \(=34198.84\)
AWCFC \(=\) WCFC \(/\) ECFC * BAFCW / AWCFCTOT97
T=MAX/AWCFC
IF(T.LT.0.014620379)THEN
FCTR=T-(1.04262/1.73)*T**1.73
ELSE IF(T.LT.0.292407578)THEN
```

    FCTR=-(1.22471/1.6)*T**1.6+(.826746/1.8)*DEXP(-1.8*T)+1.8535*T-.459368449
    ELSE IF(T.LT.0.760259704)THEN
FCTR=-(.635082/2D0)*T**2+(.604884/2.9)*T**2.9-(.403213/4.6)*T**4.6+.910343*T+.002291358
ELSE IF(T.LT.1.228111829)THEN
FCTR=-(.162181/1.7)*T**1.7+(.143632/2.7)*T**2.7-(.312012/3.4)*T**3.4+.841165*T+.011332647
ELSE IF(T.LT.1.520519407)THEN
FCTR=-(1.34084/3.5)*T**3.5+(1.09868/5D0)*T**5-(.404253/5.8)*T**5.8+1.17397*T-.222555715
ELSE IF(T.LT.2.339260627)THEN
FCTR=(.671304/.5)*DEXP(-.5*T)-(3.27076/1.4)*DEXP(-1.4*T)+.126626*T+.353367869
ELSE IF(T.LT.3.50889094)THEN
FCTR=(.0571643/.95)*DEXP(-.95*T)-(3.17633/1.8)*DEXP(-1.8*T)+.000623031*T+.996284293
ELSE IF(T.LT.4.970928832)THEN
FCTR=-(12.3148/2.25)*DEXP(-2.25*T)+.0000698013*T+.999222265
ELSE
FCTR=-(.0285502/2D0)*T**(-2D0)+1.00007094
END IF

```
WTFCTOT \(=\) FCTR \(* W C F C\)
Where
\begin{tabular}{ll} 
AWCFC & \(=\) Average covered Federal Civilian wages (OASDI plus MQGE) \\
AWCFCTOT97 & \(=\) Average covered Federal Civilian wages (OASDI plus MQGE) for 1997 \\
BAFCW & \(=\) Average Federal Civilian wages (OASDI plus MQGE) in 1\% sample data for 1997 used to produce equations \\
ECFC & \(=\) Covered Federal Civilian employment (OASDI plus MQGE) \\
FCTR & \(=\) Ratio of taxable to covered Federal Civilian wages (OASDI plus MQGE) \\
MAX & \(=\) OASDI taxable maximum \\
T & \(=\) Ratio of the OASDI taxable maximum to average covered Federal Civilian wages (OASDI plus MQGE) \\
WCFC & \(=\) Taxable Federal Civilian wages (OASDI plus MQGE) \\
WTFCTOT &
\end{tabular}

\section*{Annual MQGE wages}

BAFCW \(=50147.72\)
AWCFC \(=\) WCFC \(/\) ECFC \(*\) BAFCW \(/\) AWCFCHO97
T = MAX / AWCFC
IF(T.LT.0.019941085)THEN
FCTR=T-(0.0450661/1.47)*T**1.47
ELSE IF(T.LT.0.099705424)THEN
FCTR=-(.0518044/1.9)*T**1.9-(.0368056/2.3)*T**2.3+.99479*T+.0000248091
ELSE IF(T.LT.0.358939528)THEN
FCTR \(=-(.05907 / 1.25) * \mathrm{~T}^{* *} 1.25-(.0746657 / 2.9) * \mathrm{~T}^{* *} 2.9+1.02092 * \mathrm{~T}-.00032173\)
ELSE IF(T.LT.0.558350377)THEN
FCTR \(=-(2.4664 / 1.4)^{*} \mathrm{~T}^{* *} 1.4+(4.82919 / 2.3) * \mathrm{~T}^{* * 2.3-(3.97473 / 3) *} \mathrm{~T}^{* * 3+1.83998 * T-.026694932}\)
LSE IF(T.LT.0.797643395)THEN
FCTR=(.609091/2.1)*T**2.1-(1.16086/4)*T**4+.788373*T+. 043208139
ELSE IF(T.LT.1.196465093)THEN
FCTR=(2.35647/.4)*DEXP(-.4*T)-(3.87811/1.2)*DEXP(-1.2*T)-(1.1179/2.5)*DEXP(-2.5*T)+.738296*T-2.83402534
ELSE IF(T.LT.1.694992215)THEN
FCTR=-(.422884/1.3)*DEXP(-1.3*T)-(6.90241/3D0)*DEXP(-3*T)-.0229917*T+1.068147457
ELSE IF(T.LT.2.592341034)THEN
FCTR \(=(.557032 / 1.2) * \operatorname{DEXP}(-1.2 * \mathrm{~T})-(5.40739 / 2.2) * \operatorname{DEXP}(-2.2 * \mathrm{~T})+.0102014 * \mathrm{~T}+.960037325\)
ELSE
FCTR=-(32.3187/3.5)*DEXP(-3.5*T)+1.000030482
END IF
WTFCHO \(=\) FCTR*WCFC
Where
\begin{tabular}{ll} 
AWCFC & \(=\) Average covered Federal Civilian MQGE wages \\
AWCFCHO97 & \(=\) Average covered Federal Civilian MQGE wages for 1997 \\
BAFCW & \(=\) Average Federal Civilian MQGE wages in 1\% sample data for 1997 used to produce equations \\
ECFC & \(=\) Covered Federal Civilian MQGE employment \\
FCTR & \(=\) Ratio of taxable to covered Federal Civilian MQGE wages \\
MAX & \(=\) Ratio of the OASDI taxable maximum to average covered Federal Civilian MQGE wages \\
T & \(=\) Covered Federal Civilian MQGE wages \\
WCFC & Taxable Federal Civilian MQGE wages
\end{tabular}

\section*{Annual OASDI taxable wages}
WTFC \(=\) WTFCTOT
- WTFCHO
Where
WTFC \(=\) Annual OASDI taxable Federal Civilian wages
WTFCHO \(=\) Taxable Federal Civilian MQGE wages
WTFCTOT \(=\) Taxable Federal Civilian wages (OASDI plus MQGE)

Quarterly OASDI covered wages
```

CFCQD(1) = .98357 * TCFCD(I,1) + FCPD(I,1)
CFCQD(2) = .98909* TCFCD(I,2) + FCPD(I,2)
CFCQD(3) = 1.01833 * TCFCD(I,3) + FCPD(I,3)
CFCQD(4) = 1.00814 * TCFCD(I,4) + FCPD (I,4)
QWCFCOD(J) = CFCQD(J) * WTFC
Where
CFCQD = Proportion of annual OASDI covered Federal Civilian wages paid in each quarter
FCPD = Payday variable for Federal Civilian wages based on calendar
I = Calendar year
J = Quarter
TCFCD = Proportion of annual NIPA Federal Civilian wages paid in each quarter
QWCFCOD = Quarterly OASDI covered Federal Civilian wages
WTFC = Annual OASDI taxable Federal Civilian wages

```

\section*{Quarterly OASDI taxable wages}

IF(FCTR.LE.0.928)FCQD(2)=CFCQD(2)+.27522*(1.-FCTR)-. \(15127^{*}(1 .-\mathrm{FCTR})^{* *} 2+.35146 *(1 .-\mathrm{FCTR}) * * 3\)
IF(FCTR.LE.0.993)THEN
\(\operatorname{FCQD}(3)=\mathrm{CFCQD}(3)+.28047 *(1 .-\mathrm{FCTR})-4.73021 *(1 .-\mathrm{FCTR})^{* *} 2+25.3606 *(1 .-\mathrm{FCTR})^{* *} 3-58.1741 *(1 .-\mathrm{FCTR})^{* *} 4+45.1465 *(1 .-\mathrm{FCTR}) * * 5\)
\(\operatorname{FCQD}(4)=\mathrm{CFCQD}(4)-.75095 *(1 .-\mathrm{FCTR})+3.65109 *(1 .-\mathrm{FCTR}) * * 2-16.9355^{*}(1 .-\mathrm{FCTR}) * * 3+23.9578 *(1 .-\mathrm{FCTR}) * * 4\)
END IF
First quarter is always 100 percent taxable.
QWTFC(I,1)=QWCFC(I,1)
IF(FCTR.LE.0.928)THEN
Compute taxable for 2nd-4th quarter.
\(\mathrm{FCQ}=\mathrm{FCQD}(2)+\mathrm{FCQD}(3)+\mathrm{FCQD}(4)\)
WTFC2=WTFC-QWTFC(I,1)
\(\operatorname{FCQD}(2: 4)=\mathrm{FCQD}(2: 4) / \mathrm{FCQ}\)
QWTFC(I,2:4)=FCQD(2:4)*WTFC2
ELSE IF(FCTR.LE.0.993)THEN
Second quarter covered is completely taxable.
QWTFC(I,2)=QWCFC(I,2)
QWTFC(I,3)=FCQD(3)*WTFC
QWTFC( \(\mathrm{I}, 4\) ) =WTFC-QWTFC( \(\mathrm{I}, 1)-\mathrm{QWTFC}(\mathrm{I}, 2)-\mathrm{QWTFC}(\mathrm{I}, 3)\)
ELSE
Second and third quarter covered is completely taxable.
QWTFC(I,2)=QWCFC(I,2)
QWTFC(I,3)=QWCFC(I,3)
QWTFC(I,4)=WTFC-QWTFC(I,1)-QWTFC(I,2)-QWTFC(I,3)
END IF
Where
\begin{tabular}{ll} 
CFCQD & \(=\) Proportion of annual OASDI covered Federal Civilian wages paid in each quarter \\
FCQ & \(=\) Sum of proportions of annual OASDI covered Federal Civilian wages paid in each quarter for quarters two to four \\
FCQD & \(=\) Proportion of annual OASDI taxable Federal Civilian wages paid in each quarter \\
FCTR & \(=\) Ratio annual OASDI taxable to covered Federal Civilian wages \\
I & \(=\) Calendar year \\
TCFCD & \(=\) Proportion of annual NIPA Federal Civilian wages paid in each quarter \\
QWCFC & \(=\) Quarterly OASDI covered Federal Civilian wages \\
QWTFC & \(=\) Quarterly OASDI taxable Federal Civilian wages \\
WTFC & \(=\) Annual OASDI taxable Federal Civilian wages \\
WTFC2 & \(=\) Total OASDI taxable Federal Civilian wages paid in quarters two to four
\end{tabular}

Quarterly OASDI taxable wage liabilities
\(\begin{array}{ll}\text { WTLQFCEE }(\mathrm{I}, \mathrm{J}) & =\operatorname{QWTFC}(\mathrm{I}, \mathrm{J}) * \operatorname{TRWEE}(\mathrm{I}) \\ \text { WTLQFCER(I, J) } & =\operatorname{QWTFC}(\mathrm{I}, \mathrm{J}) * \operatorname{TRWER}(\mathrm{I})\end{array}\)
\begin{tabular}{ll} 
WTLQFC(I, J) & \(=\) WTLQFCEE(I, J) + WTLQFCER(I, J) \\
Where & \(=\) Calendar year \\
I & \(=\) Quarter \\
J & \(=\) OASDI employee tax rate \\
TRWEE & \(=\) OASDI employer tax rate \\
TRWER & \(=\) Quarterly OASDI taxable Federal Civilian combined employee-employer wage liabilities \\
WTLQFC & \(=\) Quarterly OASDI taxable Federal Civilian employee wage liabilities \\
WTLQFCEE & \(=\) Quarterly OASDI taxable Federal Civilian employer wage liabilities
\end{tabular}

\section*{Military wages}

\section*{Annual OASDI taxable wages}

BACMW \(=16439.95\)
ACMW = AWCML * BACMW / AWCML97
T = MAX / ACMW
IF(T.LT.0.060827432)THEN
MTR=T-(.712875/2)*T**2
ELSE IF(T.LT.0.182482295)THEN
MTR \(=(.71197 / 1.8)^{*} \mathrm{~T}^{* *} 1.8-(1.59752 / 2 \mathrm{D} 0){ }^{*} \mathrm{~T}^{* *} 2+.97587 * \mathrm{~T}+0.000542413\)
ELSE IF(T.LT.0.608274315)THEN

ELSE IF(T.LT.1.094893767)THEN
MTR=-(.700864/1.4)*T**1.4-(.40042/3.3)*T**3.3+(.197091/4.1)*T**4.1+1.33615*T-. 056637087
ELSE IF(T.LT.1.703168082)THEN
MTR=(21.3527/.3)*DEXP(-.3*T)-(21.1277/0.5)*DEXP(-.5*T)+(2.73027/1.1)*DEXP(-1.1*T)+4.34833*T-31.56802874
ELSE IF(T.LT.2.311442397)THEN
MTR=-(33.3894/1.2)*T**1.2+(14.9436/1.6)*T**1.6-(2.58041/2.1)*T**2.1+21.3365*T-.872981629
ELSE IF(T.LT.3.163026438)THEN
MTR \(=-(.076094 / .3) * \operatorname{DEXP}(-.3 * T)-(1.59668 / 1.4) * \operatorname{DEXP}(-1.4 * \mathrm{~T})-.0271355^{*} \mathrm{~T}+1.182946986\)
ELSE IF(T.LT.4.257920205)THEN
\(\operatorname{MTR}=(.482918 / 1.5) * \mathrm{~T}^{* *} 1.5-(9.21141 / .9) * \operatorname{DEXP}(-.9 * \mathrm{~T})+(25.93 / 1.5) * \operatorname{DEXP}(-1.5 * \mathrm{~T})-1.14706 * \mathrm{~T}+3.246003821\)
ELSE
MTR=-(9.00723/1.8)*DEXP(-1.8*T)+1.000285789
END IF
WTML=MTR*WCML
Where
\begin{tabular}{ll} 
ACMW & \(=\) Average OASDI covered military wages adjusted for level used to produce equations \\
AWCML & \(=\) Average OASDI covered military wages \\
AWCML97 & \(=\) Average OASDI covered military wages in 1997 \\
BACMW & \(=\) Average OASDI covered military wages in \(1 \%\) sample data for 1997 used to produce equations \\
MAX & \(=\) OASDI taxable maximum \\
MTR & \(=\) Ratio of OASDI taxable to covered military wages \\
T & \(=\) Ratio of the OASDI taxable maximum to average covered military wages \\
WCML & \(=\) Annual OASDI covered military wages \\
WTML & \(=\) Annual OASDI taxable military wages
\end{tabular}

Quarterly OASDI covered wages
\begin{tabular}{ll} 
CMLQD(1) & \(=\) \\
CMLQD(2) & \(=1.97978 * T C M L D(I, 1) * \operatorname{MLPD}(\mathrm{I}, 1)\) \\
CMLQD(3) & \(=1.002 * T C M L D(I, 2) * \operatorname{MLPD}(\mathrm{I}, 2)\) \\
CMLQD(4) & \(=.99689 * T C M L D(I, 3) * \operatorname{MLPD}(\mathrm{I}, 3)\) \\
QWCML & \(=\operatorname{CMLQD}(\mathrm{J}) * W C M L\)
\end{tabular}

Where
\begin{tabular}{ll} 
CMLQD & \(=\) Proportion of annual OASDI covered military wages paid in each quarter \\
I & \(=\) Calendar year \\
J & \(=\) Quarter \\
MLPD & \(=\) Payday variable for military wages based on calendar \\
QWCML & \(=\) Quarterly OASDI covered military wages \\
TCMLD & \(=\) Proportion of annual NIPA military wages paid in each quarter \\
WCML & \(=\) Annual OASDI covered military wages
\end{tabular}

\section*{Quarterly OASDI taxable wages}
```

T=MAX/AWCML
IF(MLTR.LT.0.776)QML(1)=CMLQD(1)+.393565-.018307*T-3.44641/T+15.6381/T**2-40.0168/T**3+62.0449/T**4-
57.525/T**5+30.2498/T**6-7.8664/T**7+.674629/T**8
IF(MLTR.LT.0.952)QML(2)=CMLQD(2)+.844748-.0401062*T-7.24247/T+32.4957/T**2-83.3328/T**3+129.374/T**4-
122.526/T**5+68.2737/T**6-20.1479/T**7+2.34289/T**8
IF(MLTR.LT.0.985)QML(3)=CMLQD(3)-2.62266+.125592*T+22.5832/T-105.727/T**2+300.027/T**3-540.915/T**4+622.304/T**5-
441.658/T**6+175.722/T**7-29.8987/T**8
IF(MLTR.LT.1.)QML(4)=CMLQD(4)+2.37295-.111565*T-21.1954/T+106.049/T**2-330.637/T**3+658.869/T**4-
835.626/T**5+648.641/T**6-279.392/T**7+50.9246/T**8
IF(MLTR.LT.0.776)THEN
QWTML(I,1:4)=QML(1:4)*WTML
ELSE IF(MLTR.LT.0.952)THEN
QWTML(I,1)=QWCML(I,1)
TOTWG1=WTML-QWTML(I,1)
Q1=QML(2)+QML(3)+QML(4)
QML(2:4)=QML(2:4)/Q1
QWTML(I,2:4)=QML(2:4)*TOTWG1
ELSE IF(MLTR.LT.0.985)THEN
QWTML(I,1)=QWCML(I,1)
QWTML(I,2)=QWCML(I,2)
TOTWG1=WTML-QWTML(I,1)-QWTML(I,2)
Q1=QML(3)+QML(4)
QML(2:4)=QML(2:4)/Q1
QWTML(I,2:4)=QML(2:4)*TOTWG1
ELSE IF(MLTR.LT.1.)THEN
QWTML(I,1)=QWCML(I,1)
QWTML(I,2)=QWCML(I,2)
QWTML(I,3)=QWCML(I,3)
QWTML(I,4)=WTML-QWTML(I,1)-QWTML(I,2)-QWTML(I,3)
END IF

```
Where
\begin{tabular}{ll} 
AWCML & \(=\) Average OASDI covered military wages \\
CMLQD & \(=\) Proportion of annual OASDI covered military wages paid in each quarter \\
MLTR & \(=\) Ratio of OASDI taxable to covered military wages \\
MAX & \(=\) OASDI taxable maximum \\
I & \(=\) Calendar year \\
Q1 & \(=\) Sum of proportions of annual OASDI taxable military wages paid in each quarter for last three or two quarters in year \\
QML & \(=\) Proportion of annual OASDI taxable military wages paid in each quarter \\
QWCML & \(=\) Quarterly OASDI covered military wages \\
QWTML & \(=\) Ratio of the OASDI taxable military wages \\
T & \(=\) Annual OASDI taxable military wages for all quarters except first or first and second \\
TOTWG1 & \(=\) Annual OASDI taxable military wages \\
WTML &
\end{tabular}

Quarterly OASDI taxable wage liabilities
\begin{tabular}{ll} 
WTLQMLEE(I, J) & \(=\quad\) QWTML(I, J) * TRWEE(I) \\
WTLQMLER(I, J) & \(=\) \\
WTLQML(I, J) & \(=\quad\) WTLQMLI, J) * TRWER(I) \\
WTE J) + WTLQMLER(I, J)
\end{tabular}

Where
\begin{tabular}{ll} 
I & \(=\) Calendar year \\
J & \(=\) Quarter \\
TRWEE & \(=\) OASDI employee tax rate \\
TRWER & \(=\) OASDI employer tax rate \\
WTLQML & \(=\) Quarterly OASDI taxable military combined employee-employer wage liabilities \\
WTLQMLEE & \(=\) Quarterly OASDI taxable military employee wage liabilities \\
WTLQMLER & \(=\) Quarterly OASDI taxable military employer wage liabilities
\end{tabular}

\section*{Federal}
\begin{tabular}{lll} 
WCF & \(=\) & WCFC + WCML \\
QWCF & \(=\) & QWCFC + QWCML \\
WTF & \(=\) & WTFC + WTML \\
QWTF & \(=\) & QWTFC + QWTML \\
WTLQFEE \((I, J)\) & \(=\) & QWTF(I,J) * TRWEE(I) \\
WTLQFER(I,J) & \(=\) & QWTF(I,J) \(*\) TRWER(I) \\
WTLQF(I,J) & \(=\) & WTLQFEE \((I, J)+W T L Q F E R(I, J) ~\)
\end{tabular}
\begin{tabular}{rl} 
Where & \\
I & \(=\) Calendar year \\
J & \(=\) Quarter \\
QWCF & \(=\) Quarterly OASDI covered Federal wages \\
QWCF C & \(=\) Quarterly OASDI covered Federal Civilian wages \\
QWCML & \(=\) Quarterly OASDI taxable Federal wages \\
QWTF & \(=\) Quarterly OASDI taxable Federal Civilian wages \\
QWTFC & \(=\) Annual OASDI covered Federal wages \\
QWTML & \(=\) Annual OASDI covered Federal Civilian wages \\
WCF & \(=\) Annual OASDI covered military wages \\
WCFC & \(=\) Qnnual OASDI taxable Federal wages Federal Civilian wages \\
WCML & \(=\) Quarterly OASDDI taxable Federal combined employee-employer wage liabilities \\
WTF & \(=\) Quarterly OASDI taxable Federal employee wage liabilities \\
WTFC & \(=\) Annual OASDI taxable military wages \\
WTLQF &
\end{tabular}

\section*{State and Local wages}

\section*{Annual OASDI taxable wages}

BACW \(=21583.61\)
AWCSL = WCSL / ESLC * BACW / AWCSLOD97
S = MAX / ASLC
IF(S.LT.0.02316573)THEN
SLTR=S-(1.1803/1.71)*S**1.71
ELSE IF(S.LT.0.463314609)THEN
SLTR \(=-(1.54738 / 1.6) * S^{* *} 1.6-(.421147 / 2.5) * S * * 2.5+(3.34881 / .5) * \operatorname{DEXP}(-.5 * \mathrm{~S})+4.39012 * \mathrm{~S}-6.697774474\)
ELSE IF(S.LT.0.833966296)THEN
SLTR \(=-(.756943 / 1.8) * S^{* *} 1.8+(.485982 / 2.3) * S * * 2.3-(.175681 / 3.2) * S^{* * 3.2+.88749 * S+.004652169}\)
ELSE IF(S.LT.1.945921357)THEN
SLTR=(3.4167/.3)*DEXP(-.3*S)-(7.26467/.9)*DEXP(-.9*S)+(4.57049/1.5)*DEXP(-1.5*S)+1.0378*S-6.245057503
ELSE IF(S.LT.3.243202261)THEN
SLTR \(=-(2.40293 / .2) * \operatorname{DEXP}(-.2 * \mathrm{~S})+(6.44952 / .4) * \operatorname{DEXP}(-.4 * \mathrm{~S})-(5.64852 / .6) * \operatorname{DEXP}(-.6 * \mathrm{~S})-.278204 * \mathrm{~S}+5.099074279\)
ELSE IF(S.LT.5.559775305)THEN
SLTR=-(.0434955/.6)*DEXP(-.6*S)-(4.00403/1.7)*DEXP(-1.7*S)+.00006219*S+. 997065459
ELSE IF(S.LT.18.53258435)THEN
SLTR \(=-(.0272758 / .5) * \operatorname{DEXP}(-.5 * S)+.0000671826 * S+.997657785\)
ELSE
SLTR=-(.00861948/.7)*S**(-.7)+1.000492941
END IF
WTSL=SLTR*WCSL

Where
\begin{tabular}{ll} 
AWCSL & \(=\) Average OASDI covered State and Local wages adjusted for average wage used to produce equations \\
AWCSLOD97 & \(=\) Average OASDI covered State and Local wages for 1997 \\
BACW & \(=\) Average OASDI covered State and Local wages in \(1 \%\) sample data for 1997 used to produce equations \\
ESLC & \(=\) OASDI covered State and Local employment \\
MAX & \(=\) RASDI taxable maximum \\
S & \(=\) Ratio of OASDI taxable to covered State and Local wages \\
SLTR & \(=\) OASDI covered State and Local wages \\
WCSL & \(=\) OASDI taxable State and Local wages \\
WTSL &
\end{tabular}

\section*{Quarterly OASDI covered wages}

CSLQD(1) \(=1.0131455^{*}\) TCSLD(I,1)+SLPD(I,1)
\(\operatorname{CSLQD}(2)=1.0431906 * T C S L D(1,2)+S L P D(1,2)\)
CSLQD(3)=.9060524*TCSLD(I,3)+SLPD(I,3)
\(\operatorname{CSLQD}(4)=1.0365866 * \operatorname{TCSLD}(\mathrm{I}, 4)+\operatorname{SLPD}(\mathrm{I}, 4)\)
QWCSL=CSLQD(1:4)*WCSL
Where
CSLQD \(=\) Proportion of annual OASDI covered State and Local wages paid in each quarter
I \(=\) Calendar year

QWCSL \(=\) Quarterly OASDI covered State and Local wages
SLPD \(=\) Payday variable for State and Local wages based on calendar
TCSLD \(\quad=\quad\) Proportion of annual NIPA State and Local wages paid in each quarter

\section*{Quarterly OASDI taxable wages}
```

QSL(1)=(CSLQD(1)-.24087*(1.-1./SLTR))
QSL(2)=(CSLQD(2)-1.0492*(1.-1./SLTR)+.51259*(1.-1./SLTR**2)-.07643*(1.-1./SLTR**3))
QSL(3)=(CSLQD(3)-5.99032*(1.-SLTR**2)+13.238*(1.-SLTR**3)-11.3291*(1.-SLTR**4)+3.52237*(1.-SLTR **5))
QSL(4)=(CSLQD(4)+8.99897*(1.-SLTR**.25)-5.48866*(1.-SLTR**.5))
TQSL=QSL(2)+QSL(3)+QSL(4)
QSL(2:4)= QSL(2:4)/TQSL
QWTSL(I,1)=QSL(1)*WTSL
QWTSL(I,2:4)=QSL(2:4)*(WTSL- QWTSL(I,1))
Where
CSLQD = Proportion of annual OASDI covered State and Local wages paid in each quarter
I = Calendar year
QSL = Proportion of annual OASDI taxable State and Local wages paid in each quarter
QWTSL = Quarterly OASDI taxable State and Local wages
SLTR = Ratio of OASDI taxable to covered State and Local wages
WTSL = OASDI taxable State and Local wages

```

\section*{Quarterly OASDI taxable wage liabilities}
```

WTLQSL(I,J) = QWTSL(I,J) * TRW(I)

```
Where
\begin{tabular}{lll} 
I & \(=\) Calendar year \\
J & \(=\) Quarter \\
TRW & \(=\) OASDI combined employee-employer tax rate \\
WTLQSL & \(=\) Quarterly OASDI taxable State and Local combined employee-employer wage liabilities
\end{tabular}

\section*{Private household quarterly OASDI taxable wages and liabilities}

QWTPHH \((\mathrm{I}, \mathrm{J})=\mathrm{WCPHH}(\mathrm{I}) *\) QDPHH \((\mathrm{J})\)
\(\mathrm{WTLQPHH}(\mathrm{I}, \mathrm{J})=\) QWTPHH \((\mathrm{I}, \mathrm{J}) * \operatorname{TRW}(\mathrm{I})\)

Where
I \(=\) Calendar year
= Quarter
QDPHH \(\quad=\quad\) Proportion of annual OASDI taxable private household wages paid in each quarter
QWTPHH \(=\) Quarterly OASDI taxable private household wages
TRW \(\quad=\quad\) OASDI combined employee-employer tax rate
WCPHH \(=\) Annual OASDI covered private household wages
WTLQPHH = Quarterly OASDI taxable private household combined employee-employer wage liabilities

\section*{Farm taxable wages}

\section*{Annual OASDI}

BAFMW \(=7467.91\)
AWCFM97 = ACFMW(1997)
F = MAX / (ACFMW * BAFMW / AWCFM97)
IF(F.LT.0.066953142)THEN
FMTR=F- \(1.30211 / 1.75)^{*} \mathrm{~F}^{* *} 1.75\)
ELSE IF(F.LT.0.401718855)THEN
FMTR=-(1.18244/1.35)*F**1.35+(.25412/1.75)*F**1.75+1.24681*F-. 001598087
ELSE IF(F.LT.0.669531425)THEN
FMTR \(=-(.508764 / .6) * \operatorname{DEXP}(-.6 * \mathrm{~F})-(.300083 / 2.8) * \operatorname{DEXP}(-2.8 * \mathrm{~F})+.0188542 * \mathrm{~F}+.966550312\)
ELSE IF(F.LT.1.87468799)THEN
FMTR=-(.638146/.6)*DEXP(-.6*F)-(.0322774/1.5)*DEXP(-1.5*F)-.033706*F+1.133974442
ELSE IF(F.LT.2.41031313)THEN
FMTR \(=-(2.64644 / 1.1) * \operatorname{DEXP}(-1.1 * \mathrm{~F})+(17.4638 / 2) * \operatorname{DEXP}(-2 * \mathrm{~F})-(26.4191 / 2.5) * \mathrm{DEXP}(-2.5 * \mathrm{~F})+.00686748 * \mathrm{~F}+.909154345\)
ELSE IF(F.LT.4.82062626)THEN
FMTR \(=-(1.06567 / 1.3)^{*} \mathrm{~F}^{* *} 1.3+(.073837 / 2.1)^{*} \mathrm{~F}^{* *} 2.1+1.31021 * \mathrm{~F}-.007628879\)
ELSE IF(F.LT.6.427501679)THEN
FMTR=-(.178355/.5)*DEXP(-.5*F)-(1.70356/1.3)*DEXP(-1.3*F)+.00115171*F+. 959096096
ELSE IF(F.LT.10.7125028)THEN
FMTR=-(.0474377/0.35)*DEXP(-.35*F)-(1.32456/1)*DEXP(-1*F)+.0016146*F+. 957903052
ELSE IF(F.LT.11.38203422)THEN


Quarterly OASDI wages and liabilities
\(\operatorname{QWTFM}(\mathrm{I}, \mathrm{J})=\operatorname{TTFMD}(\mathrm{I}, \mathrm{J}) * \operatorname{TFMW}\)
\(\mathrm{WTLQFM}(\mathrm{I}, \mathrm{J})=\operatorname{QWTFM}(\mathrm{I}, \mathrm{J}) * \operatorname{TRW}(\mathrm{I})\)

Where
\begin{tabular}{ll} 
I & \(=\) Calendar year \\
J & \(=\) Quarter \\
QWTFM & \(=\) Quarterly OASDI taxable farm wages \\
TFMW & \(=\) Annual OASDI taxable farm wages \\
TRW & \(=\) OASDI com \\
TTFMD & \(=\) Proportion of annual OASDI taxable farm wages paid in each quarter \\
WTLQFM & \(=\) Quarterly OASDI taxable farm combined employee-employer wage liabilities
\end{tabular}

Quarterly OASDI taxable employee tips
\begin{tabular}{|c|c|c|}
\hline \multicolumn{3}{|l|}{QWTTIPSEE(I,J) = QDTIP(J) * WTTIPSEE(I)} \\
\hline \multicolumn{3}{|l|}{QWTTIPSEE (I,2) = QWTTIPSEE(I,2) + WTTIPSSR(I)} \\
\hline \multicolumn{3}{|l|}{WTLQTIPSEE \((\mathrm{I}, \mathrm{J})=\) QWTTIPSEE \((\mathrm{I}, \mathrm{J}) *\) TRW(I)} \\
\hline \multicolumn{3}{|l|}{Where} \\
\hline I & = & Calendar year \\
\hline J & = & Quarter \\
\hline QDTIP & = & Proportion of annual OAS \\
\hline QWTTIPSEE & = & Quarterly OASDI taxab \\
\hline WTLQTIPSEE & = & Quarterly OASDI comb \\
\hline TRW & = & OASDI combined employ \\
\hline WTTIPSEE & & Annual OASDI taxable \\
\hline WTTIPSSR & & Annual OASDI taxable \\
\hline
\end{tabular}

Private non-farm OASDI taxable wages and liabilities

\section*{Annual}

WTPNF \(=\mathrm{WTER}-\mathrm{WTFC}-\mathrm{WTML}-\mathrm{WTSL}-\) TFMW \(-\mathrm{WTTIPSEE}-\mathrm{WTTIPSSR}\)
Where
\begin{tabular}{ll} 
TFMW & \(=\) Annual OASDI taxable farm wages \\
WTSL & \(=\) Annual OASDI taxable State and Local wages \\
WTFC & \(=\) Annual OASDI taxable Federal Civilian wages \\
WTPNF & \(=\) Annual OASDI taxable private non-farm wages excluding tips \\
WTTIPSEE & \(=\) Annual OASDI taxable tips received by employees reported by employers \\
WTTIPSSR & \(=\) Annual OASDI taxable tips received by employees self-reported on income tax returns \\
WTER & \(=\) Annual OASDI taxable employer wages
\end{tabular}

\section*{Quarterly}

BACW93 \(=21912.00\)
NACW = BACW93 / ACW93 * AWC
X = MAX / NACW
IF(X.LT.0.91274)THEN
TWTR=1D0+. \(990751 * \operatorname{DEXP}(\mathrm{X})^{* *}(-1) /(-1)-.013904602\)
\begin{tabular}{|c|c|c|}
\hline \multicolumn{3}{|l|}{} \\
\hline \multicolumn{3}{|l|}{TWTR=1D0+(-.003129*X+(1.} \\
\hline \multicolumn{3}{|l|}{ELSE IF(X.LT.4.791895)THEN} \\
\hline \multicolumn{3}{|l|}{TWTR=1D0+(.003962*X+(.770093* \({ }^{* * *(-1.85053)) /(-1.85053))-.06071106 ~}\)} \\
\hline \multicolumn{3}{|l|}{ELSE} \\
\hline \multicolumn{3}{|l|}{TWTR \(=1 \mathrm{D} 0+\left(.267708 *{ }^{* *}(-.94)\right) /(-.94)+.00066\)} \\
\hline \multicolumn{3}{|l|}{END IF} \\
\hline \multicolumn{3}{|l|}{IF(TWTR.LT.0.70)THEN} \\
\hline \multicolumn{3}{|l|}{QP(1)=-(-0.000575+0.18692*DLOG(TWTR)-0.23133*DLOG(TWTR)**2-} \\
\hline \multicolumn{3}{|l|}{\(0.10453 * \operatorname{DLOG}(\mathrm{TWTR}) * * 3+0.04306 * \operatorname{DLOG}(T W T R) * * 4+0.01906 * \operatorname{DLOG}(\mathrm{TWTR}) * * 5)-0.0325201+\mathrm{PD}(1)+\mathrm{TCPD}(\mathrm{I}, 1)\)} \\
\hline \multicolumn{3}{|l|}{} \\
\hline \multicolumn{3}{|l|}{\(0.0080956+\mathrm{PD}(2)+\mathrm{TCPD}(\mathrm{I}, 2)\)} \\
\hline \multicolumn{3}{|l|}{\(\mathrm{QP}(3)=-\left(0.12167+1.31142 * \mathrm{TWTR}^{* * 3-6.31672 * T W T R * * 4+8.03785 * T W T R * * 5-3.15412 * T W T R ~}{ }^{*} 6\right.\) ) \(+0.019325+\mathrm{PD}(3)+\mathrm{TCPD}(\mathrm{I}, 3)\)} \\
\hline \multicolumn{3}{|l|}{QP(4)=-(0.1548-0.41354*TWTR **5+0.25874*TWTR **7) +0.0197767+PD(4)+TCPD(I,4)} \\
\hline \multicolumn{3}{|l|}{ELSE IF(TWTR.LT.0.88)THEN} \\
\hline \multicolumn{3}{|l|}{\(\mathrm{QP}(1)=0.224763-0.237056 * \mathrm{TWTR}+\mathrm{PD}(1)+\mathrm{TCPD}(\mathrm{I}, 1)\)} \\
\hline \multicolumn{3}{|l|}{\(\mathrm{QP}(2)=0.190385-0.209676 *\) TWTR \(+0.00176 *(\) TWTR-0.7)/(0.88-0.7)+PD(2)+TCPD \((\mathrm{I}, 2)\)} \\
\hline \multicolumn{3}{|l|}{\(\mathrm{QP}(3)=-0.052523+0.05309 *\) TWTR \(+\mathrm{PD}(3)+\mathrm{TCPD}(\mathrm{I}, 3)\)} \\
\hline \multicolumn{3}{|l|}{\(\mathrm{QP}(4)=-0.354571+0.38249 *\) TWTR \(+\mathrm{PD}(4)+\mathrm{TCPD}(\mathrm{I}, 4)\)} \\
\hline \multicolumn{3}{|l|}{ELSE} \\
\hline \multicolumn{3}{|l|}{\(\mathrm{QP}(1)=0.968092-1.877574 *\) TWTR \(+0.904348 * \mathrm{TWTR}^{* * 2+P D}(1)+\mathrm{TCPD}(\mathrm{I}, 1)\)} \\
\hline \multicolumn{3}{|l|}{QP(2) \(=-0.468266+1.148107 * T W T R-0.690132 * T W T R * * 2+P D(2)+\) TCPD \((1,2)\)} \\
\hline \multicolumn{3}{|l|}{\(\mathrm{QP}(3)=-0.850885+1.824094 *\) TWTR-0.981557*TWTR **2+PD(3)+TCPD (I,3)} \\
\hline \multicolumn{3}{|l|}{\(\mathrm{QP}(4)=0.350767-1.093966 *\) TWTR \(+0.766972 *\) TWTR **2+PD(4)+TCPD(I,4)} \\
\hline \multicolumn{3}{|l|}{END IF} \\
\hline \multicolumn{3}{|l|}{\multirow[t]{2}{*}{\[
\begin{aligned}
& \text { IF(PTR.LT.0.86)THEN } \\
& \text { QP(J)=QP(J)+ADJTP(J) }
\end{aligned}
\]}} \\
\hline & & \\
\hline \multicolumn{3}{|l|}{ELSE} \\
\hline \multicolumn{3}{|l|}{\(\operatorname{IF}((\operatorname{ADJCP}(\mathrm{J})-\mathrm{ADJTP}(\mathrm{J})) . \mathrm{NE.0D0}) \mathrm{QP}(\mathrm{J})=\mathrm{QP}(\mathrm{J})+\operatorname{ADJTP}(\mathrm{J})+((\mathrm{PTR}-\mathrm{BPTR}) /(1 .-\mathrm{BPTR}))^{* *} 4^{*}(\mathrm{ADJCP}(\mathrm{J})-\mathrm{ADJTP}(\mathrm{J}))\)} \\
\hline \multicolumn{3}{|l|}{END IF} \\
\hline \multicolumn{3}{|l|}{QWTPNF( \(\mathrm{I}, \mathrm{J})=\mathrm{QP}(\mathrm{J}) * \mathrm{WTPNF}(\mathrm{I})+\mathrm{QWTTIPSEE}(\mathrm{I}, \mathrm{J})+\mathrm{QWTPHH}(\mathrm{I}, \mathrm{J})\)} \\
\hline \multicolumn{3}{|l|}{QWTPNF(I, 2) = QWTPNF(I, 2) + WTTIPSSR(I)} \\
\hline \multicolumn{3}{|l|}{Where} \\
\hline ACW93 & \(=\) & Ann \\
\hline AWC & & Ann \\
\hline BACW93 & \(=\) & Ann \\
\hline I & & Cale \\
\hline J & & Qua \\
\hline MAX & = & Ann \\
\hline NACW & & \begin{tabular}{l}
Ann \\
equa
\end{tabular} \\
\hline PD & \(=\) & Pay \\
\hline QP & & Prop \\
\hline QWTPNF & & Qua \\
\hline TCPD & \(=\) & Prop \\
\hline TWTR & & Rati \\
\hline X & & Rati \\
\hline
\end{tabular}

\section*{Quarterly OASDI wage liabilities}
```

WTLQPNF(I,J) = (QWTPNF(I,J) - QWTPHH(I,J)) * TRW(I)

```

Where
QWTPHH \(=\) Quarterly OASDI taxable private household wages
QWTPNF \(=\) Quarterly OASDI taxable private non-farm wages including tips
TRW \(=\) OASDI combined employee-employer tax rate
WTLQPNF \(=\) Quarterly OASDI tax liabilities from taxable private non-farm wages including tips, excluding private household
taxable wages

\section*{Total quarterly OASDI taxable wages and wage liabilities}
```

QWT(I,J) = QWTPNF(I,J) + QWTF(I,J) + QWTSL(I,J) + QWTFM(I,J)
WTLQ(I,J) = QWT(I,J) * TRW(I)

```
Where
\begin{tabular}{ll} 
I & \(=\) Calendar year \\
J & \(=\) Quarter \\
QWT & \(=\) Quarterly OASDI taxable wages
\end{tabular}
\begin{tabular}{ll} 
QWTF & \(=\) Quarterly OASDI taxable Federal wages \\
QWTFM & \(=\) Quarterly OASDI taxable farm wages \\
QWTPNF & \(=\) Quarterly OASDI taxable private non-farm wages including tips \\
QWTSL & \(=\) Quarterly OASDI taxable State and Local wages \\
WTLQ & \(=\) Quarterly OASDI taxable wage liabilities
\end{tabular}

\section*{2-4.12 OASDI quarterly taxable wage liability collections (WTLQC)}

\section*{OASDI taxable private non-farm wages by sub-quarterly periods}
```

PTR =WTP/WCP
MR =MAR(I)-.04346*(1.-PTR)+.08497*(1.-PTR)**2
JR =JUN(I)-.02627*(1.-PTR)-.26844*(1.-PTR)**2
SR =SEP(I)-.12321*(1.-PTR)-.02344*(1.-PTR)**2
DR =DEC(I)-.12468*(1.-PTR)-.20710*(1.-PTR)**2
MWTP(1)=QWTP(I,1)*MR
MWTP(2)=QWTP(I,1)-MWTP(1)
MWTP(3)=QWTP(I,2)*JR
MWTP(4)=QWTP(I,2)-MWTP(3)
MWTP(5)=QWTP(I,3)*SR
MWTP(6)=QWTP(I,3)-MWTP(5)
MWTP(7)=QWTP(I,4)*DR
MWTP(8)=QWTP(I,4)-MWTP(7)

```
Where
    DEC \(=\) Proportion of fourth quarter OASDI covered private non-farm wages (excluding tips and household) paid in December
    DR \(\quad=\quad\) Proportion of fourth quarter OASDI taxable private non-farm wages (excluding tips and household) paid in December
    \(\mathrm{I}=\) Calendar year
    JR \(\quad=\quad\) Proportion of second quarter OASDI taxable private non-farm wages (excluding tips and household) paid in June
    JUN \(=\) Proportion of second quarter OASDI covered private non-farm wages (excluding tips and household) paid in June
    MAR \(\quad=\quad\) Proportion of first quarter OASDI covered private non-farm wages (excluding tips and household) paid in March
    MR \(=\) Proportion of first quarter OASDI taxable private non-farm wages (excluding tips and household) paid in March
    MWTP \(=\) OASDI taxable private non-farm wages (excluding tips and household) paid in last month and in first two months of
        quarter
    PTR \(\quad=\quad\) Ratio of annual OASDI taxable private non-farm wages (excluding tips and household) to covered private non-farm wages
    QWTP \(=\) Quarterly OASDI taxable private non-farm wages (excluding tips and household)
    SEP \(\quad=\quad\) Proportion of third quarter OASDI covered private non-farm wages (excluding tips and household) paid in September
    SR \(\quad=\quad\) Proportion of third quarter OASDI taxable private non-farm wages (excluding tips and household) paid in September
    WCP \(=\) Annual OASDI covered private non-farm wages
    WTP \(=\) Annual OASDI taxable private non-farm wages (excluding tips and household)

OASDI taxable private non-farm wages collected on in same quarter wages are paid
```

TRAT =RATEE(I,5)
CA =.95
MWCP(1)=QWSCPNF(I,1)*MAR(I)
MWCP(2)=QWSCPNF(I,1)-MWCP(1)
MWCP(3)=QWSCPNF(I,2)*JUN(I)
MWCP(4)=QWSCPNF(I,2)-MWCP(3)
MWCP(5)=QWSCPNF(I,3)*SEP(I)
MWCP(6)=QWSCPNF(I,3)-MWCP(5)
MWCP(7)=QWSCPNF(1,4)*DEC(I)
MWCP(8)=QWSCPNF(I,4)-MWCP(7)
RCSM =.80
QRMREQ=750.
QRWREQ=11250.
RMF=70786.*WSP(I)/1001400.
CALL ITERNU(QRMREQ,MWTP(2),MWCP(2),TRAT,RMF,PWCS(1))
CALL ITERNU(QRWREQ,MWTP(1),MWCP(1),TRAT,RMF,PWCE(1))
CALL ITERNU(QRMREQ,MWTP(4),MWCP(4),TRAT,RMF,PWCS(2))
CALL ITERNU(QRWREQ,MWTP(3),MWCP(3),TRAT,RMF,PWCE(2))
CALL ITERNU(QRMREQ,MWTP(6),MWCP(6),TRAT,RMF,PWCS(3))
CALL ITERNU(QRWREQ,MWTP(5),MWCP(5),TRAT,RMF,PWCE(3))
CALL ITERNU(QRMREQ,MWTP(8),MWCP(8),TRAT,RMF,PWCS(4))
CALL ITERNU(QRWREQ,MWTP(7),MWCP(7),TRAT,RMF,PWCE(4))
DO J=1,4
QWTPC(I,J)=PWCS(J)+PWCE(J)*RCSM*CA
QWTPF(I,J)=QWSTXPHH(I,J)-QWTPC(I,J)
END DO

```


\section*{OASDI taxable private wages collected on in same quarter wages paid and in following quarter}
```

QWTPCQ(I,J)=QWTPC(I,J)+QWTPHHCQ(I,J)+QWTFM(I,J)
QWTPFQ(I,J)=QWTPF(I,J)+QWTPHHFQ(I,J)

```

\section*{OASDI taxable State and Local wages collected on in same quarter wages paid and in following quarter}
```

SLTR=WTSL/WCSL
LMPW(1)=MARSL(I)-.00329*(1.-SLTR**2)
LMPW(2)=JUNSL(I)-.68187*(1.-SLTR**3)+.52206*(1-SLTR**4)
LMPW(3)=SEPSL(I)-1.33596*(1.-SLTR)+1.51187*(1.-SLTR**2)-.63523*(1.-SLTR**3)
LMPW(4)=DECSL(I)-2.03892*(1.-SLTR)+1.90430*(1.-SLTR**2)-.6633*(1.-SLTR**3)
DO J=1,4
SLCR(J)=(1.-LMPW(J))+LMPW(J)*LMCRPR(I-16,J)
QWTSLC(I,J)=SLCR(J)*QWTSL(I,J)
QWTSLF(I,J)=QWTSL(I,J)-QWTSLC(I,J)
END DO

```
Where
    DECSL \(\quad=\quad\) Proportion of OASDI taxable State and Local wages paid in fourth quarter which are paid in December
    I \(=\) Calendar year
    J = Quarter
    JUNSL \(=\) Proportion of OASDI taxable State and Local wages paid in second quarter which are paid in June
    LMCRPR \(=\) Proportion of OASDI taxable State and Local wages paid in final month of quarter on which employers are to deposit
        taxes in the same quarter
    LMPW \(\quad=\quad\) Proportion of quarterly OASDI taxable State and Local wages paid in final month of quarter
    MARSL \(\quad=\quad\) Proportion of OASDI taxable State and Local wages paid in first quarter which are paid in March
    QWTSL \(=\) Quarterly OASDI taxable State and Local wages paid in quarter
    QWTSLC \(=\) Quarterly OASDI taxable State and Local wages paid in quarter on which taxes are deposited by the employer in the
        same quarter
    QWTSLF \(=\) Quarterly OASDI taxable State and Local wages paid in quarter on which taxes are deposited by the employer in the
        following quarter
    SEPSL \(\quad=\quad\) Proportion of OASDI taxable State and Local wages paid in third quarter which are paid in September
    SLCR \(\quad=\quad\) Proportion of OASDI taxable State and Local wages paid in quarter on which taxes are deposited by the employer in
        the same quarter
    SLTR \(=\) Ratio of OASDI taxable to covered State and Local wages
    WCSL \(=\) Annual OASDI covered State and Local wages
    WTSL \(=\) Annual OASDI taxable State and Local wages

OASDI taxable wages collected on in same quarter wages paid and in following quarter
WTQCQ(I,J)= QWTPCQ(I,J)+ QWTSLC(I,J)+QWTF(I,J)
\begin{tabular}{ll} 
WTQFQ(I,J)= QWTPFQ(I,J)+ QWTSLF(I,J) \\
Where & \(=\) Calendar year \\
I & \(=\) Quarter \\
J & \(=\) Quarterly OASDI taxable Federal wages \\
QWTF & \(=\) Quarterly OASDI taxable private wages collected on in same quarter wages paid \\
QWTPCQ & \(=\) Quarterly OASDI taxable private wages collected on quarter following that in which wages paid \\
QWTSLCQ & \(=\) Quarterly OASDI taxable State and Local wages collected on in same quarter wages paid \\
QWTSLFQ & \(=\) Quarterly OASDI taxable State and Local wages collected on in quarter following that in which wages paid \\
WTQCQ & \(=\) Quarterly OASDI taxable wages collected on in same quarter wages paid \\
WTQFQ & \(=\) Quarterly OASDI taxable wages collected on in quarter following that in which wages paid
\end{tabular}

\section*{Quarterly OASDI wage tax collections}
```

WTLQC(I,1) = TRW(I-1) * WTQFQ(I-1,4) + TRW(I) * WTQCQ(I,J)
DO J = 2, 4
WTLQC(I,J) = TRW(I) * (WTQFQ(I,J-1) + WTQCQ(I,J))
END DO
Where
I = Calendar year
J = Quarter
TRW = OASDI combined employee-employer tax rate
WTLQC = Quarterly OASDI wage tax collections
WTQCQ = Quarterly OASDI taxable wages collected on in same quarter wages paid
WTQFQ = Quarterly OASDI taxable wages collected on in quarter following that in which wages paid

```

2-4.13 Quarterly Self-Employed Net Income Tax Collections (SELQC)
```

DO J = 1, 4
SELQC(I,J) = SECRCY(I,J) * SEL(I) + SECRPY(I,J) * SEL(I-1)
END DO

```
Where
    \(\mathrm{I}=\) Calendar year
    J \(=\) Quarter
    SECRCY \(=\) Proportion of OASDI taxable self-employment earnings collected on in same year earned
    SECRPY \(=\) Proportion of OASDI taxable self-employment earnings collected on in year following that in which earned
    SEL \(=\) OASDI taxable self-employment earnings liabilities
    SELQC = Quarterly OASDI self-employed net income tax collections

2-4.14 SECA Appropriation Adjustments (MSECAAA)
MSECAAA \(=0\)
DO L = I-2, I-9, -1
MSECAAA \(=\) MSECAAA + SEAACO(I,J) \(*\) SELIAC(L)
END DO
Where
\(\mathrm{I}=\) Calendar year
J \(=\) Quarter
L = Liability year
MSECAAA \(=\) OASDI SECA appropriation adjustment for quarter (assigned to last month in quarter)
SEAACO \(=\) Proportion of past year's OASDI self-employment tax liability which will be reported in current quarter
SELIAC \(=\) Prior years' OASDI self-employment tax liability

\section*{Appendix 2-2}

\section*{Economic Acronyms}
\begin{tabular}{ll} 
AA & Appropriation adjustments \\
ACE & Average OASDI covered earnings \\
ACSE & Average OASDI covered self-employed income \\
ACW & Average OASDI covered wage \\
ACWC & Average economy-wide wage \\
ADJ_FSA_FC & Adjustment to lower federal civilian covered wages relative to NIPA \\
& wages due to a presumed increase in the relative amount placed into an \\
& FSA \\
AWEFC_N & Average wage for Federal civilian employees not covered under OASDI \\
AWI & Average wage index calculated by SSA; based on the average wage of all \\
& workers with wages from Forms W-2 \\
AWSE & Economy-wide average wage \\
AWSGEFC & Average wage for the Federal government enterprises \\
AWSGFC & Average wage for the Federal civilian government \\
AWSGFM & Average wage for the military \\
AWSGGEFC & Average wage for the Federal government \& government enterprises \\
AWSGGESL & Average wage for State and local government and government enterprises \\
AWSP & Average wages, private sector \\
AWSPH & Average wage in private household sector \\
AWSPL & Average wages, private sector, 2-year moving average \\
AWSSP & Average compensation, private sector \\
AWSSPBNFXGE & Average compensation, private nonfarm business, excluding government \\
& enterprises \\
AWSSPES & Average compensation, private sector, educational services \\
AWSSPF & Average compensation, private farm, wage workers \\
AWSSPHS & Average compensation, private sector, health services \\
AWSSPL & Lagged average compensation for private sector workers \\
AWSSPSS & Average compensation, private sector, social services \\
AWSUI & Average wage of workers under UI \\
AWS_MEF & Average wage for employees with any wages (covered and noncovered) \\
& posted to the MEF \\
AYF & Average proprietor income, private farm \\
AYF_K & Ratio of average self-employment income to average wage-worker \\
& compensation for the agriculture sector \\
AYNF & Average proprietor income, private nonfarm business \\
AYNF_K & Ratio of average self-employment income to average wage-worker \\
BEA & compensation for the nonagriculture sector \\
BLS & The Bureau of Economic Analysis \\
CFCGEFC & The Bureau of Labor Statistics \\
CFCGESL & Compensation of fixed capital, Federal government enterprises \\
CFCGFC & Qovernment consumption of fixed capital, Government enterprises, State \\
& \& local \\
Compensation of fixed capital, Federal civilian
\end{tabular}
\begin{tabular}{ll} 
CFCGFM & \begin{tabular}{l} 
Federal Government Consumption Expenditures, Defense Consumption \\
Expenditures \\
State \& Local Government consumption expenditures, Gross output of \\
general government, Value added, consumption of general government \\
fixed capital
\end{tabular} \\
CFCGSL & \begin{tabular}{l} 
Workers that have a combination of both OASDI covered wages and self- \\
employed income. \\
Ratio of Federal military OASDI covered wages to NIPA wages \\
Economic Sub-Process: Covered Employment and Earnings
\end{tabular} \\
CMB_TOT & \begin{tabular}{l} 
Ratio of Private OASDI Covered to NIPA wages; OASDI private coverage \\
ratio \\
CML
\end{tabular} \\
& \begin{tabular}{l} 
The Consumer Price Index for Urban Wage Earners and Clerical Workers \\
(CPI-W) is an official measure of inflation in consumer prices, published
\end{tabular} \\
CP & \begin{tabular}{l} 
by the BLS.
\end{tabular} \\
Current Population Survey, conducted monthly by the Bureau of Census \\
for the Bureau of Labor Statistics. It is the source of historical monthly \\
economic data (such as labor force, civilian noninstitutional population, \\
and unemployment) used to project US employment.
\end{tabular}

E Total employment, CPS concept (i.e., average of monthly estimates of total wage and salary workers, plus self employed, plus unpaid family workers)
E_FE
EA
EAS
EAW
EDMIL

EGFC
EGEFCPS
Civilian employment level at full employment (i.e., at potential GDP)
Total agricultural employment
Civilian Employment Level, Self employed workers: agriculture, SA
Employment by class of worker, agricultural wage workers
Total number serving in the US Armed Forces estimated by the
Department of Defense and published by the Census Bureau
Federal civilian government employment
Employment, Establishment Data, All Employees: Government, Federal Government Enterprises, U.S. Postal Service
\begin{tabular}{|c|c|}
\hline EGGEFC & Employment, Establishment Data, All Employees: Government, Federal Government, SA \\
\hline EGGESL & Employment, State \& Local government enterprises \\
\hline EMPTRDI & DI employer tax rate \\
\hline EMPTRHI & HI employer tax rate \\
\hline EMPTROASI & OASI employer tax rate \\
\hline ENA & Civilian Employment Level, Nonagricultural industries, 16 years and over, SA \\
\hline ENAS & Employment by class of worker, nonagricultural self-employed \\
\hline ENAU & Employment by class of worker, nonagricultural unpaid family workers \\
\hline ENAW & Employment by class of worker, nonagricultural wage workers \\
\hline ENAWPBXGE & Employment for private nonfarm business \\
\hline ENAWPH & Employment by class of worker, nonagricultural wage workers, private household workers \\
\hline ENAWSPBXGE & Employment for private nonfarm business and nonagricultural selfemployed \\
\hline EO & Total employment in the other immigrant population \\
\hline EO_A & Total employment in the other immigrant population who are temporarily authorized to reside or work in the US \\
\hline EO_ESF & Total employment in the other immigrant population whose reported earnings are posted to the Earnings Suspense File \\
\hline EO_MEF & Total employment in the other immigrant population whose earnings are reported and posted to the Master Earnings File \\
\hline EO_MEFC & Total employment in the other immigrant population whose earnings are reported and posted to the Master Earnings File and are OASDI-covered \\
\hline EO_NA & Total employment in the other immigrant population who have overstayed their authorization \\
\hline EO_NO & Total employment in the other immigrant population who were never authorized to reside or work in the US \\
\hline EO_UND & Total employment in the other immigrant population that is strictly in the underground economy (i.e., with no earnings reported) \\
\hline EP & Employees in Private industries \\
\hline EPES_EST & Employees by industry, Private industries, Educational services \\
\hline EPHS_EST & Employment for private health services \\
\hline EPSS_EST & Employees by industry, Private industries, Social Assistance \\
\hline ES & Self-employed workers \\
\hline ETP & Effective annual taxable payroll, equal to total employer taxable OASDI wages plus total self-employed taxable income minus one half of the multi-employer refund wages \\
\hline EU & Unpaid family workers \\
\hline EW & Wage and salaried workers \\
\hline FERS & Federal Employee Retirement System \\
\hline GDP & Gross domestic product \\
\hline
\end{tabular}
\begin{tabular}{|c|c|}
\hline GDP05 & GDP, 2005\$ \\
\hline GDPG & GDP, General Government \\
\hline GDPGE & GDP, Federal and State \& local government enterprises \\
\hline GDPGEFC & GDP, Federal civilian government enterprises \\
\hline GDPGESL & GDP, State \& local government enterprises \\
\hline GDPGF & GDP, General Government, Federal \\
\hline GDPGFC & GDP, Federal civilian \\
\hline GDPGFM & GDP, military \\
\hline GDPGGE & GDP, Federal and State \& local government enterprises \\
\hline GDPGGEFC & GDP, Federal civilian government and government enterprises \\
\hline GDPGGESL & GDP, State \& local government and government enterprises \\
\hline GDPGSL & GDP, General Government, State \& Local \\
\hline GDPPBNFXGE & GDP, private nonfarm business, excluding government enterprises \\
\hline GDPPF & GDP, private business sector, farm \\
\hline GDPPH & GDP, Private Households \\
\hline GDPPNI & GDP, Nonprofit institutions serving households \\
\hline HI & Hospital insurance \\
\hline HIFC_L & HI Employer Liability, Federal Civilian \\
\hline HIFM_L & HI Employer Liability, Federal Military \\
\hline HIP_L & HI Employer Liability, Private \\
\hline HISL_L & HI Employer Liability, State \& Local \\
\hline KGDP05 & Potential real GDP, 2005\$ \\
\hline LC & US labor force, equal to the sum of number of persons employed and number of persons seeking employment \\
\hline LFPR & Labor force participation rate, defined as the ratio of the number of persons in the US labor force to the number of persons in the US noninstitutional population. \\
\hline M & Military population \\
\hline MER & Multi-employer refund wages \\
\hline MRAZ & Military pay raise \\
\hline N & Civilian noninstitutional population \\
\hline NCE & Total noncovered employment \\
\hline NIPA & The National Income and Product Accounts, published by the BEA, providing historical estimates of quarterly earnings and output measures \\
\hline NRA & Normal retirement age \\
\hline OASDI & Old-Age, Survivors, and Disability Insurance \\
\hline OASDIFC_L & OASDI Employer Liability, Federal Civilian \\
\hline OASDIFM_L & OASDI Employer Liability, Federal Military \\
\hline OASDIP_L & OASDI Employer Liability, Private \\
\hline OASDISL_L & OASDI Employer Liability, State \& Local \\
\hline OASDHI & Old-Age, Survivors, Disability, and Health Insurance \\
\hline OLI & Employer contributions for employee pension and insurance funds \\
\hline OLI_CSRS1 & Contributions for CSRS employees’ pay \\
\hline OLI_FC & Other labor income, Federal civilian \\
\hline
\end{tabular}
\begin{tabular}{|c|c|}
\hline OLI_FERS1 & \\
\hline OLI_FERSFC & Employer contributions to Thrift Savings Plan for FERS employees \\
\hline OLI_GGE & Other labor income, government and government enterprises \\
\hline OLI_GHI & Other labor income by type, Employer contributions to pension and welfare funds, private welfare funds, Group health insurance \\
\hline OLI_GHI_FC & Employer contributions for employee pension \& insurance funds, group health insurance, Federal civilian government sector \\
\hline OLI_GHI_P & Employer contributions for employee pension \& insurance funds, group health insurance, private sector \\
\hline OLI_GHI_SL & Employer contributions for employee pension \& insurance funds, group health insurance, State \& local government sector \\
\hline OLI_GLI & Employer contributions for employee pension and insurance funds, Group life insurance \\
\hline OLI_GLI_FC & Employer contributions for employee pension \& insurance funds, group life insurance, Federal civilian government sector \\
\hline OLI_GLI_P & Employer contributions for employee pension \& insurance funds, group life insurance, private sector \\
\hline OLI_GLI_SL & Employer contributions for employee pension \& insurance funds, group life insurance, State \& local government sector \\
\hline OLI_P & Employer contributions for employee pension and insurance funds, private industries \\
\hline OLI_PPPS & Other Labor Income, Private Sector Pension and Profit Sharing \\
\hline OLI_PPS & Employer contributions for employee pension and insurance funds, Pension \& profit-sharing \\
\hline OLI_RETFC & Employer contributions for employee pension and insurance funds, Publicly administered government employee retirement plans, Federal civilian \\
\hline OLI_RETFM & Employer contributions for employee pension and insurance funds, Publicly administered government employee retirement plans, Federal military \\
\hline OLI_RETSL & Employer contributions for employee pension and insurance funds, Publicly administered government employee retirement plans, State and local \\
\hline OLI_SL & Other labor income, State and local \\
\hline OLI_SU & Employer contributions for employee pension and insurance funds, Supplemental unemployment \\
\hline OLI_WC & Employer contributions for employee pension and insurance funds, Workers' compensation \\
\hline OLI_WCP & Private employer contribution to other labor income, total for workers' compensation \\
\hline OLI_WCSL & Employer contributions to workers' compensation, State and local \\
\hline OLIF_RETFCO & Other government contributions to Federal civilian retirement \\
\hline OOH & Owner-occupied housing \\
\hline OP & Other immigrant population \\
\hline OP_A & Other immigrant population who are temporarily authorized to reside or work in the US \\
\hline
\end{tabular}
\begin{tabular}{|c|c|}
\hline OP_NA & Other immigrant population who have overstayed their authorization \\
\hline OP_NO & Other immigrant population who were never authorized to reside or work in the US \\
\hline PGDP & Gross Domestic Product Price Index, Units: 2005=100 \\
\hline PGDPAF & Deflator for farm output \\
\hline PIA & Primary insurance amount \\
\hline PIARR & PIA replacement rate, defined as the ratio of a hypothetical medium scale worker's PIA to his/her career average indexed earnings. \\
\hline PGDP & GDP price deflator \\
\hline PBNFXGE & Private nonfarm business excluding government enterprises \\
\hline RCMB & Proportion of wage workers who are also self-employed (CMB_TOT/WSW) \\
\hline RCSE & Covered self-employed ratio, defined as the ratio of total covered selfemployment income to total proprietor income (CSE_TOT/Y). \\
\hline RCWSF & Ratio of compensation to wages in the Federal government \\
\hline RCWSM & Ratio of compensation to wages in the military \\
\hline RCWSP & Ratio of compensation to wages in the private sector \\
\hline RCWSSL & Ratio of compensation to wages in the State and local sector \\
\hline RD & Disability prevalence ratio, defined as the ratio of disabled worker beneficiaries to the disability-insured population. \\
\hline RELMAX & Ratio of the TAXMAX to averaged covered earnings \\
\hline RELMAX_UI & Ratio of the aggregate weighted average of the UI taxable maximum to the average UI wage \\
\hline RET & Earnings test ratio, defined the ratio of the maximum amount of earnings before an OASDI benefit is reduced to the average wage index. \\
\hline RFS & Family size ratio, defined as the ratio of the number of children under 6 to mothers of a certain age. \\
\hline RGR_GHI & \\
\hline RHIP_L & Product of HI tax rate, private coverage ratio, and the taxable ratio \\
\hline RM & Military ratio, the ratio of the US armed forces to the noninstitutionalized population. \\
\hline RMER & Multi-employer refund wage ratio, defined as the ratio of multiemployer refund wages to total OASDI wages. \\
\hline ROASDIP_L & Product of OASDI tax rate, private coverage ratio, and the taxable ratio \\
\hline ROLI_PPPS & Ratio of employer contributions to private pension and profit-sharing to private wages \\
\hline ROLI_SU & Ratio of private employer contributions for employee pension and insurance funds, Supplemental unemployment to private wages \\
\hline ROLI_WCP & Ratio of private employer contribution to other labor income, total for workers' compensation to private wages \\
\hline RSET & Self-employed net income taxable ratio, defined as the ratio of total selfemployed taxable income to total OASDI wages. \\
\hline RSOC_UIP & Ratio of private employer contributions to social insurance, total for unemployment insurance, to private wages \\
\hline RSOC_WCP & Ratio of private employer contributions to social insurance, total for \\
\hline
\end{tabular}
\begin{tabular}{ll} 
& \begin{tabular}{l} 
workers' compensation to private wages \\
Ratio of private employer insurance contribution to the Pension Benefit \\
Guaranty Trust Corporation to private wages
\end{tabular} \\
RSOCF_PBG & \begin{tabular}{l} 
Ratio of combined Private and State \& local sector employer \\
contributions to social insurance for workers' compensation to the \\
combined Private and State and local sector employer contributions to
\end{tabular} \\
RSOCSL_WC \\
workers' compensation \\
Ratio of total employment to the sum of wage \& salary, self-employed \\
workers, and the military (TE/(EW + ES + military))
\end{tabular}
\begin{tabular}{|c|c|}
\hline SOC_UISL & State and local government employer contributions to social insurance, total for unemployment insurance \\
\hline SOC_WCP & Private employer contributions to social insurance, total for workers' compensation \\
\hline SOC_WCSL & State and Local government employer contributions to social insurance, total for workers'compensation \\
\hline SOCF_HI & Contributions for Government Social Insurance, Employer Contributions, Federal Social Insurance Funds, Hospital Insurance \\
\hline SOCF_MIFM & Contributions for Government Social Insurance, Employer Contributions, Federal Social Insurance Funds, Military Medical Insurance \\
\hline SOCF_OASDI & Contributions for Government Social Insurance, Employer Contributions, Federal Social Insurance Funds, Old-age, Survivors, And Disability Insurance \\
\hline SOCF_PBG & Contributions for Government Social Insurance, Employer Contributions, Federal Social Insurance Funds, Pension Benefit Guaranty \\
\hline SOCF_RETRR & Contributions for Government Social Insurance, Employer Contributions, Federal Social Insurance Funds, Railroad Retirement \\
\hline SOCF_UIF & Contributions for Government Social Insurance, Employer Contributions, Federal Social Insurance Funds, Federal Unemployment Tax \\
\hline SOCF_UIFC & Total federal civilian government employer contributions to unemployment insurance \\
\hline SOCF_UIFED & Contributions for Government Social Insurance, Employer Contributions, Federal Social Insurance Funds, Federal Employees’ Unemployment Insurance \\
\hline SOCF_UIFM & Total federal government employer contributions to unemployment insurance, military \\
\hline SOCF_UIS & Contributions for Government Social Insurance, Employer Contributions, Federal Social Insurance Funds, State Unemployment Insurance \\
\hline SOCF_WC & Contributions for Government Social Insurance, Employer Contributions, Federal Social Insurance Funds, Worker’s Compensation \\
\hline SOCSL_WC & Contributions for Government Social Insurance, Employer Contributions, State and Local Social Insurance Funds, Workers’ Compensation \\
\hline SSA & Social Security Administration \\
\hline TAXMAX & OASDI contribution and benefit base \\
\hline TAXPAY & Economic Sub-Process: Taxable Payroll \\
\hline TCE & Total OASDI covered employment \\
\hline TE & Total "at any time" employment \\
\hline TEFC_N & Total "at any time" employment, Federal civilian, without Federal civilian OASDI \\
\hline TEO & Total "at any time" employment in the other immigrant population \\
\hline
\end{tabular}
\begin{tabular}{|c|c|}
\hline TEO_ESF & Total "at any time" employment in the other immigrant population whose reported earnings are posted to the Earnings Suspense File \\
\hline TEO_MEF & Total "at any time" employment in the other immigrant population whose earnings are reported and posted to the Master Earnings File \\
\hline TEO_MEFC & Total "at any time" employment in the other immigrant population whose earnings are reported and posted to the Master Earnings File and are OASDI-covered \\
\hline TEO_UND & Total "at any time" employment in the other immigrant population that is strictly in the underground economy (i.e., with no earnings reported) \\
\hline TMAXUI_SL & Taxable maximum for State \& local unemployment insurance \\
\hline TRATE_UI & \\
\hline TRATIO_UI & \\
\hline TRSE & OASDI self-employed tax rate \\
\hline TRW & Combined OASDI employee-employer tax rate \\
\hline TXRP & OASDI private taxable ratio \\
\hline U & The number of persons in the labor force who are unemployed \\
\hline USEAR & Economic Sub-Process: U.S. Earnings \\
\hline USEMP & Economic Sub-Process: U.S. Employment \\
\hline WEFC_N & Wages for Federal civilian employees not covered under OASDI \\
\hline WS & Compensation of Employees, Wage and Salary Accruals \\
\hline WSC & Total OASDI covered wages \\
\hline WSD & Total wage and salary disbursements \\
\hline WSDP & Private wage and salary disbursements \\
\hline WSGEFC & Government Wages and Salaries, Federal civilian, Government Enterprises \\
\hline WSGFC & Wage and salary accruals by industry, Government, Federal civilian \\
\hline WSGFM & Wage and salary accruals by industry, Government, Federal, Military \\
\hline WSGGEFC & Wages for the Federal government \& government enterprises \\
\hline WSGGESL & Wages for State and local government and government enterprises \\
\hline WSP & Compensation of Employees, Wage and Salary Accruals \\
\hline WSPF & Wage and salary accruals by industry, Private industries, Farms \\
\hline WSPH & Wage and salary accruals by industry, Private industries, Households \\
\hline WSPNI & Wage and salary accruals by industry, Private industries, Nonprofit institutions serving households \\
\hline WSPRRB & Wages covered by Railroad Retirement Act \\
\hline WS_MEF & Total wages posted to the MEF \\
\hline WSS & Total wage worker compensation \\
\hline WSSG & Compensation for Federal and State \& local government \\
\hline WSSGE & Compensation for Federal and State \& local government enterprises \\
\hline WSSGEFC & Compensation of employees by industry, Government, Federal \\
\hline WSSGESL & Compensation of employees by industry, Government, State and local government enterprises \\
\hline WSSGF & Federal Government Consumption Expenditures, Compensation of General Government Employees \\
\hline WSSGFC & Compensation of employees by industry, Government, Federal civilian \\
\hline
\end{tabular}
\begin{tabular}{ll} 
WSSGFM & \begin{tabular}{l} 
Compensation of employees by industry, Government, Military \\
WSSGGE
\end{tabular} \\
WSSGGESL & \begin{tabular}{l} 
Compensatione wor the Sapital Consumption Adjustment, Government local government and government \\
enterprises
\end{tabular} \\
WSSGSL & \begin{tabular}{l} 
State \& Local Government Consumption Expenditures, Compensation \\
of General Government Employees
\end{tabular} \\
WSSP & \begin{tabular}{l} 
Compensation of employees by industry, Private industries \\
Compensation in private business nonfarm excluding government \\
enterprises
\end{tabular} \\
WSSPBNFXGE
\end{tabular}

\section*{Process 3:}

Beneficiaries

\section*{3. Beneficiaries}

OCACT uses the Beneficiaries process to project the fully insured and disability insured population, the number of disabled workers and their dependent beneficiaries, the number of retired worker and their dependent beneficiaries, and the number of dependent beneficiaries of deceased workers. The Beneficiaries process receives input data from the Demography and Economics sections along with data received from the Social Security Administration and other government agencies. Output data is provided to the Economics and Trust Fund Operations and Actuarial Status processes.

The Beneficiaries Process is composed of three subprocesses: INSURED, DISABILITY, and OLD-AGE AND SURVIVORS. As a rough overview, INSURED projects the number of people in the Social Security area population that have sufficient work histories for disability and retirement benefit eligibility. DISABILITY projects the number of disabled workers and their dependent beneficiaries. OLD-AGE AND SURVIVORS projects the number of retired workers, their dependent beneficiaries, and the dependent beneficiaries of deceased workers.

All programs output data on an annual basis.

\subsection*{3.1. INSURED}

\section*{3.1.a. Overview}

Insured status is a critical requirement for a worker, who has participated in the covered economy, to receive Social Security benefits upon retirement or disability. The requirement for insured status depends on the age of a worker and his (or her) accumulation of quarters of coverage (QC).

INSURED is a simulation model that estimates the percentage of the population that is fully insured (FPRO) and disability insured (DPRO) throughout the projection period. These estimates are used in conjunction with estimates of the Social Security area population (SSAPOP) to estimate the number of people that are fully insured (FINPOP) and disability insured (DINPOP). FINPOP is then used by the OLD-AGE AND SURVIVORS INSURANCE subprocess, and both FINPOP and DINPOP are used by the DISABILITY subprocess. FINPOP and DINPOP are projected by age, sex, and cohort.

For each sex and birth cohort, INSURED simulates 30,000 work histories that represent the nonother immigration population (LEGPOP). These histories are constructed from past and projected cover worker rates of the non-other immigration population, median earnings, and amounts required for crediting QC.

The equations for this subprocess are given below:
\[
\begin{align*}
& \text { FPRO }=\text { FPRO }(\cdot)  \tag{3.1.1}\\
& \text { DPRO }=\text { DPRO }(\cdot)  \tag{3.1.2}\\
& \text { FINPOP }=\text { FPRO } * \text { SSAPOP }  \tag{3.1.3}\\
& \text { DINPOP }=\text { DPRO } * \text { SSAPOP } \tag{3.1.4}
\end{align*}
\]

\section*{3.1.b. Input Data}

All data are updated annually, except those that are noted.

\section*{Long-Range OCACT Data}

\section*{Demography}
- Social Security area population as end of year (1940 - 2095) by age (0 -100, age 100 including age 100 and older), marital status (single, married, widowed, divorced) and sex (M, F). (Workflow 3.a)
- "Other immigrant" population as end of year ( 1963 - 2095) by age ( \(0-100\), age 100 including age 100 and older) and sex (M, F). (Workflow 3.c)
- Number of new "net legal immigrants" (legal immigrants - estimated legal
emigrants) entering the Social Security area each year (1940 - 2095) by age (14-84) and sex (M, F). (Workflow 3.d)
- The population granted deferred action for childhood arrivals (DACAs) as end of year (2012-2095) by age ( \(0-100\), age 100 including age 100 and older) and sex (M, F). (Workflow 3.e)
- The population granted deferred action for parental arrivals (DAPAs) as end of year (2014-2095) by age ( \(0-100\), age 100 including age 100 and older) and sex (M, F). (Workflow 3.f)
- The population that attain deferred action for childhood arrivals status (DACAATT) as beginning of year (2012-2095) by age (-1-100, age 100 including age 100 and older) and sex (M, F). (Workflow 3.g)
- The population that attain deferred action for parental arrivals status (DAPAATT) as beginning of year (2014-2095) by age (-1-100, age 100 including age 100 and older) and sex (M, F). (Workflow 3.h)

\section*{Economics}
- Annual estimates of covered workers posted to the Master Earnings File (MEF) by sex (M, F) and age ( \(0-100\) ) for years (1937-2095). (Workflow 3.b)
- Annual projection (2012 - 2095) of average wage index and median covered earnings. (Workflow 3.j)
- "Other immigrant" workers with earnings posted to the (MEF) by sex (M, F), age (16 100), and for years (1964-2095). (Workflow 3.k)

\section*{Beneficiaries}
- Disabled-worker beneficiaries at year end (2014-2095) by age (15-66), sex (M, F) and duration ( \(0-10\), duration 10 including duration 10 and above) from the previous year's Trustees Report. These data are read in from files that are generated annually from the Beneficiaries/DISABILITY (\#3.2) area. (Workflow 3.i)

\section*{Short-Range OCACT data}
- FINPOP by age ( \(14-95\), age 95 including age 95 and older) and sex (M, F) from the end of year 1969 to the end of Short-Range projection period (2025) (EOY 19692030 is provided). (Workflow 2.f)
- DINPOP by age ( \(15-66\) ) and sex (M, F) from the end of year 1969 to the end of Short-Range projection period (2025) (EOY 1969-2030 is provided). (Workflow 2.g)

\section*{Other input data}
- Historical series of annual median earnings of covered workers by age group (<20, 20-24, 25-29, 30-34, 35-39, 40-44, 45-49, 50-54, 55-59, 60-64, 65-69, 70-74, 75-79, 80-84) and sex (M, F) for years 1937-2012. Data are updated using the CWHS file from the mainframe. (Workflow 2.b)
- Number of disabled workers by age (20-69) and sex (M, F) for years 1958-2014.

Ages 66-69 are zeros. Data are updated using the data from the historical disability file "wkrben". (Workflow 2.e)
- The amount required for crediting one quarter of coverage for years 1937-2012 from an OCACT web site. (Workflow 2.d)
- Historical series of annual median earnings of all covered workers for years 19372012. Data are updated using the data in the most recent Social Security Annual Statistical Supplement Table 4.B6. (Workflow 2.c and 2.a)
- The number of all covered workers (wage/salary workers, self-employed workers) by sex and amount of earnings for 2012 in the most recent Social Security Annual Statistical Supplement Table 4.B7 \& 4.B9. These are used to produce the input data for the distribution of earnings (FRACMOD.f90). (Workflow 2.a) (No Change for the 2016 TR)
- ANNUAL factor (comparability factor between quarterly and annual reporting of earnings) by age (13-84) and sex (M, F) for years prior to 1978. (Workflow 4.d) (No Change for the 2016 TR)
- SLCT factor (adjustment factor to bring simulated fully insured rate in line with historical fully insured rate) by age (13-84) and sex (M, F). These data were updated for the 2016 TR. (Workflow 4.b)
- SRCH factor (adjustment factor to bring simulated fully insured rate in line with historical fully insured rate) by age (13-84) and sex (M, F). These data were updated for the 2016 TR. (Workflow 4.c)
- DIADJ factor (adjustment factor to bring simulated disability insured rate in line with historical disability insured rate) by age (13-69) and sex (M, F). These data were updated for the 2016 TR. (Workflow 4.e)

\section*{3.1.c. Development of Output}

Equation 3.1.1 \& 3.1.2 -

\section*{Determining the QC distribution}

There are three variables playing important roles in the simulation process starting from age 13 through 84 of a birth cohort by sex. They include historical and projected covered worker rates of the non-other population \({ }^{1}\) by age and sex, the amounts required for crediting QC, and a cumulative worker distribution by earnings level.

Covered worker rates of the non-other population (CPRO) are the ratio of the non-other covered workers to the non-other Social Security area population. A non-other population is its total population minus its other immigrant population. Historical and projected (total and other immigrant) numbers of covered workers and the Social Security area population, which are provided by the Economics and Demography sections respectively, are used to calculate the rates for ages 13 through 84 .

\footnotetext{
\({ }^{1}\) The insured model treats the DACA and DAPA population (those granted deferred action) like legal permanent resident immigrants.
}

The law specifies the amount of earnings needed to earn one QC for each year of the historical period. Its projection assumes the same growth rate as the Social Security average wage index.

The cumulative worker distribution by earnings level is 'FRAC'. It is a function of covered earnings relative to median earnings. For a given ratio of covered earnings relative to median earnings, FRAC returns the percentage of covered workers whose earnings relative to median earnings are less than the given ratio. It is constructed based on the latest historical data. It is used for each age and sex and is assumed to remain constant throughout the projection period. The program uses FRAC to estimate the percentage of covered workers that earn \(0,1,2,3\) or 4 QC in a given year. Thus, for a particular age and sex, the percentage of covered workers earning at least \(n\) QC is defined as:
\[
\text { QCDist }=1-\operatorname{FRAC}\left[\frac{n * Q C \text { amount }}{\text { median earnings }}\right], \quad \text { for } n=1,2,3,4
\] where median earnings is for that age and sex.

\section*{Simulation process - assigning QC to records}

Once the QCDist is known, the simulation process begins with 30,000 records for each sex and birth cohort. Starting with the QC distribution at age 13, INSURED randomly assigns a number of QC (1, 2, 3 or 4 ) to these records based on QCDist.

For ages 14 to 84, INSURED begins the simulation process by randomly selecting records to represent new net legal immigrants and the other immigrant population that attain DACA or DAPA status from the covered worker portion of 30,000 records. For each record, a number of QC ( \(0,1,2,3\) or 4 ) is assigned on a uniform basis. Once a record is assigned a number of QC, INSURED nullifies the previous earnings of the record. In the initial years of eligibility for DACAs (2013 and 2017-2018) and DAPAs (2017-2018), we assume that 10 percent of DACAs and 15 percent of DAPAs will retain covered earnings made prior to attaining legal status. We assume a higher rate for DAPAs since they are older and more likely to have been working.

After the records for new immigrants are selected, the rest of the records for ages 14 to 84 are either non-covered workers or covered workers. The total number of records assigned as noncovered workers is set equal to (1-covered worker rate) * 30,000. These records receive no QC. To identify records as non-covered workers, INSURED uses two parameters (SRCH, SLCT), which vary by age and sex.

SRCH sets a limit on the number of consecutive records to be searched for a non-covered worker. In general, the younger age groups have lower SRCH values. SLCT is the number of consecutive prior years in which no QC were earned that is required in order for a simulated record to be assigned as a non-covered worker. Lower SLCT values are set for the very young age groups. Sensitivity analyses show that insured percentages are negatively correlated with these two parameters. When the female covered worker rates approach the male rates, the female SRCH and SLCT values are graded toward the male values \({ }^{2}\).

\footnotetext{
\({ }^{2}\) This occurs when the female covered worker rate is greater than 80 percent of the male rate. When the female rate
}

For each sex and birth cohort, the simulation process of assigning records as non-covered workers uses the following approach. This approach is repeated until the targeted number of non-covered workers is achieved.
1. One of the records, which is designated as one for legal immigrants, is randomly selected as the starting record.
2. Beginning with the starting record, each record is examined until a record that matches the SLCT criterion is found.
3. However, if the number of records examined equals the value of SRCH and no record matches the SLCT criterion, then the record closest to the SLCT criterion is assigned no QC as a non-covered worker.

Initially, values for SRCH and SLCT are the same as those used in the prior Trustees Report. Adjustments to these values are only made when the results are not consistent with historical data.

The final step of the simulation process is to use QCDist to randomly assign QC of \(0,1,2,3\) or 4 to the remaining covered worker records, which are not new net legal immigrants, for ages 1484.

\section*{Determining Insured Status}

Once the simulation process is complete, the insured status for each record at any age can be determined based on the total QC assigned up to that age. The simulated non-other fully insured percentage (FSIM_LEG) is calculated as the percentage of the 30,000 simulated records meeting the QC requirements for insured status. The same calculation is applied to the disability-insured percentage (DSIM_LEG).

For each sex and cohort, FSIM_LEG is determined at ages 13 to 84. DSIM_LEG is determined at ages 13 to 69.

An adjustment, DINADD, is made to DSIM_LEG. This additive adjustment accounts for workers who fail to meet the requirement for disability-insured status solely because of having no earnings while receiving disability benefits. INSURED assumes that workers who have been on the disability rolls for more than 3 years would be in this situation \({ }^{3}\). Thus, DINADD is
\# of workers on the disability rolls morethan 3 years
Social Security Area population
by age, sex, and cohort.

\footnotetext{
is equal to or greater than the male rate, the female SLCT and SRCH parameters are set equal to the male parameters. The parameters are linearly interpolated when the female covered worker rate is between 80 and 100 percent of the male rate.
}

\footnotetext{
\({ }^{3}\) Those who are on the rolls for less than 4 years are assumed to meet the requirement for disability-insured status based on their earnings histories.
}

A small proportion of the other immigrant population, OTLPOP, is added to calculate the simulated fully insured rate of the Social Security Area population, FSIM. We assume that other immigrants who have their earnings posted to the MEF are three-fourths as likely to be insured as the non-other population. We project FSIM as
\[
\frac{F S I M_{-} L E G *\left(L E G P O P+A L P H A *\left(C W \_O T H E R / C P R O\right)\right)}{S S A P R O}
\]
by age, sex, and cohort, where ALPHA is equal to 0.75 .
Hence, the simulated fully insured rate of the other immigrant population FSIM _OTL is
\[
\frac{F S I M_{-} L E G^{*} A L P H A * C W \_O T H E R / C P R O}{O T L P O P} \text { by age, sex, and cohort. }
\]

FSIM_LEG is assumed to remain the same beyond age 84. FSIM_OTL is assumed to remain the same beyond age 69. For ages 70 and older, FSIM is projected by multiplying the non-other and other fully insured rates to their respective populations and dividing the sum by the Social Security Area population.

DSIM and DSIM_OTL are projected in a similar manner by using DSIM_LEG. If the simulated results for DSIM are not consistent with historical data, an additional age-sex-specific additive adjustment (DIADJ) is used to bring the simulated results in line with the historical estimates.

Finally, incorporation of Short-Range projections produces FPRO and DPRO. For the first 10 years, FPRO and DPRO are calculated by dividing the Short-Range estimates by the Social Security area population. The difference in terms of the percentage between the Long-Range (FSIM and DSIM) and Short-Range projections at the end of \(10^{\text {th }}\) year is linearly phased out during the next ten years by cohort and sex. The Long-Range projections are assumed thereafter.

\section*{Number of Fully Insured and Disability Insured Workers}

The numbers of Fully Insured and Disability Insured workers are obtained by applying FPRO and DPRO, respectively, to the Social Security area population. The result is an estimate of the number of people that are fully insured (FINPOP) and disability insured (DINPOP) by single year of age and sex, respectively. For a given age and sex, the proportion of the Social Security area population that is insured (FPRO) is assumed to be the same for each marital status.

\section*{DISABILITY}

\section*{3.2.a. Overview}

The Social Security Administration pays monthly disability benefits to disability-insured workers who meet the definition of "disability". If they meet certain requirements, spouses and children of disabled-worker beneficiaries may also receive monthly benefits.

DISABILITY projects the number of disabled-worker beneficiaries in current-payment status (DIB) at the end of each year by age at entitlement, sex, and duration from entitlement. We base the number of DIB at the end of each year on the number of disabled-worker beneficiaries who are currently entitled to benefits (CE). We calculate the number of CE at the end of year by adding the number of newly entitled CE (New Entitlements) during the year and subtracting the number of CE who leave the disability rolls (Exits) during the year to the number of CE at the end of the prior year. Disabled-worker beneficiaries who leave the disability rolls (Exits) do so by recovering from disabilities (Recoveries), by dying (Deaths), or by converting to retired worker status (Conversions). A disabled-worker beneficiary converts to retired worker status upon reaching Normal Retirement Age (NRA), the age at which a person first becomes entitled to unreduced retirement benefit.

DISABILITY also projects the number of future dependent beneficiaries of DIB by category, age, and sex. The six categories are minor child, student child, disabled adult child, young spouse, married aged spouse and divorced aged spouse. We generate the numbers of dependent beneficiaries of DIB by multiplying the relevant subset of the SSA area population (Exposures) by a series of probabilities that relate to the regulations and requirements for obtaining benefits (Linkages).
\[
\begin{align*}
& \text { New Entitlements }(\text { year })=\text { Exposure }_{\text {BOY }} \times \text { Incidence Rate }(\text { year })  \tag{3.2.1}\\
& \text { where BOY is beginning of year. } \\
& \text { Exits }(\text { year })= \text { Recoveries }(\text { year }+ \text { Deaths }(\text { year })+\text { Conversions }(\text { year })  \tag{3.2.2}\\
& \text { where Recoveries }(\text { year })=\text { CE }_{\text {BOY }} \times \text { Recovery Rate }(\text { year }) \\
& \text { where Deaths (year) }=\text { CE }_{\text {BOY }} \times \text { Death Rate }(\text { year }) .
\end{align*}
\]

\section*{3.2.b. Input Data}

\section*{Trustees Assumptions}

Each year, the Trustees set the assumption for the ultimate incidence rates and the ultimate recovery rates for the twentieth year of the projection period. The following chart shows our "target" incidence rates by age group and sex. We further adjust these rates with a \(0.6 \%\) load for
the removal of closure of the record and a \(0.8 \%\) discount for the physician review requirement mandated by the Bipartisan Budget Act of 2015.

Target Incidence Rates by Age Group
\begin{tabular}{|c|c|c|c|c|c|c|c|c|c|c|c|}
\hline \multirow[t]{2}{*}{} & \multicolumn{11}{|c|}{Age Group} \\
\hline & \(<20\) & 20-24 & 25-29 & 30-34 & 35-39 & 40-44 & 45-49 & 50-54 & 55-59 & 60-64 & 65-69 \\
\hline Male & 2.46 & 2.58 & 1.66 & 2.49 & 3.34 & 4.52 & 5.96 & 9.89 & 15.72 & 19.02 & 9.00 \\
\hline Female & 1.90 & 2.02 & 1.62 & 2.82 & 3.54 & 4.78 & 6.39 & 10.10 & 14.16 & 14.72 & 8.34 \\
\hline
\end{tabular}

Using a standard population of disability insured who are not in current pay as of December 1999, the age-sex-adjusted incidence rate (on an entitlement rate basis) for the 2016 Trustees Report is 5.79 per 1,000 . The new physician requirement reduced this rate from the 2015 Trustees Report rate of 5.84 per 1,000 . Using a standard population of DIBs as of December 1999, the age-sex-adjusted recovery rate for the 2016 Trustees Report is 9.3 per 1,000.

\section*{Long-Range OCACT Data}

All data is updated annually except those noted otherwise. Population data are as of December 31. We assume data as of December 31 of year \(z-1\) is equal to data as of January 1 of year \(z\). Workflow reference is bolded in parenthesis after each item.

\section*{Demography}
- Social Security area population by age, sex, and marital status \({ }^{1}\) (dimensioned (0:100,1:2,1:4)) for years 1970-2095. (3a)
- Probabilities of death by sex, age and year (1:2,15:148,2014:2095). (3b)
- Total children by sex of parent, age of parent and age of child (dimensioned (1:2,19:71,0:18)) for years 1970-2095. (3c)
- Total married lives by age of husband crossed with age of wife (dimensioned (14:100,14:100)) for years 1970-2095. (3d)
- Average number of children under 18 per couple with children by age group (<25, 2529, ...,60-64) of head of household (dimensioned (1:9)) for years 1970-2095. (3e)

\section*{Economics}
- Unemployment rates by age group (16-19, 20-24, ...,60-64), sex and year (1990:2027). (3h)

\section*{Beneficiaries}

INSURED subprocess \#3.1
- Disability-insured population by age, sex and year (15:69,1:2,1969:2095) from the 2016 Trustees Report. (3g)
- Disability-insured population by age, sex and year (15:69,1:2,1969:2095) from the 2015 Trustees Report. (2j)
- Fully insured population by age, sex and marital status (14:95,1:2,1:4) for years 19702095. (3f)

2015 Trustees Report DISABILITY subprocess \#3.2

\footnotetext{
\({ }^{1}\) Single, married, widowed, divorced.
}
- Death rate projection factors from the 2015 Trustees Report by age group (15-19,20\(24, \ldots, 60-64)\), sex and year (1:2,1:10,2015:2095). (3i)
- Recovery rate projection factors and recovery rates from the 2015 Trustees Report by age group (15-19,20-24, ..,60-64), sex and year (1:2,1:10,2015:2095) and (1:2,1:10,1970:2095), respectively. (3j) and (3k)

\section*{Other input data}

We update only the most recent year data annually for this category except as noted otherwise below:
- The December 2015 data from the Master Beneficiary Record (MBR) containing the number of DIB by duration of entitlement, age of entitlement, sex and time of year (BOY or EOY) (0:55,15:66,1:2,1:2). (2b)
- December 2015 and December 2014 data from the MBR containing the number of disabled workers in current-payment status by age, sex and year (15:66,1:2,1969:2015). (2a) and (2j)
- December 2015 data from the MBR containing the number of disabled workers in withheld or suspended status by age, sex and year (15:66,1:2,1969:2015). (2a)
- December data from the MBR containing the number of dependent beneficiaries by age, sex of the account holder, and year for the following beneficiary categories. (2a)
1) Minor child (0:17,1:2,1970:2015)
2) Student child (18:21,1:2,1970:2015)
3) Disabled adult child (age group \(\left.1: 9^{2}, 1: 2,1970: 2015\right)\)
4) Young spouse (19:64,1:2,1970:2015)
5) Married aged spouse (62:100,1:2,1970:2015)
6) Divorced aged spouse (62:100,1:2,1970:2015)

We also read totals for each category.
- December data from the MBR containing the number of DIB awards by age, sex and year (15:67,1:2,1970:2015). (2c)
- December data from the MBR containing (1) the number of DIB total terminations (recoveries and deaths) and (2) the number of conversions \({ }^{3}\). These data are by sex and year (1:2,1970:2015). (2d)
- December data from the MBR containing the number of DIB deaths by age, sex and year (15:67,1:2,1975:2015). (2d)
- December data from the MBR containing the number of estimated DIB recoveries by age, sex and year (15:67,1:2,1975:2015). (2d)
- December data from the MBR containing the number of old-age beneficiaries who at some point in time were converted to retired worker status. This data is by age, sex and year: \((65: 95+, 1: 2,1970: 2015)\) (2b)
- December data from the MBR containing the number of DIB entitled to the Hospital Insurance portion of Medicare by age group ( \(<25,25-29, \ldots, 60-64,65+\) ), sex and year

\footnotetext{
\({ }^{2}\) Age groups 1 through 9 are 18-19, 20-24, 25-29,..., 55-59.
\({ }^{3}\) Conversions are DIB beneficiaries who become eligible for old-age benefits due to reaching the normal retirement age.
}
(2:11,1:2,1973:2015). This file is not used for the SOSI. (2e)
- December data from the MBR containing the number of DIB awards by duration, age, sex and year (0:5,0:65,1:2,1993:2014). 2014 Awards were retrieved for this year's TR, 1993-2013 Awards have not been updated. (2j)
- Retroactive factors \({ }^{4}\) by year (1969:2014). These values are estimated using OCACT beneficiary data. This file is not used for the SOSI. (2h)

All numbers in the following categories are updated annually unless otherwise noted.
- Average incidence rates by age and sex (15:65,1:2) for the base period 2000-2009 based on awards data from 2000-2014 (also known as the base incidence rates). We update these values when time and data are available. Note that rates for ages 60 through NRA are from the 2011 TR. (2j)
- Probability of death for DIB's - in a multiple-decrement environment by duration, age and sex \((0: 10,15: 65,1: 2)\) for the base period 2006-2010. These numbers (also known as the base probabilities of death) are from Actuarial Study No. 123. We update these values when time and data are available. (2f)
- Probability of recovery for DIB's - in a multiple-decrement environment by duration, age and sex ( \(0: 10,15: 65,1: 2\) ) for the base period 2006-2010. These numbers (also known as the base probabilities of recovery) are from Actuarial Study No. 123. We update these values when time and data are available. (2g)
- Initial Incurred but not reported (IBNR) \({ }^{5}\) factors by duration, age and sex (0:10, 15: 69, 1:2) based on 2002 entitlements (awards data from 2002-2007). (2j) (Not Updated)
- Ultimate IBNR factors by duration, age and sex ( \(0: 10,15: 69,1: 2\) ) based on 20002009 entitlements (awards data from 2000-2014). We update these values when time and data are available. The IBNR calculation linearly grades between the initial and ultimate IBNR factors during the ten years from 2008-2017. The ultimate IBNR factor table applies for all years after 2017. (2j)
- For each year 2000-2105, (1) the Normal Retirement Age (NRA), (2) the proportion of DIBs who stay on the DI roll for that age, and (3) the proportion of DIBs who convert to an old-age benefit during that year for that age. We update these values only when there is a change in the NRA or in the present law. (2k)
- The following linkages for the calculations of auxiliary beneficiaries; the probability that student is in an eligible school, the probability that adult child is disabled, the probability that beneficiary is not subject to the earnings test, and the probability that beneficiary was married 10 or more years are estimated and are updated when time and data are available. (4a)
- Short-Range/Long-Range adjustment (APROJ) factors by auxiliary beneficiary category (1:7) for years 2016-2095. These seven categories are; minor child, student child, disabled adult child, young wife, young husband, age wife and aged husband.

\footnotetext{
\({ }^{4}\) Retroactive factors for each calendar year are the ratio of the total monthly payments to DIBs to the monthly DIBs in current payment status times the average DIB monthly benefit.
\({ }^{5}\) IBNR factors reflect the proportion of DIBs entitled to benefits who have been awarded since the year of their entitlement.
}

We calculate these values by comparing Short-Range and Long- Range numbers for auxiliary beneficiaries. For the 2016 Trustees Report, these factors also include adjustments to aged-spouses and disabled adult children because of demographic assumption changes (2i)
- IPROJG, DPROJG and RPROJG adjustment factors used to adjust incidence, death and recovery rates to reconcile between the long-range model and the short-range model. IPROJG (1:11, 2, 2016-2095) and DPROJG (1:10, 2, 2016-2095) adjustment factors are by age group, sex and year. RPROJG (2, 2016-2025) adjustment factors are by sex and year. (2l),(4c), and (4h)
- Ultimate RPROJG values by sex and age group (1:2, 1:10) calculated to reach a target value. We update these values when the probabilities of recovery for DIB's are updated. (4c)

\section*{3.2.c. Development of Output}

\section*{Equation 3.2.1 - New Entitlements}

We calculate new entitlements by multiplying age-sex-specific incidence rates to the exposed population at the beginning of the year. The exposed population is the disability-insured population less the currently entitled population. We calculate future age-sex-specific incidence rates by multiplying the base incidence rates by the incidence rate projection factors (IPROJGs). For the first ten years of the projection (short-range period), IPROJGs by 5-year age group and sex are obtained by using regression equations with the change in unemployment in the two prior years as the independent variables. We describe the regression equations in detail in Appendix 3.2-1. Then, we run the IPROJGs through the main model and analyze the resulting incidence rates by age group and sex. We adjust the IPROJGs by age and sex to reach "target" incidence rates in the twentieth year of the projection period. These "target" incidence rates were determined based on a recent review of historical and projected award rates by age group and sex. For projection periods between the tenth and twentieth years, we linearly interpolate the IPROJGs between the ultimate IPROJGs values and the IPROJGs values at the end of shortrange period. Additional adjustments to the IPROJGs during the short-range period may be necessary for reconciliation between the long-range model and the short-range model. For the 2016 Trustees Report, we made IPROJG adjustments in the short-range period to better project the higher claims expected due to the planned hiring of more Administrative Law Judges (ALJs) and the growing backlog of pending ALJ claims.

\section*{Equation 3.2.2 - Exits}

The long-range model projects three types of exits from the disability rolls; death, recovery and conversion to an old-age beneficiary upon reaching normal retirement age (NRA). Deaths and recoveries are projected by multiplying the beginning currently entitled population by the probabilities of death only and recovery only, \(\left(q_{x}^{(d)}\right)\) and \(\left(q_{x}^{(r)}\right)\), respectively. Projected ( \(q_{x}^{(d)}\) ) and \(\left(q_{x}^{(r)}\right)\) by age, sex, and duration are calculated by multiplying the base probabilities by the respective projection factors by age group and sex for that year.

For the first ten years, we derive the recovery projection factors (RPROJGs) by age group and sex from linear interpolation between an estimated starting level for the RPROJGs and an estimated tenth-year projection target level for the RPROJGs. For each age group and sex, we calculate the starting RPROJGs the following way:
\[
\begin{gathered}
\operatorname{RPROJG}^{\mathrm{TR16}}(2015)=\text { RPROJG }^{\mathrm{TR15}}(2015) \times \text { actual recovery rate }(2015) / \text { estimated } \\
\text { recovery rate }{ }^{\mathrm{TR15}}(2015)
\end{gathered}
\]

Because there is no apparent upward or downward trend, we use the average recovery rates for the last ten historical years as the target values for the \(10^{\text {th }}\) year (2025). Then, for each age group and sex, we calculate the tenth year's RPROJGs as follows:
\[
\begin{aligned}
\operatorname{RPROJG}^{\mathrm{TR16}}(2025)= & \operatorname{RPROJG}^{\mathrm{TR16}}(2015) \times \text { target value recovery rate }(2025) / \\
& \text { actual recovery rate }(2015)
\end{aligned}
\]

For the second 10 years of the projection period, we linearly interpolate between the ultimate RPROJG value and the RPROJG value at the end of short-range period (2025). We assume attainment of ultimate recovery rates in the twentieth year of the projection period. Ultimate recovery rates by age group and sex are determined by analyzing historical recovery rates. We may make additional adjustments to the RPROJGs to reconcile with the short-range model.

For the first year of the projection period, the death projection factors (DPROJGs) by age group and sex are determined so that they achieve a targeted death rate. The targeted death rate is determined by fitting an exponential curve to historical death rates for DIBs by age group and sex (see Appendix 3.2-1). For the rest of the projection period, we assume the DPROJGs improve at the same rate as the general population for that age group and sex. We calculate the DPROJGs for each year by 5-year age group and sex the following way:

We may make additional adjustments to the DPROJGs to reconcile with the short-range model.

\section*{Equation 3.2.3 - Disabled-Worker Beneficiaries}

The projection begins with the latest data available from the mainframe of disabled-worker beneficiaries in current-payment status. This data is from a 100 percent sample of the Master Beneficiary Record (MBR) at the end of the year. We split up disabled-worker beneficiaries by age at entitlement, sex and duration of entitlement. We convert this population to a currently entitled population by dividing each age, sex and duration cell by the appropriate duration-age-sex-year-specific IBNR factor. An iterative process begins with new entitlements added to and exits subtracted from the previous year's currently entitled population to get the following year's currently entitled population with advancement of duration within the age of entitlement. We reduce this currently entitled population by multiplying by the appropriate duration-age-sex-year-specific IBNR factor. The result is the following year's disabled-worker beneficiaries in current-payment status. The process repeats over each sex, age of entitlement and duration of
entitlement throughout the projection period.

\section*{Equation 3.2.4 - Dependent Beneficiary of Disabled Workers}

There are six dependent-beneficiary categories; minor child, student child, disabled adult child, young spouse, married aged spouse and divorced aged spouse. We disaggregate projections by age of the beneficiary and sex of the account holder. We detail below the linkages and exposures used in each category of dependent beneficiaries.

\section*{Minor Child}

Exposure: Single SSA population by single ages 0-17
Linkages: pMCAGA = Probability that parent is under NRA
pMCDIA \(\quad=\) Probability that parent is disability insured given that the parent is under NRA
pMCDPA = Probability that disability insured parent under NRA is disabled
MCRES \(\quad=\) Residual Factor
Student Child
Exposure: Single SSA population by single ages 18-19
Linkages: pSCAGA = Probability that parent is under NRA
pSCDIA \(\quad=\) Probability that parent is disability insured given that the parent is under NRA
pSCDPA \(\quad=\) Probability that disability insured parent under NRA is disabled
pSCDPC = Probability that student is in an eligible school
SCRES \(\quad=\) Residual Factor

\section*{Disabled Adult Child}

Exposure: Single SSA population by age groups 18-19, 20-24, 25-29,
30-34, 35-39, 40-44, 45-49, 50-54, 55-59
Linkages: pDCAGA
= Probability that parent is under NRA pDCDIA \(\quad=\) Probability that parent is disability insured given that the parent is under NRA
pDCDPA \(\quad=\) Probability that disability insured parent under NRA is disabled
pDCDPC = Probability that adult child is disabled DCRES \(\quad=\) Residual Factor
Young Spouse
Exposure: Married SSA population by sex and by single ages 20-64
Linkages: pYSAGA = Probability that account holder is under NRA
pYSDIA = Probability that account holder is disability insured given that the account holder is under NRA
pYSDPA = Probability that disability insured account holder under NRA is disabled
pYSETB = Probability that young spouse is not subject to
\begin{tabular}{|c|c|c|}
\hline & & earnings test \\
\hline & pYSMCB & \(=\) Probability that young spouse has a minor child beneficiary in his/her care \\
\hline & pYSDCB & \(=\) Probability that young spouse has a disabled child beneficiary in his/her care \\
\hline & YSRES & = Residual Factor \\
\hline Married Aged & Spouse & \\
\hline Exposure: & Married SSA & on by sex and by single ages 62-100 \\
\hline Linkages: & pMSAGA & \(=\) Probability that account holder is under NRA \\
\hline & pMSDIA & \(=\) Probability that account holder is disability insured given that the account holder is under NRA \\
\hline & pMSDPA & \(=\) Probability that disability insured account holder under NRA is disabled \\
\hline & pMSFIB & \(=\) Probability that beneficiary is not insured \\
\hline & MSRES & \(=\) Residual Factor \\
\hline Divorced Age & d Spouse & \\
\hline Exposure: & Divorced SSA & ation by sex and by single ages 62-100 \\
\hline Linkages: & pDSDEA & \(=\) Probability that account holder is living \\
\hline & pDSAGA & \(=\) Probability that account holder is under NRA \\
\hline & pDSDIA & \(=\) Probability that account holder is disability insured given that the account holder is under NRA \\
\hline & pDSDPA & \(=\) Probability that disability insured account holder under NRA is disabled \\
\hline & pDSFIB & \(=\) Probability that beneficiary is not insured \\
\hline & pDSDMB & \(=\) Probability that beneficiary was married 10 or more years \\
\hline
\end{tabular}

We estimate the residual factors for each of the dependent categories using a 10-year Least Squares regression formula. We then hold these residual factor values constant for the duration of the long-range period. If the 10-year Least Squares method results in a negative residual factor, we hold the last historical residual factor instead.

We develop factors for a dependent beneficiary category to match short-range results during the first 10 years of the projection period. We phase these factors out linearly over the second ten years of the projection period.

\section*{Appendix: 3.2-1}

The following information provides details about the regression equations used in determining incidence rates and IPROJG values by age group and sex for the first ten years of the projection period.

\section*{Male}

Independent Variable: Unemp Rate(t)-Unemp Rate(t-1) for ages 16-19

Independent Variable: Unemp Rate(t-1)-Unemp Rate(t-2) for ages 16-19
Dependent Variable: 15-19 incidence rates
Observation Period: 1994-2012
Adjusted R square: 0.10285531
Standard Deviation: 0.57381920
Coefficient Intercept: 2.26564896
Coefficient Slope1: 0.10592411
Coefficient Slope2: 0.03810985

Independent Variable: Unemp Rate(t)-Unemp Rate(t-1) for ages 20-24
Independent Variable: Unemp Rate(t-1)-Unemp Rate(t-2) for ages 20-24
Dependent Variable: 20-24 incidence rates
Observation Period: 1994-2012
Adjusted R square: 0.28384642
Standard Deviation: 0.37775655
Coefficient Intercept: 2.25404835
Coefficient Slope1: 0.11947935
Coefficient Slope2: 0.06076160

Independent Variable: Unemp Rate(t)-Unemp Rate(t-1) for ages 25-29
Independent Variable: Unemp Rate(t-1)-Unemp Rate(t-2) for ages 25-29
Dependent Variable: 25-29 incidence rates
Observation Period: 1994-2012
Adjusted R square: 0.61414473
Standard Deviation: 0.15873244
Coefficient Intercept: 2.01407262
Coefficient Slope1: 0.08971943
Coefficient Slope2: 0.08165595

Independent Variable: Unemp Rate(t)-Unemp Rate(t-1) for ages 30-34
Independent Variable: Unemp Rate(t-1)-Unemp Rate(t-2) for ages 30-34
Dependent Variable: 30-34 incidence rates
Observation Period: 1994-2012
Adjusted R square: 0.16129961
Standard Deviation: 0.21323977
Coefficient Intercept: 2.50960268
Coefficient Slope1: 0.05582442
Coefficient Slope2: 0.04984552

Independent Variable: Unemp Rate(t)-Unemp Rate(t-1) for ages 35-39
Independent Variable: Unemp Rate(t-1)-Unemp Rate(t-2) for ages 35-39
\begin{tabular}{|c|c|}
\hline Dependent Variable: & 35-39 incidence rates \\
\hline Observation Period: & 1994-2012 \\
\hline Adjusted R square: & 0.25199638 \\
\hline Standard Deviation: & 0.25810193 \\
\hline Coefficient Intercept: & 3.29309120 \\
\hline Coefficient Slope1: & 0.09253364 \\
\hline Coefficient Slope2: & 0.08398824 \\
\hline Independent Variable: & Unemp Rate(t)-Unemp Rate(t-1) for ages 40-44 \\
\hline Independent Variable: & Unemp Rate(t-1)-Unemp Rate(t-2) for ages 40-44 \\
\hline Dependent Variable: & 40-44 incidence rates \\
\hline Observation Period: & 1994-2012 \\
\hline Adjusted R square: & 0.45850494 \\
\hline Standard Deviation: & 0.26583255 \\
\hline Coefficient Intercept: & 4.38675576 \\
\hline Coefficient Slope1: & 0.15219498 \\
\hline Coefficient Slope2: & 0.11955849 \\
\hline Independent Variable: & Unemp Rate(t)-Unemp Rate(t-1) for ages 45-49 \\
\hline Independent Variable: & Unemp Rate(t-1)-Unemp Rate(t-2) for ages 45-49 \\
\hline Dependent Variable: & 45-49 incidence rates \\
\hline Observation Period: & 1994-2012 \\
\hline Adjusted R square: & 0.57467043 \\
\hline Standard Deviation: & 0.29758276 \\
\hline Coefficient Intercept: & 5.91429408 \\
\hline Coefficient Slope1: & 0.19279725 \\
\hline Coefficient Slope2: & 0.19066446 \\
\hline Independent Variable: & Unemp Rate(t)-Unemp Rate(t-1) for ages 50-54 \\
\hline Independent Variable: & Unemp Rate(t-1)-Unemp Rate(t-2) for ages 50-54 \\
\hline Dependent Variable: & 50-54 incidence rates \\
\hline Observation Period: & 1994-2012 \\
\hline Adjusted R square: & 0.38894650 \\
\hline Standard Deviation: & 0.71152128 \\
\hline Coefficient Intercept: & 9.78110701 \\
\hline Coefficient Slope1: & 0.18205677 \\
\hline Coefficient Slope2: & 0.44231375 \\
\hline Independent Variable: & Unemp Rate(t)-Unemp Rate(t-1) for ages 55-59 \\
\hline Independent Variable: & Unemp Rate(t-1)-Unemp Rate(t-2) for ages 55-59 \\
\hline Dependent Variable: & 55-59 incidence rates \\
\hline
\end{tabular}
\begin{tabular}{|c|c|}
\hline Observation Period: & 1994-2012 \\
\hline Adjusted R square: & 0.21737183 \\
\hline Standard Deviation: & 0.97192205 \\
\hline Coefficient Intercept: & 15.93692204 \\
\hline Coefficient Slope1: & 0.21292470 \\
\hline Coefficient Slope2: & 0.50090579 \\
\hline Independent Variable: & Unemp Rate(t)-Unemp Rate(t-1) for ages 60-64 \\
\hline Independent Variable: & Unemp Rate(t-1)-Unemp Rate(t-2) for ages 60-64 \\
\hline Dependent Variable: & 60-64 incidence rates \\
\hline Observation Period: & 1994-2012 \\
\hline Adjusted R square: & 0.09481495 \\
\hline Standard Deviation: & 0.80332283 \\
\hline Coefficient Intercept: & 16.58265440 \\
\hline Coefficient Slope1: & 0.16853055 \\
\hline Coefficient Slope2: & 0.25477139 \\
\hline \multicolumn{2}{|l|}{Female} \\
\hline Independent Variable: & Unemp Rate(t)-Unemp Rate(t-1) for ages 16-19 \\
\hline Independent Variable: & Unemp Rate(t-1)-Unemp Rate(t-2) for ages 16-19 \\
\hline Dependent Variable: & 15-19 incidence rates \\
\hline Observation Period: & 1994-2012 \\
\hline Adjusted R square: & 0.09271622 \\
\hline Standard Deviation: & 0.38678032 \\
\hline Coefficient Intercept: & 1.55310909 \\
\hline Coefficient Slope1: & 0.08459404 \\
\hline Coefficient Slope2: & 0.04389858 \\
\hline Independent Variable: & Unemp Rate(t)-Unemp Rate(t-1) for ages 20-24 \\
\hline Independent Variable: & Unemp Rate(t-1)-Unemp Rate(t-2) for ages 20-24 \\
\hline Dependent Variable: & 20-24 incidence rates \\
\hline Observation Period: & 1994-2012 \\
\hline Adjusted R square: & 0.18699211 \\
\hline Standard Deviation: & 0.31645134 \\
\hline Coefficient Intercept: & 1.67692223 \\
\hline Coefficient Slope1: & 0.14156733 \\
\hline Coefficient Slope2: & 0.04837250 \\
\hline Independent Variable: & Unemp Rate(t)-Unemp Rate(t-1) for ages 25-29 \\
\hline Independent Variable: & Unemp Rate(t-1)-Unemp Rate(t-2) for ages 25-29 \\
\hline Dependent Variable: & 25-29 incidence rates \\
\hline
\end{tabular}
\begin{tabular}{ll} 
Observation Period: & 1994-2012 \\
Adjusted R square: & 0.42165067 \\
Standard Deviation: & 0.22629668 \\
Coefficient Intercept: & 1.86090404 \\
Coefficient Slope1: & 0.16570133 \\
Coefficient Slope2: & 0.08418915 \\
& \\
Independent Variable: & Unemp Rate(t)-Unemp Rate(t-1) for ages 30-34 \\
Independent Variable: & Unemp Rate(t-1)-Unemp Rate(t-2) for ages 30-34 \\
Dependent Variable: & \(30-34\) incidence rates \\
Observation Period: & \(1994-2012\) \\
Adjusted R square: & 0.49369795 \\
Standard Deviation: & 0.21871776 \\
Coefficient Intercept: & 2.61457070 \\
Coefficient Slope1: & 0.16256661 \\
Coefficient Slope2: & 0.16009397 \\
& \\
Independent Variable: & Unemp Rate(t)-Unemp Rate(t-1) for ages 35-39 \\
Independent Variable: & Unemp Rate(t-1)-Unemp Rate(t-2) for ages 35-39 \\
Dependent Variable: & \(35-39\) incidence rates \\
Observation Period: & \(1994-2012\) \\
Adjusted R square: & 0.51774106 \\
Standard Deviation: & 0.26500379 \\
Coefficient Intercept: & 3.69748652 \\
Coefficient Slope1: & 0.24507296 \\
Coefficient Slope2: & 0.18191224 \\
& \\
Independent Variable: & Unemp Rate(t)-Unemp Rate(t-1) for ages 40-44 \\
Independent Variable: & Unemp Rate(t-1)-Unemp Rate(t-2) for ages 40-44 \\
Dependent Variable: & \(40-44\) incidence rates \\
Observation Period: & \(1994-2012\) \\
Adjusted R square: & 0.33879496 \\
Standard Deviation: & 0.38126950 \\
Coefficient Intercept: & 5.01562760 \\
Coefficient Slope1: & 0.24180929 \\
Coefficient Slope2: & 0.23030862 \\
Independent Variable: & Unemp Rate(t)-Unemp Rate(t-1) for ages 45-49 \\
Independent Variable: & Unemp Rate(t-1)-Unemp Rate(t-2) for ages 45-49 \\
Dependent Variable: & \(45-49\) incidence rates \\
Observation Period: & \(1994-2012\) \\
Cbin
\end{tabular}
\begin{tabular}{ll} 
Adjusted R square: & 0.60466765 \\
Standard Deviation: & 0.31580899 \\
Coefficient Intercept: & 6.50091729 \\
Coefficient Slope1: & 0.35195573 \\
Coefficient Slope2: & 0.36075093 \\
& \\
Independent Variable: & Unemp Rate(t)-Unemp Rate(t-1) for ages 50-54 \\
Independent Variable: & Unemp Rate(t-1)-Unemp Rate(t-2) for ages 50-54 \\
Dependent Variable: & \(50-54\) incidence rates \\
Observation Period: & \(1994-2012\) \\
Adjusted R square: & 0.50948702 \\
Standard Deviation: & 0.51007087 \\
Coefficient Intercept: & 10.09608301 \\
Coefficient Slope1: & 0.26638909 \\
Coefficient Slope2: & 0.64403697 \\
& \\
Independent Variable: & Unemp Rate(t)-Unemp Rate(t-1) for ages 55-59 \\
Independent Variable: & Unemp Rate(t-1)-Unemp Rate(t-2) for ages 55-59 \\
Dependent Variable: & \(55-59\) incidence rates \\
Observation Period: & \(1994-2012\) \\
Adjusted R square: & 0.18540396 \\
Standard Deviation: & 0.66371367 \\
Coefficient Intercept: & 14.52275485 \\
Coefficient Slope1: & -0.02600911 \\
Coefficient Slope2: & 0.63080228 \\
Independent Variable: & Unemp Rate(t)-Unemp Rate(t-1) for ages 60-64 \\
Independent Variable: & Unemp Rate(t-1)-Unemp Rate(t-2) for ages 60-64 \\
Dependent Variable: & \(60-64\) incidence rates \\
Observation Period: & \(1994-2012\) \\
Adjusted R square: & 0.28209765 \\
Standard Deviation: & 0.46761182 \\
Coefficient Intercept: & 13.09340058 \\
Coefficient Slope1: & 0.17638034 \\
Coefficient Slope2: & 0.36960010 \\
& \\
Cla
\end{tabular}

The following information provides details about the exponentially fitted equations used in determining death rates by age group and sex for the first year of the projection period.

\section*{Male}

Independent Variable: Year
\begin{tabular}{|c|c|}
\hline Independent Variable: & 15-19 death rates \\
\hline Observation Period: & 2006-2015 \\
\hline Adjusted R square: & -0.11860 \\
\hline Standard Deviation: & 0.95019 \\
\hline Coefficient Intercept: & 47.36040 \\
\hline Coefficient Slope1: & -0.02238 \\
\hline Independent Variable: & Year \\
\hline Independent Variable: & 20-24 death rates \\
\hline Observation Period: & 2006-2015 \\
\hline Adjusted R square: & -0.06056 \\
\hline Standard Deviation: & 0.09962 \\
\hline Coefficient Intercept: & 17.90187 \\
\hline Coefficient Slope1: & -0.00765 \\
\hline Independent Variable: & Year \\
\hline Independent Variable: & 25-29 death rates \\
\hline Observation Period: & 2006-2015 \\
\hline Adjusted R square: & 0.51024 \\
\hline Standard Deviation: & 0.04494 \\
\hline Coefficient Intercept: & 34.41332 \\
\hline Coefficient Slope1: & -0.01594 \\
\hline Independent Variable: & Year \\
\hline Independent Variable: & 30-34 death rates \\
\hline Observation Period: & 2006-2015 \\
\hline Adjusted R square: & 0.46008 \\
\hline Standard Deviation: & 0.04629 \\
\hline Coefficient Intercept: & 32.68360 \\
\hline Coefficient Slope1: & -0.01501 \\
\hline Independent Variable: & Year \\
\hline Independent Variable: & 35-39 death rates \\
\hline Observation Period: & 2006-2015 \\
\hline Adjusted R square: & 0.88368 \\
\hline Standard Deviation: & 0.02330 \\
\hline Coefficient Intercept: & 45.68097 \\
\hline Coefficient Slope1: & -0.02136 \\
\hline Independent Variable: & Year \\
\hline Independent Variable: & 40-44 death rates \\
\hline
\end{tabular}
\begin{tabular}{|c|c|}
\hline Observation Period: & 2006-2015 \\
\hline Adjusted R square: & 0.76029 \\
\hline Standard Deviation: & 0.03283 \\
\hline Coefficient Intercept: & 42.46895 \\
\hline Coefficient Slope1: & -0.01965 \\
\hline Independent Variable: & Year \\
\hline Independent Variable: & 45-49 death rates \\
\hline Observation Period: & 2006-2015 \\
\hline Adjusted R square: & 0.92655 \\
\hline Standard Deviation: & 0.01543 \\
\hline Coefficient Intercept: & 39.75705 \\
\hline Coefficient Slope1: & -0.01818 \\
\hline Independent Variable: & Year \\
\hline Independent Variable: & 50-54 death rates \\
\hline Observation Period: & 2006-2015 \\
\hline Adjusted R square: & 0.90723 \\
\hline Standard Deviation: & 0.01694 \\
\hline Coefficient Intercept: & 38.81440 \\
\hline Coefficient Slope1: & -0.01759 \\
\hline Independent Variable: & Year \\
\hline Independent Variable: & 55-59 death rates \\
\hline Observation Period: & 2006-2015 \\
\hline Adjusted R square: & 0.90754 \\
\hline Standard Deviation: & 0.01033 \\
\hline Coefficient Intercept: & 25.20895 \\
\hline Coefficient Slope1: & -0.01075 \\
\hline Independent Variable: & Year \\
\hline Independent Variable: & 60-64 death rates \\
\hline Observation Period: & 2006-2015 \\
\hline Adjusted R square: & 0.56940 \\
\hline Standard Deviation: & 0.01905 \\
\hline Coefficient Intercept: & 18.86489 \\
\hline Coefficient Slope1: & -0.00753 \\
\hline
\end{tabular}

\section*{Female}

Independent Variable: Year Independent Variable: 15-19 death rates
\begin{tabular}{ll} 
Observation Period: & \(2006-2015\) \\
Adjusted R square: & -0.09575 \\
Standard Deviation: & 1.12833 \\
Coefficient Intercept: & 117.83769 \\
Coefficient Slope1: & -0.05741 \\
& \\
Independent Variable: & Year \\
Independent Variable: & \(20-24\) death rates \\
Observation Period: & \(2006-2015\) \\
Adjusted R square: & 0.09785 \\
Standard Deviation: & 0.10587 \\
Coefficient Intercept: & -30.46411 \\
Coefficient Slope1: & 0.01639 \\
& \\
Independent Variable: & Year \\
Independent Variable: & \(25-29\) death rates \\
Observation Period: & \(2006-2015\) \\
Adjusted R square: & -0.12440 \\
Standard Deviation: & 0.06779 \\
Coefficient Intercept: & 3.28215 \\
Coefficient Slope1: & -0.00049 \\
& \\
Independent Variable: & Year \\
Independent Variable: & \(30-34\) death rates \\
Observation Period: & \(2006-2015\) \\
Adjusted R square: & -0.07011 \\
Standard Deviation: & 0.05976 \\
Coefficient Intercept: & 10.89892 \\
Coefficient Slope1: & -0.00421 \\
& \\
Independent Variable: & Year \\
Independent Variable: & \(35-39\) death rates \\
Observation Period: & \(2006-2015\) \\
Adjusted R square: & 0.43111 \\
Standard Deviation: & 0.01559 \\
Coefficient Intercept: & 12.23822 \\
Coefficient Slope1: & -0.00480 \\
Independent Variable: & Year \\
Independent Variable: & \(40-44\) death rates \\
Observation Period: & \(2006-2015\) \\
& \\
Coper
\end{tabular}
\begin{tabular}{ll} 
Adjusted R square: & 0.57780 \\
Standard Deviation: & 0.02696 \\
Coefficient Intercept: & 24.55290 \\
Coefficient Slope1: & -0.01083 \\
& \\
Independent Variable: & Year \\
Independent Variable: & \(45-49\) death rates \\
Observation Period: & \(2006-2015\) \\
Adjusted R square: & 0.81660 \\
Standard Deviation: & 0.02002 \\
Coefficient Intercept: & 31.36496 \\
Coefficient Slope1: & -0.01412 \\
& \\
Independent Variable: & Year \\
Independent Variable: & \(50-54\) death rates \\
Observation Period: & \(2006-2015\) \\
Adjusted R square: & 0.88405 \\
Standard Deviation: & 0.01190 \\
Coefficient Intercept: & 25.11935 \\
Coefficient Slope1: & -0.01093 \\
& \\
Independent Variable: & Year \\
Independent Variable: & \(55-59\) death rates \\
Observation Period: & \(2006-2015\) \\
Adjusted R square: & 0.82912 \\
Standard Deviation: & 0.01735 \\
Coefficient Intercept: & 28.91360 \\
Coefficient Slope1: & -0.01277 \\
Independent Variable: & Year \\
Independent Variable: & \(60-64\) death rates \\
Observation Period: & \(2006-2015\) \\
Adjusted R square: & 0.86626 \\
Standard Deviation: & 0.01526 \\
Coefficient Intercept: & 29.38372 \\
Coefficient Slope1: & -0.01293 \\
& \\
\hline
\end{tabular}

\subsection*{3.3. Old-Age and Survivors Insurance}

\section*{3.3.a. Overview}

Every month, the Social Security program pays benefits to retired workers and their dependents. It also provides benefits to eligible dependents of deceased workers. The OLD-AGE AND SURVIVORS subprocess projects the number of people expected to receive benefits over the next 75 years. The projection method is very similar to the method used for dependent beneficiaries of disabled workers in the DISABILITY subprocess. We compute the projection of beneficiaries by multiplying a subset of the Social Security area population by a series of probabilities of the conditions that a person must meet to receive benefits. The main program receives all necessary input data and performs all preliminary calculations. It then calls each individual beneficiary type subroutine where it makes all beneficiary calculations.

We categorize retired workers and their dependent beneficiaries as follows:
- retired workers ( \(R W N\) ) by age (62-95+), sex, and marital status (single, married, widowed, divorced)
- aged spouses of retired workers (ASRWN), by age (62-95+), sex of the account holder, and marital status of the beneficiary (married, divorced)
- young spouses of retired workers (YSRWN) by age-group (under 25, 25-29,..., 65-69) and sex of the account holder
- minor, student, and disabled adult children of retired workers (MCRWN, SCRWN, and \(D C R W N\), respectively) by age of the child (0-17 for minor, 18-19 for student, age groups \(18-19,20-24, \ldots, 55-59,60+\) for disabled adult) and sex of the account holder

Dependent beneficiaries of deceased workers include:
- aged spouses of deceased workers, ASDWN, by age (60-95+), sex of the account holder, marital status (widowed, divorced) and insured status (insured, uninsured)
- disabled spouses of deceased workers ( \(D S D W N\) ) by age (50-69), sex of the account holder and marital status (widowed, divorced)
- young spouses of deceased workers (YSDWN) by age-group (under 25, 25-29,..., 65-69), sex of the account holder and marital status of the beneficiary (widowed, divorced)
- minor, student, and disabled adult children of deceased workers (MCDWN, SCDWN, and \(D C D W N\), respectively) by age of the child (0-17 for minor, 18-19 for student, age groups 18-19, 20-24,..., 55-59, 60+ for disabled adult) and sex of the account holder

Lastly, we estimate the number of deaths of insured workers (LUMSUM) by 5-year age group (20-24, 25-29,..., 80-84, 85+) and sex.

Equations 3.3.1-13 indicates the flow of calculations of beneficiaries.
\[
\begin{align*}
A S D W N & =A S D W N(\cdot)  \tag{3.3.1}\\
R W N & =R W N(\cdot)  \tag{3.3.2}\\
A S R W N & =A S R W N(\cdot)  \tag{3.3.3}\\
D S D W N & =D S D W N(\cdot)  \tag{3.3.4}\\
M C R W N & =M C R W N(\cdot)  \tag{3.3.5}\\
M C D W N & =M C D W N(\cdot)  \tag{3.3.6}\\
S C R W N & =S C R W N(\cdot)  \tag{3.3.7}\\
S C D W N & =S C D W N(\cdot)  \tag{3.3.8}\\
D C R W N & =D C R W N(\cdot)  \tag{3.3.9}\\
D C D W N & =D C D W N(\cdot)  \tag{3.3.10}\\
Y S R W N & =Y S R W N(\cdot)  \tag{3.3.11}\\
Y S D W N & =Y S D W N(\cdot)  \tag{3.3.12}\\
L U M S U M & =L U M S U M(\cdot) \tag{3.3.13}
\end{align*}
\]

The appendix 3.3-1 at the end of this section provides a listing with explanation of the acronyms used in this documentation.

\section*{3.3.b. Input Data}

We update all data annually unless otherwise noted. Timing of data received is denoted 'BOY' (beginning of year) or 'EOY' (end of year). Workflow reference is bolded in parentheses after each item.

\section*{Long-Range OCACT Data}

\section*{Demography}
- Social Security area population by year (EOY 1970-2095), single year of age (0-100+), sex, and marital status (single, married, widowed, divorced) (3a)
- Deaths by year (during years 2015-2095), age group (20-24,...,80-84, 85+) and sex (3b)
- Average number of children per family by year (EOY 1970-2095), and age group of the householder (20-24,...,60-64) (3c)
- Children by year (EOY 1970-2095), single year of age (0-18), sex of primary account holder (parent), status of primary (62+ or deceased), and age of the other parent (15-19,20-24,...,65-69, 70+, total ages) (3d)
- Married couples by year (EOY 1970-2095), age of husband (62-95+) and age of wife (62-95+) (3e)
- Persons with an aged spouse by year (EOY 1970-2095), age group (15-24, 25-29,...,6569) and sex (3f)
- Probabilities of death by year (EOY 1941-2100), single year of age (-1,100), and sex (3g)

\section*{Economics}
- Covered wages and employment in the Federal Civilian and State and Local Sectors (during years 1998-2095) (2c)
- Labor force participation rates for age 62 by year (during years 1970-2095) and sex (2d, 3h)

\section*{Beneficiaries}
- Fully insured persons by year (EOY 1969-2095), age (14-95+), sex, and marital status (single, married, widowed, divorced) (3i)
- Disabled-worker beneficiaries in current pay by year (EOY 1970-2095), age (62-66) and sex (3j)
- Converted DI to OAI beneficiaries by year (EOY 1970-2095), age(65-95+) and sex (3j)
- Disability prevalence rates by year (EOY 1970-2095), age (50-66) and sex (3k)

\section*{Short-Range OCACT Data}
- Insured aged spouses of deceased workers by year (EOY 1974-2015), age (60-95+) and sex (2b)
- Retired worker beneficiaries in-current-pay status by age (62-70, 70+) and sex for EOY 2014-2025 (2e)
- We receive the following for EOY 2015 (2b):
a. Aged spouses of deceased workers by age (60-95+), sex and marital status (widowed, divorced)
b. Retired workers by age (62-95+) and sex
c. Aged spouses of retired workers by age (62-95+), sex and marital status (married, divorced)
d. Disabled widow(er)s by age (50-65), sex and marital status (widowed, divorced)
e. Minor children by age (0-17), sex of parent and status of parent (retired, deceased)
f. Student children by age (18-21), sex of parent and status of parent (retired, deceased)
g. Disabled adult children by age group (18-19, 20-24,...,55-59, 60+), sex of parent and status of parent (retired, deceased)
h. Young spouses of retired workers by age group (under 25, 25-29,...,60-64, 65-69) and sex
i. Young spouses of deceased workers by age group (under 25, 25-29,...,65-69), sex and marital status (widowed, divorced)
j. Total parent beneficiaries

Note: Each year, we append data for the most recent historical year.
- We receive the following for EOY 2025 (2e):
a. Retired workers by age group (62-64, 65-69) and sex
b. Insured widows by age group ( \(60-64, \ldots, 80-84,85+\) )
c. Uninsured widows by age group (60-64, 65+)
d. Total disabled widows
e. Female young spouses of deceased workers
f. Female aged spouses of retired workers by age group (62-64, 65-67, 68-70, 71+)
g. Female young spouses of retired workers
h. Minor children by status of parent
i. Student children by status of parent
j. Disabled adult children by status of parent
- Total amount of lump-sum death payments during 2014 (2a)

\section*{Other Input Data}
- For EOY 1970-2004, obtained from the MBR10PER dataset on the mainframe:
a. Aged spouses of deceased workers by age (60-95+), sex and marital status (widowed, divorced)
b. Retired workers by age (62-95+) and sex
c. Aged spouses of retired workers by age (62-95+), sex and marital status (married, divorced)
d. Disabled widow(er)s by age (50-64), sex and marital status (widowed, divorced)
e. Minor children by age (0-17), sex of parent and status of parent (retired, deceased)
f. Student children by age (18-19), sex of parent and status of parent (retired, deceased)
g. Disabled adult children by age (20-95+), sex of parent and status of parent (retired, deceased)
h. Young spouses of retired workers by age group (under 25, 25-29,...,60-64) and sex
i. Young spouses of deceased workers by age group (under \(25,25-29, \ldots, 65-69\) ), sex and marital status (widowed, divorced)
j. Total parent beneficiaries

Note: We will not update this data.
- Number of beneficiaries with benefits withheld due to receipt of a significant government pension by sex and marital status (married, widowed) for EOY 2014 from the 2015 Annual Statistical Supplement (2c)
- Age distribution of beneficiaries with benefits withheld due to receipt of a significant government pension by age (60-95+) and sex, computed as an average from the 2011 through 2015 WEP 100-percent sample (2g)
- Proportions of disabled adult children of retired and deceased workers (proportioned by age and sex of the child) from the 2003 MBR ten-percent sample. (Note: The RSB program calculates disabled adult children by sex of the primary account holder, not by sex of the child. The RSB program outputs a file, which we use for Annual Update \#9, which calculates beneficiaries by sex. Therefore, we apply the 2003 proportions to estimate the breakdown of disabled adult children by sex of the child. We will not update this input.). Not used for SOSI.
- Schedule of normal retirement age (NRA), delayed retirement credit, and actuarial
reduction factors for ages more than 3 years below NRA and less than 3 years below \(N R A\) for years 1970-2095 from the Social Security website (Note: these values are only updated when there is a Social Security law change regarding the NRA) (2h)
- Prevalence rate regression coefficients (slopes and y-intercept value by sex) (2d)
- Regressed prevalence rate by sex for the most recent historical year (2d)
- Adjustment factors which account for the difference between estimated and actual historical retired worker prevalence rates by year (EOY 1970-2095), age (63-69) and sex (2d, 2f)
- Adjustment factors which account for the difference between projected beneficiary values for the tenth year of the projection period made by the Long-Range and Short-Range offices. Factors are computed for:
a. Retired workers by age group (62-64, 65-69) and sex (2f)
b. Insured widows by age group ( \(60-64, \ldots, 80-84,85+\) ) (2e)
c. Uninsured widows by age group (60-64, 65+) (2e)
d. Total disabled widows (2e)
e. Female young spouses of deceased workers (2e)
f. Female aged spouses of retired workers by age group (62-64, 65+) (2e)
g. Female young spouses of retired workers (2e)
h. Minor children by status of parent (2e)
i. Student children by status of parent (2e)
j. Disabled adult children by status of parent (2e)
- Adjustment factors for auxiliary beneficiary categories due to the 2014 immigration executive action by sex for years 2016-2095 (2j). Factors are computed for:
a. Aged spouses of deceased workers (widow(er)s)
b. Aged spouses of retired workers
c. Disabled adult children of retired workers
d. Minor children of deceased workers
e. Disabled adult children of deceased workers
f. Young spouses of deceased workers
- Number of two-earner couples by year of birth of the spouse where the primary account holder files for and may or may not suspend their retired worker benefit and their spouse receives a spousal benefit for years 2006-2015 from a 100 percent MBR sample. (2i)
- Number of one-earner couples by year of birth of the spouse where the primary account holder files for and suspends their retirement worker benefit and their spouse receives a spousal benefit for years 2008-2015 from a 1 percent MBR sample. (2i)
- Number of one-earner couples by year of birth of the primary account holder where the primary account holder files for and suspends their retirement worker benefit and their spouse receives a spousal benefit for years 2008-2015 from a 1 percent MBR sample. (2i)

\section*{3.3.c. Development of Output}

We use several acronyms to describe the equations presented below. Acronyms not preceded by a subscript generally refer to the number of beneficiaries. For example, RWN refers to the number of retired workers. Acronyms preceded with a ' \(p\) ' refer to probabilities. For example, \(p R W_{F I A}\) refers to the probability that a person is fully insured.

\section*{Equation 3.3.1 - Aged Spouses of Deceased Workers (ASDWN)}

Aged Spouses of Deceased Workers
Exposures: SSA population by age (60-95+), sex and marital status (widowed and divorced)
Linkages: \(p A S D W_{D E A}=\) probability that the primary account holder \((\mathrm{PAH})\) is deceased
\(p A S D W_{\text {FIA }}=\) probability that the PAH was fully insured at death \(p A S D W_{\text {MBB }}=\) probability that the widow(er) is not receiving a youngspouse benefit for the care of a child \(p A S D W_{\text {FIB }}=\) probability that the aged-widow(er) is or is not fully insured
\(p A S D W_{G P B}=\) probability that the aged-widow(er)'s benefits are not withheld or offset totally because of receipt of a significant government pension based on earnings in noncovered employment
\(p A S D W_{\text {RES }}=\) probability that a widow(er) eligible to receive his/her own retired-worker benefits would instead apply for and receive widow(er) benefits

We project the number of aged spouses of deceased workers (widow(er)s), along with all linkage factors, by age, sex of the account holder ( \(s a=1\) for male, \(s a=2\) for female), marital status and insured status. Age ranges from 60 to 95+, marital status includes widowed \((\mathrm{mb}=1)\) and divorced \((\mathrm{mb}=2)\), and insured status includes insured (in=1) and uninsured (in=2). Note that all variables preceded by the letter \(p\) refer to calculated probabilities. We calculate the projected number of insured aged spouses of deceased workers age 60 to 70, and uninsured aged spouses of deceased workers age 60 to \(95+\) as follows:
\[
\begin{align*}
A S D W N & =A S D W_{P O P} \times p A S D W_{D E A} \times p A S D W_{F I A} \times p A S D W_{M B B} \\
& \times p A S D W_{F I B} \times p A S D W_{G P B} \times p A S D W_{R E S} \tag{3.3.1}
\end{align*}
\]

For those with a marital status of widowed ( \(\mathrm{mb}=1\) ) we apply an additional adjustment factor (ImmEAFact(typ)) to the above calculation for the 2014 immigration executive order by type of beneficiary (typ=1: insured with male PAH; 2: insured with female PAH; 15: uninsured with male PAH; 16: uninsured with female PAH):
\[
A S D W N=A S D W N \times \operatorname{ImmEAFact}(t y p)
\]

For each sex we calculate the projected number of insured aged spouses of deceased workers over age 70 by applying mortality rates to the population already receiving such benefits:
\[
A S D W N_{N, Y E A R}=A S D W N_{N-1, Y E A R-1} \times\left(1-q x_{N-1, Y E A R}\right)
\]

Where N is the age, and \(q x_{N-1, Y E A R}\) is the death rate for age \(\mathrm{N}-1\) in the given year.
\(\boldsymbol{A S D W}_{\text {Pop }}\) represents the subset of the population from which we draw these beneficiaries. We set this equal to the Social Security area population \(\left(S S A P O P_{m b}\right)\) for each possible marital status.
\[
A S D W_{P O P}=S S A P O P_{m b}
\]
\(\boldsymbol{p A S D} \boldsymbol{W}_{\text {DEA }}\) represents the probability that the primary account holder (PAH) is deceased. For the widowed population, we set this factor equal to one. For the divorced population, we set this factor equal to the portion of the total widowed \(\left(S S A P O P_{\text {wid }}\right)\) and married \(\left(S S A P O P_{m a r}\right)\) population who are widowed.
\[
p A S D W_{D E A}= \begin{cases}1, & m b=1(\text { widowed }) \\ \frac{S S A P O P_{\text {wid }}}{S S A P O P_{\text {mar }}+S S A P O P_{\text {wid }}}, & m b=2(\text { divorced })\end{cases}
\]
\(\boldsymbol{p A S D W} \boldsymbol{W}_{\text {FIA }}\) represents the probability that the PAH was fully insured at death. For a given age of widow, \(A W\), we assume that the age of her deceased husband, \(A H\), ranges from \(A W-6\) to \(A W+12\) with a lower and upper bound of 60 and \(95+\). Further, we assume that the more likely age of the husband is \(A W+3\). For each age, we calculate \(p A S D W_{\text {FIA }}\) as a weighted average of the portion of the Social Security area population who are fully insured at each possible age of the husband \(\left(\right.\) FINS \(\left._{A H}\right)\). For example, for a widow age 70, we assume that the age of her husband is between 64 and 82 , therefore we calculate the weighted average of the portion of the population who are fully insured males, applying the highest weight of 10 to age 73, and a linearly reduced weight to zero for each age above and below 73 . We use the same concept for widow(er)s with the assumption that the age of his deceased wife ranges from \(A H-12\) to \(A H+6\), with a greater likelihood of her age being \(A H-3\). Let WEIGHT represent the specific weight applied to each potential age of the spouse.
\[
\begin{gathered}
W^{W} I G H T_{A H}=10-|A W+3-A H| \\
W_{E I G H T_{A W}=10-|A H-3-A W|} \\
p A S D W_{F I A}=\left\{\begin{array}{lc}
\frac{\sum_{A H=A W-6}^{A W+12} W E I G H T_{A H} \times F I N S_{A H}}{\sum_{A H=A W-6}^{A W+12} W E I G H T_{A H}}, & s a=1 \\
\frac{\sum_{A W=A H-12}^{A H+6} W_{A E I G H T}^{A W}}{} \times F I N S_{A W} \\
\sum_{A W=A H-12}^{A H+6} W E I G H T_{A W}
\end{array}\right. \\
s a=2
\end{gathered}
\]
\(\boldsymbol{p A S D} \boldsymbol{W}_{\text {MBB }}\) represents the probability that the widow(er) is not receiving a young-spouse benefit for the care of a child. A widow(er) can receive a young-spouse benefit up to age 69 if he/she meets all other eligibility requirements. Since the minimum age requirement to receive a widow(er) benefit is 60, it is necessary to remove those receiving a young-spouse benefit (YSDWN \({ }^{a b}\) ), where ab represents the 5 -year age bracket \({ }^{1}\). We assume a uniform breakdown to divide the age groups into single-age estimates.

For in = 1 (insured):
\[
p A S D W_{M B B}=1
\]

For in = 2 (uninsured):
\[
\begin{aligned}
& \text { FACTOR }_{A G E}=\left\{\begin{array}{cr}
1, & 65 \leq \text { age and } N R A \geq \text { age }+1 \\
N R A-\text { age }, & \text { age }<N R A<\text { age }+1 \\
0, & \text { elsewhere }
\end{array}\right. \\
& p A S D W_{M B B}=\left\{\begin{array}{lr}
1-\frac{0.2 \times Y S D W N^{60-64}}{A S D W_{P O P} \times p A S D W_{D E A} \times p A S D W_{F I A}}, & \text { age }=60-64 \\
1-\frac{Y S D W N^{65-69}}{A S D W_{P O P} \times p A S D W_{D E A} \times p A S D W_{F I A}} \times \frac{F^{-4 C T O R_{\text {age }}}}{\sum_{65}^{69} F A C T O R_{\text {age }}}, & 65 \leq \text { age } \leq N R A \\
1, & \text { age }>N R A
\end{array}\right.
\end{aligned}
\]
\(\boldsymbol{p A S D} W_{\text {FIB }}\) represents the probability that the aged widow(er) is fully insured. For insured widow(er)s, pASDW \({ }_{F I B}\) is the portion of the Social Security area population that is fully insured (FINS) at each age, sex, and marital status. For uninsured widow(er)s, pASDW FIB is simply one minus the probability for insured widow(er)s.
\[
p A S D W_{F I B}=\left\{\begin{aligned}
\frac{F I N S}{S S A P O P_{m b}}, & \text { in }=1 \\
1-\frac{F I N S}{S S A P O P_{m b}}, & \text { in }=2
\end{aligned}\right.
\]

Where in represents the insured status of the account holder.
\(\boldsymbol{p A S D} \boldsymbol{W}_{\text {GPB }}\) represents the probability that the aged-widow(er)'s benefits are not withheld or completely offset because of receipt of a significant government pension based on earnings in noncovered employment. According to the 1977 amendments, Social Security benefits are

\footnotetext{
\({ }^{1}\) There are no young spouses at NRA or above.
}
subject to reduction by up to two-thirds of non-covered government pension. GPWHLD represents the total number of widow(er) beneficiaries (for all ages) expected to receive a significant government pension. rGPOAGE represents the ratio of the total for each given age. If a person is insured, this implies that he/she is eligible to receive Social Security benefits based on his/her own earnings regardless of a government pension. Therefore, we do not apply a factor.

For in = 1 (insured):
\[
p A S D W_{G P B}=1
\]

For in = 2 (uninsured):
\[
p A S D W_{G P B}= \begin{cases}1, & \text { year } \leq 1978 \\ 1-\frac{r G P O A G E \times G P W H L D}{A S D W_{P O P} \times p A S D W_{D E A} \times p A S D W_{F I A} \times p A S D W_{M B B} \times p A S D W_{F I B}}, & \text { year }>1978\end{cases}
\]
\(\boldsymbol{p A S D} \boldsymbol{W}_{\text {RES }}\) represents the probability that a widow(er), who is eligible to receive widow(er)'s benefits, will actually receive benefits. In particular, for in \(=1\), this factor is equivalent to the probability that a widow(er) eligible to receive his/her own retired-worker benefits would instead apply for and receive widow(er) benefits. For all historical years, we calculate \(p A S D W_{R E S}^{\text {year }}\) as the ratio of \(A S D W N\), the actual number of widow(er)s, to the number of persons meeting all previously mentioned requirements by age, sex, insured status, and marital status.
\(p A S D W_{R E S}^{\text {year }}=\frac{A S D W N}{A S D W_{P O P} \times p A S D W_{D E A} \times p A S D W_{F I A} \times p A S D W_{M B B} \times p A S D W_{F I B} \times p A S D W_{G P B}}, \quad\) year \(<\) TRYR

Where TRYR is the Trustees Report year.
For each age, sex, insured status, and marital status, we use a least squares regression over the last ten years of historical data to determine a starting value in TRYR-1 for \(p A S D W_{\text {RES }}^{\text {year }}\) from which we project future values. In addition, for each sex, insured status, and marital status, we graduate the regressed values of \(p A S D W_{R E S}^{\text {TRYR-1 }}\) over age using a weighted minimized third-difference formula to produce ESTRES \({ }^{\text {ASDW }}\). ESTRES \({ }^{\text {ASDW }}\) are the preliminary estimates of \(p A S D W_{R E S}^{\text {TRYR }+9}\), the values in the tenth year of the projection period. In addition, we apply adjustments to the widows ( \(\mathrm{sa}=1\) ) by age group (60-64, 65-69 for insured; 60-64, 65+ for uninsured), \(S R A D J^{A S D W}\), to the tenth year of the projection period in order to match the projections made by the Short-Range office. We linearly interpolate the values of \(p A S D W_{R E S}^{\text {year }}\) for intermediate years between \(p A S D W_{R E S}^{T R Y R-1}\) and \(p A S D W_{R E S}^{T R Y R+9}\) (equal to \(E S T R E S^{A S D W} * S R A D J^{A S D W}\) ). After the \(10^{\text {th }}\) year, we linearly grade these adjustment factors to one over the 10 years beyond the end of the short-range period, thus gradually eliminating the effect of the short-range adjustment factors, so that we ultimately return to long-range projections.

We also apply adjustments to the 70+ insured widows by age group (70-74, 75-79, 80-84, \(85+\) ). The adjustments for the intermediate years between the TRYR and TRYR +9 are applied on a cohort basis to the number of aged spouses of deceased workers (ASDWN) using the following formula:
\[
A S D W N_{N, Y E A R}=A S D W N_{N, Y E A R} \times \operatorname{SRADJ}\left(\frac{1}{h}\right)
\]

Where h is the number of years the adjustment is applied for the cohort. For age 70, we linearly grade these adjustment factors to one over the 10 years beyond the end of the shortrange period, thus gradually eliminating the effect of the short-range adjustment factors, so that we ultimately return to long-range projections.

\section*{Equation 3.3.2 - Retired Workers (RWN)}

\section*{Retired Workers}

Exposure: SSA population by age (62-95+), sex and marital status (single, married, widowed and divorced)
Linkages: \(p R W_{\text {FIA }}=\) probability that the primary account holder (PAH) is insured \(p R W_{D B B}=\) probability that the PAH is not receiving a disabled-worker benefit
\(p R W_{\text {WBB }}=\) probability that the PAH is not receiving a widow(er) benefit \(p R W_{\text {RES }}=\) retirement prevalence rate; probability that a fully insured worker not receiving disability or widow(er)'s benefits would receive a retired-worker benefit

We project the numbers of retired-worker beneficiaries, along with all linkage factors, by age, sex, and marital status. Age ranges from 62 to \(95+\), and marital status includes single, married, widowed, and divorced ( \(m s=1\) to 4 ). We calculate the projected number of retiredworker beneficiaries as follows:
\[
\begin{equation*}
R W N=R W_{P O P} \times p R W_{F I A} \times p R W_{D B B} \times p R W_{W B B} \times p R W_{R E S} \tag{3.3.2}
\end{equation*}
\]
\(\boldsymbol{R} \boldsymbol{W}_{\boldsymbol{P O P}}\) represents the subset of the population from which we draw these beneficiaries. We set this equal to the Social Security area population \(\left(S S A P O P_{m s}\right)\) for \(m s=1\) to 4 .
\[
R W_{P O P}=S S A P O P_{m s}
\]
\(\boldsymbol{p} \boldsymbol{R} \boldsymbol{W}_{\text {FIA }}\) represents the probability that the primary account holder ( PAH ) is insured. We set this factor equal to the portion of the Social Security area population that is fully insured (FINS) for \(m s=1\) to 4 .
\[
p R W_{F I A}=\frac{F I N S}{R W_{P O P}}
\]
\(\boldsymbol{p} \boldsymbol{R} \boldsymbol{W}_{\text {DBB }}\) represents the probability that the PAH is not receiving a disabled-worker or disability-conversion benefit. We set this factor equal to the portion of fully insured workers who are neither disabled-worker beneficiaries nor converted from disabled-worker
beneficiaries (DIBCON). ASDWN represents the number of aged spouses of deceased workers.
\[
p R W_{D B B}= \begin{cases}1-\frac{D I B C O N}{R W_{P O P} \times p R W_{F I A}}, & m s=1-2 \\ \left(1-\frac{D I B C O N+A S D W N}{R W_{P O P} \times p R W_{F I A}}\right)^{\left(\frac{D I B C O N}{D I B C O N+A S D W N}\right)}, & m s=3-4\end{cases}
\]
\(\boldsymbol{p} \boldsymbol{R} \boldsymbol{W}_{\text {WBB }}\) represents the probability that the PAH is not receiving a widow(er) benefit. We set this factor equal to the portion of fully insured workers that is not aged spouses of deceased workers.
\[
p R W_{W B B}= \begin{cases}1, & m s=1-2 \\ \left(1-\frac{D I B C O N+A S D W N}{R W_{P O P} \times p R W_{F I A}}\right)^{\left(\frac{A S D W N}{D I B C O N+A S D W N}\right)}, & m s=3-4\end{cases}
\]
\(\boldsymbol{p} \boldsymbol{R} \boldsymbol{W}_{\text {RES }}{ }^{N, y e a r}\) represents the retirement prevalence rate, which is the probability that a fully insured worker not receiving disability or widow(er)'s benefits would receive retired-worker benefits as of the given age, \(N\), for the given year. In order to estimate the future prevalence rate, the program first calculates the historical values of \(p R W_{R E S}^{N, \text { year }}\).

For each historical year and sex, we calculate \(p R W_{R E S}^{N, \text {,year }}\) as the ratio of \(R W N\), the actual number of retired workers, to the number of persons meeting all previously mentioned requirements by age, sex, and marital status.
\[
p R W_{R E S}^{N, y e a r}=\frac{R W N}{R W_{P O P} \times p R W_{F I A} \times p R W_{D B B} \times p R W_{W B B}}, \quad N=62-95+\text { and } y e a r<\text { TRYR }
\]

Historical prevalence rates at age 62 follow an inverse relationship with (1) labor force participation rates ( \(L F P R^{\text {year }}\) ) at age 62, by sex, and (2) increases in the normal retirement age over the historical period. We assume this relationship holds in the projection period, and therefore we used it to calculate \(R E G P R^{\text {year }}\), the regressed prevalence rate based on the projected \(L F P R^{\text {year }}\) at age 62 for each year and sex, and the number of months from age 62 to the normal retirement age (monthNRA \({ }^{\text {year }}\) ). Note that we calculate prevalence rates on a cohort basis \({ }^{2}\). The regression equation used to estimate the prevalence rates is:
\[
\begin{gathered}
R E G P R^{\text {year }}=-1.02646 \times L F P R^{\text {year }}+-0.00364 \times \text { month } N R A^{\text {year }}+1.08627 \text { for male with } \\
\text { an } \mathrm{R}^{2} \text { value of } 0.899087, \\
\text { and } \\
R E G P R^{\text {year }}=-0.52481 \times L F P R^{\text {year }}+-0.00705 \times \text { monthNRA }{ }^{\text {year }}+0.94229 \text { for female with }
\end{gathered}
\]

\footnotetext{
\({ }^{2}\) For example, to calculate the projected number of 65 year olds in a given year, the prevalence rate at age 62 is needed. This is actually the prevalence rate that occurred three years ago at age 62.
}
an \(R^{2}\) value of 0.905583
We then set the future prevalence rate at age 62, \(p R W_{R E S}^{62, y e a r}\), equal to the sum of the regressed prevalence rate ( \(\left.R E G P R^{\text {year }}\right)\) and \(E R R O R\), the difference between the actual prevalence rate and the regressed prevalence rate in the most recent historical year, which we phase out linearly over 10 years.
\[
p R W_{R E S}^{62, \text { year }}=R E G P R^{\text {year }}+(E R R O R) \times \max \left(0, \frac{T R Y R+9-\text { year }}{10}\right), \quad N=62 \text { and year } \geq \text { TRYR- } 1
\]

To compute \(p R W_{R E S}^{N, \text { year }}\) for ages 63 to 69 in the projection period, we must calculate several preliminary variables. These include:
- \(\quad\) MBAPIA \(A_{N}\), for \(N=62,70\) (same for both sexes),
- \(E S T P R_{N}^{\text {year }}\), for \(N=63,69\) and by sex, and
- \(\quad D I F F A D J_{N}\), for \(N=63,69\) and by sex.
\(M B A P I A_{N}\) is the ratio of the monthly benefit amount (MBA) to the primary insurance amount (PIA) at age \(N\) and is calculated on a cohort basis for \(N=62,70\). We base the calculation of \(M B A P I A_{N}\) on the normal retirement age (NRA), delayed retirement credits (DRC), and actuarial reduction factors, ARFLE3 when the difference between \(N R A\) and age at retirement is less than 3, and ARFGT3 when the difference is greater than 3 within each cohort. If a person retires after \(N R A\), his/her benefits are increased by \(D R C\) for each year the age exceeds \(N R A\). If a person retires before \(N R A\), his/her benefits are decreased by ARFLE3 for each of the first three years that \(N R A\) exceeds the age, and further decreased by ARFGT3 for any remaining years.
\[
M B A P I A_{N}=\left\{\begin{array}{lr}
1+(N-N R A) \times D R C, & N \geq N R A \\
1-(N R A-N) \times A R F L E 3, & N R A-3 \leq N \leq N R A \\
1-3 \times A R F L E 3-(N R A-3-N) \times A R F G \square 3, & N<N R A-3
\end{array}\right.
\]
\(E S T P R_{N}{ }^{\text {year }}\), the estimated prevalence rate at age \(N\), is then calculated as the prevalence rate at age \(62\left(p R W_{R E S}^{62 \text { year }-(N-62)}\right)\) plus an estimate on the expected portion of the remaining probability ( \(1-p R W_{R E S}^{62, \text { year }-(N-62)}\) ), that a potential retired worker will actually retire by that given age. We base this estimate on \(M B A P I A_{N}\), assuming that the retirement decision by a worker is totally and completely influenced by the expected change in the portion of PIA that is payable at each age relative to the potential change after the initial eligibility.
\[
E S T P R_{N}^{\text {year }}=p R W_{R E S}^{62, \text { year }-(N-62)}+\left(1-p R W_{R E S}^{62, y e a r-(N-62)}\right) \times \frac{M B A P I A_{N}-M B A P I A_{62}}{M B A P I A_{70}-M B A P I A_{62}}, \quad N=63-69
\]

In the first year of the projection period, an adjustment \(\left(D I F F A D J_{N}\right)\) is made which accounts for the difference between the actual and estimated prevalence rate at each age in the most recent historical years. For ages 63 to 69, the value used beginning in TRYR is the average of the last 5 years' differences between the actual and estimated PR. With the exception of
age 66, we hold this value constant throughout the projection period.
For years in which the NRA increases, \(D I F F A D J_{N}\) is greater for a period of time both before and after the NRA transitions to a new age. This affects DIFFADJ 66 due to the increase in NRA from age 65 to age 66 between the years 2000-2005, and the scheduled increase from age 66 to 67 over the years 2017-2022. Thus for the years preceding the change in NRA from age 66 to 67, we use the average from the last five historical years, which includes the increase in the difference between actual and estimated prevalence rates from when the NRA changed from age 65 to 66 , in order to calculate \(D I F F A D J_{66}\). After the NRA increases to age 67 in 2026, we use \(D I F F A D J 2_{66}\) which is the average difference between actual and estimated prevalence rates from the years 1995-1999 since at that time it is expected to decrease to the level it was prior to the increase in NRA to age 66 in the year 2000. In the years during which the NRA transitions to age 67 (2021-2026) we use a linear interpolation based on the number of months that the NRA has increased in order to phase out DIFFADJ 66 and phase in DIFFADJ2 \({ }_{66}\). It will remain at DIFFADJ2 \({ }_{66}\) for the rest of the projection period.
\[
p R W_{R E S}^{N, y e a r}=E S T P R_{N}+D I F F A D J_{N}
\]

For age 70, we assume that the values of the latest actual \(p R W_{\text {RES }}^{N, \text { year }}\) by sex change linearly to the ultimate level of 0.995 for male and 0.99 for female over the first 20 years of the projection period.

For year \(\geq\) TRYR:
\[
p R W_{R E S}^{70, \text { year }}= \begin{cases}0.995-\left[0.995-p R W_{R E S}^{70, T R Y R-1}\right] \times \max \left(0, \frac{T R Y R+19-y e a r}{20}\right), & s a=1 \\ 0.99-\left[0.99-p R W_{R E S}^{70, T R Y R-1}\right] \times \max \left(0, \frac{T R Y R+19-y e a r}{20}\right), & s a=2\end{cases}
\]

For ages 71 and older, we assume \(p R W_{R E S}^{N, y e a r}\) stays constant at the level when the age was 70 because there is no incentive to delay applying for benefits beyond age 70 .
\[
p R W_{R E S}^{N, y e a r}=p R W_{R E S}^{70, \text { year }-(N-70)}, \text { for } N=71-95+\text { and year } \geq \text { TRYR }
\]

In addition, we apply adjustments by age group (62-64, 65-69) and sex, \(S R A D J^{R W}\), to the tenth year of the projection period in order to match the projections made by the Short-Range office. We also apply these adjustments to \(p R W_{\text {RES }}^{N, \text {,year }}\) for all years after TRYR +9 . The values of \(p R W_{R E S}^{N, \text { year }}\) for intermediate years are linearly interpolated between \(p R W_{R E S}^{N, T R Y R-1}\) and \(p R W_{R E S}^{N, T R Y R+9}\).

In response to the elimination of the file and suspend and deemed filing claiming strategies for aged spouses by the Bipartisan Budget Act of 2015 (BBA), we expect that the married
and divorced retired workers who would have delayed receiving their worker benefits will instead file for benefits earlier. As such, we use the unsuspRW(ag) factor to increase the projected number of married workers for those who would have suspended their benefits but will now begin receiving them earlier.
\[
\operatorname{unsusp} R W(a g)=\frac{\operatorname{suspRW}(a g)+R W N_{a g, M, M a r}+R W N_{a g, F, M a r}}{R W N_{a g, M, M a r}+R W N_{a g, F, M a r}}-1
\]

Where \(\operatorname{suspRW}(\mathrm{ag})\) is the number of retired workers who would have suspended their benefits for the file and suspend strategy. Since the utilization of this filing strategy had been expected to continue to grow in popularity had it not been eliminated, we apply a behavioral increase factor, RWBehavioralInc(ag), to unsuspRW(ag) to reflect the increasing number of retired workers who will no longer delay their benefits now that the filing strategy is no longer available.
\[
\text { RWBehavioralInc }(a g)=\left\{\begin{array}{lr}
1.5, & \text { for those born in } 1947 \\
1.03, & \text { for those born in } 1948 \\
1.0, & \text { for those born after } 1948
\end{array}\right.
\]

We use the \(\operatorname{DFRetWrkr}(\mathrm{ag})\) factor to increase the number of married and divorced retired workers for the insured aged spouses who will no longer be able to delay receiving their worker benefit while receiving their aged spouse benefit. We calculate these factors as:
\[
\begin{aligned}
& \text { DFRetWrkr }(a g) \\
& \quad=\frac{\text { DeemdFiler }(a g)+R W N_{a g, M, M a r}+R W N_{a g, F, M a r}+R W N_{a g, M, D i v}+R W N_{a g, F, D i v}}{R W N_{a g, M, M a r}+R W N_{a g, F, M a r}+R W N_{a g, M, D i v}+R W N_{a g, F, D i v}}-1
\end{aligned}
\]

Where DeemdFiler(ag) is the number of insured aged spouses who would have used the deemed filing strategy. We assume that 80 percent of these deemed filers will file for worker benefits earlier now that the strategy is no longer available, with \(1 / 3\) claiming at age 66 , and \(1 / 9\) at each age 67 through 69.

As the NRA transitions from 66 to 67, we reduce both the unsuspRW(66) and DFRetWrkr(66) by one minus the increase in NRA expressed as a fraction of a year with each factor being set to zero once the NRA is 67 .

For each age, retired workers are further broken down by age at entitlement, \(A E\), by multiplying the number of retired workers at age \(N\) by the ratio of the incidence rate at \(A E\) ( \(N-A E\) years prior) to the prevalence rate at age \(N\).
\[
{ }_{A E}^{N} R W N^{\text {year }}={ }^{N} R W N^{\text {year }} \times \frac{A E^{I N C R A T E} \text { year }-(N-A E)}{p R W_{R E S}^{N, y e a r}}, \quad A E \leq \mathrm{N}
\]
where we calculate the incidence rate for a given \(A E=N\) and year as the change in the prevalence rate at age \(N\) to the prevalence rate at age \(N-1\) in the previous year.
\[
{ }_{A E} I^{I N C R A T E} \begin{array}{lr}
p R W_{R E S}^{N, y e a r}, & N=A E=62 \\
p R W_{R E S}^{N, y e a r}-p R W_{R E S}^{N-1, \text { year }-1}, & 63 \leq N=A E \leq 69 \\
1-p R W_{R E S}^{N-1, \text { year }-1,} & N=A E=70
\end{array}
\]

Additional adjustments are applied to the incidence rates in order to account for the increase in retired worker beneficiaries due to the elimination of the deemed filing and file and suspend strategies. Since many of the beneficiaries utilizing these strategies delayed receiving their worker benefit until age 70, we assume that the retired workers added through the unsuspRW (ag) and DFRetWrkr (ag) factors will survive until age 70. As such, our incidence rate calculations are adjusted to reflect the age at which the new retired workers claim benefits and to ensure that they remained on the rolls until age 70.

\section*{Equation 3.3.3 - Aged Spouses of Retired Workers (ASRWN)}

Aged Spouses of Retired Workers
Exposures: SSA population by age (62-95+, sex, and marital status (married and divorced)
Linkages: \(p A S R W_{\text {DEA }}=\) probability that the primary account holder ( PAH ) is not deceased
\(p A S R W_{A G A}=\) probability that the PAH is of the required age \(p A S R W_{\text {FIA }}=\) probability that the PAH is fully insured \(p A S R W_{\text {CPA }}=\) probability that the PAH is receiving benefits \(p A S R W_{M B B}=\) probability that the beneficiary is not receiving a youngspouse
\[
\left.\left.\begin{array}{rl}
p A S R W_{F I B}= & \begin{array}{l}
\text { benefit } \\
p A S R W_{G P B}=
\end{array} \\
p \text { probability that the aged-spouse is not fully insured } \\
\text { probability that the aged-spouse's benefits are not withheld } \\
\text { because of receipt of a significant government pension based } \\
\text { on earnings in noncovered employment }
\end{array}\right] \begin{array}{l}
\text { probability that a couple will engage in a filing } \\
\text { strategy in which the PAH files for and suspends their benefit } \\
\text { and an uninsured spouse receives an aged spouse benefit }
\end{array}\right\} \begin{aligned}
& \text { probability that a couple will engage in a filing } \\
& \text { strategy in which the PAH files for and may or may not } \\
& \text { suspend their benefit and the fully insured spouse files only for } \\
& \text { their aged spouse benefit. }
\end{aligned}
\]

We project the number of aged spouses of retired workers, along with all linkage factors, by
age, sex of the account holder \((s a=1,2)\), and marital status of the beneficiary. Age ranges from 62 to \(95+\), and marital status includes married \((m b=1)\) and divorced ( \(m b=2\) ).
Additionally we apply an adjustment factor (ImmEAFact(typ)) for the 2014 immigration executive order by beneficiary type (typ = 3: married aged wife of retired worker; 4: married aged husband of retired worker; 5: divorced aged wife of retired worker; 6: divorced aged husband of retired worker). We calculate the projected number of aged spouses of retired workers as follows:
\[
\begin{align*}
& A S R W N=A S R W_{P O P} \times p A S R W_{D E A} \times p A S R W_{A G A} \times p A S R W_{F I A} \times p A S R W_{C P A} \times \\
& \quad p A S R W_{M B B} \times p A S R W_{F I B} \times p A S R W_{G P B} \times p A S R W_{F S 1} \times p A S R W_{F S 2} \times \\
& \quad p A S R W_{R E S} \times I m m E A F a c t(t y p) \tag{3.3.3}
\end{align*}
\]

ASR \(\boldsymbol{W}_{\text {POP }}\) represents the subset of the population from which these beneficiaries are drawn, and we set it equal to the Social Security area population \(\left(S S A P O P_{m b}\right)\) for \(m b=1,2\).
\[
A S R W_{P O P}=S S A P O P_{m b}
\]
pASR \(W_{\text {DEA }}\) represents the probability that the PAH is not deceased. For the married population, we do not apply a factor. For the divorced population, we set the factor equal to the portion of the total married and widowed population who are married.
\[
p A S R W_{D E A}= \begin{cases}1, & m b=1(\text { married }) \\ \frac{S S A P O P_{\operatorname{mar}}}{S S A P O P_{\operatorname{mar}}+S S A P O P_{\text {wid }}}, & m b=2(\text { divorced })\end{cases}
\]
\(\boldsymbol{p A S R} \boldsymbol{W}_{\text {AGA }}\) represents the probability that the PAH is of the required age, and we set it equal to the portion of the married population with a spouse ( PAH ) at least age 62.
\[
p A S R W_{A G A}=\frac{\sum_{A A=62}^{95} \operatorname{mar}(A A, A S)}{S S A P O P_{A S, \text { mar }}}
\]

Where mar(AA,AS) is the number of married couples where the age of the account holder is AA and the spouse of the account holder is age AS.
\(\boldsymbol{p A S R} \boldsymbol{W}_{\text {FIA }}\) represents the probability that the PAH is fully insured, and we set it equal to the portion of married couples of the required age where the PAH is fully insured ( \(F I \_P A H\) ). For example, when the program estimates the number of female aged spouse of retired workers, this factor will find the portion where their spouse, the male PAH, is fully insured.
\[
p A S R W_{F I A}=\frac{\sum_{A A=62}^{95}\left[\operatorname{mar}(A A, A S) \times F I_{-} P A H_{A A, \operatorname{mar}}\right]}{\sum_{A A=62}^{95} \operatorname{mar}(A A, A S)}
\]
\(\boldsymbol{p A S R} \boldsymbol{W}_{\text {CPA }}\) represents the probability that the PAH is receiving benefits. We set this factor equal to the portion of eligible married couples where the PAH is receiving benefits
(RETIRED). If the beneficiary is divorced, we do not apply a factor, since it is not required for the retired worker to be receiving benefits for the divorced aged spouse to receive benefits.
\[
p A S R W_{C P A}=\left\{\begin{array}{lr}
1, & y e a r \geq 1985 \text { and } \\
& m b=2 \\
\frac{\sum_{A A=62}^{95}\left[\operatorname{mar}(A A, A S) \times F I_{-} P A H_{A A, \text { mar }} \times R E T I R E D\right]}{\sum_{A A=62}^{95} \operatorname{mar}(A A, A S) \times F I_{-} P A H_{A A, \text { mar }}}, & \text { elsewhere }
\end{array}\right.
\]
\(\boldsymbol{p A S R} \boldsymbol{W}_{\text {MBB }}\) represents the probability that the beneficiary is not receiving a young-spouse benefit. If the beneficiary is age 70 or older or if the beneficiary is divorced, we do not apply a factor. Otherwise, we set this factor equal to the portion of potentially eligible widow(er)s where the spouse of the PAH is not receiving a young-spouse benefit (YSRWN \({ }^{a b}\) ), where ab represents the 5 -year age group. \({ }^{3}\)

For \(m b=1\) (married):
\[
F_{A C T O R_{A G E}=}=\left\{\begin{array}{lr}
1, & 65 \leq \text { age and } N R A \geq \text { age }+1 \\
N R A-\text { age }, & \text { age }<N R A<\text { age }+1 \\
0, & \text { elsewhere }
\end{array}\right.
\]
\[
p A S R W_{M B B}= \begin{cases}1-\frac{0.2 \times Y S R W N^{60-64}}{A S R W_{P O P} \times p A S R W_{D E A} \times p A S R W_{A G A} \times p A S R W_{F I A} \times p A S R W_{C P A}}, & \text { age }=62-64 \\ 1-\frac{Y S R W N^{65-69}}{A S R W_{P O P} \times p A S R W_{D E A} \times p A S R W_{A G A} \times p A S R W_{F I A} \times p A S R W_{C P A}} \times \\ \frac{F A C T O R_{a g}}{\sum_{65}^{69} F A C T O R_{a g}}, & 65 \leq \text { age } \leq N R A \\ 1, & \text { elsewhere }\end{cases}
\]

For \(m b=2\) (divorced):
\[
p A S R W_{M B B}=1
\]
\(\boldsymbol{p A S R} \boldsymbol{W}_{\text {FIB }}\) represents the probability that the aged spouse is not fully insured, and is

\footnotetext{
\({ }^{3}\) There are no young spouses at NRA or above.
}
therefore not receiving a retired-worker benefit based on his/her own earnings. We set this factor equal to the portion of the married and divorced population that is not fully insured. For example, when the program estimates the number of female aged spouse of retired workers, this factor will find the portion of female beneficiaries that is fully insured.
\[
p A S R W_{F I B}=1-\frac{F I N S}{S S A P O P}, \quad m b=1-2
\]
\(\boldsymbol{p A S R} \boldsymbol{W}_{\boldsymbol{G P B}}\) represents the probability that the aged-spouse's benefits are not withheld because of receipt of a significant government pension based on earnings in noncovered employment. GPWHLD represents the total number of aged spouse of retired-worker beneficiaries (for all ages) expected to receive a significant government pension. rGPOAGE represents the ratio of the total for each given age.

\(\boldsymbol{p A S R} \boldsymbol{W}_{\boldsymbol{F S} 1}\) represents the probability that a couple engages in a filing strategy in which the PAH files for and suspends their worker benefit and the uninsured aged spouse files for a spousal benefit. Due to the Bipartisan Budget Act of 2015 (BBA), this filing strategy is no longer available to married couples if the worker files and suspends their worker benefits after April 29, 2016. We assume that any couple with a worker age 66 and older who was planning on using this strategy will do so before it is eliminated. For the couples where the worker reaches age 66 after April 29, 2016, we assume that the worker will instead opt to receive their worker benefit when they would have suspended so that their spouse can receive the aged spouse benefit.

For \(m b=1\) (married):
Since we assume the one-earner couples will begin receiving benefits at the same time regardless of whether the file and suspend strategy is available, we expect that the projected prevalence rate, \(p A S R W_{\text {RES }}\), will be similar to that in the historical period. Thus, we set the \(p A S R W_{F S 1}\) to 1 for all years less than TRYR.

Additionally, we set the factor to 1 for all ages in the year 2020, the year in which the last group of workers able to make use of this strategy would reach age 70, and later. For the years prior to 2020, we calculate \(p A S R W_{F S 1}\) as the percent of new couples engaging in this strategy over the total number of married aged spouses receiving benefits in the last historical year.
\[
p A S R W_{F S 1}^{\text {year }}=1+\frac{\text { FSCouple } 1(a g)-\text { lastyrFSCouple } 1(a g)}{A S R W N_{a g, M, M a r}+A S R W N_{a g, F, M a r}}
\]

Where FSCouple1(ag) is the number of one-earner couples expected to use the file and suspend strategy, and lastyrFSCouple1(ag) is the number of one-earner couples using this
strategy in the last historical year. We then phase out \(p A S R W_{F S 1}\) for those who turn 66 after April 29, 2016, based upon the age distribution of the workers newly suspending their benefits in order make use of this strategy.

For mb = 2 (divorced):
\[
p A S R W_{F S 1}^{\text {year }}=1
\]
\(\boldsymbol{p A S R} \boldsymbol{W}_{\text {FS2 }}\) represents the probability that a couple engages in a filing strategy in which the PAH files for and may or may not suspend their worker benefit and their fully insured spouse receives only a spousal benefit thus delaying receiving their worker benefit. For historical years 2006 to TRYR-1, we set this equal to 1 plus the ratio of couples ages NRA to 70 where the PAH has filed for benefits and their fully insured spouse is receiving only a spousal benefit (FSCouple2(ag, sa, mb)) to the actual number of aged spouse beneficiaries. Due to the BBA, those turning 62 in 2016 or later will no longer be able to use this filing strategy. For the projection period prior to the elimination of this filing strategy, we set \(p A S R W_{F S 2}^{\text {year }}\) equal to \(p A S R W_{F S 2}^{\text {year }-1}\) multiplied by a behavioral increase factor (BehaviorInc(ag)). This factor accounts for the increase in the number of couples engaging in this filing strategy due to knowledge of the strategy becoming more widespread as well as the growth in the age 62 insured married and divorced population (insur62). We set \(p A S R W_{F S 2}\) to 1 in the projection period as those age 62 in 2016 reach NRA and above.

For year < TRYR:
\[
p A S R W_{F S 2}^{\text {year }}=\left\{\begin{array}{lr}
1+\frac{F S C o u p l e}{}(a g, s a, m b) \\
A S R W N_{a g, s a, m b}^{\text {year }}, & 2006 \leq y e a r \text { and } N R A \leq a g \leq 70 \\
1, & \text { elsewhere }
\end{array}\right.
\]

For year \(\geq\) TRYR:
\[
\begin{gathered}
\text { BehaviorInc }(\mathrm{ag})=\left\{\begin{array}{lr}
1.4 \times \text { insur } 62, & \text { for those born in } 1946 \\
1.3 \times \text { insur } 62, & \text { for those born in } 1947 \\
1.2 \times \text { insur } 62, & \text { for those born in } 1948 \\
1.0 \times \text { insur } 62, & \text { for those born after } 1948
\end{array}\right. \\
p A S R W_{F S 2}^{\text {year }}=\left\{\begin{array}{r}
1+\left(p A S R W_{F S 2}^{\text {year }-1}-1\right) \times \text { Behavior } \operatorname{lnc}(\text { ag }), \quad N R A \leq a g \leq 70 \text { and } \\
\text { year }-2016+62<a g \\
1,
\end{array}\right. \text { elsewhere }
\end{gathered}
\]
\(\boldsymbol{p A S R} \boldsymbol{W}_{\text {RES }}\) represents the probability that a person who is eligible to receive aged-spouse benefits actually receive the benefits. For all historical years, we calculate \(p A S R W_{R E S}^{\text {year }}\) as the
ratio of \(A S R W N\), the actual number of aged spouses receiving benefits, to the number of persons meeting all previously mentioned requirements by age, sex, and marital status.

For year < TRYR:
\(p A S R W_{\text {RES }}^{\text {year }}=\)
\(\frac{A S R W N}{A S R W_{P O P} \times p A S R W_{D E A} \times p A S R W_{A G A} \times p A S R W_{F I A} \times p A S R W_{C P A} \times p A S R W_{M B B} \times p A S R W_{F I B} \times p A S R W_{G P B} \times p A S R W_{F S 1} \times p A S R W_{F S 2}}\)

For each age, sex , and marital status, we use a least squares regression over the last ten years of historical data to determine a starting value in TR-1 for \(p A S R W_{R E S}^{\text {year }}\) from which we project future values. In addition, for each sex and marital status, we graduate the regressed values of \(p A S R W_{R E S}^{\text {TRYR-1 }}\) over age using a weighted minimized third-difference formula to compute ESTRES \({ }^{\text {ASRW }}\). ESTRES \({ }^{A S R W}\) are the preliminary estimates of \(p A S R W_{R E S}^{\text {TRYR+9 }}\), the values in the tenth year of the projection period. For female spouses, we apply additional adjustments by age group (62-64, 65+), SRADJ \({ }^{\text {ASRW }}\), to the tenth year of the projection period in order to match the projections made by the Short-Range office. We linearly interpolate the values of \(p A S R W_{R E S}^{\text {year }}\) for intermediate years interpolated between \(p A S R W_{R E S}^{\text {TRYR- } 1}\) and \(p A S R W_{R E S}^{\text {TRYR }+9}\) (equal to \(E S T R E S^{\text {ASRW }} \times S R A D J^{\text {ASRW }}\) ). After the \(10^{\text {th }}\) year of the projection period, we linearly grade these adjustment factors to one over the 10 years beyond the end of the short-range period, thus gradually eliminating the effect of the shortrange adjustment factors, so that we ultimately return to the long-range projections.

\section*{Equation 3.3.4 - Disabled Spouses of Deceased Workers (DSDWN)}

\section*{Disabled Spouses of Deceased Workers}

> Exposure: SSA population by age (50-69), sex and marital status (widowed and divorced)
> Linkages: \(p D S D W_{D E A}=\) probability that the primary account holder (PAH) is deceased
> \(p D A D W_{\text {FIA }} \quad=\quad\) probability that the PAH was fully insured at death \(p D S D W_{S S B}=\) probability that the spouse is indeed disabled \(p D S D W_{\text {DEB }}=\) probability that the disabled spouse is not receiving another type of benefit
> \(p D S D W_{\text {RES }}=\) probability that a person who is eligible to receive disabled-
> spouse benefits actually receive the benefits

We project the number of disabled spouses of deceased workers, along with all linkage factors, by age, sex of the account holder ( \(s a=1\) for male, \(s a=2\) for female), and marital status. Age ranges from 50 to 69 , and marital status includes widowed ( \(m b=1\) ) and divorced \((m b=2)\). We calculate the projected number of disabled spouses of deceased workers as follows:
\[
\begin{align*}
D S D W N= & D S D W_{P O P} \times p D S D W_{D E A} \times p D S D W_{F I A} \times \\
& p D S D W_{S S B} \times p D S D W_{D E B} \times p D S D W_{R E S} \tag{3.3.4}
\end{align*}
\]
\(\boldsymbol{D S D W}_{\mathbf{P O P}}\) represents the subset of the population from which we draw these beneficiaries, and we set it equal to the Social Security area population \(\left(S S A P O P_{m b}\right)\) for \(m b=1,2\).
\[
D S D W_{P O P}=S S A P O P_{m b}
\]
\(\boldsymbol{p} \boldsymbol{D S D} W_{\text {DEA }}\) represents the probability that the primary account holder is deceased. For the widowed population, we do not apply a factor. For the divorced population, we set this factor equal to the portion of the total widowed and married population that is widowed.
\[
p D S D W_{D E A}= \begin{cases}1, & m b=1(\text { widowed }) \\ \frac{S S A P O P_{\text {wid }}}{{S S S A P O P_{\text {wid }}+S S A P O P_{\text {mar }}},} & m b=2(\text { divorced })\end{cases}
\]
\(\boldsymbol{p} D S D W_{\text {FIA }}\) represents the probability that the PAH was fully insured at death. Given the age of the widow, \(A W\), we assume that the age of her deceased husband, \(A H\), ranges from \(A W-6\) to \(A W+12\) with a lower and upper bound of 50 and \(95+\). Further, we assume that the more likely age of the husband is \(A W+3\). For each age, we calculate \(p D S D W_{\text {FIA }}\) as a weighted average of the portion of the Social Security area population that is fully insured (FINS), at each possible age of the husband. For example, if the widow is age 65, we assume that the age of the husband is between 59 and 77. Therefore, when we calculate the weighted average of the portion of the population who are fully insured males, we apply the highest weight of ten to age 68 and linearly reduce the weight to zero for each age above and below 68. We use the same concept for widow(er)s with the assumption that the age of his deceased wife ranges from \(A H-12\) to \(A H+6\), with a greater likelihood of her age being \(A H-3\). Let WEIGHT represent the specific weight applied to each age.
\[
\begin{gathered}
{W E I G H T_{A H}=10-|A W+3-A H|}^{W_{E I G H T_{A} \square}=10-|A H-3-A W|} \\
p D S D W_{F I A}=\left\{\begin{array}{lr}
\frac{\sum_{A H=A W-6}^{A W+12} W E I G H T_{A H} \times F I N S_{A H}}{\sum_{A H=A W-6}^{A W+12} W E I G H T_{A H}}, & s a=1 \\
\frac{\sum_{A W=A H-12}^{A H+6} W E I G H T_{A W} \times F I N S_{A W}}{\sum_{A W=A H-12}^{A H+6} W E I G H T_{A W}}, & s a=2
\end{array}\right.
\end{gathered}
\]
pDSDW \({ }_{\text {SSB }}\) represents the probability that the spouse is indeed disabled. We set this factor equal to the disability prevalence rates (DISPREV) by age and sex received from the DISABILITY subprocess.
\[
p D S D W_{S S B}=D I S P R E V
\]
\(\boldsymbol{p D S D} \boldsymbol{W}_{\text {DEB }}\) represents the probability that the disabled spouse is not dually eligible for another type of benefit. We assume this factor remains at a constant level by sex.
\[
p D S D W_{D E B}= \begin{cases}0.85, & s a=1 \\ 0.06, & s a=2\end{cases}
\]
\(\boldsymbol{p D S D W}_{\text {RES }}\) represents the probability that a person who is eligible to receive disabled-spouse benefits actually receive the benefits. For all historical years, we calculate \(p D S D W_{\text {RES }}^{\text {year }}\) as the ratio of \(D S D W N\), the actual number of disabled spouses of deceased workers receiving benefits, to the number of persons meeting all previously mentioned requirements by age, sex, and marital status.
\[
p D S D W_{R E S}^{\text {year }}=\frac{D S D W_{N}}{D S D W_{P O P} \times p D S D W_{D E A} \times p D S D W_{F I A} \times p D S D W_{S S B} \times p D S D W_{D E B}}, \quad \text { year }<\mathrm{TRYR}
\]

For ages 50 to 64, and each sex, and marital status, we use the average \(p D S D W_{R E S}^{\text {year }}\) over the last five years of historical data to determine starting values for \(p D S D W_{\text {RES }}^{\text {year }}\) from which we project future values. We phase in these averages over five years using a linear interpolation between the five-year averages and \(p D S D W_{R E S}^{T R Y R-1}\). For female disabled spouses, we apply an adjustment, \(S R A D J^{D S D W}\), to the tenth year of the projection period in order to match the projections made by the Short-Range office. We exponentially interpolate the values of \(p D S D W_{R E S}^{y e s r}\) for intermediate years between \(p D S D W_{R E S}^{T R Y R-1}\) and \(p D S D W_{R E S}^{T R Y R+9}\) (equal to \(\left.p D S D W_{R E S}^{\text {year }} \times S R A D J^{D S D W}\right)\). After the \(10^{\text {th }}\) year of the projection period, we linearly grade the adjustment factors to one over the 10 years beyond the end of the short-range period, thus gradually eliminating the effect of the short-range adjustment factors, so that we ultimately return to the long-range projections.

For the projection period, for ages 65 to 69 where age is less than the NRA when the beneficiary was age \(60, p D S D W_{R E S}^{\text {year }}\) is equal to \(p D S D W_{R E S}^{\text {year }}\) at age 64 times an adjustment that accounts for the additional ages as NRA changes.
\[
\begin{gathered}
F A C T O R_{\text {age }}= \begin{cases}1, & \text { NRA at age } 60 \geq \text { age }+1 \\
N R A \text { at age } 60-\text { age }, & \text { age }<N R A \text { at age } 60<\text { age }+1\end{cases} \\
p D S D W_{R E S}^{\text {year }}=p D S D W_{R E S}^{\text {year, } 64} \times\left(\frac{p_{D S D W_{R E S}}^{\text {year }, 64}}{p D S D W_{R E S}^{\text {Ear }, 63}}\right)^{(\text {age }-64)} \times \text { FACTOR }_{\text {age }}, \quad 65 \leq \text { age } \leq 69 \text { and } \\
\text { age }<\text { NRA at age } 60
\end{gathered}
\]

Equation 3.3.5 - Minor Children of Retired Workers (MCRWN)
We project the number of minor children of retired workers, MCRWN, by age of the minor
( \(\underline{a m}=0\) to 17) and sex of the account holder ( \(s a=1\) for male, \(s a=2\) for female).
For children of male retired workers:
\[
\begin{equation*}
M C R W N_{M, s a}^{\text {year }}=m c r w C h i_{M, a m}^{\text {year }} \times \frac{r w_{\_} s m_{\text {year }}}{\text { pop_sum }} \text { year } \quad \times M C R W_{R E S}^{\text {year }} \tag{3.3.5.1}
\end{equation*}
\]

For the number of minor children of male retired workers, we multiply the number of children under the age of 18 with a father who is at least 62 years old and who's other parent is not deceased ( \(\operatorname{mcrwChi} i_{M, a m}^{\text {year }}\) ) by the total number of male retired workers ages 62 to 71 ( \(r w_{-}\)sum \(_{\text {year }}\) ) divided by the total number of males in the population ages 62 to 71 (pop_sum year).
\(\boldsymbol{p} \boldsymbol{M C R} \boldsymbol{W}_{\text {RES }}\) represents the probability that a child who is eligible to receive minor-child benefits actually receives them. In the historical period, we calculate this as the ratio of the total number of minor children of retired workers actually receiving benefits to the total eligible to receive benefits.
\[
p M C R W_{R E S}^{\text {year }}=\frac{M C R W N_{M}^{\text {year }}}{\operatorname{mcrwChi}_{M, a m}^{\text {year }} \times \frac{r w_{-} \text {Sum }_{\text {year }}}{\text { pop_Sum }_{\text {year }}}}
\]

For the projection period, we take the average of these ratios over the last ten historical years and phase it in using a linear interpolation between the average and \(p M C R W_{R E S}^{T R Y R-1}\). We apply an adjustment, \(S R A D J^{M C R W}\), to the tenth year of the projection period in order to match the projections made by the Short-Range office. The adjustment factor is phased in over the first 10 years of the projection period with the full factor applied in the tenth year. After the 10th year of the projection period, we linearly grade these adjustment factors to one over the 10 years beyond the end of the short-range period, thus gradually eliminating the effect of the short-range adjustment factors, so that we ultimately return to the long-range projections.

For children of female retired workers:
\[
\begin{equation*}
M C R W N_{F, s a}^{\text {year }}=\frac{M C R W N_{\text {Total, }, ~} \times m \text { crwPrcntFem }}{\text { year }} \times \frac{M C R W N_{\text {antF }}^{T R Y R-1}}{1-\text { mcrwPrcntFem }} \tag{3.3.5.2}
\end{equation*}
\]
\(\boldsymbol{m c r w P r} \boldsymbol{n t F e m}\) represents the average over the last five historical years of the percentage of the total number of minor children of female retired workers, ages 0 to 17, to the total number of minor children for both male and female retired workers.
\[
\frac{M C R W N_{\text {Total }, F}^{\text {year }}}{M C R W N_{\text {Total }, F}^{\text {year }} \times M C R W N_{\text {Total }, M}^{\text {year }}}
\]

We maintain the total \(M C R W N_{0-17, F}^{Y R}\) at this percentage over the entire projection period by multiplying our estimate of \(M C R W N_{0-17, M}^{Y R}\) by mcrwPrcntFem divided by its complement.
\[
M C R W N_{0-17, F}^{Y R}=\frac{M C R W N_{0-17, M}^{Y R} \times \text { mcrwPrcntFem }}{1-\text { mcrwPrcntFem }}
\]

In order to distribute \(M C R W N_{0-17, F}^{Y R}\) among the ages 0 to 17 , we multiply \(M C R W N_{0-17, F}^{Y R}\) by the proportion of beneficiaries at each age in the last historical year.

\section*{Equation 3.3.6 - Minor Children of Deceased Workers (MCDWN )}

\section*{Minor Children of Deceased Workers}

Exposure: SSA population by age (0-17) and sex of the account holder
Linkages: \(p M C D W_{D E A}=\) probability that the parent is either retired or deceased \(p M C D W_{\text {FIA }}=\) probability that the primary account holder (PAH) is fully insured
\(p M C D W_{\text {RES }}=\) probability that a child who is eligible to receive minorchild benefits actually receive the benefits

We project the number of minor children of deceased workers, along with all linkage factors, by age of the minor ( \(a m=0\) to 17) and sex of the account holder ( \(s a=1\) for male, \(s a=2\) for female). Additionally we apply an adjustment factor (ImmEAFact(typ)) for the 2014 immigration executive order by beneficiary type (typ = 9: minor child of deceased father; 10 : minor child of deceased mother). We calculate it as follows:
\[
\begin{equation*}
M C D W N=M C D W_{P O P} \times p M C D W_{D E A} \times p M C D W_{F I A} \times p M C D W_{R E S} \times \operatorname{ImmEAFact}(t y p) \tag{3.3.6}
\end{equation*}
\]
\(M_{C D} W_{P O P}\) represents the subset of the population from which we draw these beneficiaries, and we set it equal to the Social Security area population (SSAPOP).
\[
M C D W_{P O P}=S S A P O P
\]
\(\boldsymbol{p} \boldsymbol{M C D} \boldsymbol{W}_{\text {DEA }}\) represents the status of the parent ( PAH ). This is set equal to the portion of the minor population where at least one parent is deceased. CHI_DEA represents the number of children having at least one deceased parent.
\[
p M C D W_{D E A}=\frac{C H I_{-} D E A}{M C D W_{P O P}}
\]
\(\boldsymbol{p} \boldsymbol{M C D} \boldsymbol{W}_{\text {FIA }}\) represents the probability that the parent (PAH) is fully insured. We set this equal to the portion of the population aged \(25+a m\) to \(35+a m\) where the PAH is fully insured (FI_PAH).
\[
p M C D W_{F I A}=\frac{\sum_{25+a m}^{35+a m}\left[S S A P O P \times F I_{-} P A H\right]}{\sum_{25+a m}^{35+a m} S S A P O P}
\]
\(\boldsymbol{p} \mathbf{M C D W}_{\text {RES }}\) represents the probability that a child who is eligible to receive minor-child benefits actually receive the benefits. For all historical years, we calculate \(p M C D W_{R E S}^{\text {year }}\) as
the ratio of MCDWN, the actual number of minor children of deceased workers receiving benefits, to the number of persons meeting all previously mentioned requirements by age and sex of the parent.
\[
p M C D W_{R E S}^{\text {year }}=\frac{M C D W N}{M C D W_{P O P} \times p M C D W_{D E A} \times p M C D W_{F I A}}, \quad \text { year }<\mathrm{TRYR}
\]

For each age and sex of parent, we use a least squares regression over the last ten years of historical data to determine a starting value in TR-1 for \(p M C D W_{\text {RES }}^{\text {year }}\) from which we project future values. We apply an adjustment, \(S R A D J^{M C D W}\), to the tenth year of the projection period in order to match the projections made by the Short-Range office. We linearly interpolate the values of \(p M C D W_{R E S}^{\text {year }}\) for intermediate years between the regressed values for \(p M C D W_{R E S}^{\text {TRYR- } 1}\) and \(p M C D W_{R E S}^{\text {TRYR }+9} \times S R A D J^{M C D W}\). After the \(10^{\text {th }}\) year of the projection period, we linearly grade the adjustment factors to one over the 10 years beyond the end of the short-range period, thus gradually eliminating the effect of the short-range adjustment factors, so that we ultimately return to the long-range projections.

\section*{Equations 3.3.7-8 - Student Children of Retired and Deceased Workers (SCRWN and SCDWN)}

\section*{Student Children of Retired Workers}

Exposure: SSA population by age of the student (18-21) and sex of the account holder
Linkages: \(p S C R W_{D E A}=\) probability that at least one parent is retired \(p S C R W_{A G A} \quad=\quad\) probability that the primary account holder \((\mathrm{PAH})\) is
age 62
\[
\begin{array}{ll} 
& \text { or older } \\
p S C R W_{F I A} & =\text { probability that the PAH is fully insured } \\
p S C R W_{C P A} & =\text { probability that the PAH is receiving benefits } \\
p S C R W_{S S B}= & \text { probability that the child is indeed attending school } \\
p S C R W_{\text {RES }}= & \text { probability that a child who is eligible to receive student- } \\
& \text { child benefits actually receives the benefits }
\end{array}
\]

\section*{Student Children of Deceased Workers}

Exposure: SSA population by age of the student (18-21) and sex of the account holder
Linkages: \(p S C D W_{D E A}=\) probability that at least one parent is deceased
\(p S C D W_{A G A}=\) probability that the PAH is age 62 or older (set to 1 )
\(p S C D W_{\text {FIA }}=\) probability that the PAH is fully insured
\(p S C D W_{C P A} \quad=\quad\) probability that the PAH is receiving benefits (set to 1 )
\(p S C D W_{\text {SSB }}=\) probability that the child is indeed attending school
\(p S C D W_{\text {RES }}=\) probability that a child who is eligible to receive studentchild benefits actually receives the benefits

We project the number of student children of retired and deceased workers, along with all linkage factors, by age of the student ( \(a s=18\) to 19 ) and sex of the account holder ( \(s a=1\) for male, \(s a=2\) for female). We calculate the projected number of student children of retired and deceased workers as follows:
\[
\begin{align*}
S C R W N & =S C R W_{P O P} \times p S C R W_{D E A} \times p S C R W_{A G A} \times p S C R W_{F I A} \\
& \times p S C R \square_{C P A} \times p S C R W_{S S B} \times p S C R W_{R E S}  \tag{3.3.7}\\
S C D W N= & S C D W_{P O P} \times p S C D W_{D E A} \times p S C D W_{A G A} \times p S C D W_{F I A} \\
\times & p S C D W_{C P A} \times p S C D W_{S S B} \times p S C D W_{R E S} \tag{3.3.8}
\end{align*}
\]
\(S_{C R} W_{P O P}\) and \(S_{C D} \boldsymbol{W}_{\mathbf{P O P}}\) represent the subset of the population from which these beneficiaries are drawn, and we set them equal to the Social Security area population (SSAPOP).
\[
S C R W_{P O P}=S C D W_{P O P}=S S A P O P
\]
\(\boldsymbol{p S C R} \boldsymbol{W}_{\text {DEA }}\) and \(\boldsymbol{p S C D W _ { D E A }}\) represent the status of the parent (PAH). For student children of retired workers, we set this equal to the proportion of the subset of the population where neither parents are deceased. For student children of deceased workers, we set this equal to the proportion of the subset of the population where at least one parent is deceased.
CHI_DEA represents the number of student children having at least one deceased parent.
\[
\begin{gathered}
p S C R W_{D E A}=1-\frac{C H I_{-} D E A}{S C R W_{P O P}} \\
p S C D W_{D E A}=\frac{C H I_{-} D E A}{S C D W_{P O P}}
\end{gathered}
\]
\(\boldsymbol{p S C R} \boldsymbol{W}_{\text {AGA }}\) and \(\boldsymbol{p S C D} \boldsymbol{W}_{\text {AGA }}\) represent the probability that the PAH is age 62 or older. For student children of retired workers, we set this equal to the proportion of the student population that has one parent age 62 or older, CHI _62+. For student children of deceased workers, we set the factor equal to one.
\[
\begin{gathered}
p S C R W_{A G A}=\frac{C H I_{-} 62+}{S C R W_{P O P}} \\
p S C D W_{A G A}=1
\end{gathered}
\]
\(p S C R W_{\text {FIA }}\) and \(\boldsymbol{p S C D W} \boldsymbol{W}_{\text {FIA }}\) represent the probability that the PAH is fully insured. For student children of retired workers, we set this equal to the portion of the population aged 62 to \(64+a s\) where the PAH is fully insured (FI_PAH). For student children of deceased workers, we calculate the factor similarly with the population being aged 25+as to 35+as.
\[
p S C R W_{F I A}=\frac{\sum_{62}^{64+a s}\left[S S A P O P \times F I_{-} P A H\right]}{\sum_{62}^{64+a s} S S A P O P}
\]
\[
p S C D W_{F I A}=\frac{\sum_{25+a S}^{35+a s}\left[S S A P O P \times F I_{-} P A H\right]}{\sum_{25+a s}^{35+a s} S S A P O P}
\]
\(\boldsymbol{p S C R} \boldsymbol{W}_{\boldsymbol{C P A}}\) and \(\boldsymbol{p S C D} \boldsymbol{W}_{\boldsymbol{C P A}}\) represent the probability that the PAH is receiving benefits. For student children of retired workers, we set this factor equal to the portion of the population aged 62 to \(64+\) as where the PAH is receiving benefits (RETIRED). For student children of deceased workers, we set this factor equal to one.
\[
\begin{gathered}
p S C R W_{C P A}=\frac{\sum_{62}^{64+a s}\left[S S A P O P \times F I_{-} P A H \times R E T I R E D\right]}{\sum_{62}^{64+a s}\left[S S A P O P \times F I_{-} P A H\right]} \\
p S C D W_{C P A}=1
\end{gathered}
\]
\(p S C R W_{\text {SSB }}\) and \(\boldsymbol{p S C D} W_{\text {SSB }}\) represent the probability that the child is indeed attending school (full-time elementary or secondary school). This factor is dependent upon the age of the child, and we calculate it as follows.
\[
p S C R W_{S S B}=p S C D W_{S S B}= \begin{cases}\frac{1}{a s-16}, & \text { year } \leq 1981 \\ \frac{0.5}{a s-16}, & \text { year }>1981\end{cases}
\]
\(\boldsymbol{p S C R} \boldsymbol{W}_{\text {RES }}\) and \(\boldsymbol{p S C D} \boldsymbol{W}_{\text {RES }}\) represent the probability that a child who is eligible to receive student-child benefits actually receive the benefits. For all historical years, we calculate \(p S C R W_{R E S}^{\text {year }}\) and \(p S C D W_{R E S}^{\text {year }}\) as the ratio of SCRWN and SCDWN, the actual number of student children receiving benefits, to the number of persons meeting all previously mentioned requirements by age and sex of the parent.
\[
\begin{aligned}
& p S C R W_{R E S}^{\text {year }}=\frac{S C R W N}{S C R W_{P O P} \times p S C R W_{D E A} \times p S C R W_{A G A} \times p S C R W_{F I A} \times p S C R W_{C P A} \times p S C R W_{S S B}}, y e a r<\mathrm{TRYR} \\
& p S C D W_{R E S}^{y e a r}=\frac{S C D W N}{S C D W_{P O P} \times p S C D W_{D E A} \times p S C D W_{A G A} \times p S C D W_{F I A} \times p S C D W_{C P A} \times p S C D W_{S S B}}, \text { year }<\mathrm{TRYR}
\end{aligned}
\]

For each age and sex of parent, we use a least squares regression over the last ten years of historical data to determine a starting value in TR-1 for \(p S C R W_{R E S}^{\text {year }}\) from which we project future values. We apply an adjustment, \(S R A D J^{S C R W}\), to the tenth year of the projection period in order to match the projections made by the Short-Range office. We linearly interpolate the values of \(p S C R W_{R E S}^{\text {year }}\) for intermediate years between the regressed values for \(p S C R W_{R E S}^{\text {TRYR-1 }}\) and \(p S C R W_{R E S}^{\text {TRYR+9 }} \times S R A D J^{S C R W}\). After the \(10^{\text {th }}\) year of the projection period, we linearly grade the adjustment factors to one over the 10 years beyond the end of the short-
range period, thus gradually eliminating the effect of the short-range adjustment factors, so that we ultimately return to the long-range projections. We calculate the values of \(p S C D W_{R E S}^{\text {year }}\) similarly.

Equations 3.3.9-10 - Disabled Adult Children of Retired and Deceased Workers (DCRWN and DCDWN)

\section*{Disabled Adult Children of Retired Workers}

Exposure: SSA population by age of the adult child (18-95)
Linkages: \(p D C R W_{A G A}=\) probability that the primary account holder (PAH) is age 62 or older
\(p D C R W_{\text {DEA }}=\) probability that the parent is retired
\(p D C R W_{\text {FIA }}=\) probability that the PAH is fully insured
\(p D C R W_{C P A}=\) probability that the PAH is receiving benefits
\(p D C R W_{S S B} \quad=\quad\) probability that the child is indeed disabled
\(p D C R W_{\text {RES }} \quad=\quad\) probability that a child who is eligible to receive
disabled-
child benefits actually receive benefits
Disabled Adult Children of Deceased Workers
Exposure: SSA population by age of the adult child (18-95)
Linkages: \(p D C D W_{A G A}=\) probability that the PAH is age 62 or older (set to 1 )
\(p D C D W_{D E A}=\) probability that the parent is deceased
\(p D C D W_{F I A}=\) probability that the PAH is fully insured
\(p D C D W_{C P A}=\) probability that the PAH is receiving benefits (set to 1 )
\(p D C D W_{S S B}=\) probability that the child is indeed disabled
\(p D C D W_{\text {RES }}=\) probability that a child who is eligible to receive disabledchild benefits actually receive the benefits

We project the number of disabled adult children of retired and deceased workers, along with all linkage factors, by age group of the disabled adult child ( \(a d=1-10\) ) and sex of the account holder ( \(s a=1\) for male, \(s a=2\) for female). The age groups are 18-19, 20-24, \(\ldots\), 5559, 60+. Additionally we apply an adjustment factor (ImmEAFact(typ)) for the 2014 immigration executive order by beneficiary type (typ \(=7\) : disabled adult child of retired male worker; 8: disabled adult child of female retired worker; 11: disabled adult child of deceased male worker; 12: disabled adult child of deceased female worker). We calculate the projected number of disabled adult children of retired and deceased workers as follows:
\[
\begin{align*}
& \text { DCRWN }=\mathrm{DCRW}_{\text {POP }} \times \mathrm{pDCRW}_{\text {AGA }} \times \mathrm{pDCRW} \mathrm{DEA} \times \mathrm{pDCRW}_{\text {FIA }} \\
& \times p D C R W_{C P A} \times p D C R W_{S S B} \times p D C R W_{\text {RES }} \times \operatorname{ImmEAFact}(t y p)  \tag{3.3.9}\\
& D C D W N=D C D W_{P O P} \times p D C D W_{A G A} \times p D C D W_{D E A} \times p D C D W_{F I A} \\
& \times p D C D W_{C P A} \times p D C D W_{S S B} \times p D C D W_{R E S} \times \operatorname{ImmEAFact}(t y p) \tag{3.3.10}
\end{align*}
\]

We calculate all factors similarly to those for student children with the exception of the following.
\(\boldsymbol{p D C R} \boldsymbol{W}_{\text {DEA }}\) is set equal to the proportion of the married and widowed population who are married (for ages of the parent that are reasonable based on the given age range of the disabled child). We calculate \(p D C D W_{D E A}\) similarly for disabled children of deceased workers.
\[
\begin{gathered}
p D C R W_{D E A}= \begin{cases}\frac{S S A P O P_{\text {mar }}}{S S A P O P_{\text {mar }}+S S A P O P_{\text {wid }}}, & a d=1-9 \\
\frac{0.25 \times S S A P O P_{\text {mar }}}{S S A P O P_{\text {mar }}+S S A P O P_{\text {wid }}}, & a d=10\end{cases} \\
p D C D W_{D E A}= \begin{cases}\frac{S S A P O P_{\text {wid }}}{S S A P O P_{\text {mar }}+S S A P O P_{\text {wid }}}, & a d=1-9 \\
\frac{0.25 \times S S A P O P_{\text {wid }}}{S S A P O P_{\text {mar }}+S S A P O P_{\text {wid }}}+0.75, & a d=10\end{cases}
\end{gathered}
\]
\(\boldsymbol{p D C R} \boldsymbol{W}_{\text {SSB }}\) and \(\boldsymbol{p D C D} \boldsymbol{W}_{\text {SSB }}\) represent the probability that the adult child is indeed disabled. DCPREM is the preliminary calculation of this factor and we assume it to remain constant. For the projection period, for ad=6-10, we set \(p D C R W_{S S B}\) and \(p D C D W_{S S B}\) equal to the preliminary factor, plus an adjustment which accounts for the year.
\[
\begin{aligned}
& \text { DCPREM }= \begin{cases}0.012, & a d=1-2 \\
0.009, & a d=3 \\
0.007, & a d=4 \\
0.006, & a d=5 \\
0.005, & a d=6 \\
0.004, & a d=7-10\end{cases} \\
& p D C R W_{S S B}=p D C D W_{S S B}= \begin{cases}\min [0.005, D C P R E M+0.0001 \times(\text { year }-T R Y R)], & \text { ad }=7-10 \text { and } \\
\text { year }>T R Y R+1 \\
\text { DCPREM, }, & \text { elsewhere }\end{cases}
\end{aligned}
\]
\(\boldsymbol{p D C R} W_{\text {RES }}\) and \(\boldsymbol{p D C D} \boldsymbol{W}_{\text {RES }}\) represent the probability that a child who is eligible to receive disabled-child benefits actually receive the benefits. For all historical years, we calculate \(p D C R W_{R E S}^{\text {year }}\) and \(p D C D W_{R E S}^{\text {year }}\) as the ratio of \(D C R W N\) and \(D C D W N\), the actual number of disabled children receiving benefits, to the number of persons meeting all previously
mentioned requirements by age and sex of the parent.
For year < TRYR:
\[
\begin{aligned}
& p D C R W_{R E S}^{\text {year }}=\frac{D C R W N}{D C R W_{P O P} \times p D C R W_{D E A} \times p D C R W_{A G A} \times p D C R W_{F I A} \times p D C R W_{C P A} \times p D C R W_{S S B}} \\
& p D C D W_{R E S}^{\text {year }}=\frac{D C D W N}{D C D W_{P O P} \times p D C D W_{D E A} \times p D C D W_{A G A} \times p D C D W_{F I A} \times p D C D W_{C P A} \times p D C D W_{S S B}}
\end{aligned}
\]

We apply an adjustment, \(S R A D J^{D C R W}\), to the tenth year of the projection period in order to match the projections made by the Short-Range office. We linearly interpolate the values of \(p D C R W_{R E S}^{\text {year }}\) for intermediate years between \(p D C R W_{R E S}^{\text {TRYR-1 }}\) and \(p D C R W_{R E S}^{\text {TRY }+9} \times S R A D J^{D C R W}\). After the \(10^{\text {th }}\) year of the projection period, we linearly grade the adjustment factors to one over the 10 years beyond the end of the short-range period, thus gradually eliminating the effect of the short-range adjustment factors, so that we ultimately return to the long-range projections. We calculate the values of \(p D C D W_{\text {RES }}^{\text {year }}\) similarly.

Equations 3.3.11-12 - Young Spouses of Retired and Deceased Workers (YSRWN and YSDWN)
Young Spouses of Retired Workers
Exposure: SSA population by age (15-69), sex of account holder, and marital status (married)
Linkages: \(p Y S R W_{A G A}=\) probability that the primary account holder \((\mathrm{PAH})\) is of the required age.
\(p Y S R W_{E C B}=\) probability that the young spouse has an entitled child in their care
\(p Y S R W_{F S B}=\) probability that the young spouse is not already receiving benefits based on another child in their care \(p Y S R W_{\text {RES }}=\) probability that a person who is eligible to receive youngspouse benefits actually receive the benefits

Young Spouses of Deceased Workers
Exposure: SSA population by age (15-69), sex of the account holder and marital status (widowed and divorced)
Linkages: \(p Y S D W_{D E A}=\) probability that the PAH is deceased \(p Y S D W_{E C B}=\) probability that the young spouse has an entitled child in their care
\(p Y S D W_{F S B}=\) probability that the young spouse is not already receiving benefits based on another child in their care \(p Y S D W_{R M B}=\) probability that the young spouse is not remarried \(p Y S D W_{\text {RES }}=\) probability that a person who is eligible to receive youngspouse benefits actually receive the benefits

We project the number of young spouses of retired and deceased-workers, along with all linkage factors, by age group ( \(a b=1-10\) ) of the young spouse and sex of the account holder ( \(s a=1\) for male, \(s a=2\) for female). We also project young spouses of deceased workers by marital status of the young spouse ( \(m b=1\) for widowed and \(m b=2\) for divorced). For the young spouses of deceased worker we apply an additional adjustment factor
(ImmEAFact(typ)) for the 2014 immigration executive order by beneficiary type (typ = 13: male PAH; 14: female PAH). The age groups are under 25, 25-29 ..., 65-69. We calculate the projected number of young spouses of retired and deceased-workers as follows:
\[
\begin{gather*}
Y S R W N=Y S R W_{P O P} \times p Y S R W_{A G A} \times p Y S R W_{E C B} \times p Y S R W_{F S B} \times p Y S R W_{R E S}  \tag{3.3.11}\\
Y S D W N=Y S D W_{P O P} \times p Y S D W_{D E A} \times p Y S D W_{E C B} \times p Y S D W_{F S B} \\
\times p Y S D W_{R M B} \times p Y S D W_{R E S} \times \operatorname{ImmEAFact}(t y p) \tag{3.3.12}
\end{gather*}
\]

YSR \(_{P O P}\) and YSDW \(_{\text {POP }}\) represent the subset of the population from which we draw these beneficiaries. We set \(Y S R W_{P O P}\) equal to the married Social Security area population \(\left(S S A P O P_{m a r}\right)\) and we set \(Y S D W_{P O P}\) equal to \(S S A P O P_{m b}\) for \(m b=1-2\).
\[
\begin{aligned}
& Y S R W_{P O P}=S S A P O P_{m a r} \\
& Y S D W_{P O P}=S S A P O P_{m b}
\end{aligned}
\]
\(\boldsymbol{p Y S R} \boldsymbol{W}_{A G A}\) represents the probability that the PAH is of the required age. We set \(p Y S R W_{A G A}\) equal to the portion of the married population who has an aged spouse (AGSP).
\[
p Y S R W_{A G A}=\frac{A G S P}{Y S R W_{P O P}}
\]
\(\boldsymbol{p Y S D} \boldsymbol{W}_{\text {DEA }}\) represents the probability that the PAH is deceased. For mb=1, we do not apply any factor. For \(\mathrm{mb}=2\), we set this factor equal to the portion of young spouses that is widowed.
\[
p Y S D W_{D E A}= \begin{cases}1, & m b=1(\text { widowed }) \\ \frac{S S A P O P_{\text {wid }}}{S S A P O P_{\text {wid }}+S S A P O P_{\text {mar }}}, & m b=2(\text { divorced })\end{cases}
\]
\(\boldsymbol{p Y S R} \boldsymbol{W}_{\boldsymbol{E C B}}\) and \(\boldsymbol{p Y S D} \boldsymbol{W}_{\boldsymbol{E C B}}\) represent the probability that the young spouse has an entitled child in their care. We set \(p Y S R W_{E C B}\) equal to the portion of persons meeting the previously mentioned requirements who have a minor or disabled adult child in their care. We set \(p Y S D W_{E C B}\),by marital status, equal to the portion of persons meeting the previously mentioned requirements who have a minor or disabled adult child in their care. MCRWN \({ }^{a b}\) and \(D C R W N^{a b}\) represent the total number of minor and disabled adult children of retired workers where the other parent (young spouse) is in the age bracket ab.
\[
\begin{gathered}
p Y S R W_{E C B}^{a b}=\frac{M C R W N^{a b}+D C R W N^{a b}}{Y S R W_{P O T} \times p Y S R W_{A G A}} \\
p Y S D W_{E C B}^{m b}=\frac{\left(M C D W N^{a b}+D C D W N^{a b}\right) \times\left[\frac{Y S D W_{P O P}^{m b} \times p Y S D W_{D E A}^{m b} \times p Y S D W_{A G A}^{m b}}{Y S D W_{P O P}^{t o t a l} \times p Y S D W_{D E A}^{t o t a l} \times p Y S D W_{A G A}^{\text {total }}}\right]}{Y S D W_{P O P} \times p Y S D W_{D E A}}
\end{gathered}
\]
\(\boldsymbol{p Y S R} \boldsymbol{W}_{\boldsymbol{F S B}}\) and \(\boldsymbol{p Y S D} \boldsymbol{W}_{\boldsymbol{F S B}}\) represent the probability that the young spouse is not already receiving benefits based on another child in their care. We set this factor equal to one divided by the number of children in the average family \(\left(A S O F_{a b}\right)\) for the given age bracket of the spouse. For young spouses of retired workers, we do not apply a factor for sa=2.
\[
\begin{gathered}
p Y S R W_{F S B}= \begin{cases}\frac{1}{A S O F_{a b}}, & s a=1 \\
1, & s a=2\end{cases} \\
p Y S D W_{F S B}=\frac{1}{A S O F_{a b}}
\end{gathered}
\]
\(\boldsymbol{p Y S D} \boldsymbol{W}_{\boldsymbol{R M B}}\) represents the probability that the spouse is not remarried. We assume this factor remains constant at 0.600 .
\[
p Y S D W_{R M B}=0.600
\]
\(\mathbf{p Y S R}_{\text {RES }}\) and \(\mathbf{p Y S D W}\) RES represent the probability that a person who is eligible to receive young-spouse benefits actually receive the benefits. For all historical years, we calculate \(p Y S R W_{R E S}^{\text {year }}\) as the ratio of YSRWN, the actual number of young spouses of retired workers receiving benefits, to the number of persons meeting all previously mentioned requirements by age, sex, and marital status. We calculate \(p Y S D W_{R E S}^{\text {year }}\) similarly, using the number of young spouses of deceased workers.
\[
\begin{gathered}
p Y S R W_{R E S}=\frac{Y S R W N}{Y S R W_{P O P} \times p Y S R W_{A G A} \times p Y S R W_{E C B} \times p Y S R W_{F S B}}, \quad \text { year }<T R Y R \\
p Y S D W_{R E S}=\frac{Y S D W N}{Y S D W_{P O P} \times p Y S D W_{D E A} \times p Y S D W_{E C B} \times p Y S D W_{F S B} \times p Y S D W_{R M B}}, \text { year }<T R Y R
\end{gathered}
\]

For each age, sex , and marital status, we use a least squares regression over the last ten years
of historical data to determine a starting value in TR-1 for \(p Y S R W_{R E S}^{\text {year }}\). In addition, for each sex and marital status we graduate the regressed values of \(p Y S R W_{R E S}^{T R Y R-1}\) over age using a weighted minimized third-difference formula to compute ESTRES \({ }^{\text {YSRW }}\). ESTRES \({ }^{\text {YSRW }}\) are the preliminary estimates of \(p Y S R W_{R E S}^{\text {TRYR }+9}\), the values in the tenth year of the projection period. For female young spouses, we apply an adjustment, \(S R A D J^{\text {YSRW }}\), to the tenth year of the projection period in order to match the projections made by the Short-Range office. We exponentially interpolate the values of \(p Y S R W_{R E S}^{\text {year }}\) for intermediate years between \(p Y S R W_{R E S}^{\text {TRYR-1 }}\) and \(p Y S R W_{R E S}^{\text {TRYR+9 }}\) (equal to \(E S T R E S^{\text {YSRW }} \times S R A D J^{\text {YSRW }}\) ). After the \(10^{\text {th }}\) year of the projection period, we linearly grade the adjustment factors to one over the 10 years beyond the end of the short-range period, thus gradually eliminating the effect of the shortrange adjustment factors, so that we ultimately return to the long-range projections. We calculate the values of \(p Y S D W_{R E S}^{\text {year }}\) similarly.

\section*{Equation 3.3.13 - Number of Deaths of Insured Workers (LUMSUM \({ }_{a b}\) )}

We project the number of deaths of insured workers by sex and 5-year age group ( \(a b=1-14\) ). Age groups include 20-24, 25-29,...,80-84, 85+. We calculate EXPOSURE \({ }_{a b}\), the estimated number of lump-sum payments paid during the year for age group \(a b\), as the number of total deaths during the year times the probability that the deceased was fully insured and has a surviving spouse or child. We calculate BASE as the ratio of the actual total amount of lump-sum death payments paid in TRYR-1 to the estimated total amount of lump-sum payments paid in TRYR-1. We then calculate \(L U M S U M_{a b}\) for each year in the projection period.
\[
\begin{equation*}
\operatorname{LUMSUM}_{a b}=E X P O S U R E_{a b} \times B A S E \tag{3.3.13}
\end{equation*}
\]

\section*{Appendix 3.3-1: Glossary}
\(A B\) : age group of the beneficiary
\(A D\) : age of the disabled child
AGSP: married population where at least one spouse is age 62 or older
AM: age of the minor child
ARFGT3: actuarial reduction factor for ages more than 3 years below normal retirement age
ARFLE3: actuarial reduction factor for ages less than 3 years below normal retirement age
AS: age of the student child
ASDW: aged spouse of deceased worker by linkage factor, age (60-95+), sex of the account holder, marital status (widowed, divorced) and insured status (insured, uninsured). Linkage factors are:
\(A S D W_{P O P}\) : population of potential aged spouse of retired workers
\(p A S D W_{\text {DEA }}\) : probability that the primary account holder (PAH) is deceased
\(p A S D W_{F I A}\) : probability that the PAH was fully insured at death
\(p A S D W_{\text {MBB }}\) : probability that the widow(er) is not receiving a young-spouse benefit for
the care of a child
\(p A S D W_{F I B}\) : probability that the aged widow(er) is fully insured
\(p A S D W_{G P B}\) : probability that the aged-widow(er)'s benefits are not withheld or offset
totally because of receipt of a significant government pension based on earnings in noncovered employment
\(p A S D W_{\text {RES }}\) : probability that a widow(er) eligible to receive his/her own retired-worker benefits would instead apply for and receive widow(er) benefits
ASDWN: final number of aged spouse of deceased workers (product of all linkage factors)
ASOF: average number of children in a family, by age group (under 25, 25-29 ..., 60-64)
ASRW: aged spouse of retired worker by linkage factor, age (62-95+), sex of the account holder, and marital status of the beneficiary (married, divorced). Linkage factors are:
\(A S R W_{P O P}\) : population of potential aged spouse of retired worker beneficiaries
\(p A S R W_{\text {DEA }}\) : probability that the primary account holder ( PAH ) is not deceased
\(p A S R W_{A G A}\) : probability that the PAH is of the required age
\(p A S R W_{F I A}\) : probability that the PAH is fully insured
\(p A S R W_{\text {CPA }}\) : probability that the PAH is receiving benefits
pASRW \({ }_{\text {MBB }}\) : probability that the beneficiary is not receiving a young-spouse benefit \(p A S R W_{\text {FIB }}\) : probability that the aged spouse is not fully insured
\(p A S R W_{G P B}\) : probability that the aged-spouse's benefits are not withheld because of
receipt of a significant government pension based on earnings in noncovered employment
\(p A S R W_{F S 1:} \quad\) probability that a couple will engage in a filing strategy in which the PAH files for and suspends their benefit and an uninsured spouse receives an aged spouse benefit
\(p A S R W_{F S 2:} \quad\) probability that a couple will engage in a filing strategy in which the PAH files for and may or may not suspend their benefit and the fully insured for their aged spouse benefit.
\(p A S R W_{\text {RES }}\) : probability that a person who is eligible to receive aged-spouse benefits actually receive the benefits
ASRWN: final number of aged spouse of retired workers (product of all linkage factors)
AH: age of husband
AW: age of wife
BASE: ratio of actual to estimated total amount of lump-sum death payments paid in TRYR-1
BBA: Bipartisan Budget Act of 2015. A provision of this act closes unintended loopholes by eliminating (1) the ability to receive only a retired-worker benefit or an aged-spouse benefit when eligible for both, for those attaining age 62 in 2016 and later, and (2) the ability of a family member other than a divorced spouse to receive a benefit based on the earnings of a worker with a voluntarily suspended benefit, for voluntary suspensions requested after April 29, 2016.
CHI_62+: number of children having at least one parent aged 62 or older
CHI_DEA: number of children having at least one deceased parent
CON: number of persons converted from disabled-worker beneficiaries
DCDW: disabled child of deceased workers by linkage factor, age group of the child (18-19, 20\(24, \ldots, 55-59,60+\) ) and sex of the account holder. Linkage factors are same as SCDW.
\(D C D W_{P O P}\) : population of potential disabled children
\(p D C D W_{A G A}\) : probability that the PAH is age 62 or older
\(p D C D W_{D E A}\) : probability that the parent is either retired or deceased
\(p D C D W_{F I A}\) : probability that the PAH is fully insured
\(p D C D W_{\text {CPA }}\) : probability that the PAH is receiving benefits
\(p D C D W_{\text {SBB }}\) : probability that the child is indeed disabled
\(p D C D W_{\text {RES }}\) : probability that a child who is eligible to receive disabled-child benefits actually receive the benefits
DCDWN: final number of disabled children of deceased workers (product of all linkage factors)
DCPREM: preliminary calculation of the probability that a child is disabled, by age
DCRW: disabled child of retired workers by linkage factor, age group of the child (18-19, 20-
\(24, \ldots, 55-59,60+\) ) and sex of the account holder. Linkage factors are same as those for DCDW.
DCRWN: final number of disabled children of retired workers (product of all linkage factors)
DIB: number of disabled-worker beneficiaries
DIFFADJ: adjustment that accounts for the difference between the actual and estimated prevalence rate at each age in the most recent historical years
DISPREV: disability prevalence rate by age and sex
DRC: delayed retirement credit
DSDW: disabled spouse of deceased worker by linkage factor, age (50-69), sex of the account holder, and marital status (widowed, divorced). Linkage factors are:
\(D S D W_{P O P}\) : population of potential beneficiaries
\(p D S D W_{\text {DEA }}\) : probability that the primary account holder (PAH) is deceased
\(p D S D W_{\text {FIA }}\) : probability that the PAH was fully insured at death
\(p D S D W_{\text {SSB }}\) : probability that the spouse is indeed disabled
\(p D S D W_{D E B}\) : probability that the disabled spouse is not receiving another type of benefit
\(p D S D W_{\text {RES }}\) : probability that a person who is eligible to receive disabled-spouse
benefits actually receive the benefits
DSDWN: final number of disabled spouse of deceased workers (product of all linkage factors)
ERROR: actual prevalence rate minus the regressed prevalence rate in the most recent historical year
ESTPR: preliminary estimate of the prevalence rate for retired workers
ESTRES: preliminary estimate of the RES factor for the tenth year of the projection period
EXPOSURE: estimated number of lump-sum payments by age group (20-24, 25-29,...,80-84,
85+)
FACTOR: adjustment for calculation of MBB factor of aged spouse of deceased worker
FINS: portion of the SSA population that is fully insured
FI_PAH: portion of married population where one spouse is fully insured
FP: status of the parent (retired, deceased)
GPOAGE: portion, by age, of the total beneficiaries expected to receive a significant government pension
GPWHLD: total number of beneficiaries (for all ages) expected to receive a significant government pension
IMMEAFACT: Adjustment factors applied to auxiliary beneficiary categories due to the 2014
immigration executive action by sex for years 2016-2095. Factors are computed for aged spouses of deceased workers (widow(er)s), aged spouses of retired workers, disabled adult children of retired workers, minor children of deceased workers, disabled adult children of deceased workers, and young spouses of deceased workers.
\(I N\) : insured status of the beneficiary
LFPR: labor force participation rates for age 62, by sex
LUMSUM: number of deaths of insured workers by sex and age group (20-24,...,80-84,85+)
MAR62PLUS: number of couples where both husband and wife are age 62 and over
MS: marital status of the primary account holder
MB: marital status of the beneficiary
MBAPIA: ratio of the monthly benefit amount (MBA) to the primary insurance amount (PIA) by age (62-70) and sex
MCDW: minor children of deceased workers by linkage factor, age of the child (0-17) and sex of the account holder. Linkage factors are:
\(M C D W_{\text {POP }}\) : population of potential minor children
\(p M C D W_{\text {DEA }}\) : probability that the parent is either retired or deceased
\(p M C D W_{F I A}\) : probability that the PAH is fully insured
\(p M C D W_{R E S}\) : probability that a child who is eligible to receive minor-child benefits
actually receive the benefits
MCDWN: final number of minor children of deceased workers (product of all linkage factors) MCRW: minor children of retired workers by linkage factor, age of the child ( \(0-17\) ) and sex of the account holder.
MCRWN: final number of minor children of retired workers (product of all linkage factors)
NRA: normal retirement age
PAH: primary account holder
REGPR: regressed prevalence rate for retired workers
RETIRED: number of retired workers receiving benefits
\(\boldsymbol{R W}\) : retired workers by linkage factor, age (62-95+), sex, and marital status (single, married, widowed, divorced). Linkage factors are:
\(R W_{P O P}\) : population of potential retired-worker beneficiaries
\(p R W_{\text {FIA }}\) : probability that the primary account holder (PAH) is insured
\(p R W_{\text {DBB }}\) : probability that the PAH is not receiving a disabled-worker benefit
\(p R W_{\text {WBB }}\) : probability that the PAH is not receiving a widow(er) benefit
\(p R W_{\text {RES }}\) : retirement prevalence rate; probability that a fully insured worker (not receiving disability or widow(er)'s benefits) would receive a retiredworker benefit
\(\boldsymbol{R W N}\) : final number of retired workers (product of all linkage factors)
SA: sex of the account holder
SCDW: student children of deceased workers by linkage factor, age of the student (18-21) and sex of the account holder. Linkage factors are:
\(S_{\text {SOP }}\) : population of potential student children
\(p S C D W_{D E A}\) : probability that the parent is either retired or deceased
\(p S C D W_{A G A}\) : probability that the PAH is age 62 or older
\(p S C D W_{\text {FIA }}\) : probability that the PAH is fully insured
\(p S C D W_{C P A}\) : probability that the PAH is receiving benefits
\(p S C D W_{\text {SSB }}\) : probability that the child is indeed attending school
\(p S C D W_{\text {RES }}\) : probability that a child who is eligible to receive student-child benefits actually receive the benefits
SCDWN: final number of student children of deceased workers (product of all linkage factors)
SCRW: student children of retired workers by linkage factor, age of the student (18-21) and sex of the account holder. Linkage factors are same as SCDW.
SCRWN: final number of student children of retired workers (product of all linkage factors)
SRADJ: adjustment to match short-range projections in \(10^{\text {th }}\) year of projection period
SSAPOP: Social Security area population by age (0:100), sex, and marital status (single, married, widowed, divorced)
\(\boldsymbol{S X}\) : sex of the beneficiary
TRYR: first year of the projection period
WEIGHT: estimated probability applied to each possible age of the spouse, given the age of the primary account holder
YSDW: young spouse of deceased worker by linkage factor, age group (under 25, 25-29,...,6569), sex of the account holder and marital status (widowed, divorced). Linkage factors are:
\(Y S D W_{P O P}\) : population of potential young spouse of deceased workers \(p Y S D W_{D E A}\) : probability that the primary account holder ( PAH ) is of the required age \(p Y S D W_{E C B}\) : probability that the young spouse has an entitled child in their care \(p Y S D W_{F S B}\) : probability that the young spouse is not already receiving benefits based on another child in their care
\(p Y S D W_{\text {RMB }}\) : probability that the young spouse is not remarried
\(p Y S D W_{\text {RES }}\) : probability that a person who is eligible to receive young-spouse benefits actually receive the benefits
YSDWN: final number of young spouse of deceased workers (product of all linkage factors) YSRW: young spouse of retired worker by linkage factor, age group (under 25, 25-29,...,65-69) and sex of the account holder. Linkage factors are:

YSR \(W_{P O P}\) : population of potential young spouse of retired workers
\(p Y S R W_{A G A}\) : probability that the primary account holder (PAH) is of the required age
\(p Y S R W_{\text {ECB }}\) : probability that the young spouse has an entitled child in their care
\(p Y S R W_{F S B}\) : probability that the young spouse is not already receiving benefits based on another child in their care
\(p Y S R W_{\text {RES }}\) : probability that a person who is eligible to receive young-spouse benefits actually receive the benefits
YSRWN: final number of young spouse of retired workers (product of all linkage factors)
\[
\begin{gathered}
\text { Process 4: } \\
\text { Trust Fund } \\
\text { Operations \& } \\
\text { Actuarial Status }
\end{gathered}
\]

\section*{4. Trust Fund Operations and Actuarial Status}

OCACT uses the Trust Fund Operations and Actuarial Status Process to project (1) the annual flow of income from payroll taxes, taxation of benefits, and interest on assets in the trust fund and (2) the annual flow of cost from benefit payments, administration of the program, and railroad interchange. The annual flows are projected for each year of the 75-year projection period. In addition, this subprocess produces annual and summarized values to help access the financial status of the Social Security program.

The Trust Fund Operations and Actuarial Status Process is composed of three subprocesses: TAXATION OF BENEFITS, AWARDS, and COST. As a rough overview, TAXATION OF BENEFITS projects, for each year during the 75-year projection period, the amount of income from taxation of benefits as a percent of benefits paid. AWARDS projects information needed to determine the benefit levels of newly awarded retired workers and disabled workers by age and sex. COST uses information from the AWARDS and TAXATION OF BENEFITS subprocesses, as well as information from other processes, to project the annual flow of income and cost to the trust funds. In addition, COST produces annual and summarized measures of the financial status of the Social Security program.

\subsection*{4.1. TAXATION OF BENEFITS}

\section*{4.1.a. Overview}

The 1983 Social Security Act specifies including up to 50 percent of the Social Security benefits to tax return filer's adjusted gross income (AGI) for tax liability if a tax return filer's adjusted gross income plus half of his (or her) Social Security benefits is above the specified income threshold amount of \(\$ 25,000\) as a single filer (or \(\$ 32,000\) as a joint filer). Subsequently, the 1993 OBRA (Omnibus Budget Reconciliation Act) provided for taxation of up to 85 percent if a tax return filer's adjusted gross income plus half of his (or her) Social Security benefits is above the specified income threshold amount of \(\$ 34,000\) as a single filer (or \(\$ 44,000\) as a joint filer).

The proceeds from taxing up to 50 percent of the OASDI benefits, as a result of the 1983 Act, are credited to the OASI and DI Trust Funds, while additional taxes on the OASDI benefits, as a result of the 1993 Act, are credited to the HI Trust Fund.

Income to the Trust Funds from such taxation is estimated by using ratios of taxes on benefits to benefits for the OASI and DI programs separately. These ratios, called "RTBs", are applied to projected OASI and DI benefit amounts to estimate tax revenues to the OASI and DI Trust Funds.

For the short range period (first 10 years of the projection), the Cost sub-process (4.3) uses OTA (Office of Tax Analysis)'s projected estimates for (1) the percent of benefits taxable and (2) the average marginal tax rates applicable to those taxable OASI and DI benefits. The multiplication of (1) and (2) produces projected RTBs under the 1983 Act (up to 50 percent of benefits taxable) and the 1993 OBRA (additional up to 35 percent of benefits taxable).

For the long range period ( \(11^{\text {th }}\) through \(75^{\text {th }}\) year of the projection period), the RTB ratios for OASI benefits and those for DI benefits under the 1983 Act and the combined \(1983+\) 1993 Act are computed with the following formula for each projection year.

RTB(yr) \(=\) RTB(tryr+9) \(*\{\operatorname{AWI}(t r y r+9) / A W I(y r)\} \wedge P+\) RTB(ultimate)*\{1- AWI(tryr+9)/AWI(yr) \({ }^{\wedge}{ }^{\wedge}\),
where

> tryr = first year of the projection period (year of the Trustees Report)

RTB(ultimate) = ratio of taxes on benefits to benefits assuming income thresholds equal zero.

AWI \(=\) SSA average wage index series
\(\mathrm{P}=\) exponential parameter for a trend curve line.
Finally, the Cost sub-process (4.3) applies the projected RTB ratios to its own estimates
of the projected OASI and DI benefit payments to produce taxation of benefit revenues to the OASI and DI Trust Funds.

\section*{4.1.b. Input Data}

OCACT Data
Economics--process 2
- Projected SSA wage index series by year, updated yearly
- Projected COLAs and average wage index levels under the intermediate assumptions of the prior Trustees Report

Beneficiaries--process 3
- Projected OASI beneficiaries by age and sex under the intermediate assumptions of the prior Trustees Report
- Projected DI beneficiaries by age and sex under the intermediate assumptions of the prior Trustees Report

\section*{Trust Fund Operations--process 4}
- Aggregate OASI and DI benefit ratios (as a ratio of total OASDI benefits) under the intermediate assumptions of the prior Trustees Report

Other
- Aggregate OASDI benefit payments for calendar year ni-3 (ni = initial projection year, e.g, 2013 for the 2016 Trustees Report)

Other input Data
- OTA's projected percent of benefits taxable and average marginal tax rates by type of benefit (OASI and DI) for the short range period (updated yearly).
- OTA's ultimate ratios of taxes on benefits to benefits (i.e., with income thresholds, assumed equal to 0). Such ultimate ratios are provided on a combined OASDI benefit basis, and are expected to be updated annually based on OTA's update.
- Current Population Survey (CPS) data for year ni-3
- Personal income tax brackets for tax year ni-3
- Marginal income tax rates for tax year ni-3
- General filing requirement amounts (standard deduction amounts) for personal income tax purposes for tax year ni-3
- Ratios of taxable income to adjusted gross income by income level (IRS data) for tax year ni-3
- Treasury's aggregate taxable benefit amount (IRS data) for tax year ni-3
- OTA's estimated taxes on benefits for the OASDI and HI Trust Funds for tax year ni-3

\section*{4.1.c. Development of Output}

For the short range period, the Cost sub-process (4.3) uses OTA's projected RTBs for OASI and DI benefits under the 1983 Act, to project taxation of benefit revenues to the OASI and DI Trust Funds.

For the long range period, formula 4.1.1 computes projected ratios of taxes on OASI benefits to OASI benefits and projected ratios of taxes on DI benefits to DI benefits under the 1983 Act (up to 50 percent of benefits taxable). This formula essentially provides more weight to the ultimate RTB ratios as time progresses, using the ratio of AWI ( \(10^{\text {th }}\) year) to AWI (projection year) as the "weight." Additionally, an exponential parameter P value to the AWI "weights" is set judgmentally such that the estimate continues the short range trend into the transitional \(11^{\text {th }}\) through \(20^{\text {th }}\) projection years before it approaches the ultimate RTB ratio. For the RTB ratios for up to 50 percent of benefits taxable, the P values were set at 0.98 and 0.93 to project smooth transitional RTB ratios for OASI and DI benefits, respectively.

The ultimate RTB ratios used in the projection are based on OTA's ultimate ratios, reduced by about 4.0 percent. This reduction reflects estimates of the effect of the higher proportion of "old elderly" beneficiaries in the 2090 OASDI beneficiary population distribution relative to the 2025 OASDI beneficiary population distribution, due to improved mortality.

For the 2016 Trustees Report, the ultimate RTB ratios for up to 50 percent of OASI and DI benefits taxable were set at 0.058 and 0.029 , as compared to 0.056 and 0.028 for the 2015 TR ultimate ratios. Slight increases in the ultimate RTB ratios reflect OTA’s updated zero threshold estimates.

Lastly, the Cost sub-process (4.3) applies these projected RTB ratios to projected OASI and DI benefit payments to develop estimated taxation of benefit revenues to the OASI and DI Trust Funds.

\subsection*{4.2. AWARDS}

Each year over 2 million workers begin receiving either retired-worker or disabled-worker benefits. The monthly benefits for these new awards are based on their primary insurance amount (PIA). The PIA is computed using the average indexed monthly earnings (AIME) and the PIA benefit formula as specified in the 1977 amendments. The AIME depends on the worker's number of computation years, \(Y\), and the earnings in each year. For retired-worker beneficiaries who have attained or will attain age 62 in 1991 or later, \(Y=35\).

The AWARDS subprocess (AWARDS) selects records from a 10 percent sample of newly entitled worker beneficiaries obtained from the Master Beneficiary Record (MBR). \({ }^{1}\) The selected sample, referred to as "sample", contains 293,892 beneficiary records, and each record, r , includes a worker's history of taxable earnings under the OASDI program as well as additional information such as sex, birth date, month of initial entitlement, and type of benefit awarded. To estimate the benefit levels of future newly entitled worker beneficiaries, AWARDS modifies the earnings records in the sample to reflect the expected work histories and earnings levels of future beneficiaries (equation 4.2.1). After the modifications, AWARDS computes an AIME for each record in the future sample of beneficiaries (equation 4.2.2). AWARDS subdivides the AIME value of each record into bend point subintervals \({ }^{2}\) (equation 4.2.3). As input to the Cost subprocess, the AIME values are used to calculate aggregate percentages of AIME in each bend point subinterval for each age at entitlement, sex and trust fund (equation 4.2.4). Equations 4.2.1 through 4.2.4 outline the overall structure and solution sequence. The subscript \(n\) refers to the bend point subinterval and \(r\) refers to the sample record.
\[
\begin{align*}
\text { Projected Earnings } & =\text { Projected Earnings }(\cdot)  \tag{4.2.1}\\
\operatorname{AIME}(r) & =\frac{\sum \text { Highest } Y \text { Indexed Earnings }(r)}{Y * 12}  \tag{4.2.2}\\
\text { AIME }_{n}(r) & =\operatorname{AIME}_{n}(\cdot)  \tag{4.2.3}\\
\operatorname{PAP}_{n} & =\frac{\sum_{r} \operatorname{AIME}_{n}(r)}{\sum_{r} \operatorname{bp}_{n}} \tag{4.2.4}
\end{align*}
\]
where \(\mathrm{bp}_{n}\) is the length of the \(n\)th bend point subinterval,
Y is the number of computation years, and
AIME \(_{n}(r)\) is the AIME amount contained within the \(n\)th interval for record \(r\).

\footnotetext{
\({ }^{1}\) A record is selected if the year of initial entitlement equals 2013 as of the December 2015 MBR extract file date, and the beneficiary is not in death status as of the December 2013 MBR extract file date. Retired beneficiaries over age 70 and disability beneficiaries under age 20 are excluded.
\({ }^{2}\) The current law PIA formula has two bend points. For the purposes of PAP, the Awards subprocess instead uses 30 subintervals.
}

\section*{4.2.b. Input Data}

\section*{Long-Range OCACT Projection Data}

Demography-
- Total Social Security area population (as of July) by sex and age.
o From 1951 to 2095
o Updated annually
- Other-than-legal population (as of July) by sex and age
o From 1964 to 2095
o Updated annually
- Deferred Action for Childhood Arrivals (DACA) population (as of July) by sex and age.
o From 2012 to 2095
o Updated annually
- Deferred Action for Parental Accountability (DAPA) population (as of July) by sex and age.
o From 2014 to 2095
o Updated annually
Economics -
- Covered workers by sex and age—with earnings posted to the Master Earnings File (MEF) only. - used with CWHS data to project future earnings levels
o From 1951 to 2095
o Updated annually
- Covered workers not in the other-than-legal population by sex and age -with earnings posted to the Master Earnings File (MEF) only
o From 1951 to 2095
o Updated annually
- Average Wage Index (AWI), projected values.
o From 2014 to 2095
o Updated annually
- Total taxable earnings and number of workers with taxable earnings by age, sex, and year from the Continuous Work History Sample (CWHS).
o From 1951 to 2013
o Updated annually
- Historical Average Taxable Earnings (ATE) —with earnings posted to the Master Earnings File (MEF) only - used with CWHS data to project future earnings levels
o From 1951 to 2014 (only 2013-2014 data used in SOSI)
o Updated annually
- Projected Average Taxable Earnings (ATE) -with earnings posted to the Master Earnings File (MEF) only - used with CWHS data to project future earnings levels
o From 2015 to 2095
o Updated annually
- Projected Covered Worker Rate based on total covered workers/total population (not used in SOSI)
o From 2015 to 2095
o Updated annually
- COLA (Cost Of Living Adjustment) - not used in SOSI
o From 2015 to 2095
o Updated annually
- Projected Wage Base (current law)
o From 2017 to 2095
o Updated annually
Fully Insured -
- Historical and projected fully insured rates by sex and single year of age 14-95
o From 1969 to 2095
o Updated annually

\section*{Beneficiaries -}
- Distribution of newly entitled retired worker beneficiaries-projected weighted ratio of beneficiary for retired worker by year, sex and single year of age 62-70
o From 2015 to 2095
o Updated annually

\section*{Other input data}
- \(10 \%\) Awards Sample from the MBR and Master Earnings File
o Newly entitled OASI / DI beneficiaries, whose initial entitlement year was 2013 as of the December 2015 MBR extract file date, and the beneficiary is not in death status as of the December 2013 MBR extract file date.
- SSN
- Type of benefit
- Type of claim (retirement or disability)
- Sex
- Date of birth
- Date of initial entitlement
- Date of disability onset
- PIA amount
- Type of dual entitlement
- Dual entitlement status code
- PIFC
- LAF
- Eligibility year
- Trust fund
- Earnings histories for each worker from 1951 to 2013
o Generally updated annually, pending validation of the sample
- AWI, Average Wage Index, historical values
o From 1951 to 2014
o Data obtained from OCACT internet site.
o Updated annually
- Wage base
o From 1951 to 2016
o Data obtained from OCACT internet site.
o Updated annually
- COLA, cost of living adjustment - not used in SOSI
o From 1975 to 2014
o Data obtained from OCACT internet site.
o Updated annually
- Amount of earnings needed to earn one quarter of coverage
o From 1951 to 2016
o 1978-2016 data obtained from OCACT internet site. 1951-1977 values estimated by applying projection methodology backwards from 1978.
o Updated annually
- Windfall Elimination Provision (WEP) factors, the \% of sample cases affected by the WEP which will no longer be affected by the WEP, by sex and projection year
o From 2015 to 2095
o Data obtained from OCACT internal calculations
o Updated annually
- PIA bend points - not used in SOSI (expect for 1979 bend points)
o From 1979 to 2016
o Data obtained from OCACT internet site.
o Updated annually
- Hypothetical Wage Base (to reflect relative changes in relative taxable maximum levels over time)
o From 1951 to 2016
o Updated annually
- Projected DAPA disabled worker and retired worker entitlements
o From 2017 to 2095
o Based on application of ratio of DAPA entitlements divided by DAPA population from 2015TR intermediate assumptions, applied to DAPA population based on 2016TR intermediate assumptions

\section*{4.2.c. Development of Output}

All equations described below are projected separately for the OASI and DI program.

\section*{Equation 4.2.1 - Projected Earnings}

In order to estimate future benefit levels, the work histories and earnings levels in the current sample must be modified to represent those for a sample of worker beneficiaries who are newly entitled in future years. Three distinct modifications are made to the earnings records. For each future year, changes are made to the earnings records in order to reflect:
- Changes in Wage Bases.

For some years, the projected wage base (contribution and benefit base), on an AWI discounted basis, is higher than the historical wage base. Therefore, the taxable earnings of future beneficiaries may need to include covered earnings above the
reported historical wage base. Thus, for each record with reported taxable earnings at the wage base in a given year, AWARDS imputes his/her covered earnings.
- Changes in Covered Worker Rates.

Adjustments are made to work histories to be consistent with the projected changes in the economy-wide covered worker rates. Economy-wide covered worker rates are defined as the ratio of "legal" covered workers (from Economics subprocess) to the Social Security area "legal" population (from Demography subprocess). See the below detailed section "Changes in Covered Worker Rates" for more information.
- Earnings Experience in the CWHS \({ }^{3}\).

Earning levels are modified to capture the changes to date that are reflected in the average taxable earnings reported in the CWHS by age and sex and the changes expected in the future.

\section*{Change in Wage Bases}

The earnings posted in the sample are limited by the historical wage base (contribution and benefit base). Prior to 1975, the maximum annual amount of earnings on which OASDI taxes were paid was determined by ad hoc legislation. After 1974, however, the annual maximum level was legislated to be determined automatically, based on the increase in the Social Security Average Wage Index (AWI). Prior to these automatic wage base increases, a relatively large portion of workers earned amounts above the base. Additional legislation raising the annual maximum taxable amount occurred in 1979, 1980, and 1981 to improve the financial future of the OASDI Trust Funds. In addition, the AWI used in the automatic calculation of the annual taxable maximum was modified in the early 1990s to include deferred compensation amounts.

Therefore, for each record in the sample with earnings at the wage base, AWARDS imputes covered earnings above the historical wage base in order to reflect higher maximum taxable amounts imposed on future newly entitled beneficiaries. Please refer to appendix 4.2-1 at the end of this subprocess for details of this imputation. Then, these projected covered earnings are capped at the wage base values that would be in effect for future samples of retired workers (using the "projected wage base" input file) to determine the taxable earnings to use in the benefit calculations.

\section*{Change in Covered Worker Rates}

The sample's covered worker rate by age group and sex is defined as the ratio of (1) the number of those beneficiaries with covered earnings in the sample to (2) the total number of beneficiaries in the sample. For both males and females, the work histories are modified to reflect changes in the covered worker rates that would apply to a future sample of

\footnotetext{
\({ }^{3}\) This file is a \(1 \%\) sample of individuals who had covered earnings at some point in their work histories.
}
beneficiaries. These changes in the covered worker rates are based on changes in the economywide covered worker rates. The economy-wide covered worker rate is defined for an age-sex group in a particular period which represents a future sample cohort as the ratio of (1) the number of "legal" workers in the economy in this group that have some earnings in this period, to (2) the total midyear "legal" population in this group in this period. \({ }^{4}\) Economy-wide covered worker rates are calculated separately for each age-sex group and each historical and projected calendar year based on input data from the Economics, Demography, and Fully Insured subprocesses.

In projecting sample covered worker rates, examination is done of the change in adjusted economy-wide covered worker rates, by age group, between the "base period" (representing individuals retiring in the sample year) and the "projection period" (representing individuals retiring in a year later than the sample year). The adjusted economy-wide covered worker rates in the base year take into account \(60 \%\) of the change in projected age- 62 fully insured rates relative to the base year, based on analyses of historical data. That is, for each projection year, the method estimates what the base year economy wide covered worker rates would have been if the fully insured rates in the base year matched those in the projection year. Details of how this change is used to estimate the change in a covered worker rate for retired workers from a current period in the sample to a future period are given below. The method used is the same for males and females in projecting sample covered worker rates. For additional explanation of this calculation, refer to example 1.1 in appendix 4.2-2 of this subprocess.

Projected sample covered worker rate equals (for increasing economy-wide rates):
- \(\quad\) The ratio of (1) the potential difference in the economy-wide male (or female) covered worker rate in the projection year to (2) the potential difference in the adjusted economy-wide male (or female) covered worker rate in the sample year (i.e., 1 - adjusted economy-wide male (or female) covered worker rate), multiplied by
- \(\quad\) The corresponding potential difference in the sample's male (or female) covered worker rates (i.e.. 1 - sample male (or female) covered worker rate)).
- \(\quad\) The above result is subtracted from 1 to get the projected sample covered worker rate.

This presentation above presumes that economy-wide covered rates increase over time, which is very common for females but not always true for males. The calculation of the change in covered worker rate differs if there is a reduction in relevant economy-wide covered worker rates. Example 1.2 gives an example of the calculations done for males and females if economy-wide covered worker rates decline.

\footnotetext{
\({ }^{4}\) For this purpose, we define the "legal" population as the total SSA area population minus the other-than-legal population (those in the U.S. illegally and those in the U.S. on a temporary basis, e.g., individuals with specific non-immigrant worker or student visas) plus the DACA population and the DAPA population (DACA and DAPA individuals are included in the other-than-legal data but have been effectively made legal, for Social Security purposes, by legislation or executive order).
}

A similar procedure exists for projecting sample covered worker rates for disabled workers, except that the calculations are further broken down by entitlement age group. See Example 1.3 for an example, for male disabled workers.

Once the covered worker rates for the future sample of beneficiaries are determined, modifications to work histories of the sample to attain these rates are generally done by randomly removing or adding earnings. \({ }^{5}\) For males, the procedure is to select records randomly. However, for females, an additional selection criterion is included in order to achieve a specified distribution of the number of years of earnings for retired female beneficiaries. Female records with 10 or fewer years of earnings are not modified. A distribution limit is set for those female workers with 11 to 25 total years of career earnings within the projection year. This distributional limit changes each projected year. In the first year after the sample year, the distribution limit for females is equal to the male distribution plus \(97 \%\) of the difference between the initial male and female distributions within the sample. In each subsequent year, the percentage decreases by three percent until it reaches \(0 \%\). Thus, the females' years of earnings distribution for those with 11 to 25 years of earnings is adjusted to approach that of the males.

If a record is selected for adding earnings in a particular year, the amount of earnings added is based on the career earnings pattern of the selected record. When earnings are added to a record, AWARDS calculates the ratio of (1) the record's Average Indexed Earnings, AIE \({ }^{6}\), to (2) the AIE of a hypothetical worker, \(w\), whose year of birth and sex are the same as the record and whose annual earnings are set equal to average taxable earnings. For this purpose, average taxable earnings are determined by averaging the earnings over all records in the sample with the same sex and year of birth. Then, the preliminary amount of earnings \({ }^{7}\) in year \(t\) that is added to the record is
\[
\text { Earnings }(r, t)={ }^{\operatorname{Pre}} \operatorname{ATE}_{\mathrm{f}}(\operatorname{sex}, t) * \frac{\operatorname{AIE}(r)}{\operatorname{AIE}(w)},
\]
where \({ }^{\text {Pre }} \mathrm{ATE}_{\mathrm{f}}(\) sex, \(t)\) is the average taxable earnings in year \(t\), for those in the sample with the same sex as that of the record, for those retiring in year \(f\).

For additional explanation of this calculation, refer to example 2 in appendix 4.2-2 of this subprocess. Note that all earnings levels get further adjusted by earnings experience in the CWHS for recent workers, as discussed in the below section.

\section*{Earnings Experience in the CWHS}

\footnotetext{
\({ }^{5}\) Individuals in the sample affected by the Windfall Elimination Provision are less likely to have earnings removed or added by this process.
\({ }^{6}\) AIE is the average indexed annual earnings, average over the highest \(Y\) years of earnings (similar to AIME, but an annual amount).
\({ }^{7}\) In this subprocess, earnings histories of projected beneficiaries are all reflected as wage-indexed earnings histories in the 2013 sample.
}

For historical years beginning with 1951, AWARDS uses average taxable earnings by age and sex (cwhs \(\mathrm{ATE}_{\text {as }}\) ) and numbers of covered workers by age and sex ( \(\mathrm{cwhs}^{\mathrm{CW}} \mathrm{as}_{\text {as }}\) ) as tabulated from the most recent CWHS file \({ }^{8}\). To estimate ATE levels for the first projection year, the AWARDS uses the average values of normalized average taxable earnings calculated from the last five historical years of the most recent CWHS file. These computed normalized values take into account changes in aggregate ATEs from the Economics subprocess between each of the five historical years and the first projection year; this allows the comparison of corresponding values from different years in a way that reduces time series effects. For future years, AWARDS projects these values (ATEs and taxable earnings) from this base year (the first projection year).

Projections are made for each year after the base year through the end of the 75-year projection period using projected economy-wide number of covered workers by age and sex and aggregate annual average taxable earnings (ATE) from the Economics process \({ }^{9}\). The first step is to determine preliminary cwhs ATE \(_{\text {as }}^{\prime}\) (preliminary average taxable earnings by age and sex) by using the annual growth rate in the total economy-wide ATE and projected number of covered workers from the Economics process. A further multiplicative adjustment is made to each cwhs \({ }^{\text {ATE }}\) 'as such that the resulting aggregate average taxable earnings, determined by combining the projected values of covered workers from the Economics process and \(\mathrm{cwhs} \mathrm{ATE}_{\text {as }}\) for the year, produces the same aggregate ATE level for that year as projected by the Economics process.

For additional explanation of this calculation, refer to example 3 in appendix 4.2-2 of this subprocess.

The historical and projected \({ }_{\text {cwhs }}\) ATE \(_{\text {as }}\) are then used to change the earnings histories of the sample of newly entitled beneficiaries so that the earnings better represent newly entitled beneficiaries in future years. For a given sex, trust fund, and earnings year, the expected annual average taxable earnings of a future sample is denoted as ATE \(_{f}\) '. ATE \(_{f}\) ' equals the sample's average taxable earnings for a specific earnings year, multiplied by the comparable changes in the cwhs \(\mathrm{ATE}_{\text {as }}\), that is, the wage-indexed changes in the cwhs ATE \(_{\text {as }}\) between the year of earnings in the sample of new beneficiaries and the year of earnings in the projected sample. For OASI, the projected \({ }_{\mathrm{CWHS}} \mathrm{ATE}_{\text {as }}\) are determined using the age distribution of newly entitled retired worker beneficiaries in the projection year as weights. For DI, the projected cwhs ATE \(_{\text {as }}\) are determined using the historical and projected number of covered workers for all ages corresponding to the earnings year as weights. Refer to example 3.1 in appendix 4.2-2 for additional explanation.

ATE \(_{f}\) ' is then compared to the average taxable earnings of the sample (after adjustments to the records’ earnings levels for changes in wage bases and covered worker rates), denoted as ATE \(_{f}\) and computed by sex, trust fund, and earnings year. The difference between these values is the amount by which the average annual earnings levels are adjusted. Let

\footnotetext{
\({ }^{8}\) These historical values are tabulated by the Economic subprocess.
\({ }^{9}\) These values are based on earnings posted to the Master Earnings File (MEF), excluding earnings posted to the suspense file.
}
\[
\delta(t)=\mathrm{ATE}_{\mathrm{f}}{ }^{\prime}-\mathrm{ATE}_{\mathrm{f}}
\]
for each year \(t\). Denote the total workers in the sample in year \(t\) as TotalWorkers \((t)\). Then, \((\delta(t)\) * TotalWorkers \((t)\) ) is the total amount of earnings which the model distributes for a given sex and age group in a way so that the average taxable earnings after distribution is ATE \(_{f}\) '.

For additional explanation of the calculation \(\delta(t)\), refer to example 4 in the appendix 4.2-2.
When \(\delta(t)\) is negative, earnings for the year are decreased. To achieve ATE \(_{\mathrm{f}}{ }^{\prime}\) for the given sex and age-group, AWARDS multiplies CoveredEarnings \((r, t)\) by a ratio,
\[
\operatorname{ratio}(t)=1+\frac{\delta(t)}{\operatorname{ATE}_{\mathrm{f}}(t)}+\alpha
\]

The term, \(\alpha\), is an additional adjustment necessary because covered earnings near or above the wage base, may have either a partial effect or no effect on modifying ATE \(_{f}\) to ATE \(_{f}{ }^{\prime}\). These \(\alpha\) values vary by sex, trust fund, and whether earnings increase or decrease. In the 2016 Trustees Report, for OASI, \(\alpha\) is set equal to -0.075 for males and -0.05 for females when the earnings are decreasing; and \(\alpha\) is 0.03 for males and 0.05 for females when the earnings are increasing. Similarly, for DI, \(\alpha\) is set equal to -0.075 for males and -0.08 for females when the earnings are decreasing; and \(\alpha\) is 0.065 for males and 0.005 for females when the earnings are increasing. These \(\alpha\) values are set to best target ATE \(_{f}\) ' while making adjustments throughout as many of the sample records as possible.

As AWARDS applies ratio \((t)\) to Earnings \((r, t)\) by each record, it makes sure that the total earnings adjustment in a year does not exceed \(\delta(t) *\) TotalWorkers \((t)\). For additional explanation of this calculation, refer to example 5 in appendix 4.2-1 of this subprocess.

\section*{Equation 4.2.2 - Average Indexed Monthly Earnings (AIME)}

\section*{Step 1: Index Earnings}

To compute an individual's AIME, all taxable earnings after 1950 are considered. First, the earnings are indexed up to the index year, \(i\), which is defined as the year of attaining age 60 for retired-worker beneficiaries (eligible for benefits at age 62). For disabled-worker beneficiaries, \(i\) is set to be 2 years before the sample year. Thus,
\[
\text { IndexedEarnings }(\mathrm{r}, \mathrm{t})= \begin{cases}\operatorname{Earnings}(r, t) * \frac{\text { AverageWage }(i)}{\text { AverageWage }(t)}, & \text { if } \mathrm{t}<i \\ \operatorname{Earnings}(\mathrm{r}, \mathrm{t}), & \text { if } \mathrm{t} \geq i\end{cases}
\]

Step 2: Determine Computation Years

For each record, the number of computation years, Y, is determined. For a retired-worker beneficiary in the sample, Y is 35.

For a disabled-worker beneficiary, Y is calculated as follows:
- Determine the number of elapsed years, which is equal to the year of disability onset (not later than the year the worker turned age 62) minus the greater of 1951 or the year the disabled worker turned age 22.
Elapsed Years \(=\min \{\) Year of disability onset, Year attained age 62\(\}-\max \{1951\), Year attained age 22\}
- Divide the elapsed years by five and truncate. Subtract this number (cannot exceed five) from the number of elapsed years.
\[
Y=\text { Elapsed } Y \text { ears }-\min \left\{\left(\frac{\text { ElapsedYears }}{5}\right\rfloor, 5\right\}
\]
- Y must be at least 2 .

Step 3: Determine AIME
Finally, an individual's AIME is computed by summing the highest Y indexed earnings and dividing by the number of months in those years. Hence, for each record,
\[
\operatorname{AIME}(\mathrm{r})=\frac{\sum \text { Highest } \mathrm{Y} \text { Indexed Earnings }(\mathrm{r})}{\mathrm{Y} * 12}
\]

\section*{Equation 4.2.3- AIME \(_{n}(r)\)}

The Possible AIME value is divided into 30 intervals (bend point subintervals). The length of each interval in 1979 dollars is given below:
\[
\mathrm{bp}_{n}= \begin{cases}\$ 45, & \text { if } 0<n \leq 13 \\ \$ 100, & \text { if } 14 \leq n \leq 18 \\ \$ 200, & \text { if } 19 \leq n \leq 28 \\ \$ 1000, & \text { if } 29 \leq n \leq 30\end{cases}
\]

Thus, the interval points of AIME division given below in 1979 dollars, \(\mathrm{y}_{\mathrm{k}}\), are equal to \(\sum_{n=1}^{k} b p_{n}\) and
\[
\mathrm{y}_{k}= \begin{cases}\$ 180, & \text { if } k=4 \\ \$ 1085, & \text { if } k=18 \\ \$ 5085, & \text { if } k=30\end{cases}
\]

For each record (r), the values for \(\mathrm{bp}_{n}\) are indexed from 1977 to his/her indexing year \(i\) using the Social Security average wage index (AWI). So for \(n=1\) to 30 ,
\[
\mathrm{bp}_{n}(r)=\mathrm{bp}_{n} * \frac{\mathrm{AWI}(\mathrm{i})}{\mathrm{AWI}(1977)}
\]

Next the record's AIME amount, AIME ( \(r\) ), is compared to the indexed intervals. If
\[
\sum_{n=1}^{k-1} \mathrm{bp}_{n}(r)<\operatorname{AIME}(r) \leq \sum_{\mathrm{n}=1}^{\mathrm{k}} \mathrm{bp}_{n}(r)
\]
then AIME (r) falls within the \(k\) th interval. And for \(n=1\) to 30 ,
\[
\operatorname{AIME}_{\mathrm{n}}(r)= \begin{cases}\mathrm{bp}_{n}(r), & \text { if } n<k \\ \operatorname{AIME}(r)-\sum_{n=1}^{k} \mathrm{bp}_{n}(r), & \text { if } n=k \\ 0, & \text { if } n>k\end{cases}
\]

\section*{Equation 4.2.4 - Potential AIME Percentages (PAPS)}

Finally, for \(n=1\) to 30 , AWARDS sums the values of AIME \(_{n}\) and \(\mathrm{bp}_{n}\) across all the records for years 20154 to 2095 by sex, age (20-65 for disabled workers, and 62-70 for retired workers), and trust fund. The ratio of these values gives the average potential AIME percentages (PAPS)
\[
\operatorname{PAP}_{n}=\frac{\sum_{\mathrm{r}} \operatorname{AIME}_{\mathrm{n}}(r)}{\sum_{\mathrm{r}} \mathrm{bp}_{\mathrm{n}}(r)}
\]

For an example of this calculation, refer to example 6 in appendix 4.2-2 of this subprocess.
In November 2014, President Obama announced a number of executive actions dealing with immigration. The DAPA program, allows legal work authorization for parents who (a) have a child that is a citizen or a legal permanent resident and (b) have been present in the country since January 1, 2010. Because individuals affected by DAPA presumably did not have prior legal work authorization as of late 2014, it is likely that such individuals have limited or no
covered earnings before 2015. As a result, DAPA individuals are not well represented by the current historical 2013 10\% sample of initial entitlements used in the Awards area.

To account for the lack of records in the sample presenting DAPA individuals, the AWARDS subprocess generated the hypothetical records to represent them. In these hypothetical records, the covered earnings begin in 2017 and continue until disability or retirement. In some cases, earnings prior to 2017 are required to obtain fully insured status. The number of hypothetical records is based on projected number of their entitlements, which are based on the projected DAPA population and "incidence rates" based on 2015TR projected DAPA entitlements by trust fund/age/sex/year relative to the corresponding DAPA population by age/sex/year.

\section*{Appendix 4.2-1}

This appendix provides additional details on how the AWARDS process imputes covered earnings above the historical wage base.

To do this, AWARDS first computes the cumulative distribution, \(F\), of the workers in the sample by their earnings level. Each historical wage base is divided into 20 equal intervals, \(n\), and each interval length in year \(t\) is
\[
\text { IntervalLength }(t)=\frac{\text { WageBase }(\mathrm{t})}{20}
\]

The cumulative distribution \(\mathrm{F}(n, t)\) is the proportion of workers whose earnings are less than IntervalLength \((t)^{*} n\), for \(n=1\) to 20. Let NumberWorkers \((n, t)\) be the number of workers whose earnings in year \(t\) fall within the \(n\)th interval, that is the earnings are greater or equal to IntervalLength \((t)^{*}(n-1)\) and less than IntervalLength \((t)^{*} n\). Also, let TotalWorkers( t\()\) be the total number of workers in the sample with earnings in year \(t\). Then for any \(n, 1 \leq n \leq 20\),
\[
\mathrm{F}(n, t)=\frac{\sum_{m=1}^{n} \text { NumberWorkers }(m, t)}{\operatorname{TotalWorkers}(t)} .
\]

Once \(\mathrm{F}(n, t)\) is computed for \(n=1\) to 20, AWARDS extends the function for those who had earnings at the wage base. To extrapolate F past the historical base (define \(\mathrm{F}(n, t)\) for \(n>20\) ), AWARDS groups the maximum earners in each year in the sample based on the number of years they had earnings at the wage base during the next four years \((0,1,2,3,4)\). Under the assumption of uniform distribution within each group, AWARDS assigns an \(\mathrm{F}\left(n_{r}, t\right)\) value to each record with earnings at the tax maximum beginning with the group that has no other earnings at the tax maximum during the next four years and ending with the group that has maximum earnings in each of the next four years. Note that for these beneficiaries \(\mathrm{F}\left(n_{r}, t\right)>\) \(\mathrm{F}(20, t)\). Once \(\mathrm{F}\left(n_{r}, t\right)\) is computed for these beneficiaries, values for \(\mathrm{F}(n, t)\), where \(n>20\) are estimated.

To find \(\mathrm{F}(n, t)\), where \(n>20\), the log odds transformation is utilized. The odds ratio,
\[
\mathrm{T}(n, t)=\frac{\mathrm{F}(n, t)}{1-\mathrm{F}(n, t)} \text {, where } n \leq 20
\]
is the ratio of (1) the proportion of beneficiaries with earnings levels below the \(n\)th interval to (2) the proportion of beneficiaries with earnings levels above the \(n\)th interval. Next, the natural logarithm of the odds ratio is computed, giving the log odds transformation,
\[
\mathrm{Y}(n, t)=\ln [\mathrm{T}(n, t)]=\ln \left(\frac{\mathrm{F}(n, t)}{1-\mathrm{F}(n, t)}\right) \text {, where } n \leq 20
\]

Utilizing the most linear portion of the function at the upper values of \(n\), AWARDS regresses Y on those values. The regression line of Y has the form
\[
\hat{\mathrm{Y}}(n, t)=\beta_{0}+\beta_{1} * n
\]

Finally, the amount of covered earnings of a record that has earnings at the taxable maximum is determined based on the \(\mathrm{F}\left(n_{r}, t\right)\) value assigned to the record. The \(\mathrm{F}\left(n_{r}, t\right)\) value for this record is used in the above equations to determine \(\mathrm{T}\left(n_{r}, t\right), \mathrm{Y}\left(n_{r}, t\right)\), and then \(n_{r}\), the noninteger value for \(n\) in the regression equation of \(\hat{Y}\) above. Thus, if earnings \((r, t)=\) wage base in year \(t\) then
\[
\text { CoveredEarnings }(\mathrm{r}, \mathrm{t})=n_{r} * \text { IntervalLength }(\mathrm{t})+\text { error }^{10},
\]
where \(n_{r}\) is the record's non-integer value for \(n\) in the regression equation of \(\hat{\mathrm{Y}}\) above.
If earnings \((r, t)<\) wage base in year \(t\), then CoveredEarnings \((r, t)=\) earnings \((r, t)\)
At this point, AWARDS defines the expected taxable earnings of a future sample as,
\[
\text { Earnings }(r, t)= \begin{cases}\text { CoveredEarnings }(\mathrm{r}, \mathrm{t}), & \text { CoveredEarnings }(\mathrm{r}, \mathrm{t})<\text { future wage base }(\mathrm{t}) \\ \text { future wage base }(\mathrm{t}), & \text { CoveredEarnings }(\mathrm{r}, \mathrm{t}) \geq \text { future wage base }(\mathrm{t})\end{cases}
\]

\footnotetext{
\({ }^{10}\) The difference between \(Y(20, t)\) and \(\hat{Y}(20, t)\).
}

\section*{Appendix 4.2-2}

This appendix provides examples to help understand the calculations described in the model documentation of the AWARDS subprocess. These examples do not necessarily reflect actual values.

\section*{Example 1.1: (OASI-Male with increasing economy-wide covered worker rates)-same method applies for females}

Task: In projecting the 2013 sample of newly entitled male beneficiaries to represent newly entitled male beneficiaries in 2035, an adjustment to the earnings histories for those males age 30-34 is needed to reflect higher covered worker rates expected for males in this age group.

This example illustrates the calculation of the projected covered worker rate for males who are age 30-34 in the projection period. We will be comparing the group of males age 30-34 in the base period with its counterpart group of males age 30-34 in the projection period.

\section*{Information given:}
- Newly entitled retired male beneficiaries represented in the 2013 sample are age 30-34 in the base period, 1973-1985, and the counterpart group of males retiring in 2035 is age 30-34 in the projection period, 1995-2007.
- Based on the 2013 sample, the covered worker rate for males age 30-34 in the base period \(=84.64 \%\).
- Fully insured rate for age 62 in \(2012=95.66 \%\)
- Fully insured rate for age 62 in \(2034=95.40 \%\)
- Unadjusted economy-wide covered worker rate for males age 30-34 = 88.59\% in the base period (2012).
- Economy-wide covered worker rate for males age \(30-34=88.92 \%\) in the projection period (2035).

\section*{Calculations:}
1. Adjusted economy wide covered worker rate in the base period \(=\) Unadjusted economy-wide covered worker rate in the base period * (1+ (Fully insured rate 2035/Fully insured rate 2012-1) * . 6 ) \(=88.45 \%\)
2. The potential difference in the economy-wide covered worker rate for males age \(30-34\) in the projection period is \(100.0 \%-88.92 \%\) or \(11.08 \%\).
3. The potential difference in the adjusted economy-wide covered worker rate for males age \(30-34\) in the base period is \(100.0 \%-88.45 \%\) or \(11.55 \%\).
4. The ratio from steps 2 and 3 is \(95.91 \%\).
5. The potential difference in the sample's covered worker rate for the males age \(30-34\) in the base period is \(100.0 \%-84.64 \%\) or \(15.36 \%\).
6. The ratio from step 4 is multiplied by the potential difference in the sample's covered worker rate for males age 30-34 in the base period to yield \(14.73 \%\) ( \(95.91 \%\) * 15.36\%).
7. The amount in step 6 (14.73\%) would be subtracted from 1 to yield the sample's covered worker rate for males who are age 30-34 in the projection period (85.27\%).

\section*{Example 1.2: (OASI-Male with decreasing economy-wide covered worker rates) - same method applies for females}

Task: In projecting the 2013 sample of newly entitled male beneficiaries to represent newly entitled male beneficiaries in 2050, an adjustment to the earnings histories for those males age 40-44 is needed to reflect lower covered worker rates expected for males in this age group.

This example illustrates the calculation of the projected covered worker rate for males who are age 40-44 in the projection period. We will be comparing the group of males age 40-44 in the base period with its counterpart group of males age 40-44 in the projection period.

\section*{Information given:}
- Newly entitled retired male beneficiaries represented in the 2013 sample are age 40-44 in the base period, 1983-1995, and the counterpart group of males retiring in 2050 is age 40-44 in the projection period, 2020-2032.
- Based on the 2013 sample, the covered worker rate for males age 40-44 in the base period \(=86.03 \%\).
- Fully insured rate for age 62 in \(2012=95.66 \%\)
- Fully insured rate for age 62 in \(2049=94.37 \%\)
- Unadjusted economy-wide covered worker rate for males age \(40-44=\) 85.91\% in the base period.
- Economy-wide covered worker rate for males age \(40-44=83.36 \%\) in the projection period.

\section*{Calculations:}
1. Adjusted economy wide covered worker rate in the base period \(=\) Unadjusted economy-wide covered worker rate in the base period * (1+ (Fully insured rate 2049/Fully insured rate 2012-1) *.6 ) = 85.21\%
2. The economy-wide covered worker rate for males age 40-44 in the projection period is \(83.36 \%\).
3. The ratio from steps 2 and 1 is \(.8336 / .8521\) or \(97.82 \%\).
4. The ratio from step 3 is multiplied by the sample's covered worker rate for males age 40-44 in the base period to yield 84.16\% (97.82 \% * \(86.03 \%=84.16 \%)\).

The amount in step 4 (84.16\%) would be the sample covered worker rate for males who are age 40-44 in the projection period.

\section*{Example 1.3: (DI-Male with decreasing economy-wide covered worker rates) - same method applies for females}

Task: In projecting the 2013 sample of newly entitled male DI beneficiaries age 60-64 to represent newly entitled male beneficiaries in 2050, an adjustment to the earnings histories for those males age 40-44 is needed to reflect lower covered worker rates expected for males in this age group.

This example illustrates the calculation of the projected covered worker rate for males who are newly entitled at age 60-64 in 2050 and are age 40-44 in the projection period. We will be comparing the group of males age 40-44 in the base period with its counterpart group of males age 40-44 in the projection period. The formulas are similar to Example 1.2 in relating the age-group of earnings to a specific entitlement age-group, but with no fully insured adjustment to economy-wide covered worker rates in the base year.

\section*{Information given:}
- Newly entitled male DI beneficiaries who are age 60-64 represented in the 2013 sample are age 40-44 in the base period, 1989-1997, and the counterpart group of males DI beneficiaries who are age 60-64 in 2050 is age 40-44 in the projection period, 2026-2034.
- Based on the 2013 sample, the covered worker rate for males DI beneficiary age 40-44 in the base period \(=90.34 \%\).
- The economy-wide covered worker rate for males age 40-44 in the base period (2012) is \(85.65 \%\).
- The economy-wide covered worker rate for males age 40-44 in the projection period (2049) is \(83.62 \%\).

\section*{Calculations:}
1. The economy wide covered worker rate for male age 40-44 in the base period is \(85.65 \%\)
2. The economy wide covered worker rate for male age 40-44 in the projection period is \(83.62 \%\).
3. The sample covered worker rate for males DI beneficiary age \(40-44\) is 90.34\%.
4. The ratio from steps 2 and 1 is \(97.63 \%\).
5. The ratio from step 4 is multiplied by the sample's covered worker rate for males age 40-44 in the base period to yield 88.20\% (97.63 \% * 90.34\%).

\section*{Example 2:}

Task: In projecting the 2013 sample of newly entitled male OASI beneficiaries to represent newly entitled male OASI beneficiaries in 2040, an adjustment to the earnings histories for those males age 40-44 is needed to reflect higher covered worker rates expected for males in this age group. To achieve this target, the desired number of records with zero reported earnings in this age group are randomly selected and assigned earnings.

This example illustrates the calculation of earnings to be assigned to a randomly chosen newly entitled retired male record with zero taxable earnings in the base year.

\section*{Information given:}
- Newly entitled retired male beneficiaries represented in the 2013 sample are age 40-44 in the base period, 1983-1995, and the counterpart group of males retiring in 2040 will be age 40-44 in the projection period, 20102022.
- Based on the 2013 sample, a male record, \(r=59,950\), has been randomly selected to replace his zero taxable earnings reported in the base year 1990 at age 40 with an amount based on his career earnings pattern. This record is age 63 in 2013, and his year of birth is 1950.
- A beneficiary retiring at age 63 in 2040 will have a year of birth of 1977. And at age 40, the corresponding projection year to the base year of 1990 is the projection year 2017.
- \(\quad\) The Average Indexed Earnings for this record, AIE \((59,950)\), is computed to be \(\$ 74,344\). Note: This value is calculated by (1) using the record's annual taxable earnings reported each year through 2012, (2) converting them to 2012 year dollars, and then (3) summing the highest 35 years of earnings and dividing by 35 .
- \(\quad\) The Average Indexed Earnings for a hypothetical worker, \(\operatorname{AIE}(w)\) whose year of birth is 1950 is \(\$ 66,741\). This value is calculated as above given the hypothetical worker earned the average taxable earnings in each of the base years for males retiring at age 63 in the 2013 sample.
- \(\quad\) The projected Average Taxable Earnings of males retiring in 2040 for the projection year 2017 (converted to 1990 dollars \({ }^{11}\) )
is \({ }^{\text {Pre }} \mathrm{ATE}_{2040}(\) male,2017 \()=\$ 29,869\).

\section*{Calculations:}

\footnotetext{
\({ }^{11}\) All projection year dollars are converted back to 'sample year' dollar amounts, using the Average Wage Index.
}
1. The ratio of the Average Indexed Earnings for record number 59,950, AIE \((59,950)\) to the Average Indexed Earnings of a hypothetical male worker born in 1950 and retiring at age 63 , \(\operatorname{AIE}(w)\) is \(\$ 74,344 / \$ 66,741\) or 1.114.
2. The amount in step 1 (1.114) would be multiplied by \({ }^{\text {Pre }} \mathrm{ATE}_{2040}\) (male,2017), which is given as \(\$ 29,869\). This yields the amount of earnings assigned to record number 59,950 in the projection year 2017 representing a sample retiring in 2040, Thus, Earnings \((59,950,2017)=1.114 * \$ 29,869\) which equals \(\$ 33,272\).

Note that, at this stage of the process, the average taxable earnings, denoted \({ }^{\text {Pre }}\) ATE \(_{2040}\) (male,2017), have been computed using projected earnings adjusted for changes in the wage base. Adjustments to earnings for the earnings experience in the CWHS have not yet been applied. See the earlier section "Earnings Experience in the CWHS" and example 4 below for a description of this adjustment process.

\section*{Example 3:}

Task: The AWARDS subprocess estimates projected values of Average Taxable Earnings by age and sex using the values \({ }^{12}\) in the 2013 CWHS file supplied by the Economic subprocess as the base year on which to build our projections.

This example illustrates the calculation of the projected Average Taxable Earnings of the CWHS in 2017 for 42 year old females, cwhs ATE 42 ,female \((2017)\). We will be using the number of female covered workers age 42 and the total taxable earnings for females age 42 as given the in 2013 CWHS data.

\section*{Information given:}
- \(\quad\) The average taxable earnings calculated from last 5 historical years in the CWHS for females age 42 are:
\[
\begin{aligned}
& \text { ATE_cwhs(2009) }=33,899.55 \\
& \text { ATE_cwhs(2010) }=34,682.18 \\
& \text { ATE_cwhs(2011) }=35,845.87 \\
& \text { ATE_cwhs(2012) }=37,309.87 \\
& \text { ATE_cwhs(2013) }=37,945.75
\end{aligned}
\]
- The economy-wide average taxable earnings 2009-2017 for females age 42 are:

ATE_econ(2009) = 33,203.01
ATE_econ(2010) \(=33,536.98\)

\footnotetext{
\({ }^{12}\) Because not all earnings are posted for the most recent years for a given CWHS file, adjustment factors, based on historical trends, are applied by the Economic subprocess to complete these earnings. For the 2013 CWHS, adjustment factors were applied to data in years 2009 through 2013.
}

ATE_econ(2011) = 34,292.49
ATE_econ(2012) \(=35,199.98\)
ATE_econ(2013) \(=35,898.04\)
ATE_econ(2014) \(=36,946.23\)
ATE_econ(2015) \(=37,531.18\)
ATE_econ(2016) \(=38,512.75\)
ATE_econ(2017) \(=40,500.83\)

\section*{Calculations:}
1. The normalized ATE for females age 42 in 2017, using 2009-2013 CWHS data brought forward to 2017 ATE levels. The calculation is shown in 2 steps :

Step a
ATE_cwhs(2009)*ATE_econ(2017)/ATE_econ(2009) +
ATE_cwhs(2010)*ATE_econ(2017)/ATE_econ(2010)
\(+\ldots . . .+\) ATE_cwhs(2013)*ATE_econ(2017)/ATE_econ(2013)
Step b
Sum of Step a (\$211,308.72), then take the average (\$211,308.72
/ 5 =\$42,261.74)
2. Repeat the same steps in (1) to calculate the normalized ATE for each age (15 to 80). Multiply the normalized ATE by number of econ-wide covered workers by age and sex in year 2017. This gives us the taxable earnings. Sum up the taxable earnings for all ages and sex (both male and female). Then divide the total taxable earnings (age and sex combined) by total number of economy-wide covered workers in 2017 (age and sex combined). The resulting ATE for 2017 using economywide covered workers and 2009-2013 CWHS data brought forward to 2017 ATE levels is \(\$ 39,786.29\).
3. The value from step \(1(\$ 42,261.74)\) is multiplied by a final multiplicative adjustment factor to match the aggregate ATE level produced by the Economic subprocess. The adjustment is a ratio of economy-wide ATE to CWHS projected aggregate average taxable earnings (\$40,500.83 / \$39,786.29 from step 2 above) yielding for females age 42 in 2017 an ATE value of \(\$ 43,020\).

\section*{Example 3.1:}

This example illustrates the calculation of the projected Average Taxable Earnings in 2027 for 34 to 42 year old females who retire at age 62 to 70 in 2055. We will be using the historical and projected ATE and distribution of retired workers in the projection
year to calculate the weighted CWHS ATE for that year.

\section*{Information given:}
- \(\quad\) The historical ATE from CWHS sample for age 34 to 42 in 1985, who retire in sample year 2013, is listed as follows.
ATE (1985,female,34) \(=12,667\)
ATE \((1985\), female,35 \()=12,718\)
ATE \((1985\), female,36 \()=12,641\)
ATE (1985,female,37) \(=12,840\)
ATE \((1985\),female,38) \(=12,821\)
ATE(1985,female,39) \(=13,076\)
ATE \((1985\), female, 40\()=12,861\)
ATE \((1985\), female, 41\()=12,761\)
\(\operatorname{ATE}(1985\), female,42 \()=12,806\)
- \(\quad\) The projected ATE, using historical CWHS sample data as a guide and projected covered worker data from the Economics group for age 34 to 42 in 2027, who retire in sample year 2055, is listed as follows.
ATE(2027,female,34) \(=60,722\)
ATE (2027,female,35) = 61,907
ATE (2027,female,36) \(=62,752\)
ATE (2027,female,37) \(=63,459\)
ATE (2027,female,38) \(=64,038\)
ATE (2027,female,39) \(=64,561\)
ATE (2027,female,40) = 65,173
ATE (2027,female,41) \(=65,458\)
ATE (2027,female,42) \(=65,601\)
- The distribution of retired worker for age 62 to 70 in 2055, is listed as follows.
weight(2055,female,62) \(=.2519\)
weight(2055,female,63) = . 1346
weight(2055,female,64) \(=.0605\)
weight(2055,female,65) \(=.1277\)
weight(2055,female,66) \(=.1632\)
weight(2055,female,67) \(=.1296\)
weight(2055,female,68) \(=.0465\)
weight(2055,female,69) \(=.0421\)
weight(2055,female,70) \(=.0437\)
- Average wage index in \(1985=\$ 16,822.51\)
- \(\quad\) Average wage index in \(2027=\$ 77,326.64\)
- \(\quad\) Sample average taxable earnings in \(1985=\$ 13,529.41\)

\section*{Calculations:}
1. The calculated weighted historical ATE in 1985 is the sum of historical ATE multiplied by the corresponding weight from the 2055 retired worker distribution. That is \(\$ 12,789\) equals the sum of \(12,667^{*} .2519+\) 12,718*. \(1346+12,641^{*} .0605+\) \(\qquad\) \(+12,806 * .0437\)
2. The calculated weighted projected ATE in 2027 is the sum of projected ATE multiplied by the corresponding weight from the 2055 retired worker distribution. That is \(\$ 63,000\) equals the sum of \(60,722^{*} .2519+\) 61,907*. \(1346+62,752^{*} .0605+\ldots . .+65,601^{*} .0437\)
3. The ratio of AWI (1985) to AWI (2027) \(=.2176\) (\$16,822.51/\$77,326.64)
4. The adjusted sample ATE in 2027 (in 1985 dollars) = sample ATE in 1985 * projected weighted ATE in 2027/ projected weighted ATE in 1985 * AWI (1985) / AWI (2027) = \$14,499 (13,529.41 * (63,000 / \(12,789)\) *.2176)

\section*{Example 4:}

Task: In projecting the 2013 sample of newly entitled female beneficiaries to represent newly entitled female beneficiaries in 2055, the projected Average Taxable Earnings of females in the sample \(\left(\mathrm{ATE}_{\mathrm{f}}\right)^{13}\) for year \(t=2027\) must be adjusted by an amount, \(\delta(2027)\), to meet a targeted Average Taxable Earnings(ATE \({ }_{f}\) ') for 2027.

This example illustrates the calculation of \(\delta(2027)\) for the female cohort retiring at ages 62 to 70 in the projection year 2055. The cohort is age 34 to 42 in the projection year 2027. The value \(\delta(2027)\) is the dollar amount in which the average annual earnings levels are adjusted for females in this age group in the year 2027 and retiring in 2055. We will be comparing this group of females age 34-42 in the projection 2027 year with its counterpart group of females age 34-42 in the base year 1985.

\section*{Information given:}
- A cohort of newly entitled retired female beneficiaries retiring at ages 62 to 70 represented in the 2013 sample are ages 34-42 in the base year, 1985, and the counterpart group of females retiring in 2055 are ages \(34-\) 42 in the projection year, 2027.
- Based on the 2013 sample, the average taxable earnings for females ages \(34-42\) in the base year 1985 is \(\$ 13,529\).
- For a sample projected to be retiring in 2055, the average taxable earnings ( \(\mathrm{ATE}_{f}\) ) for females in the projection year 2027 is \(\$ 12,826^{14}\), after applying adjustments to the records' earning levels for changes in the wage base and in covered worker rates.

\footnotetext{
\({ }^{13}\) The average taxable earnings have been computed using projected earnings adjusted for changes in the wage base and for changes in covered worker rates.
\({ }^{14}\) Amount is in 1985 dollars, 'sample year dollars'.
}
- The average taxable earnings (ATE \({ }_{f}\) ') of future sample (2055) in 2027 as shown in Example 3.1 is \(\$ 14,499\).

\section*{Calculations:}
- \(\quad\) The difference in ATE \(_{f}{ }^{\prime}(\$ 14,499)\) and ATE \(_{f}(\$ 12,826)\) yields the \(\delta(2027)\) value \(\$ 1,673\) ( \(\$ 14,499-\$ 12,826=\$ 1,673)\).

\section*{Example 5:}

Task: In projecting the 2013 sample of newly entitled female OAB beneficiaries to represent newly entitled female OAB beneficiaries in 2055, for year \(t=2027, \delta(2027)\) is positive indicating an adjustment to earnings histories is needed to reflect higher average taxable earnings by this cohort for the projection year 2027.

This example illustrates the calculation of the ratio (2027) in projection year 2027 for the female cohort retiring at ages 62 to 70 in the projection year 2055. The value, ratio (2027), is the adjustment ratio that will be applied to this cohort projected covered earnings in 2027 in order to achieve the targeted Average Taxable Earnings of this cohort for 2027.

\section*{Information given:}
- Earnings in the projection year 2027 for a group of newly entitled female beneficiaries retiring in 2055 is the counterpart corresponding to earnings in the base year 1985 for the group of newly entitled female beneficiaries in the 2013 sample.
- The targeted average taxable earnings ATE \(_{2027}{ }^{\prime}\) for the projection year 2027 is \(\$ 14,499\) (shown in Example 4).
- For newly entitled females retiring in 2055, the average taxable earnings (ATE \({ }_{2027}\) ) for the projection year 2027 is \(\$ 12,826^{15}\) (shown in Example 4).
- \(\quad \delta(2027)\), the difference in the targeted average taxable earnings ATE \(_{2027^{\prime}}\) and ATE \(_{2027}\), is calculated to be \(\$ 1,673\) ( \(\$ 14,499-\$ 12,826\) ).
- \(\quad \alpha\) is an additive constant to the ratio(2027). For female OABs the constant \(\alpha\) when \(\delta(t)\) is positive is .05 .

\section*{Calculations:}
- The ratio(2027) multiplied to the covered earnings in 2027 for females retiring in 2055 is \(1+\delta(2027) /\) ATE \(_{2027}+\alpha=(1+(\$ 1,673 / \$ 12,826)+\) 0.05 ), or 1.18 .

\footnotetext{
\({ }^{15}\) Amount is in 1985 dollars, 'sample year dollars'.
}

\section*{Example 6:}

Task: The AWARDS subprocess calculates the Average Indexed Monthly Earnings (AIME) of each beneficiary in the sample. The AIME values are then divided into 30 intervals.

This example illustrates the division of a possible AIME value into intervals.

\section*{Information given:}
- An OAB beneficiary retired at age 64 in 2013
- This OAB beneficiary is record \#150000 in the sample \((r=150000)\)
- The AIME for this individual is \(\$ 3,000\)
- The initial eligibility year is 2011, the year the individual turned age 62.
- The length of each interval \(\left(\mathrm{bp}_{n}\right)\) in 1979 dollars is given in Equation 4.2.3. The length of each interval in 2011 dollars is given by the equation
\[
\mathrm{bp}_{\mathrm{n}}(\mathrm{r})=\mathrm{bp}_{n} * \operatorname{AWI}(2009) / \mathrm{AWI} \text { (1977) }
\]
where \(\mathrm{bp}_{n}\) is the length of interval n in 1979 dollars
- \(\quad\) The average wage index (AWI) for year 2009 is \(\$ 40,711.61\)
- \(\quad\) The AWI for year 1977 is \(\$ 9,779.44\)
- When converting the intervals from 1979 dollars to 2011 dollars, there is a 2-year lag in AWI values.
- \(\quad \operatorname{AIME}_{\mathrm{n}}\) (150000) is the AIME value in interval n for Record \#150000

\section*{Calculations:}
- \(\quad\) The AIME for this individual \((\$ 3,000)\) is compared to the indexed intervals. It falls within the \(15^{\text {th }}\) interval.
- The AIME 15 is the residual of \(\$ 3,000\) subtracting the cumulative indexed bend points up to \(14^{\text {th }}\) interval \((\$ 2,851.64)\). The AIME for this individual in \(15^{\text {th }}\) interval is \(\$ 148.36\)
- \(\quad \operatorname{AIME}_{\mathrm{n}}\) (150000) for interval 1 through14 equals \(\mathrm{bp}_{\mathrm{n}}\) (150000) for the corresponding intervals, such that \(\mathrm{PAP}_{\mathrm{n}}=\mathrm{AIME}_{\mathrm{n}} / \mathrm{bp}_{\mathrm{n}}=1\) for these intervals
- \(\quad \operatorname{AIME}_{15}(150000)=\$ 148.36\), such that PAP \(_{15}=\$ 148.36 / \$ 416.30=\) 0.3564
- \(\quad \operatorname{AIME}_{\mathrm{n}}\) (150000) for interval 16 through 30 equals 0 , such that \(\mathrm{PAP}_{\mathrm{n}}=0\) for these intervals.
- \(\quad\) The following table details these results.
\begin{tabular}{|c|c|c|c|c|c|}
\hline n & \[
\begin{array}{r}
\mathrm{bp}_{n} \text { in } \\
1979 \\
\text { dollars }
\end{array}
\] & \[
\begin{array}{r}
\mathrm{bp}_{n}(\mathrm{r}) \text { in } \\
2011 \\
\text { dollars }
\end{array}
\] & \[
\begin{array}{r}
\sum_{k=1}^{n} \mathrm{bp}_{\mathrm{n}}(r) \\
\text { in } 2011 \\
\text { dollars }
\end{array}
\] & \[
\begin{array}{r}
\text { AIME }_{n}(r) \\
\text { in } 2011 \\
\text { dollars }
\end{array}
\] & \[
\begin{array}{r}
\text { PAP }_{n} \\
\text { in } 2011
\end{array}
\]
dollars \\
\hline 1 & \$45 & \$187.33 & \$187.33 & \$187.33 & 1 \\
\hline 2 & 45 & 187.33 & 374.67 & 187.33 & 1 \\
\hline 3 & 45 & 187.33 & 562.00 & 187.33 & 1 \\
\hline 4 & 45 & 187.33 & 749.34 & 187.33 & 1 \\
\hline 5 & 45 & 187.33 & 936.67 & 187.33 & 1 \\
\hline 6 & 45 & 187.33 & 1,124.00 & 187.33 & 1 \\
\hline 7 & 45 & 187.33 & 1,311.34 & 187.33 & 1 \\
\hline 8 & 45 & 187.33 & 1,498.67 & 187.33 & 1 \\
\hline 9 & 45 & 187.33 & 1,686.01 & 187.33 & 1 \\
\hline 10 & 45 & 187.33 & 1,873.34 & 187.33 & 1 \\
\hline 11 & 45 & 187.33 & 2,060.67 & 187.33 & 1 \\
\hline 12 & 45 & 187.33 & 2,248.01 & 187.33 & 1 \\
\hline 13 & 45 & 187.33 & 2,435.34 & 187.33 & , \\
\hline 14 & 100 & 416.30 & 2,851.64 & 416.30 & 1 \\
\hline 15 & 100 & 416.30 & 3,267.94 & 148.36 & 0.3564 \\
\hline 16 & 100 & 416.30 & 3,684.24 & 0 & 0 \\
\hline 17 & 100 & 416.30 & 4,100.53 & 0 & 0 \\
\hline 18 & 100 & 416.30 & 4,516.83 & 0 & 0 \\
\hline 19 & 200 & 832.60 & 5,349.43 & 0 & 0 \\
\hline 20 & 200 & 832.60 & 6,182.02 & 0 & 0 \\
\hline 21 & 200 & 832.60 & 7,014.62 & 0 & 0 \\
\hline 22 & 200 & 832.60 & 7,847.22 & 0 & 0 \\
\hline 23 & 200 & 832.60 & 8,679.81 & 0 & 0 \\
\hline 24 & 200 & 832.60 & 9,512.41 & 0 & 0 \\
\hline 25 & 200 & 832.60 & 10,345.00 & 0 & 0 \\
\hline 26 & 200 & 832.60 & 11,177.60 & 0 & 0 \\
\hline 27 & 200 & 832.60 & 12,010.20 & 0 & 0 \\
\hline 28 & 200 & 832.60 & 12,842.79 & 0 & 0 \\
\hline 29 & 1000 & 4,162.98 & 17,005.77 & 0 & 0 \\
\hline 30 & 1000 & 4,162.98 & 21,168.75 & 0 & 0 \\
\hline
\end{tabular}

\subsection*{4.3. Cost}

\section*{4.3.a. Overview}

The COST subprocess projects the trust fund operations for each year of the long-range 75 -year period. The COST subprocess projects the income and cost for each trust fund (OASI and DI). The two components of non-interest income are payroll tax contributions and taxation of benefits. \({ }^{1}\) The other component of income is interest earned on the trust fund assets. The three components of cost are scheduled benefits, administrative expenses, and the railroad interchange. Each of these components is projected for each trust fund (OASI and DI). The end-of-year assets is computed by taking the beginning-of-year assets (ASSETS), adding payroll contributions (CONTRIB), taxation of benefits (TAXBEN), and interest income (INT), and subtracting scheduled benefits (BEN), administrative expenses \((A D M)\), and the railroad interchange \((R R)\).

Equations 4.3.1 through 4.3.6 outline this overall structure and sequence.
\[
\begin{align*}
\text { CONTRIB } & =\operatorname{CONTRIB}(\cdot)  \tag{4.3.1}\\
\text { BEN } & =\operatorname{BEN}(\cdot)  \tag{4.3.2}\\
\text { TAXBEN } & =\operatorname{TAXBEN}(\cdot)  \tag{4.3.3}\\
\text { ADM } & =\operatorname{ADM}(\cdot)  \tag{4.3.4}\\
R R & =R R(\cdot)  \tag{4.3.5}\\
I N T & =\operatorname{INT}(\cdot)  \tag{4.3.6}\\
\text { ASSETS }_{\text {EOY }} & =\text { ASSETS }_{\text {BOY }}+\text { CONTRIB }+ \text { TAXBEN }+ \text { INT }- \text { BEN }- \text { ADM }- \text { RR }
\end{align*}
\]

The COST subprocess produces annual values that help assess the financial status of the OASI, DI, and combined funds. These include the annual income rate (ANN_INC_RT), annual cost rate (ANN_COST_RT), and trust fund ratio (TFR) as outlined below.
```

ANN_INC_RT = ANN_INC_RT (')
ANN_COST_RT = ANN_COST_RT(`)
TFR = TFR(\cdot)
ANN_COST_RT = ANN_COST_RT ( $\cdot$ )
$T F R \quad=\quad \operatorname{TFR}(\cdot)$

```

The COST subprocess also produces summarized values. These values are computed for the entire 75 -year projection periods, as well as 25 - and 50 -year periods. These include

\footnotetext{
\({ }^{1}\) As noted throughout the Trustees Report a third component of non-interest income is reimbursements from the General Fund of the Treasury. In the cost program such reimbursements are simply treated as payroll tax contributions.
}
the actuarial balance (ACT_BAL), unfunded obligation (UNF_OBL), summarized income rate (SUMM_INC_RT), summarized cost rate (SUMM_COST_RT), and closed group unfunded obligation (CLOSEDGRP_UNFOBL).
```

ACT_BAL = ACT_BAL(\cdot)
UNF_OBL = UNF_OBL(`) SUMM_INC_RT = SUMM_INC_RT(\cdot) SUMM_COST_RT = SUMM_COST_RT(`)
CLOSEDGRP_UNFOBL = CLOSEDGRP_UNFOBL(`)
SUMM_COST_RT = SUMM_COST_RT(•)
CLOSEDGRP_UNFOBL = CLOSEDGRP_UNFOBL $(\cdot)$

```

The following notation is used throughout this documentation:
- ni represents the first year of the projection period-2016 for the 2016 TR
- ni+74 represents the final year of the projection period-2090 for the 2016 TR
- \(n f\) represents the last year the cost program will project-2095 for the 2016 TR
- nim1 is equal to ni-1
- nim2 is equal to ni-2
- \(n s\) is equal to \(n i+9\)
- nbase, the year of the sample, is equal to 2013

\section*{4.3.b. Input Data}

Data received as input from the short-range office are presented first. Then data from long range and all other sources are identified separately for each equation.

\section*{Short-range OCACT Data}
- Estimates for the first ten years of the projection period for the first six equations mentioned above.
- Assets at the beginning of year ni.

All of this information is updated annually.

\section*{Long-range OCACT and other Data}
i. Equation 4.3.1 - Tax Contributions (CONTRIB)

Economics-Process 2
- \(\quad\) Projected effective taxable payroll for years nim1 through \(n f\), updated yearly

Other
- Projected employee/employer payroll tax rate, by trust fund and year, for years 1981 through \(n f\), updated as needed (e.g., as required due to legislative changes)

\section*{ii. Equation 4.3.2 - Scheduled Benefits (BEN)}

\section*{Demography-Process 1}
- Projected number of married and divorced people in the Social Security area population by age for end of years nim1 through 2100, updated yearly

\section*{Economics-Process 2}
- Historical COLA for year nim2, updated yearly
- Historical CPI for years 1990 through nim1, updated yearly
- Projected cost of living adjustment (COLA) for years nim1 through nf, updated yearly
- Historical SSA average wage index for year nim2, updated yearly
- Projected percent increases in the average wage index for years nim1 through \(n f\), updated yearly

\section*{Beneficiaries-Process 3}
- Initial and ultimate incurred but not reported (IBNR) DI beneficiary distribution by age, duration of disability ( 0 through 9 years and 10+ years) and sex, updated every few years (subprocess \#3.2). Factors are read in such that there is a ten-year linear phase-in from initial factors to ultimate factors. The number of DI beneficiaries in current-payment status is equal to the number of currently entitled DI workers times the IBNR factor.
- Projected number of disabled workers in current-pay status by sex, age in current-pay, and duration of disability ( 0 through 9 and 10+) for years nim1 through \(n f\), updated yearly from subprocess 3.2
- Projected number of retired worker beneficiaries in current-pay status by sex, age in current-pay, and age at entitlement for years nim1 through \(n f\), updated yearly from subprocess 3.3
- Projected number of auxiliary beneficiaries (by benefit category) of retiredworker, deceased-worker, and disabled-worker beneficiaries for years nim1 through \(n f\), updated yearly from subprocess 3.3
- Projected number of disability insurance beneficiaries who convert to retirement insurance status upon the attainment of normal retirement age by age in current pay, for years nim1 through \(n f\), updated yearly from subprocess 3.2 and 3.3
- Retired Workers 65+ by sex, and marital status (single, married, widowed, and divorced)
- Retired Workers 62+ by sex, and marital status (single, married, widowed, and divorced)

Other
- Total (aggregate) PIA and MBA, not actuarially reduced, of DI male and female workers in current payment status, updated yearly from the Table 1-A Supplement
- Total (aggregate) PIA and MBA, actuarially reduced, of DI male and female workers in current payment status, updated yearly from the Table 1-A

Supplement
- Total (aggregate) PIA and MBA, not actuarially reduced, of newly awarded DI male and female workers, updated yearly from the Table 1-A Supplement
- Cumulative distribution of AIME dollars for newly entitled retired-worker beneficiaries by age (62 through 70) and sex, for years nim1 through nf, updated yearly from subprocess 4.2
- Cumulative distribution of AIME dollars for newly entitled disabled-worker beneficiaries by age (20 through 65) and sex, for years nim1 through \(n f\), updated yearly from subprocess 4.2. Ages 15 through 19 are assumed to have the same distribution of dollars as does age 20. Future age 66 disabled workers are assumed to have the same distribution of dollars as age 65 workers do
- Age distribution of newly entitled retired-worker beneficiaries in the AIME awards sample by sex, updated in years that the sample changes, from subprocess 4.2
- Starting average PIA matrix for retired-worker benefits for the year nim1, by age at entitlement, age in current-pay and sex, updated yearly
- Starting average PIA matrix for disabled-worker benefits, for the year nim1, by age in current-pay, duration and sex, updated yearly
- Starting average PIA matrix for beneficiaries who convert from disabled worker to retirement worker status for the year nim1, by age in current-pay and sex, updated yearly
- Starting average MBA matrix for retired-worker benefits for the year nim1, by age at entitlement, age in current-pay and sex, updated yearly
- Starting average MBA matrix for beneficiaries who convert from disabled worker to retirement worker status for the year nim1, by age in current-pay and sex, updated yearly
- Benefit relationships between worker and auxiliary benefits (linkages) for the year nim1, for all benefit categories and worker account holders of both sexes, updated yearly from qlink16.xls
- Benefit relationships between workers and aged spouses for years nim1 through \(n f\) with the effect of the 'Bipartisan Budget Act of 2015' on "claiming strategies" taken into account
- Retroactive payment loading factors for auxiliary beneficiary categories for all years, for each benefit category and both sexes, updated yearly
- Initial and ultimate post entitlement factors for retired workers by sex and duration updated yearly. Factors are read in such that there is a twenty-year linear phase-in from initial factors to ultimate factors.
- Initial and ultimate post entitlement factors for disabled workers by sex and duration updated yearly. Factors are read in such that there is a twenty-year linear phase-in from initial factors to ultimate factors.
- Initial and ultimate post entitlement factors for DI conversion workers by sex and duration updated yearly. Factors are read in such that there is a twentyyear linear phase-in from initial factors to ultimate factors.
- Average excess MBA amounts for dually entitled wives and widows, updated
yearly from Statistical Supplement
- Initial Windfall Elimination Provision (WEP) factors by sex and age for attributed year, updated every 2 years
- Ultimate WEP factors by sex read in as a percentage of the way from initial factors to one, updated every 2 years
- Year in which ultimate WEP factor is reached by age at initial entitlement (62-70), updated every 2 years
- Trendline by which WEP factors are phased-in from initial value to ultimate value, updated every 2 years
- Workers' Compensation cumulative factors by duration that adjust benefits to account for decreasing offsets (i.e. - Workers Comp offsets decrease as duration increases), updated yearly
- Workers' Compensation reduction factors (used in retroactive category) to reflect offsets starting and stopping in the year of DI entitlement
- Workers' Compensation parameter to account for offsets that begin and end in the year of entitlement, updated yearly
- Dual entitlement regression coefficients for the number of wives, husbands, widows, and widowers, and the average excess amounts for wives and widows, updated yearly
- Dual entitlement widower excess amount as a percent of widow excess amount
- Dual entitlement husband excess amount as a percent of wife excess amount
- Dual entitlement average excess amounts and percentages of exposure population for December, year ni-1 for wives, husbands, widows, and widowers, updated yearly
- Target values for ratios relating to the above dual entitlement categories, updated yearly
- Number of years in which the difference between the results from the regression coefficients and targeted values are phased in for the five above dual entitlement categories, updated yearly
- Adjustment factors for average retired and disabled worker benefit amounts (PIA and MBA) in current-pay at durations 0 and 1, by sex and age, updated yearly
- Number of months retroactive benefits are received by a worker who is paid retroactively in their year of entitlement, by trust fund
- Adjustment factors for average retired worker retroactive benefit amounts by sex, updated yearly
- Adjustment factors for disabled worker retroactive benefit amounts in currentpay at durations 0,1 and 2+, by sex and age, updated yearly
- Adjustment factors to apply to long-range average benefit levels in the first 20 years of the projection period by sex and trust fund, updated yearly
- Average retired worker PIA in the \(75^{\text {th }}\) year of the projection period by sex, updated yearly

\section*{iii. Equation 4.3.3 - Taxation of Benefits}

\section*{Trust Fund Operations and Actuarial Status}
- Taxation of benefits as a percentage of scheduled benefits by trust fund for years nim1 through nf, updated yearly from subprocess \#4.1
iv. Equation 4.3.4 - Administrative Expenses

\section*{Economics-Process 2}
- Average wage indexes for years nim1 through nf, updated yearly
- Ultimate value of productivity factor for the period ni through nf updated yearly

\section*{Beneficiaries-Process 3}
- Total number of beneficiaries in current-pay status by trust fund for years nim1 through \(n f\), updated yearly

\section*{v. Equation 4.3.5 - Railroad Interchange}

\section*{Economics-Process 2}
- Increase in the average wage index for years nim1 through \(n f\), updated yearly
- Ultimate value of productivity factor for the period ni through nf updated yearly

\section*{Trust Fund Operations and Actuarial Status}
- Taxation of benefits as a percent of the amount of benefits scheduled to be paid, by trust fund for years nim1 through \(n f\), updated yearly (use same factors as in equation 4.3.3)

\section*{Other input data}
- Nominal annual yield rate on the combined OASDI trust fund for years nim1 through \(n f\), updated as the ultimate real interest rate and ultimate CPI are changed by the Trustees
- Regression coefficients to project annual prescribed interest rates, related to railroad interchange, updated annually
- Ratio of railroad retirement OASI and DI average benefits to overall OASI and DI average benefits, updated yearly
- Number of railroad beneficiaries (retirement and disability) for December of year nim3, updated yearly
- Average taxable earnings in railroad employment for year nim2, updated yearly
- Historical data on average railroad employment, 1960 through nim2
- Average worker benefit by sex and trust fund, updated yearly from the December nim1 Table I-A Supplement
- Expected railroad new awards as a percent of the average of historical employment data
- Auxiliary loading factor by trust fund, updated yearly using 10 years of historical financial interchange benefit data
- Railroad initial mortality rate by trust fund for year nim2, updated yearly using 10 years of historical financial interchange benefit data
- Railroad mortality improvement rate by trust fund, updated yearly using 10 years of historical financial interchange benefit data
- Fiscal Year Railroad transfer amount in millions of dollars for year nim2
vi. Equation 4.3.6 - Interest Income

\section*{Economics-Process 2}
- Annual increase in the CPI for years ni through nf, updated yearly

\section*{Trustees assumptions}
- Ultimate real interest rate, updated annually

\section*{Other input data}
- Factors for exposure to interest rate for benefits, payroll, and taxation of benefits, updated yearly
- Factors for exposure to railroad interchange and administrative expenses, updated periodically
vii. Equations 4.3.7 through 4.3.13 - Annual Values and Summarized Values

All inputs for equations 4.3 .7 through 4.3.13 are estimated internally in the Cost program.
viii. Equation 4.3.14 - Closed Group Unfunded Obligation

\section*{Demographics-Process 1}
- Single year population and mortality rate data for years 1941 through 2101, updated yearly

\section*{Economics-Process 2}
- Historical and projected single-year COLA data and average wage indexing series (AWI) data for years 1975 through 2095 (for COLA) and 1951 through 2095 (AWI), updated yearly
- Historical and projected number of covered workers by single year of age 099 and 100+ from year ni-21 through 2095, updated yearly
- Projected number of covered workers by single year of age \(100+\) from year ni through 2095, derived from total age 100+ projections from year ni-21 through 2095
- Ultimate assumed annual average wage increase, wg_ult, updated yearly

\section*{Beneficiaries-Process 3}
- Total projected disabled workers by age for years nim1 to 2095, updated yearly
- Total projected aged spouses, divorced aged spouses, surviving aged spouses and divorced surviving aged spouses by sex, single year of age (up to 95+) and for years nim1 to 2095.

\section*{Awards-Process 4}
- Projected number of workers and total taxable earnings by single year of age (15-80) and sex from nim1 to 2095, updated yearly

\section*{Other}
- Total count of beneficiaries in 20 of the 28 beneficiary categories (excluding retired workers, disabled workers, aged spouses (married, divorced, and dually entitled excess), and aged widow(er)s (married, divorced, and dually entitled excess) from the December 2015 Master Beneficiary Record (MBR) \({ }^{2}\) —updated yearly
- Total benefits paid in 20 of the 28 beneficiary categories (excluding retired workers, disabled workers, aged spouses (married, divorced, and dually entitled excess), and aged widow(er)s (married, divorced, and dually entitled excess) from the December 2015 MBR—updated yearly
- Consumer Price Index data from 1951-1974 from Bureau of Labor Statistics
- Number of covered workers and average taxable earnings by single year of age 1-14 for years 1993-2012 from 1 percent Continuous Work History Sample (CWHS), updated yearly to include year ni-4
- Number of covered workers and average taxable earnings by single year of age 75-99 for years 1993-2012 from 1 percent CWHS, updated yearly to include year ni-4
- Distribution of assumed age differentials between aged spouses and workers ranging from -12 years to 15 years seniority for the worker by sex of the beneficiary, age of the beneficiary 62-74, and marital status obtained from the December 2012 MBR—updated every 3 years
- Distribution of assumed age differentials between widow(er)s and workers ranging from -6 years to 12 years seniority for the male (generally not updated)
- Factors to apply to the 95+ "in current pay" counts of retired workers, aged spouses, surviving aged spouses, divorced spouses, and dually entitled spouses expanding the single age counts through 119, updated yearly

\section*{4.3.c. Development of Output}

\section*{i. Equation 4.3.1 - Payroll Tax Contributions (CONTRIB)}

\footnotetext{
\({ }^{2}\) For disabled adult children of deceased workers and lump-sum beneficiaries, data were extracted from a 1- percent sample of the December 2015 MBR, mainframe dataset ACT.TAPEL.CAN1215. For the other 18 auxiliary beneficiary categories, data was extracted from the 100 percent December 2015 MBR, mainframe dataset ACT.TAPEH.MBR100.D1512.CANSORT.
}

It would be natural to estimate the payroll tax contributions by trust fund by multiplying the applicable employer/employee tax rate by effective taxable payroll. However, tax contributions are reported on a cash basis. That is, tax contribution amounts are attributed to the year in which they are actually received by the trust funds, while taxable payroll is attributed to the year in which earnings are paid. In other words, the lag between the time the tax liability is incurred and when the taxes are actually collected must be reflected. If lag represents the proportion of incurred payroll taxes estimated to be received by the trust fund ( \(t f\) ) in year \(y r\), then tax contributions (CONTRIB) are given by the formula
\[
\begin{aligned}
\operatorname{CONTRIB}(t f, y r)= & \operatorname{lag} \times \operatorname{tax} \text { rate }(t f, y r) \times \operatorname{payroll}(y r) \\
& +(1-\operatorname{lag}) \times \operatorname{tax} \text { rate }(t f, y r-1) \times \operatorname{payroll}(y r-1)
\end{aligned}
\]
for \(y r \geq n s\).

The value of lag is estimated from the combined OASI and DI tax contributions estimated to be collected in the final year of the short-range period, \(n s\), and is given by
\[
\operatorname{lag}=\frac{\sum_{t f=1}^{2} \operatorname{CONTRIB}(t f, n s)-\sum_{t f=1}^{2}(\operatorname{taxate}(t f, n s-1) \times \operatorname{payroll}(n s-1))}{\sum_{t f=1}^{2}(\text { taxrate }(t f, n s) \times \operatorname{payroll}(n s)-\operatorname{taxrate}(t f, n s-1) \times \operatorname{payroll}(n s-1))} .
\]

For the first ten years of the long-range period, tax contributions are set equal to those provided by the short-range office. The same value of lag is used for all years, and both trust funds, thereafter.

\section*{ii. Equation 4.3.2 - Scheduled Benefits (BEN)}
(1) Disabled-Worker Benefits

\section*{Disabled Worker Beneficiary Matrix}

The number of disabled-worker beneficiaries for a given year and sex is provided from the subprocess 3.2. For each projection year, two matrices are provided - one for males and one for females. The structure of each matrix is as follows:
- 11 columns. The columns are indexed by duration of disability (0-9 and \(10+\) ).
- 52 rows. These rows correspond to the age in current pay, ages 15 through 66.

The COST subprocess, however, only uses 10 durations ( \(0-8\) and \(9+\) ), and 47 ages (ages 20 through 66). This requires a manipulation of the matrix of DI beneficiaries in current-pay status from subprocess 3.2. For ages in current pay greater than or equal to 30 , the duration 9 and \(10+\) columns of this matrix are added to give the total
number of duration 9+ beneficiaries. For ages (ag) between 20 and 30 inclusive, the number of beneficiaries in current-pay aged \(a g\) and duration \(a g-20\) is the value provided by the DISABILITY subprocess added to the number of people in current pay aged \(a g-j\) and duration \(a g-20\) for \(j=1, . ., 5\). (For example, the number of people aged 20 of duration 0 is combined with the number of people aged 15, 16, 17, 18 and 19 of duration 0 ; the number of people aged 21 of duration 1 is combined with the number of beneficiaries in current-payment status aged 16, 17, 18, 19, and 20 of duration 1, and so on. In other words, the five nonzero diagonal of the matrix provided by the DISABILITY subprocess is "combined with" the diagonal directly below it and then zeroed out.)

\section*{Building the Average Benefit Matrix for Disabled Workers}

In each projection year, the COST subprocess produces an average benefit matrix for each sex. Each matrix is a 47 by 10 matrix whose entries are the average benefit amounts of disabled worker beneficiaries whose age in current pay is indexed by the rows (ages 20 through 66) and whose duration of disability is indexed by the columns (durations 0 through 8 and 9+).

The 100 percent Master Beneficiary Record (100\% MBR) extract is processed by a side model. The final product of the side model is two matrices of average benefit levels, one for males and one for females, for December nim1 (2015 for the 2016 TR).

For a given year of the projection period, a new average benefit matrix is obtained by moving the average benefit matrix from the previous year one year forward. The next few paragraphs describe this procedure.

In general, for each age in current-pay, the age and duration are incremented by 1 and the previous PIA amount is given a cost of living adjustment. In addition, the beneficiaries are given a workers' compensation adjustment and a post-entitlement adjustment. For each duration \(j=0,1, \ldots, 7\), and \(8+\) and sex, let the workers' compensation offset factor be denoted \(w k \operatorname{comp}(y r, s x, d u r)\). We have, for durations 0 through 8, that
\[
\begin{aligned}
\operatorname{avgmba}(y r, s x, a g, d u r)= & \operatorname{avgmba}(y r-1, s x, a g-1, d u r-1) \times(1+\operatorname{COLA}(y r)) \\
& \times(1+\operatorname{wkcomp}(y r, s x, d u r)) \times \operatorname{PEadj}(y r, s e x, d u r) .
\end{aligned}
\]

A more careful explanation of the factors, \(w k \operatorname{comp}(y r, s x, d u r)\) and PEadj(yr, sex, dur), is given later in this document. See the section titled Average PIAs and MBAs for Disabled-Worker Beneficiaries, below. To move duration 8 average benefits to duration 9+ average benefits, both average benefits are given a cost of living adjustment and a post-entitlement adjustment (see section "PostEntitlement Adjustments"). The resulting duration 9+ average benefit is the weighted average of the adjusted prior year duration 8 and 9+ average benefits, weighted by the
prior year's numbers of beneficiaries in current-pay status for durations 8 and 9+ respectively.

The only column that does not follow this procedure is the duration 0 column. The duration 0 column corresponds to newly entitled disabled-worker beneficiaries. The following sections describe how average benefits are determined for this group of beneficiaries.

\section*{Average Benefits for Newly Entitled Disabled-Worker Beneficiaries}

The potential AIME percentage values for newly entitled disabled-worker benefits (DPAPs) are obtained from the AWARDS subprocess. The two bendpoints of the PIA formula, BP1 and BP2, are indexed by the increase in the average wage index. In 1979 dollars, the values of BP1 and BP2 are \(\$ 180\) and \(\$ 1,085\) respectively. The AIME dollars between 0 and BP1 are divided into four intervals (each of length \(\$ 45\) in 1979 dollars). The AIME dollars between BP1 and BP2 are divided into fourteen intervals (nine of length \(\$ 45\) and five of length \(\$ 100\), in 1979 dollars). Twelve additional intervals are added beyond BP2 (ten of length \(\$ 200\) and two of length \(\$ 1,000\), in 1979 dollars).

To determine the average PIA for newly entitled beneficiaries, the DPAP values for each of the thirty intervals of AIME dollars are multiplied by the dollar amounts attributable to each interval (the length of the interval) and by the associated PIA factors. The distribution of prior year disability onset and current year disability onset is taken into consideration. It is assumed that this distribution is 6 months for prior year disability onset and 6 months for current year disability onset. In the formulas below, \(j=1\) signifies current year disability onset and \(j=2\) signifies prior year disability onset.

Let:
- Wage_Idx \((s x, a g, y r)=\frac{\operatorname{avgwg}(y r-\max (a g-60,1+j))}{\operatorname{avgwg}(1977)}\) for \(j=1,2\).
- \(\operatorname{Cum}_{-} \operatorname{COLA}_{1}(a g, y r)=\left\{\begin{array}{cc}(1+\operatorname{COLA}(y r-1)) \times(1+\operatorname{COLA}(y r)) & a g<64 \\ \prod_{k=62}^{a g}(1+\operatorname{COLA}(y r-(k-62)) & 64 \leq a g \leq 66 .\end{array}\right.\)
\(\operatorname{Cum}_{-} \operatorname{COLA}_{2}(a g, y r)=\left\{\begin{array}{cc}1+\operatorname{COLA}(y r) & a g<63 \\ \prod_{k=63}^{a g}(1+\operatorname{COLA}(y r-(k-62)) & 63 \leq a g \leq 66 .\end{array}\right.\)
- \(w_{j}=\frac{6}{12}=\frac{1}{2}, j=1,2\).
- PIA_factor \({ }_{i}\) represent the PIA factor for interval \(i\) (equal to 0.90 for intervals \(i=1 . ., 4,0.32\) for intervals \(i=5, \ldots, 18\), and 0.15 for intervals \(i=19, \ldots, 30\) ).
- AIME_dollars \({ }_{i}\) represent the length of interval \(i\), expressed in 1979 dollars.
- \(\operatorname{dpap}_{i}(y r, s x, a g)\) represent the DPAP value for newly entitled disabled
workers in year \(y r\) whose sex is \(s x\) and age is \(a g\).
To take into account the workers' compensation offset to disability benefits, administrative data is reviewed, from which a factor is developed and applied to the average award benefit. We now describe how this factor, facm2p(yr,sx) is computed. The table 1-A supplement, for each month in 2015, contains total award PIA and MBA data for disabled workers, by sex, for beneficiaries both nonactuarially reduced and actuarially reduced. Let totmba_DIB_nar ( \(y r, s x\) ) and totpia _DIB_nar ( \(y r, s x\) ) be the total annual MBA and PIA respectively for DIBs that are not actuarially reduced as found in the table 1-A data. In the historical period 2000-nim1 we define \(f a c m 2 p(y r, s x)\) to be the ratio of the total MBA to the total PIA for those not actuarially reduced. In other words,
\[
\operatorname{facm} 2 p(y r, s x)=\frac{\text { totmba_DIB_nar }(y r, s x)}{\text { totpia_DIB_nar }(y r, s x)} .
\]

In the period ni through \(n s+9\), facm2p(yr,sx) is defined as follows. Let
\[
\begin{aligned}
& y 1=f a c m 2 p(n i m 1, \text { sex }) \\
& y 2=(1.0-y 1) / 3.0 \\
& \text { facm } 2 p(y r, s x)=\max \left(y 1-y 2, \min \left(y 1+y 2, f a c m 2 p(y r-1, s x) \times\left(\frac{f a c m 2 p(y r-1, s x)}{f a c m 2 p(y r-11, s x)}\right)^{1 / 20}\right)\right) .
\end{aligned}
\]

Projected values of facm2p are therefore held within a delta of y2 from the last historical year of facm2p.
This value is further adjusted by the variable facm2p_param to reflect the offset amounts that end within the first entitlement year. For the 2016 TR the data suggests this factor should be .34. As a result, for \(y r=n i, \ldots, n s+9\),
\[
\begin{aligned}
\text { facm } 2 p(y r, s x) & =x+(1-x) \times 0.34 \\
& =0.34+0.66 \times x .
\end{aligned}
\]

The factor reaches its ultimate value in years \(n s+10\) and later.
The preliminary average PIA for newly entitled disabled worker beneficiaries may now be defined. It is equal to
\[
\begin{aligned}
L R_{-} \operatorname{awdpia}(s x, a g, y r)=\sum_{i=1}^{30} P I A_{-} & \text {factor }_{i} \times \text { AIME_dollars }_{i} \times \operatorname{dpap}_{i}(y r, s x, a g) \\
& \times\left(\sum_{j=1}^{2} w_{j} \times \text { Wage_Idx }_{j}(s x, a g, y r) \times C u m_{-} C O L A_{j}(a g, y r)\right) \\
& \times f a c m 2 p(y r, s x) .
\end{aligned}
\]

Once these average PIAs of newly entitled disabled-worker beneficiaries are computed, their values are filled into the average PIA matrices for duration 0 for the
appropriate entitlement age.

\section*{Average PIAs and MBAs for Disabled-Worker Beneficiaries}

An overall average PIA of newly entitled disabled worker beneficiaries for each sex and projection year is computed by taking the weighted average of awdpia(sx,ag,yr), the weights being the number of disabled workers in current payment status of duration zero. This value is denoted awdpia(sx,yr).

In addition, an overall average PIA and MBA for all disabled worker beneficiaries in current-payment status is computed by finding the weighted average of the average PIAs for each age in current-pay and duration with the number of people in current pay for each of these ages and durations. The average PIAs were already reduced by a workers' compensation offset factor, as briefly described above; a more careful description is given in this section. To get the average MBAs, we apply a factor that reflects the differences in average MBAs and PIAs for disabled workers, isolating only the trend in cases with an actuarial reduction. We also provide a relatively small reduction to reflect offsets starting and stopping in the year of DI entitlement that are not captured by the current method. There is an additional adjustment to the weighted average disabled worker PIA and MBA amounts in current pay at durations 0 and 1 to account for benefit level differences that are phased out by duration 2 .

\section*{Workers' Compensation Offset Factors}

For each duration \(j=1, \ldots, 7\), and \(8+\) and sex we define a workers' compensation factor. This factor is applied to the average worker PIA matrix as mentioned above. It is denoted \(w k \operatorname{comp}(y r, s x, d u r)\). Let \(f a c m 2 p_{~} p c t(d u r)\) be defined as in the following table.
\begin{tabular}{|c|c|}
\hline Duration & \begin{tabular}{c} 
Cumulative product above set at \(\mathrm{x} \%\) of \\
way between original facm2p and 1
\end{tabular} \\
\hline 1 & 0.487328 \\
\hline 2 & 0.672613 \\
\hline 3 & 0.754520 \\
\hline 4 & 0.805062 \\
\hline 5 & 0.846726 \\
\hline 6 & 0.871565 \\
\hline 7 & 0.890246 \\
\hline \(8+\) & 0.909091 \\
\hline
\end{tabular}

Then \(w k \operatorname{comp}(y r, s x, d u r)\) is defined so that
\[
f a c m 2 p_{-} p c t(d u r)=f a c m 2 p(y r-d u r, s x) \times \prod_{j=1}^{d u r}(1+w k \operatorname{comp}(y r, s x, j))
\]

This is an iterative process that first computes \(w k \operatorname{comp}(y r, s x, 1)\) by solving the above
equation with dur set equal to 1 . The remaining factors for higher durations are then computed recursively using the above formula.

\section*{Trend in Average MBA to Average PIA}

This trend is captured in a factor denoted Fam2p(yr,sx). The table 1-A supplement as of the end of December 2015 contains total in-current pay PIA and MBA data for disabled workers, by sex, for beneficiaries both non-actuarially reduced and actuarially reduced. Let totmba_nar \((y r, s x)\) and totpia_nar \((y r, s x)\) be the total MBA and PIA respectively for DIBs that are not actuarially reduced as found in the table 1A data. Similarly, let totmba_ar \((y r, s x)\) and totpia_ar \((y r, s x)\) be the total MBA and PIA respectively for cases that are actuarially reduced. In the historical period 2000nim1 we define \(f a m 2 p(y r, s x)\) to be the ratio of the total MBA to the total annual PIA for those not actuarially reduced. In other words,
\[
\operatorname{Fam} 2 p(y r, \operatorname{sex})=\frac{\text { totmba_ar }(y r, s e x)+\text { totpia_nar }(y r, s e x)}{\text { totpia_ar }(y r, s e x)+\text { totpia_nar }(y r, s e x)}
\]

In the period \(n i\) through \(n s+10, f a c m 2 p(y r, s x)\) is defined as follows:
\[
\begin{aligned}
& y 1=\operatorname{fam} 2 p(\operatorname{nim1} 1, \text { sex }) \\
& y 2=(1.0-y 1) / 3.0 \\
& \operatorname{fam} 2 p(y r, s x)=\max \left(y 1-y 2, \min \left(y 1+y 2, \operatorname{fam} 2 p(y r-1, s x) \times\left(\frac{\operatorname{fam} 2 p(y r-1, s x)}{f a m 2 p(y r-11, s x)}\right)^{1 / 20}\right)\right) .
\end{aligned}
\]

The factor reaches its ultimate value in years \(n s+10\) and later.

\section*{More Workers' Compensation Offsets}

As mentioned above, we also we provide a relatively small reduction to retroactive benefits to reflect offsets starting and stopping in the year of DI entitlement that are not captured by the current method. Based on historical administrative data, we set these factors by duration as follows:
\begin{tabular}{|c|c|}
\hline Duration & Percentage Reduction \\
\hline 0 & \(0.2406 \%\) \\
\hline 1 & \(0.1549 \%\) \\
\hline 2 & \(0.0503 \%\) \\
\hline 3 & \(0.0237 \%\) \\
\hline 4 & \(0.0051 \%\) \\
\hline \(5+\) & \(0 \%\) \\
\hline
\end{tabular}

We define wkcomp_red(dur) to be 1 minus these percentage reductions.
By law, disabled workers are no longer subject to the workers' compensation offset at the attainment of a defined age. For those born prior to 1951, the defined age is 65. For those born in 1951 and later, the defined age is the Normal Retirement Age
(NRA). \({ }^{3}\) Therefore, all DI worker benefit levels are adjusted at either age 65 (for those born prior to 1951) or at conversion to a retired worker benefit (for those born 1951 and later) to eliminate the effect of the offset.

\section*{Adjustment to Average Benefit Levels at Durations 0 and 1}

Average disabled worker PIA and MBAs are adjusted further at durations 0 and 1 by a factor, DI_RI_facs, designed to account for average benefit level changes that occur when beneficiaries come on the rolls retroactively. This is different than projected retroactive payments discussed elsewhere in the documentation. These adjustments are phased out by duration 2 , when the vast majority of disabled workers have started receiving benefits.

\section*{Computation of Average MBA for DI Workers}

The disabled worker PIA as presented in the average benefit matrix was already incremented by age and duration using a COLA and a workers' compensation adjustment. The average PIA by year, age and duration, is denoted \(\operatorname{avgpia}(y r, a g, s x, d u r)\). The overall average MBA by year and sex is the weighted average of \(\operatorname{avgpia}(y r, a g, s x, d u r) \times \operatorname{Fam} 2 p(y r, s x, d u r)\), the weights being the number of DI workers in current payment status by age, sex, and duration.

\section*{Post-Entitlement Adjustments}

As cohorts of beneficiaries age, their average benefit level will likely change for reasons other than just the COLA increase. The two primary reasons for this are postentitlement work, which could lead to a re-calculation of one's benefit, and a known correlation between greater lifetime earnings and lower mortality rates. The Cost process uses post-entitlement factors by sex and duration to account for the expected dynamic benefit levels. One percent December MBR data from the most recent 10 historical years are used to calculate post-entitlement factors.

For disabled workers we calculate separate factors for those in current pay (ICP) who are younger than 50 and those ICP who are age 50 or older. We use separate factors for each sex and each duration \(\left(0-9^{+}\right)\). Initial and Ultimate factors are calculated before being read into the cost program with initial factors set to the most recent 3year historical average and ultimate factors at the most recent 10-year average.

All initial factors are phased-in linearly to the ultimate factors over the first 20 years of the projection period (reaching the ultimate values in ni+19).
(2) Retired-Worker Benefits

\footnotetext{
\({ }^{3}\) The NRA was 65 for individuals born before 1938. It increased to 66 at the rate of 2 months per year for individuals born 1938-1943. Under current law, the NRA will increase to 67 for individuals born from 1955-1960, again at the rate of 2 months per year.
}

The number of retired-worker beneficiaries for a given year and sex is provided from subprocess 3.3. Two matrices are provided - one for males and one for females. The structure of each matrix is as follows:
- 10 columns. The first 9 columns are the age at entitlement, ages 62 through 70. The last column is the number of disabled workers who are projected to convert to retired-worker beneficiary status (DI conversions) at normal retirement age.
- 34 rows. These rows correspond to the age in current pay, ages 62 through 94 and ages 95+.

Note that the entries on the diagonal at ages 62 through 70 (where age in current-pay equals age at entitlement) are the number of new entitlements projected for that year.

\section*{Building the Average Benefit Matrices for Retired Workers}

In each projection year, the COST subprocess produces four average benefit matrices. For each sex there are two matrices, an average monthly benefit amount (MBA) matrix and the average primary insurance amount (PIA) matrix. Each matrix has the same structure as the beneficiary matrices. In other words, each matrix is a 34 by 10 matrix whose entries are the average benefit amounts of retired worker beneficiaries whose age in current pay is indexed by the rows and whose age at initial entitlement is indexed by the columns. The final column simply gives the average benefits for DI conversions at the various ages in current pay.

The \(100 \%\) MBR extract is processed by a side model. This side model computes a starting matrix for year ni-1. This starting matrix contains the four initial benefit matrices, constructed using the most recent data. For a given year of the projection period, the average benefit matrix is updated from its previous year's value incrementing each benefit amount (PIA or MBA) by one year of age and increasing it by a cost of living adjustment (COLA) and by the appropriate post entitlement factor (see section "Post-entitlement adjustments") for males and females. Adjusted age 94 benefits and age 95+ benefits are averaged, based on the respective number of beneficiaries in current pay in the prior year, to get the new average benefit for age \(95+\). DI conversion benefits are handled similarly, except the average conversion benefit for each age 65 through 66 is combined (as a weighted average) between the DI conversion average benefits computed in subprocess 3.2 and the DI conversions of the corresponding age already receiving benefits (provided by subprocess 3.3).

The entries along the diagonal, the average benefits of newly entitled beneficiaries by age, must still be computed. The remainder of this section will explain how these average benefits are computed. Once these are computed, all entries are computed and the average benefit matrix for the year is complete.

\section*{Average Benefits for Newly Entitled Retired Worker Beneficiaries}

The potential AIME percentage (OPAPs) values for newly entitled retired-worker benefits are obtained from subprocess 4.2. The OPAPs for the projection of average benefits for newly entitled retired-worker beneficiaries are modified by the COST subprocess. The reason is that the age distribution of newly entitled retired worker beneficiaries differs between the projection year and the year of the awards sample. To incorporate the change in the age distribution of projected newly entitled retiredworker beneficiaries for determining their average benefits, we use a "shuttling method" that allocates appropriate portions of the OPAPs by different ages to represent a specific age in the projection year. The resulting new potential AIME percentages, by age and sex, are denoted OPAP1. Additional details about the shuttling method are given in Appendix 4.3-1.

Average benefits for newly entitled retired-worker beneficiaries are calculated by sex and single years of age 62 through 69, and ages 70+. The two bendpoints of the PIA formula, BP1 and BP2, are indexed by the increase in the average wage index. In 1979 dollars, the values of BP1 and BP2 are \(\$ 180\) and \(\$ 1,085\) respectively. The AIME dollars between 0 and BP1 are divided into four intervals (each of length \(\$ 45\) in 1979 dollars). The AIME dollars between BP1 and BP2 are divided into fourteen intervals (nine of length \(\$ 45\) and five of length \(\$ 100\), in 1979 dollars). Twelve additional intervals are added beyond BP2 (ten of length \(\$ 200\) and two of length \(\$ 1,000\), in 1979 dollars).

To determine the average PIA for newly entitled beneficiaries, the OPAP1 values for each interval of AIME dollars are multiplied by the dollar amounts attributable to each interval (the length of the interval). More precisely, let
- PIA_factor \({ }_{i}\) be the PIA factor for subinterval \(i\) (equal to 0.90 for intervals \(i=1,4,0.32\) for intervals \(i=4, \ldots, 18\), and 0.15 for intervals \(i=19, \ldots, 30\) ), AIME_dollars \({ }_{i}\) be the length of subinterval \(i\),
- opap \(1_{i}(y r, s x, a g)\) be the modified PAP value for retired workers newly entitled in year \(y r\) whose sex is \(s x\) and whose age is \(a g\).
- \(\quad w f f(y r, s x, a g e)\) be a reduction factor to account for the Windfall Elimination Provision. \({ }^{4}\)
- Wage_Idx \((a g, y r)=\frac{\operatorname{avgwg}(y r-(a g-62))}{\operatorname{avgwg}(n b a s e-2)}\)
- COLA_Idx \((a g, y r)=\prod_{k=62}^{a g}(1 \times \operatorname{COLA}(y r-(k-62)))\)

\footnotetext{
\({ }^{4}\) The Windfall Elimination Provision (WEP) reduces the first PIA formula factor from \(90 \%\) to as low as \(40 \%\) for individuals who receive a pension based on specified categories of non-covered employment, primarily non-covered state and local government employees and federal workers receiving a pension under the Civil Service Retirement System. The cost process uses initial factors by sex and age, ultimate factors, years in which ultimate factors are reached and phase-in trend lines to the ultimate factor, all supplied by a side model.
}

Then the average PIA for these newly entitled retired worker beneficiaries is equal to
\[
\begin{aligned}
& L R_{-} \text {awdpia }(s x, a g, y r)=\text { Wage_Idx }(a g, y r) \times C O L A_{-} I D X(a g, y r) \times w f f(y r, s x, a g e) \\
& \quad \times \sum_{i=1}^{30} \text { PIA_ } \text { factor }_{i} \times \text { AIME _dollars }_{i} \times \operatorname{opap}_{i}(y r, s x, a g) .
\end{aligned}
\]

This formula incorporates the fact that the PAP values are the estimated cumulative distribution of AIME dollars. The average award MBA for a worker beneficiary is then the average newly entitled PIA multiplied by the appropriate actuarial reduction factors and delayed retirement credits, \(\operatorname{arfdrc}(a g, y r)\), based on age at initial entitlement.

Once these average benefits of newly entitled retired-worker beneficiaries are computed, their values are filled into the appropriate average benefit matrices.

For summary purposes, the COST subprocess computes an average PIA and MBA for all male and female newly entitled retired-worker beneficiaries. These are just the respective weighted averages of the average PIAs and MBAs by age and sex, the weights being the number of newly entitled retired-worker beneficiaries. Similarly, average PIA and MBA for all retired worker beneficiaries in current pay are computed, by sex.

\section*{DI Conversions}

Disabled-worker beneficiaries convert to retired-worker beneficiary status (called DI conversions) at normal retirement age (NRA). The average new DI conversion benefit for a given sex at age NRA is the weighted average of the average DI worker benefits from the prior year for that sex and age NRA-1, weighted by the number of people in current pay in each duration and then increased by the current year COLA and adjusted by the appropriate Post-Entitlement factor. The average DI conversion benefit for a given sex and single age NRA+1 through 95+ is the average DI conversion benefit from the previous year for the same sex and age cohort increased by the current year COLA and adjusted by the appropriate Post-Entitlement factor. Both the average conversion benefit for each sex and single age NRA through 95+ and the number of people in current pay for these ages are used in the computation of average retired worker benefits.

\section*{Post-Entitlement Adjustments}

As discussed in the previous section, the Cost process uses post-entitlement factors by sex and duration to account for changes in benefit levels aside from the cost-of-living adjustment. For retired workers we calculate separate factors for those ICP who converted from DI status and those ICP who came on the rolls as a retired worker. We use separate factors for each sex and each duration \(\left(0-12^{+}\right)\). Initial and ultimate factors are calculated before being read into the cost program with initial factors set to
the most recent 3-year historical average and ultimate factors at a 10-year average. For females the ultimate post-entitlement factors are adjusted further to reflect the trend that female retired workers are starting to have earnings and benefit levels more similar to men. Therefore female ultimate post-entitlement factors are calculated in the program as \(90 \%\) of the male 10 -year average plus \(10 \%\) of the female 10 -year average.

All initial factors are phased-in linearly to the ultimate factors over the first 20 years of the projection period (reaching the ultimate values in ni +19 ).

\section*{Adjustment to Average Benefit Levels at Durations 0 and 1}

Average retired worker PIA and MBAs are adjusted at durations 0 and 1 by a factor, OA_RI_facs, designed to account for average benefit level changes that occur when beneficiaries come on the rolls retroactively. This is different than projected retroactive payments discussed elsewhere in the documentation. These adjustments are phased out by duration 2, when the vast majority of retired workers have started receiving benefits.

\section*{(3) Annualizing Benefits}

Scheduled benefits are calculated by trust fund and projection year. For each year, scheduled benefits for each trust fund are found by adding up the appropriate benefit categories.

This section applies to all benefit amounts except the "dual entitlement excess amount." If a retired worker beneficiary is also entitled to auxiliary spouse or widow(er) benefits and these auxiliary benefits are greater, then the amount by which the auxiliary benefit exceeds the worker's MBA is the dual entitlement excess amount. The four categories of excess amounts (dually entitled wives, widows, husbands, and widowers) are projected separately. More information is found in subsection (4).

The first step is to determine average benefits by category. A list of the beneficiary categories follows. An odd category number refers to the male account holder, while an even category number refers to the female account holder. As an example, for category 4, the aged married spouse is the aged married husband of the retired female worker.
\begin{tabular}{lc}
\begin{tabular}{l} 
Category \\
\# (cat)
\end{tabular} & Beneficiary Type \\
\(1 \& 2\) & Old-Age Insurance Beneficiaries \\
\(1 \& 2\) & Retired worker (includes DI conversions) \\
\(3 \& 4\) & Aged married spouse \\
\(5 \& 6\) & Aged divorced spouse \\
\(9 \& 10\) & \(\quad\) Young spouse with child \\
\(11 \& 12\) & Child \(<18\) \\
\(13 \& 14\) & Student child \\
\(15 \& 16\) & Disabled adult child
\end{tabular}
\begin{tabular}{l} 
Disability Insurance Beneficiaries \\
Disabled worker \\
Aged married spouse \\
Aged divorced spouse \\
Young spouse with child \\
Young child \\
Student child \\
Disabled adult child \\
Survivors Insurance Beneficiaries \\
\hline Aged married widow \\
Aged divorced widow \\
Young married disabled widow \\
Young divorced disabled widow \\
Aged parent \\
Young married widow with child \\
Young divorced widow with child \\
Young child \\
Student child \\
Disabled adult child \\
Lump sum death benefit (\$255)
\end{tabular}

For the worker categories, the prior sections describe the computation of average benefit levels at the end of each year. For a specific auxiliary beneficiary category, the average monthly benefit at the end of each year (avgben) is determined by multiplying:
- The linkage factor (the assumed relationship between an auxiliary beneficiary's benefit level and the corresponding worker benefit level) by
- The relevant average PIA or the average monthly benefit of the primary account holder (the worker beneficiary account on which the auxiliary beneficiary is entitled to receive the benefit).

As of the 2016 Trustees Report, the linkage factor for aged spouse categories comes not from the qlink workbook, but from file 'BAsps’. This file is created by a side model with the purpose of accounting for the expected changes to average aged spouse benefit levels due to the Bipartisan Budget Act of 2015.

In order to annualize benefits for each beneficiary category, two values are used. The beginning-of-year average benefit equals the average monthly benefit in December of the prior year. The end-of-year benefit equals the monthly average benefit of the worker beneficiary for December of the current year without the cost of living adjustment (COLA). The average benefit by category for each month is found by taking a weighted average of the benefits at the beginning and end of the year, the weights being the fractions of the year the prior and current year's beneficiaries have been exposed. If \(c p(c a t, y r)\) is the number of beneficiaries in category cat for year \(y r\), and avgben(cat,yr) is the average monthly benefit for category cat for year \(y r\), then the amount of aggregate benefits paid in year \(y r\) is given by the formula:
\[
\begin{aligned}
& \operatorname{AggBen}(y r, c a t) \\
& =\sum_{i=0}^{11}\left[\frac{(12-i)}{12} \times c p(c a t, y r-1) \times \operatorname{avgben}(c a t, y r-1)+\frac{i}{12} \times c p(c a t, y r) \times \operatorname{avgben}(c a t, y r)\right] .
\end{aligned}
\]

For all beneficiary categories expect for the lump-sum benefit, the aggregate benefit amount is increased by the retroactive payments that were projected to be paid during the year. See section (5), below.

\section*{(4) Dually Entitled Beneficiaries and Benefits}

\section*{Number of Dually Entitled Beneficiaries}

There are four categories of dually entitled beneficiaries. They are the dually entitled wives (1), widows (2), widowers (3), and husbands (4). To project the number of dually entitled beneficiaries for each category we combine a regression equation with two coefficients each, a slope of \(a_{1}^{(k)}\) and a y-intercept of \(b^{(k)}\), with a third factor, \(c^{(k)}(y r)\) derived from a process we describe as "add factoring":
\[
\operatorname{PctExp}^{(k)}(y r)=a_{1}^{(k)} \frac{\operatorname{PIA}(\mathrm{yr}, \mathrm{M})-\operatorname{PIA}(\mathrm{yr}, \mathrm{~F})}{\operatorname{PIA}(\mathrm{yr}, \mathrm{M})}+b^{(k)}+c^{(k)}(y r)
\]
( \(\mathrm{k}=1,2,3,4\) ) project the percentage of the exposed population entitled to wife (1), widow (2), widower (3), and husband (4) benefits.

PIA(yr,sex) is the average PIA of all retired worker beneficiaries in current pay by sex, wage-indexed to the year of the sample and \(\operatorname{PctExp}(y r)\) is the percentage of the entitled population in the category that is dually entitled. Only the male (widower and husband) dually entitled regression equations warrant the usage of a non-zero y-intercept. We use the "add factoring" method with variable \(c^{(k)}(y r)\) to account for the expected future comparative work history changes that will affect dual entitlement populations.

To derive \(c^{(k)}(y r)\), suppose that \(u l t^{(k)}\) is the value obtained from the regression equation without add-factoring in the final year of the projection period. Therefore
\[
u l t^{(k)}=a_{1}^{(k)} \frac{\operatorname{PIA}(n i+74, \mathrm{M})-\mathrm{PIA}(n i+74, \mathrm{~F})}{\operatorname{PIA}(n i+74, \mathrm{M})}+b^{(k)}
\]

Let \(\operatorname{targ}{ }^{(k)}\) be the target value we estimate for the final year of the projection period. Let phaseyrs be the number of years it takes to fully phase in the target value. Then we have
\[
c^{(k)}(y r)=\min (y r-2014, \text { phaseyrs }) \times \frac{\operatorname{targ}^{(k)}-u l t^{(k)}}{\text { phaseyrs }} .
\]

The following table displays the coefficients, target values, and phase-in years for each type of beneficiary.
\begin{tabular}{|c|l|c|c|c|c|}
\hline\(k\) & Type & \(a_{1}^{(k)}\) & \(b^{(k)}\) & \begin{tabular}{c} 
Target Value \\
targ \(^{(k)}\)
\end{tabular} & \begin{tabular}{c} 
Add-factoring Phase-in \\
Years \\
phaseyrs
\end{tabular} \\
\hline 1 & Wife & 0.85940 & 0.0 & 0.205 & 33 \\
\hline 2 & Widow & 1.87591 & 0.0 & 0.575 & 31 \\
\hline 3 & Widower & -0.24220 & 0.13696 & 0.10 & 45 \\
\hline 4 & Husband & -0.04491 & 0.02075 & 0.0085 & 32 \\
\hline
\end{tabular}

In the above equations, the average PIA of newly entitled retired worker beneficiaries by sex has already been computed (see subsection (2) above).

\section*{Average Excess Amount for Dually Entitled Beneficiaries}

The projection of the average excess amounts for two categories of dually entitled beneficiaries (wives and widows) is similar to that of the number of dually entitled beneficiaries. The structure of the equations used to project these amounts is similar to the equations used to project the percentage exposures.

The equations used to project the average excess amount each have two terms. Each equation \(k\) has one coefficient, \(b_{1}^{(k)}\), calculated in a side model for the cost process based on historical series of the other terms in the equations and an add factor adjustment (denoted " \(c^{(k) ")}\) ) similar to the process for the number of dually entitled beneficiaries. A target value in the \(75^{\text {th }}\) year of the projection period is used in deriving the "add-factor" adjustment. The two equations
\[
\operatorname{AvgExcPct}^{(k)}(y r)=b_{1}^{(k)} \frac{\operatorname{PIA}(\mathrm{yr}, \mathrm{M})-\operatorname{PIA}(\mathrm{yr}, \mathrm{~F})}{\operatorname{PIA}(\mathrm{yr}, \mathrm{M})}+c^{(k)}(y r)
\]
( \(k=1,2\) ) project the average excess benefit amounts of wife (1) and widow (2) beneficiaries as a percentage of the male retired worker benefit.

The derivation of \(c^{(k)}(y r)\) is similar to the case of the exposure percentages as described above. Suppose that \(u t^{(k)}\) is the value obtained from the regression equation without add-factoring in the final year of the projection period. Therefore
\[
u l t^{(k)}=b_{1}^{(k)} \frac{\operatorname{PIA}(n i+74, \mathrm{M})-\operatorname{PIA}(n i+74, \mathrm{~F})}{\operatorname{PIA}(n i+74, \mathrm{M})}
\]

Let \(\operatorname{targ}{ }^{(k)}\) be the target value we estimate for the final year of the projection period. Let
phaseyrs be the number of years it takes to fully phase in the target value. Then we have
\[
c^{(k)}(y r)=\min (y r-2014, \text { phaseyrs }) \times \frac{\operatorname{targ}^{(k)}-u l t^{(k)}}{\text { phaseyrs }} .
\]

The table below shows the regression coefficients and other relevant adjustments in the 2016 Trustees Report.
\begin{tabular}{|l|l|l|c|c|}
\hline\(k\) & Type & \(b_{1}^{(k)}\) & \begin{tabular}{c} 
Targeted Value (2090) in \\
Nominal Dollars \\
targ \(^{(k)}\)
\end{tabular} & \begin{tabular}{c} 
Add-factoring \\
Phase-in Years \\
phaseyrs
\end{tabular} \\
\hline 1 & Wife & 0.43987 & 4,035 & 28 \\
\hline 2 & Widow & 1.23033 & 10,626 & 31 \\
\hline
\end{tabular}

The average excess amount of widower beneficiaries is estimated to be a fixed percentage (50\%) of the average excess amount of widow beneficiaries. That is, \(\operatorname{AvgExcPct}{ }^{(3)}(y r)=0.50 \times \operatorname{AvgExcPct}^{(2)}(y r)\).

The average excess amount of husband beneficiaries is estimated as a percentage of average excess amount of dually entitled wives. This percentage is measured to be \(76.24 \%\) in 2014 and is assumed to linearly phase to \(80 \%\) in the \(18^{\text {th }}\) year of the projection period (2033 for 2016TR).

\section*{Annualizing Excess Amounts}

The process to annualize excess amounts is very similar to the process for annualizing auxiliary benefits.

For each dual entitlement category, the number of beneficiaries is simply \(c p(y r, c a t)=\operatorname{PctExp}^{(c a t)}(y r) \times \operatorname{ExposedPop}^{(c a t)}(y r)\). With this method, however, no linkage factor is used. Instead, the projected average excess amount, as explained above, is used. Therefore,
\[
\begin{aligned}
& \text { AggExcess(yr,cat) } \\
& =\sum_{i=0}^{11}\left[\frac{(12-i)}{12} \times c p(c a t, y r-1) \times \operatorname{AvgExcAmt~}^{(c a t)}(y r-1)+\frac{i}{12} \times c p(c a t, y r) \times \operatorname{AvgExcAmt}^{(c a t)}(y r)\right] .
\end{aligned}
\]

\section*{(5) Retroactive Payments}

Frequently, beneficiaries start receiving payments later than their actual entitlement date, such that they receive a "catch-up" lump-sum amount for the time delay, in addition to
regular monthly benefits going forward. These lump-sum amounts, which we call retroactive payments, apply for all beneficiary categories (except for the one-time \$255 death benefit) but are more significant for disabled workers because of the frequent and sometimes lengthy time lag in getting an allowance on their application, as well as the 12 months of retroactivity allowed at the time of benefit filing. This section discusses how retroactive benefit payments are projected for all beneficiary categories.

\section*{Disabled Workers}

For each age and sex the DI area (subprocess 3.2) provides two cumulative distributions (one initial and one ultimate), by duration, of incurred but not reported cases. Denote this by ibnr (yr, sex, age, dur) where for years ni-10 to 2008 ibnr factors are set to the initial distribution, for years 2018 and after ibnr factors are set to the ultimate distribution, and for years 2009 to 2017 we linearly interpolate. The number of disability beneficiaries who will eventually receive benefits (by age, sex, and at duration 0 ) is
\(\operatorname{dibce}(y r, \operatorname{sex}, a g e)=\operatorname{dibcp}(y r, \operatorname{sex}, a g e, 0) / i b n r(y r, \operatorname{sex}, a g e, 0)\).
The associated frequency distribution of incurred but not reported cases is ibnr _ freq(yr, sex, age, dur). In other words, we have
\(i b n r \quad\) freq \((y r\), sex, age, 0\()=i b n r(y r\), sex, age, 0\()\) for duration 0 and
\(i b n r\) _ freq(yr, sex, age, dur) \(=\operatorname{ibnr}(y r-d u r\), sex, age, dur \() ~-i b n r(y r-d u r\), sex, age, dur -1)
for higher durations, \(1 \leq d u r \leq 9\) and \(10+\).
Let dur be a duration, 0 through 9 or 10+. Define
- Cum_COLA(dur) \(=\prod_{j=0}^{d u r}(1+\operatorname{COLA}(y r-j))\).
- For \(i=0, \ldots, 10\), Num_Months \((i)=\left\{\begin{array}{cc}2 & d u r=0 \\ 5 & d u r>0 \text { and } i=0 \\ 12 & d u r>0 \text { and } 0<i<d u r \\ 6 & d u r>0 \text { and } i=d u r\end{array}\right.\)

Then the aggregate retroactive payments for disabled workers, in millions, are defined to be

Retro_DIB(sex,yr)

dibpia(sex, age, dur) is further altered in this formula by the adjustment factors in array:
\(D I ~ \_R I \_\)facs(sex,age,dur,2)
These adjustments to dibpia are designed to account for the small differences between
average benefit levels for all beneficiaries and levels for those receiving retroactive credit.

Retro_DIB(sex,yr) is simply added to the disabled worker benefit category by year and sex.

\section*{Retired Worker Beneficiaries}

Retired worker beneficiaries are assumed to have, on average, 0.7 months of retroactive payments as determined annually by a side model. Hence
\[
\text { retro }{ }_{-} O A B(\text { sex, } y r)=\sum_{\text {age }=62}^{70} \frac{0.7 \times \frac{\text { oabicp }(y r, \text { sex, age, age })}{1000} \times \text { oabmba }(y r, \text { sex, age age })}{1+\operatorname{COLA}(y r)} .
\]

In the above formula, oabicp ( \(y r\), sex, age, age) is the number of newly entitled beneficiaries at age equal to age (age in current pay equals age at entitlement equals age) and oabmba (yr, sex, age, age) is the corresponding average benefit.
oabmba(year, sex, age, age) in the above formula is further altered by the adjustment factors:
retrom_adj \((\operatorname{sex})=\{0.9231,1.0201\}\)
These adjustments to oabmba are designed to account for the small differences between average benefit levels for all beneficiaries and levels for those receiving retroactive credit.

The aggregate retroactive benefits for retired worker beneficiaries are simply added to the retired worker benefit category by year and sex.

\section*{Auxiliary Beneficiary Categories}

Retroactive payments for auxiliary beneficiaries (determined by a side model) are treated as a loading of the aggregate annual benefits by auxiliary category. That is, each auxiliary benefit category has a loading factor to represent retroactive payments on top of regular monthly benefits, and the aggregate annual benefits by category are increased by this loading factor.
(6) Aggregate Scheduled Benefits (BEN)

Aggregate benefits by trust fund, \(B E N(t f, y r)\), are computed as follows. For each year of the 75 -year long-range projection period, the aggregate benefits by category (including retroactive payments, as described above) are summed up to give total annual scheduled benefits by trust fund. In the short-range period, the long-range values are overridden by the values estimated by the short-range office. The difference between long-range
scheduled benefits and short-range benefits in the \(10^{\text {th }}\) year of the short-range period is called the scheduled benefits adjustment. In the 10 years after the end of the short-range period, the long-range scheduled benefits are adjusted by linearly grading the scheduled benefits adjustment to zero. From the \(20^{\text {th }}\) year forward, the projection is the pure longrange value.

\section*{iii. Equation 4.3.3 - Taxation of Benefits (TAXBEN)}

The short-range office provides taxation of benefits levels by trust fund in the short-range period. These implicitly give, for each year, an estimated taxation of benefits factor, by trust fund, equal to the estimated taxation of benefits as a percentage of benefits scheduled to be paid. The long-range office projects these factors independently for every year of the projection period, also by trust fund. (See subprocess 4.1.) The difference in the factors between the two offices at the end of the short-range period is phased out linearly over the next ten years. The long-range projection of taxation of benefits is estimated by multiplying the projected taxation of benefits factors by the benefits scheduled to be paid, by trust fund. If taxben_factor \((t f, y r)\) is the percentage of scheduled benefits for the year, by trust funds, estimated to be collected as taxation on benefits, then
\[
\text { TAXBEN }(t f, y r)=\text { taxben }_{-} \text {factor }(t f, y r) \times B E N(t f, y r)
\]
for \(y r \geq n s+10\).

\section*{iv. \(\quad\) Equation 4.3.4 - Administrative Expenses (ADM)}

Administrative expenses are estimated separately by trust fund. In the short-range period, the short-range office provides the estimates of administrative expenses by trust fund. Thereafter, administrative expenses are computed by multiplying the previous year's administrative expenses by three factors: annual changes in total beneficiaries, annual changes in AWI, and one minus annual productivity growth. As a formula, if ticp(tf,yr) is the total estimated number of beneficiaries in current-pay status by trust fund and year, \(A W I(y r)\) is the average wage index in year \(y r\), and prod is the ultimate assumed annual growth in productivity, then
\[
\begin{aligned}
A D M(t f, y r)= & A D M(t f, y r-1) \times[t i c p(t f, y r) / t i c p(t f, y r-1)] \times[A W I(y r) / A W I(y r-1)] \\
& \times(1-\operatorname{prod})
\end{aligned}
\]
for \(y r>n s\).

\section*{v. Equation 4.3.5 - Railroad Interchange (RR)}

Railroad interchange is disaggregated by trust fund and projection year. The long-range office does a projection for each year in the 75-year period. In the short-range period (first 10 years of the 75 -year projection period), the short-range office provides the estimates of railroad interchange by trust fund and the long-range projection is overridden in these years. Over the next five years of the projection period, the estimate of the
railroad interchange is a linear interpolation between the short-range projection at the end of the short-range period and the long-range projection five years hence. During the final 60 years of the projection period, the projection is as estimated by the long-range office.

By trust fund, the total cashflow in year \(y r\), \(r r_{-}\)cashflow \((t f, y r)\), is broken down into two positive components; railroad benefits in year \(y r\) and railroad administrative expenses in year \(y r\), and two negative components; railroad contributions in year \(y r\) and railroad taxation of benefits in year \(y r\). A positive cashflow in this calculation represents a net cost to the trust fund. Cashflows are calculated on a fiscal year basis.

Projections of numbers of newly entitled retired workers are determined by analyzing the ratio of new entitlements to previous levels of railroad employment using 1997-2013 new entitlement data in the analysis. After initial entitlement, a mortality rate of \(5.6 \%\) for 2014 is assumed based on analysis of recent Railroad Retirement Board financial interchange data, with mortality assumed to improve thereafter. For projections of newly entitled disabled workers a similar trend analysis based on prior employment is used. A "mortality" rate (deaths plus recoveries plus conversions to retired worker benefits) is determined using Railroad Retirement Board (RRB) data, with the rate assumed constant thereafter.

Assuming a 90/10 male/female split, the average benefit level for an OASI railroad worker is calculated as a ratio to the average OASI retired worker benefit. This ratio is constant throughout the projection period and is derived by comparing MBR and Railroad Board data. Additionally a constant loading factor based on the same data is applied to aggregate worker benefits to determine the aggregate benefit amounts for auxiliary OASI beneficiaries. This same approach is used to determine similar constants for DI railroad benefits. The aggregate disabled worker railroad benefits (and beneficiaries) are estimated in the same way.

It is assumed that the ratio of OASI taxation of benefits to OASI benefits and DI taxation of benefits to DI benefits are both the same for railroad taxation of benefits. The railroad taxation of benefits is estimated by multiplying the railroad benefits by these ratios.

Administrative expenses for railroad are computed separately by trust fund. They are set at levels determined by short range in the short-range period. For years ni +10 to nf, they are computed similarly to OASDI administrative expenses. Administrative expenses in yr-1 are multiplied by (a) the change in the total number of worker beneficiaries, (b) the annual change in average wage, and (c) one minus the ultimate productivity growth.

Railroad contributions are estimated, by trust fund, to be total railroad employment, multiplied by average railroad earnings, multiplied by the combined OASDI employer/employee tax rate. Railroad earnings are assumed to grow with the increase in the average wage index, and railroad employment is assumed to decrease over time, both of which are in line with the Railroad Retirement Board's own "most likely" projections.

The interchange amount is calculated both on a fiscal year basis and at the date of the
actual transfer from SSA to the Railroad Retirement Board (usually in the first week of June in the following year). These interchange amounts can be represented by \(R R \_T r a n s f e r \_F Y(t f, y r)\) and \(R R \_T r a n s f e r \_C Y(t f, y r+1)\). \(R R \_T r a n s f e r \_F Y(t f, y r)\) is calculated by adding the cashflow amount to an interest amount, rr_interest (tf,yr), which includes interest accrued by the cashflow components along with a reconciliation of interest amounts from the previous fiscal year calculated transfer amount and the current year June interchange. If irate_presc(tf,year) is the prescribed interest rate used by the Railroad Retirement Board the interchange amount in June is:
\(R R_{-}\)Transfer_CY(tf ,yr)\(=R R_{-}\)Transfer_FY(tf,yr-1)×(1.0+(2.0/3.0)×Irate_Presc(yf,yr))

\section*{vi. Equation 4.3.6 - Interest Income (INT)}

In the short-range period, the projection of interest income by trust fund is provided by the short-range office. In each year of the short-range period, the annual yield rate is defined as the ratio of interest earned by a fund to the average level of assets held by the fund during the year.

The ultimate annual yield rate on each trust fund is equal to the nominal yield, which is the real interest rate increased for inflation. As a formula,
\[
\text { ultimate yield rate }=(1+\text { real interest rate }) \times(1+\text { inflation rate })-1 \text {. }
\]

To get the yield rate for each year between the end of the short-range period (ns) and 5 years later, ns +5 , when the ultimate yield rate is assumed to be reached, the program linearly interpolates between the values for years \(n s\) and \(n s+5\).

The projection of interest income in a given year is the yield rate for that year multiplied by the average level of assets. As a formula,
\[
I N T(t f, y r)=y i e l d(t f, y r) \times a v g \_\operatorname{assets}(t f, y r)
\]

The amount of assets in a trust fund at the end of a given year is estimated from the level of assets at the beginning of the year by:
- Increasing the level for the tax contributions and taxation of benefits income received during the year (each exposed to the point in the year in which they are estimated to be received, on average), and
- Decreasing the level for scheduled benefits, railroad interchange, and administrative expenses paid during the year (each exposed to the point in the year in which they are estimated to be disbursed, on average).

For all years of the projection period, tax contributions are given an exposure of 0.518 , taxation of benefits are given an exposure of 0.625 , railroad interchange is given an exposure of \(0.58 \overline{3}\), and administrative expenses are given an exposure of 0.5 . For scheduled benefits, separate OASI and DI exposures are determined through a side model. The exposure, ben_exp(yr), is larger than 0.5 in the early years of the projection
period for both trust funds, whereas in later years OASI benefit exposures fall below 0.5 with DI remaining above 0.5 . The reason is that in the past, benefits were always paid on the \(3^{\text {rd }}\) of each month. Now benefits are paid out throughout the month, based on the birth date of the beneficiary. The reason for the differences between trust funds is that benefits are paid on the third of the month (exempting check cycling \({ }^{5}\) ) for a higher proportion of DI beneficiaries due primarily to (1) concurrent receipt of SSI benefits, or (2) state payment of Medicare premiums. The average assets held by the trust funds for a given year is estimated by the formula
\[
\begin{aligned}
\operatorname{avg} \_\operatorname{assets}(t f, y r)= & \operatorname{ASSETS}_{\text {ВоҮ }}(t f, y r)+0.518 \times \text { CONTRB }(t f, y r) \\
& +0.625 \times \operatorname{TAXBEN}(t f, y r)-\text { ben_exp }(t f, y r) \times B E N(t f, y r) \\
& -0.58 \overline{3} \times R R(t f, y r)-0.5 \times \text { ADM }(t f, y r) .
\end{aligned}
\]

\section*{vii. Equations 4.3.7, 4.3.8 and 4.3.9 - Annual Values}

The annual income rate for a trust fund is computed as the sum of payroll tax contributions plus taxation of benefits as a percentage of taxable payroll.
\[
A N N_{-} I N C_{-} R T(t f, y r)=\frac{\operatorname{CONTRB}(t f, y r)+T A X B E N(t f, y r)}{\text { payroll }(y r)} .
\]

The annual cost rate for a trust fund is computed as the total cost of providing scheduled benefits from that fund as a percentage of taxable payroll. If
\[
\operatorname{COST}(t f, y r)=B E N(t f, y r)+R R(t f, y r)+A D M(t f, y r),
\]
then
\[
A N N_{-} C O S T \_R T(t f, y r)=\frac{\operatorname{COST}(t f, y r)}{\operatorname{payroll}(y r)}
\]

The trust fund ratio measures the amount of beginning of year assets that can be used to pay total cost. It is expressed as a percentage:
\[
T F R(t f, y r)=\frac{\operatorname{ASSETS}_{B O Y}(t f, y r)}{\operatorname{COST}(t f, y r)}
\]

\section*{viii. Equations 4.3.10, 4.3.11, 4.3.12, and 4.3.13 - Summarized Values}

Present values of cash flows during the year are computed using the yield rate on the combined OASDI trust fund for that year. Each component of trust fund operations is exposed, with interest, to the point in the year in which, on average, it is received or

\footnotetext{
\({ }^{5}\) Under check cycling many Social Security beneficiaries filing for benefits after April 1997 are paid on either the \(2^{\text {nd }}\), \(3^{\text {rd }}\), or \(4^{\text {th }}\) Wednesday of each month. In the past, benefits were always paid on the \(3^{\text {rd }}\) day of each month.
}
disbursed. These exposure levels, ben_exp(tf,yr), are the same as described above. These exposed levels are then discounted to January 1 of the year of the Trustees Report, ni. If yield \((j)\) is the annual yield rate on the combined OASDI trust funds for year \(j\) and \(v(y r)\) is the discounting factor for the year, then
\[
v(y r)=\prod_{j=n i}^{y r} \frac{1}{[1+\text { yield }(j)]} .
\]

For a given year, and trust fund,
\[
\begin{aligned}
P V \_T A X(t f, y r) & =(1+0.518 \times y \text { ield }(y r)) \times T A X(t f, y r) \times v(y r), \\
P V \_T A X B E N(t f, y r) & =(1+0.625 \times \text { yield }(y r)) \times T A X B E N(t f, y r) \times v(y r), \\
P V \_B E N(t f, y r) & =\left(1+b e n_{-} \exp (t f, y r) \times y i e l d(y r)\right) \times B E N(t f, y r) \times v(y r), \\
P V \_R R(t f, y r) & =(1+0.58 \overline{3} \times \text { yield }(y r)) \times R R(t f, y r) \times v(y r),
\end{aligned}
\]
and \(P V_{-} A D M(t f, y r)=(1+0.5 \times\) yield \((y r)) \times A D M(t f, y r) \times v(y r)\).
The target fund for a year is next year's cost. Its present value is computed as
\[
P V_{-} T A R G(t f, y r)=[B E N(t f, y r+1)+R R(t f, y r+1)+A D M(t f, y r+1)] \times v(y r),
\]

Taxable payroll is exposed to the middle of the year when computing present values:
\[
P V_{-} P A Y R O L L(y r)=(1+0.5 \times \operatorname{yield}(y r)) \times \operatorname{payroll}(y r) \times v(y r)
\]

We also define
\[
P V \_I N C(t f, y r)=P V \_T A X(t f, y r)+P V \_T A X B E N(t f, y r)
\]
and
\[
P V_{-} C O S T(t f, y r)=P V_{-} B E N(t f, y r)+P V_{-} R R(t f, y r)+P V_{-} A D M(t f, y r)
\]

Summarized rates are calculated using beginning of period assets and a target fund. Let \(\mathrm{yr}_{1}=\) the first year of the valuation period and \(\mathrm{yr}_{2}=\) the ending year of the valuation. Then the summarized income rate is:
\[
S U M M_{-} I N C_{-} R T\left(t f, y r_{1}, y r_{2}\right)=\frac{\operatorname{ASSETS}_{B O Y}\left(t f, y r_{1}\right)+\left(\sum_{j=y r_{1}}^{y r_{2}} P V_{-} I N C(t f, j)\right)}{\sum_{j=y r_{1}}^{y r_{2}} P V_{-} P A Y R O L L(j)} .
\]

The summarized cost rate is similarly computed:
\[
S U M M_{-} C O S T \_R T\left(t f, y r_{1}, y r_{2}\right)=\frac{\left(\sum_{j=y r_{1}}^{y r_{2}} P V_{-} \operatorname{COST}(t f, j)\right)+P V \_T A R G\left(t f, y r_{2}\right)}{\sum_{j=y r_{1}}^{y r_{2}} P V \_\operatorname{PAYROLL}(j)}
\]

The 75 -year actuarial balance is computed for a period beginning January 1 of the Trustees Report year, ni. It includes both beginning of period assets and a target fund. Therefore,
\[
A C T B A L_{75 y r}(t f)=S U M M_{\_} I N C_{\_} R T(t f, n i, n i+74)-S U M M_{-} C O S T_{-} R T(t f, n i, n i+74) .
\]

In general, an actuarial balance may be computed for any given subperiod of the projection period. In general, actuarial balances for a subperiod beginning on January 1 of year ni and continuing through the end of year \(y r\) are computed using
\[
A C T B A L_{n i, y r}(t f)=S U M M_{-} I N C_{-} R T(t f, n i, y r)-S U M M_{-} C O S T_{-} R T(t f, n i, y r) .
\]

The unfunded obligation of a trust fund for a given period is the excess of the present value of the net cash deficits for each year of that period over the trust fund balance at the beginning of the period. The unfunded obligation for the period beginning on January 1 of year ni and continuing through the end of year \(y r\) is computed using
\[
U N F \_O B L(t f, y r)=\sum_{j=n i}^{y r}\left[P V_{-} C O S T(t f, j)-P V_{-} I N C(t f, j)\right]-\text { ASSETS }_{\text {BOY }}(t f, n i) .
\]

Note that the unfunded obligation excludes the target fund.

\section*{ix. Equation 4.3.14—Closed Group Unfunded Obligation}

The closed group is defined as individuals who attain specified ages in the first year of the projection period (ni). The Statement of Social Insurance displays unfunded obligations for closed groups (1) attaining 15 or later in 2016, (2) attaining 62 or older in 2016, and (3) attaining 15 to 61 in 2016. For each year of the projection period, closed group calculations attribute a portion of the items in equations 4.3.1 through 4.3.6 to individuals falling in the defined closed group. The calculation of the closed-group unfunded obligation, then, uses the equation above but only considering the present values of cost and income attributable to the closed group.

The following information, developed elsewhere in the "Cost" process, is used for developing closed group unfunded obligation amounts:
- Total number of workers and total taxable earnings by single year of age 0119 and sex, years 1951 through \(n f\), updated yearly
- Taxable payroll, years ni through nf, updated yearly
- Payroll tax income, years \(n i\) through \(n f\), updated yearly
- Income from taxation of benefits, years ni through \(n f\), updated yearly
- Scheduled benefits by beneficiary category, years ni through \(n f\), updated yearly
- Railroad interchange, years ni through \(n f\), updated yearly
- Administrative expenses, years ni through nf, updated yearly
- Yield rate on the combined OASDI trust funds, years ni through \(n f\), updated yearly
- Population counts for all retired workers, spouses, divorced spouses, and widow(er)s by year, sex, and age 95-119 (read in as a percentage of 95+ counts)
- Distribution of assumed age differentials between aged spouses and workers ranging from -12 years to 15 years seniority for the worker by sex of the beneficiary, age of the beneficiary 62-74, and marital status
- Distribution of assumed age differentials between aged widow(er)s and deceased workers from -6 years to +12 years

It is important to note that, for dependent beneficiaries, the age of the worker, on whose account the benefits are based, determines whether that beneficiary would fall in the closed group. For instance, if the closed group were defined as individuals attaining age 15 or later in 2016, a 3-year-old minor child receiving benefits in 2016 on the account of a retired worker aged 63 would be considered part of this closed group because the account holder was at least age 15 in 2016. The following describes how the various components of income and cost are allocated to the defined closed group in question:

\section*{Payroll Tax Contributions}

Closed group taxable payroll is defined as the percentage of OASDI taxable payroll attributable to the closed group in question. An input file of closed group payroll factors, containing these percentages by year from 2016 through 2115, is used by the cost program to compute payroll tax contributions attributable to the closed group. For each year, the closed group payroll factors are determined as follows:
- \(\quad\) The number of projected workers by single year of age (ages 0-119) and sex are multiplied by the associated average earnings by age/sex.
- Then, the portion of total taxable earnings attributable to the closed group is calculated.

For each year of the projection period, the number of workers and average taxable earnings by single year of age and sex are determined as follows:
- For ages 0-99, the number of projected workers comes directly from Economics group projections.
- For ages 100-119, the total age 100+ amounts from Economics are distributed among ages 100 through 119 for each projection year using mortality rates derived by looking back 10 years and multiplying these populations by 10 years
of mortality increases by sex.
- For ages \(0-14\), the average taxable earnings are obtained by analyzing historical 1993-2012 data of the average earnings at each age relative to age 15 average earnings, and judgmentally assigning a ratio (to age 15 average earnings) for each age.
- For ages \(15-80\), the average taxable earnings by age and sex come directly from the AWARDS subprocess.
- For ages 81-99, the average taxable earnings are calculated as a weighted average of the average earnings for ages 76-80, weighted by the ratios of earnings between these ages and each age 81 through 99 for earnings years 1993 through 2012. This is similar to the method to calculate average earnings for ages 0 through 14, though here we are using five anchor years instead of one. For ages 100-119, average taxable earnings are set equal to the average earnings calculated for age 99.

\section*{Benefits}

Methodologies for computing benefits attributable to the closed groups differ among benefit categories, as described below:

\section*{Retired Workers}

For each age in current pay, the number of beneficiaries is multiplied by the corresponding average benefit amount across all ages of entitlement. The same applies for DI conversion cases. Retroactive benefits for the current year, by age, are then calculated. The closed-group factors for old-age benefits for each year are found by summing the benefit amounts attributable to the specified closed group, as a proportion of total retired worker benefits for all ages. This process is done separately by sex.

\section*{Disabled Workers}

For each age from 20 to the year before normal retirement age, the program adds the products of the number of beneficiaries for each duration and the PIA for that duration. Retroactive benefits for the current year, by age, are then calculated. The closed-group factors for disability benefits are found by summing the total benefit amounts attributable to the closed group, as a proportion of total disabled worker benefits for all ages. This process is done separately by sex.

\section*{Aged Spouses and Divorced Aged Spouses}

Closed Group calculations are done separately (although in the same manner) for aged wives, aged divorced wives, aged husbands and aged divorced husbands (combined), dually entitled aged wives, and dually entitled aged husbands. The number of aged spouse beneficiaries in each beneficiary category in current pay status (no dual entitlement) is provided by single year of age (up through 119). Then, for
each single year of age, the program allocates total numbers of workers by age, from 12 years younger to 15 years older than the aged spouse using an assumed categoryspecific distribution. Next, for each age of worker in current pay, the number of workers is multiplied by the weighted average retired worker benefit for that age; this is done for all ages. The closed group factors, then, are obtained by determining the proportion of total benefit amounts attributable to the given closed group (based on the worker's age).

\section*{Aged Widows and Divorced Aged Widows}

Closed Group calculations are done separately (although in the same manner) for aged widows, aged divorced widows, aged widowers, and aged divorced widowers. The number of aged widow(er) beneficiaries in current pay status (no dual entitlement) is provided by age from 60 to 119 . For each single year of age, the program allocates total number of aged widows by age of the deceased husband (age the husband would have been if he had not died), from 6 years younger to 12 years older than the aged widow using an assumed distribution. The same distribution is used in reverse order for aged widowers leading to an age range of 12 years younger to 6 years older for the deceased wife. For each age of deceased spouses aged 119 or younger, a real wage growth factor is applied to reflect ultimate real wage growth taken to the power of the number of years younger than age 119.
\[
\text { Benefitadj } \left.=(1+\text { wg _ult }-1.026)^{(119-d e c e a s e d ~ w o r ~ k e r a g e ~}\right)
\]

This exponent is intended to reflect differences in average levels of benefits, with younger deceased spouses having higher benefits based on real wage growth. The closed group factors, then, are obtained by determining the proportion of total benefit amounts attributable to the closed group.

For aged widow(er)s in dual entitlement status (i.e. aged widow(er)s with a smaller worker benefit) we do similar calculations with the assumption that the age distributions are equal to the combined distribution of non-dually entitled aged widow(er)s and aged divorced widow(er)s.

\section*{Other Beneficiary Categories}

For the 20 other dependent beneficiaries of retired workers, disabled workers, and deceased workers, an input file of closed group benefit factors is created, which represents the proportion of total (open-group) projected benefits in that category attributable to the given closed group age and year. This file is used by the cost program to compute amounts from each beneficiary category attributable to the closed group. The file, separately created for each closed group run, contains closed group benefit factors for ages 0 through 150 for each of the 20 beneficiary categories by sex of the account holder (worker). These input files are created by examining a recent sample of Master Beneficiary Record (MBR) data \({ }^{1}\) for each of the beneficiary
categories by age of the worker, and projecting future distributions by age of the worker based on population and, for survivor benefits, projected deaths by age. Then, adjustments are made for real wage growth to reflect different benefit levels by birth cohort.

\section*{Taxation of Benefits}

Since taxation of benefits is related to benefits, the closed-group taxation of benefit amounts are computed by multiplying the total (open-group) taxation of benefit amounts by Trust Fund, by the corresponding total closed-group benefit factors by Trust Fund.

\section*{Administrative Expenses}

Since administrative expenses are also assumed to be related to benefits, the closed-group administrative expense amounts are computed by multiplying the total (open-group) administrative expenses by Trust Fund), by the corresponding total closed-group benefit factors by Trust Fund.

\section*{Railroad Interchange}

Since the railroad interchange has both a payroll tax and benefit component, each component is multiplied by its corresponding closed-group factor. That is, total payroll tax contributions arising from railroad interchange are multiplied by the closed group payroll factor discussed above in the "Tax on Contributions" section. Total railroad benefits, by Trust Fund, are multiplied by the aggregate closed-group benefit factors by Trust Fund. Closed-group railroad administrative expenses and closed-group railroad taxation of benefits are also estimated by applying aggregate closed group benefit factors by Trust Fund. The final amount is then the difference in the components (closed group railroad income less closed group railroad cost).

\section*{Appendix 4.3-1}

Shuttling Method
In this appendix, we discuss the "shuttling method".
The shuttling method as presented in the COST model attempts to reorganize the PAPs obtained from subprocess 4.2, which maintains a static age-sex distribution of newly entitled beneficiaries, in such a way that captures the changing age-sex distribution of newly entitled beneficiaries provided by subprocess 3.3. The age-sex distribution of the sample (subprocess 4.2) and those newly entitled from subprocess 3.3 are aligned in the sample year. This alignment persists throughout all years in the long-range period. When we refer to the age-sex distribution from subprocess 3.3 in what follows below, we refer to the aligned age-sex distribution.

Let oadsrs be the age/sex distribution of the sample from subprocess 4.2. Let oabicp be the number of newly entitled beneficiaries by age and sex from subprocess 3.3. Let total be the total number of newly entitled beneficiaries by sex. The ratio \(\frac{\operatorname{oabicp}(\operatorname{sex}, a g e)}{\operatorname{total}(\text { sex })}\), for age \(=62, \ldots, 70\) gives the age-sex distribution from
subprocess 3.3. The array rsb_oadscp_sampleyr is the age-sex distribution from subprocess 3.3 from the sample year (2013 for the 2016 Trustees Report).

The value of oadscp is defined, by age and sex, to be:
oadscp \((\) sex, age \()=\frac{\text { oabicp }(\operatorname{sex}, \text { age })}{\operatorname{total}(\operatorname{sex})}+\) oadsrs \((\) sex, age \()-r s b_{-}\)oadscp _sampleyr \((\)sex, age \()\).
Starting with the 2010 Trustees Report, we set all values of rsb_oadscp_sampleyr to the corresponding value of oadsrs. Therefore, the value of oadscp is now defined to be:
\(\operatorname{oadscp}(\) sex, age \()=\frac{\operatorname{oabicp}(\operatorname{sex}, \text { age })}{\operatorname{total}(\operatorname{sex})}\). Despite the fact that we eliminated the actual
alignment we will still refer to this as the aligned age-sex distribution as obtained from subprocess 3.3 (with the assumption that the alignment left the age-sex distribution unchanged).

For each sex, we construct a matrix oads. Construction of this matrix uses two different age-sex distributions: the age-sex distribution of the awards sample and the aligned age-sex distribution of newly entitled beneficiaries from subprocess 3.3. The matrix oads is a \(9 \times 9\) matrix whose rows and columns are indexed consecutively by the ages \(62, \ldots, 70\). We index the rows of this matrix by the age of entitlement of a worker in the projection (ageentRSB) and the columns of this matrix by the age of entitlement of a worker in the Awards sample (ageentAWD).

More precisely, for a given sex, let oadsrs be the age distribution of the sample; oadscp be the aligned age distribution of newly entitled beneficiaries from subprocess 3.3. Both arrays oadsrs and oadscp are indexed by age, ages \(=62, \ldots, 70\). For a
given sex, the matrix oads is constructed with the following properties:
oadscp \((\) ageentRSB \()=\sum_{\text {ageentAWD=62 }}^{70}\) oads \((\) ageentRSB, ageentAWD \()\)
and
\(\operatorname{oadsrs}(\) ageentAWD \()=\sum_{\text {ageentRSB=62 }}^{70}\) oads \((\) ageentRSB, ageentAWD \()\).
In other words, the matrix oads has the following properties:
- The sum of the entries in any row of the matrix is the value of the distribution of the projection for the age corresponding to the row.
- The sum of the entries in any column of the matrix is the value of the distribution of the Awards sample for the age corresponding to the column.

Let opap be the original potential AIME percentages (PAPs) passed to subprocess 4.3 from subprocess 4.2 . This is a \(9 \times 30\) matrix whose rows are indexed by ages at entitlement \(62, \ldots, 70\) and whose columns are indexed by benefit interval \(1, \ldots, 30\) (represented by the variable \(i\) in the formulas below). The values of opap are modified and the results are the PAPs used by process 4.3, called opap1. As a formula, opap1(ageentRSB,i)
\[
=\frac{\sum_{\text {ageentAWD }=62}^{70} \text { opap }(\text { ageentAWD }, i) \times \text { oads }(\text { ageentRSB, ageentAWD })}{\sum_{\text {ageentAWD }=62}^{70} \text { oads }(\text { ageentRSB, ageentAWD })} .
\]

This formula may be rewritten as follows.
opap1(ageentRSB,i)
\[
=\sum_{\text {ageentAWD=62 }}^{70} \text { opap }(\text { ageentAWD,i }) \times \frac{\text { oads }(\text { ageentRSB, ageentAWD })}{\sum_{\text {ageenAWD }=62}^{70} \text { oads }(\text { ageentRSB, ageentAWD })} .
\]

It follows that opap1 may be interpreted as a reweighting of opap. We have
\[
\text { opap1 }(\text { ageentRSB, } i)=\sum_{\text {ageenAAWD }=62}^{70} w_{\text {ageentRSB, ageentAWD }} \times o p a p(\text { ageentAWD, } i)
\]
with weights \(w_{\text {ageentRSB,ageentawD }}=\frac{\text { oads }(\text { ageentRSB, ageentAWD })}{\text { oadscp }(\text { ageentRSB })}\).
As matrices, opap \(1=w \times\) opap (and is another \(9 \times 10\) matrix).

Consider the following example. In this example, the projection year is 2040, and the sex is males.

The oadsrs vector (from subprocess 4.2) is as follows.
\begin{tabular}{c|ccccccccc} 
& 62 & 63 & 64 & 65 & 66 & 67 & 68 & 69 & 70 \\
\hline Male & 0.3303 & 0.1198 & 0.0630 & 0.0947 & 0.3221 & 0.0291 & 0.0128 & 0.0089 & 0.0194
\end{tabular}

The unaligned age-sex distribution (from subprocess 3.3) is as follows.
\begin{tabular}{c|ccccccccc} 
& 62 & 63 & 64 & 65 & 66 & 67 & 68 & 69 & 70 \\
\hline Male & 0.2486 & 0.1208 & 0.0468 & 0.1062 & 0.1889 & 0.1607 & 0.0485 & 0.0416 & 0.0380
\end{tabular}

The age-sex distribution (from subprocess 3.3) in 2013, the year of the sample is assumed the same as the oadsrs vector and therefore is as follows.
\begin{tabular}{c|ccccccccc} 
& 62 & 63 & 64 & 65 & 66 & 67 & 68 & 69 & 70 \\
\hline Male & 0.3303 & 0.1198 & 0.0630 & 0.0947 & 0.3221 & 0.0291 & 0.0128 & 0.0089 & 0.0194
\end{tabular}

Hence the aligned age-sex distribution (from subprocess 3.3), that is, the oadscp vector, is as follows.
\begin{tabular}{c|ccccccccc} 
& 62 & 63 & 64 & 65 & 66 & 67 & 68 & 69 & 70 \\
\hline Male & 0.2486 & 0.1208 & 0.0468 & 0.1062 & 0.1889 & 0.1607 & 0.0485 & 0.0416 & 0.0380
\end{tabular}

The matrix oads, computed in this subprocess (4.3) is as follows. An explanation of how this matrix is generated appears below.
\begin{tabular}{c|ccccccccc|c} 
ageentRSB \(\backslash\) ageentAWD & 62 & 63 & 64 & 65 & 66 & 67 & 68 & 69 & 70 & Total \\
\hline 62 & 0.2486 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.2486 \\
63 & 0.0817 & 0.0391 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.1208 \\
64 & 0.0000 & 0.0468 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0468 \\
65 & 0.0000 & 0.0339 & 0.0630 & 0.0093 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.1062 \\
66 & 0.0000 & 0.0000 & 0.0000 & 0.0854 & 0.1035 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.1889 \\
67 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.1607 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.1607 \\
68 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0485 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0485 \\
69 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0094 & 0.0291 & 0.0031 & 0.0000 & 0.0000 & 0.0416 \\
70 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0097 & 0.0089 & 0.0194 & 0.0380 \\
\hline Total & 0.3303 & 0.1198 & 0.0630 & 0.0947 & 0.3221 & 0.0291 & 0.0128 & 0.0089 & 0.0194 & 1.0000
\end{tabular}

Note that the column total is oadsrs and the row total is oadscp. The oads matrix is determined using these row and column sum constraints. The nonzero entries of the oads matrix zigzag down and to the right. Starting at the upper left hand corner, the lesser of oadsrs and oadscp is put there. So oads \((62,62)=0.2486\).

In this case (as is usually the case), oadscp is less. Since, the difference oadsrs (62) - oadscp(62) \(=0.3303-0.2486=0.0817<0.1208=\operatorname{oadscp}(63)\), by the column sum constraint we are forced to have oads \((63,62)=0.0817\).

Now we have that the column sum is oadsrs(63). We want the row sum to be oadscp(63). We move one spot to the right to oads(63,63). Since oadscp(63) - oads \((63,62)=0.1208-0.0817=0.0391<0.1198=\) oadsrs(63), the entry oads \((63,63)=0.0391\).

With the row sum constraint met, we move one down. Since
\(\operatorname{oadscp}(64)=0.0468<0.1198-0.0391=0.0807\), the entry oads \((64,63)=0.0468\).

The row sum is now oadscp(64), and we move one down. Since \(0.1198-0.0468-0.0391=0.0339<\) oadscp \((65)=0.1062\), the entry oads \((65,63)=0.0339\).

Since \(\operatorname{oadscp}(65)-\operatorname{oads}(65,63)=0.1062-0.0339=0.0723>0.0630=\operatorname{oadsrs}(64)\), by the column sum constraint we are forced to have oads \((65,64)=0.0630\).

The column constraint is now met, and we move right one to oads(65,65). For the row sum to be \(\operatorname{oadscp}(65)\) we have \(\operatorname{oads}(65,65)=0.1062-0.0339-0.0630=0.0093\).

Since the row constraint is now met, we move down to oads \((66,65)\). oadsrs \((65)-\) oads \((65,65)=0.0947-0.0093=0.0854<\) oadscp \((66)=.1889\), so by the column constraint oads \((66,65)=0.0854\).

Now we can get the row total to match oadscp(66) by setting oads \((66,66)=0.1889-0.0854=0.1035\).

Now that this row constraint is met, we move one row down to oads \((67,66)\). Since oadsrs \((66)-\) oads \((66,66)=0.3221-0.1035=0.2186>0.1607=\) oadscp(67), by the row sum constraint we are forced to have oads \((67,66)=0.1607\)

Now that this row constraint is met, we move one row down to oads \((68,66)\). Since oadsrs \((66)-\) oads \((67,66)-\) oads \((66,66)=0.3221-0.1607-0.1035=0.0579>0.0485=\operatorname{oadscp}(68)\), by the row sum constraint we are forced to have \(\operatorname{oads}(68,66)=0.0485\).

Continuing with this logic, oads \((69,66)\) is set to oads \((69,66)=0.0579-0.0485=0.0094\), and we move right one to oads \((69,67)\). By the column sum constraint we are forced to have oads \((69,67)=0.0291=\) oadsrs \((67)\).

Now the column constraint is satisfied and we move right to oads(69,68). By the row sum constraint, this entry is forced to be oads \((69,68)=0.0416-0.0291-0.0094=0.0031\), .

Moving down one to oads \((70,68)\), we are bound by the column sum constraint to set, \(\operatorname{oads}(70,68)=0.0128-0.0031=0.0097\).

Moving to column 69, by the column constraint, the entry oads \((70,69)\) is forced to be 0.0089.

Finally, by the row and column constraints, the last entry, oads(70,70), is 0.0194.

To obtain the \(w\) matrix, normalize the rows by dividing by the row sum.
\begin{tabular}{c|ccccccccc} 
ageentRSB \(\backslash\) ageentAWD & 62 & 63 & 64 & 65 & 66 & 67 & 68 & 69 & 70 \\
\hline 62 & 1.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 \\
63 & 0.6763 & 0.3237 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 \\
64 & 0.0000 & 1.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 \\
65 & 0.0000 & 0.3192 & 0.5932 & 0.0876 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 \\
66 & 0.0000 & 0.0000 & 0.0000 & 0.4521 & 0.5479 & 0.0000 & 0.0000 & 0.0000 & 0.0000 \\
67 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 1.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 \\
68 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 1.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 \\
69 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.2260 & 0.6995 & 0.0745 & 0.0000 & 0.0000 \\
70 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.2553 & 0.2342 & 0.5105
\end{tabular}

Then, as matrices, opap \(1=w \times\) opap, as one may verify. For display purposes the transposes of opap and opap1 are shown.
\begin{tabular}{rrrrrrrrrr} 
opap & & 62 & 63 & 64 & 65 & 66 & 67 & 68 & 69 \\
1 & 0.9993 & 0.9990 & 0.9984 & 0.9978 & 0.9989 & 0.9974 & 0.9986 & 0.9975 & 0.9979 \\
2 & 0.9865 & 0.9816 & 0.9812 & 0.9738 & 0.9903 & 0.9786 & 0.9705 & 0.9632 & 0.9833 \\
3 & 0.9567 & 0.9467 & 0.9510 & 0.9347 & 0.9758 & 0.9466 & 0.9238 & 0.9140 & 0.9518 \\
4 & 0.9198 & 0.9109 & 0.9210 & 0.8971 & 0.9596 & 0.9179 & 0.8963 & 0.8739 & 0.9238 \\
5 & 0.8826 & 0.8777 & 0.8939 & 0.8652 & 0.9442 & 0.8981 & 0.8778 & 0.8475 & 0.9032 \\
6 & 0.8473 & 0.8445 & 0.8656 & 0.8360 & 0.9289 & 0.8835 & 0.8562 & 0.8226 & 0.8873 \\
7 & 0.8129 & 0.8132 & 0.8388 & 0.8105 & 0.9136 & 0.8644 & 0.8304 & 0.7994 & 0.8725 \\
8 & 0.7795 & 0.7817 & 0.8148 & 0.7838 & 0.8976 & 0.8477 & 0.8088 & 0.7838 & 0.8540 \\
9 & 0.7456 & 0.7518 & 0.7921 & 0.7605 & 0.8816 & 0.8310 & 0.7896 & 0.7663 & 0.8367 \\
10 & 0.7110 & 0.7214 & 0.7689 & 0.7352 & 0.8651 & 0.8156 & 0.7706 & 0.7497 & 0.8236 \\
11 & 0.6766 & 0.6907 & 0.7418 & 0.7111 & 0.8478 & 0.7987 & 0.7524 & 0.7370 & 0.8086 \\
12 & 0.6413 & 0.6574 & 0.7119 & 0.6842 & 0.8287 & 0.7823 & 0.7371 & 0.7249 & 0.7981 \\
13 & 0.6045 & 0.6235 & 0.6806 & 0.6526 & 0.8070 & 0.7656 & 0.7226 & 0.7164 & 0.7874 \\
14 & 0.5398 & 0.5584 & 0.6272 & 0.5968 & 0.7649 & 0.7246 & 0.6826 & 0.6849 & 0.7616 \\
15 & 0.4553 & 0.4746 & 0.5442 & 0.5207 & 0.7009 & 0.6608 & 0.6282 & 0.6380 & 0.7197 \\
16 & 0.3796 & 0.4009 & 0.4688 & 0.4548 & 0.6341 & 0.6142 & 0.5900 & 0.5999 & 0.6812 \\
17 & 0.3076 & 0.3308 & 0.3989 & 0.3902 & 0.5645 & 0.5667 & 0.5549 & 0.5641 & 0.6417 \\
18 & 0.2439 & 0.2651 & 0.3362 & 0.3278 & 0.4962 & 0.5144 & 0.5245 & 0.5250 & 0.6060 \\
19 & 0.1654 & 0.1855 & 0.2462 & 0.2459 & 0.4024 & 0.4476 & 0.4637 & 0.4715 & 0.5524 \\
20 & 0.0893 & 0.1069 & 0.1553 & 0.1604 & 0.2937 & 0.3671 & 0.3872 & 0.4130 & 0.4752 \\
21 & 0.0416 & 0.0529 & 0.0874 & 0.0959 & 0.1991 & 0.2839 & 0.3002 & 0.3442 & 0.3907 \\
22 & 0.0124 & 0.0186 & 0.0370 & 0.0411 & 0.1049 & 0.1762 & 0.1951 & 0.2482 & 0.2940 \\
23 & 0.0007 & 0.0014 & 0.0044 & 0.0038 & 0.0147 & 0.0338 & 0.0441 & 0.0698 & 0.1174 \\
24 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 \\
25 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 \\
26 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000
\end{tabular}
\begin{tabular}{llllllllll}
27 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 \\
28 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 \\
29 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 \\
30 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000
\end{tabular}
\begin{tabular}{rrrrrrrrrr} 
opap1 & 62 & 63 & 64 & 65 & 66 & 67 & 68 & 69 & 70 \\
1 & 0.9993 & 0.9992 & 0.9990 & 0.9985 & 0.9984 & 0.9989 & 0.9989 & 0.9978 & 0.9980 \\
2 & 0.9865 & 0.9849 & 0.9816 & 0.9807 & 0.9828 & 0.9903 & 0.9903 & 0.9807 & 0.9753 \\
3 & 0.9567 & 0.9535 & 0.9467 & 0.9482 & 0.9572 & 0.9758 & 0.9758 & 0.9516 & 0.9357 \\
4 & 0.9198 & 0.9169 & 0.9109 & 0.9157 & 0.9313 & 0.9596 & 0.9596 & 0.9258 & 0.9050 \\
5 & 0.8826 & 0.8810 & 0.8777 & 0.8862 & 0.9085 & 0.9442 & 0.9442 & 0.9071 & 0.8836 \\
6 & 0.8473 & 0.8464 & 0.8445 & 0.8563 & 0.8869 & 0.9289 & 0.9289 & 0.8918 & 0.8641 \\
7 & 0.8129 & 0.8130 & 0.8132 & 0.8282 & 0.8670 & 0.9136 & 0.9136 & 0.8731 & 0.8445 \\
8 & 0.7795 & 0.7802 & 0.7817 & 0.8015 & 0.8461 & 0.8976 & 0.8976 & 0.8562 & 0.8259 \\
9 & 0.7456 & 0.7476 & 0.7518 & 0.7765 & 0.8268 & 0.8816 & 0.8816 & 0.8395 & 0.8081 \\
10 & 0.7110 & 0.7144 & 0.7214 & 0.7508 & 0.8064 & 0.8651 & 0.8651 & 0.8236 & 0.7926 \\
11 & 0.6766 & 0.6812 & 0.6907 & 0.7228 & 0.7860 & 0.8478 & 0.8478 & 0.8065 & 0.7774 \\
12 & 0.6413 & 0.6465 & 0.6574 & 0.6921 & 0.7634 & 0.8287 & 0.8287 & 0.7896 & 0.7653 \\
13 & 0.6045 & 0.6107 & 0.6235 & 0.6599 & 0.7372 & 0.8070 & 0.8070 & 0.7719 & 0.7541 \\
14 & 0.5398 & 0.5458 & 0.5584 & 0.6026 & 0.6889 & 0.7649 & 0.7649 & 0.7307 & 0.7233 \\
15 & 0.4553 & 0.4616 & 0.4746 & 0.5199 & 0.6194 & 0.7009 & 0.7009 & 0.6676 & 0.6770 \\
16 & 0.3796 & 0.3865 & 0.4009 & 0.4459 & 0.5530 & 0.6341 & 0.6341 & 0.6170 & 0.6387 \\
17 & 0.3076 & 0.3151 & 0.3308 & 0.3764 & 0.4857 & 0.5645 & 0.5645 & 0.5653 & 0.6012 \\
18 & 0.2439 & 0.2508 & 0.2651 & 0.3128 & 0.4201 & 0.4962 & 0.4962 & 0.5110 & 0.5660 \\
19 & 0.1654 & 0.1719 & 0.1855 & 0.2268 & 0.3316 & 0.4024 & 0.4024 & 0.4385 & 0.5106 \\
20 & 0.0893 & 0.0950 & 0.1069 & 0.1403 & 0.2334 & 0.2937 & 0.2937 & 0.3519 & 0.4380 \\
21 & 0.0416 & 0.0453 & 0.0529 & 0.0771 & 0.1524 & 0.1991 & 0.1991 & 0.2658 & 0.3565 \\
22 & 0.0124 & 0.0144 & 0.0186 & 0.0315 & 0.0761 & 0.1049 & 0.1049 & 0.1613 & 0.2578 \\
23 & 0.0007 & 0.0009 & 0.0014 & 0.0034 & 0.0098 & 0.0147 & 0.0147 & 0.0302 & 0.0874 \\
24 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 \\
25 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 \\
26 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 \\
27 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 \\
28 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 \\
29 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 \\
30 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000 & 0.0000
\end{tabular}```


[^0]:    ${ }^{1}$ The Social Security area population consists of all persons who are potentially eligible to receive benefits under the Social Security program or who have the potential to work in covered employment. This population consists of residents of the U.S. and its territories, citizens living abroad, and beneficiaries living abroad.

[^1]:    ${ }^{2}$ The ages provided include $10-14,15,16,17, \ldots, 48,49-54$. Births at ages less than 14 are treated as having occurred at age 14 and ages reported to mothers older than 49 are treated as having occurred at age 49.

[^2]:    ${ }^{3}$ The average is calculated by giving each age in the group equal weight without regard to population. The age groups calculated are: 14-19, 20-24, 25-29, 30-34, 35-39, 40-44, and 45-49.
    ${ }^{4}$ For each year of the projection, the percentage of the slopes added to each ${ }_{5} b_{x}^{z}$ is reduced linearly such that $0 \%$ of the slope is used in the ultimate year. Thus, $0 \%$ is used for 2028 and later.

[^3]:    ${ }^{5}$ Data needed in order to project central death rates by cause of death were obtained from Vital Statistics tabulations for years since 1979. For the years 1979-1998, adjustments were made to the distribution of the numbers of deaths by cause. The adjustments were needed in order to reflect the revision in the cause of death coding that occurred in 1999, making the data for the years 1979-1998 more comparable with the coding used for the years 1999 and later. The adjustments were based on comparability ratios published by the National Center for Health Statistics.

[^4]:    ${ }^{6}$ Age groups are: less than $15,15-49,50-64,65-84$, and $85+$
    ${ }^{7}$ The five causes of death are: Cardiovascular Disease, Cancer, Violence, Respiratory Disease, and Other
    ${ }^{8}$ Age groups are: under 24 hours, 1-2 days, 3-6 days, $7-27$ days, 28 days- 1 month, 2 months, 3 months,..., 11 months, 1 year, 2 years, 3 years, and 4 years

[^5]:    ${ }^{9}$ Age groups are: $0,1-4,5-14,15-24, \ldots, 75-84$, and $85+$
    ${ }^{10}$ Age groups are: 15-17, 18-19, 20-24, 25-29,..., 40-44, 45-54, ..., 65-74, and 85+
    ${ }^{11}$ Age groups are: 15-19, 20-24, 25-29, ..., 90-94, and 95+
    ${ }^{12}$ Age groups for years prior to 1980 are: $0,1-4,5-9, \ldots, 80-84$, and $85+$. For years 1980 and later, the age groups are: $0,1-4,5-9, \ldots$, 90-94, and 95+.

[^6]:    ${ }^{13}$ Age groups are: $0,1-4,5-9,10-14, \ldots, 90-94$, and $95+$
    ${ }^{14}$ The five causes of death are: Cardiovascular Disease, Cancer, Violence, Respiratory Disease, and Other.

[^7]:    ${ }^{15}$ The federal fiscal year begins on October 1 of the previous calendar year and ends on September 30 of the specified calendar year.

[^8]:    ${ }^{16}$ The geometric mean, as used in this document, is the square root of the product of two numbers.

[^9]:    ${ }^{17}$ Data for 1980 is not available and is excluded from the calculations.

[^10]:    ${ }^{18}$ Data for 1988 was used to estimate the number of Puerto Rico and Virgin Island divorces for 1989-1997.

[^11]:    ${ }^{19}$ Using the Whittaker-Henderson method of graduation.

[^12]:    ${ }^{20}$ The midyear populations exposed to marriage are the unmarried populations (sum of those single, widowed, and divorced).

[^13]:    ${ }^{1}$ Group disaggregation includes age and gender. Some groups are additionally disaggregated by marital status and by presence of children.

[^14]:    ${ }^{2}$ More details on the hypothetical scaled workers are provided in Actuarial Note \#2005.3, located at the following internet address: www.socialsecurity.gov/OACT/NOTES/ran3/index.html.

[^15]:    Restricted Ordinary Least Squares
    QUARTERLY data for 101 periods from 1982Q4 to 2007 Q4

